## ROC/GIRO Working Party

## Best Estimates and Reserving Uncertainty

## Paper for GIRO 2007

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Many thanks also to those other members of the profession who have participated in the work of this working party for some of its life, and to those who responded to the survey, provided data and provided other useful insights.

Many thanks to Insureware for providing a working group and a submission to this working party.
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## 1. Summary and Introduction

### 1.1. A Word of Thanks

A great many people have contributed a lot of hard work to this working party beyond the members named, including former working party members, their colleagues and other members of the profession. I am very grateful to all who have taken part and have helped contribute to a wide ranging paper.

### 1.2.This Report

This report has four main components:

### 1.2.1 Best Estimates

Some brief thoughts on what we mean by best estimate. This is not intended to be a comprehensive treatise on the subject but rather to point out that when thinking about reserving uncertainty, and the range of potential outcomes for the ultimate claims payments, it is as well to understand what one means by a best estimate. Our thoughts on this subject are contained in Section 2, and we provide a suggested standard definition, based on a subjective estimate of the mean of possible outcomes.

We note that some of the commonly used reserving methods, such as the basic chain ladder, do not provide an estimate of the statistical mean, although it can be tempting to describe the estimates so derived as though they were estimates of the mean. We have not probed this matter further. It has been addressed in previous papers and there is another ROC working party looking at the performance of reserving methods.

### 1.2.2 Survey

We have carried out a survey of the profession to determine what methods are being used by members now in investigating reserving uncertainty. This is described in Section 3 and Appendix E. We see that the over-dispersed Poisson (ODP or Bootstrap) method is widely used. We note that there was a previous informal survey of the profession at GIRO in 1993 and we show high level results of that survey for an interesting comparison.

### 1.2.3 Applying Methods to Real Data

The third area of work, and the one which has involved dozens of volunteers, is in the application of a numbers of methods and models to some real claims development data, which was kindly provided to the working party for this purpose. The working party split into a number of sub-groups who each carried out testing. We also opened the doors to submissions from any professional service firm who wanted to test a method or model of their choosing and we received one such submission from Insureware.

Sections 4 to 8 deal with this phase of the work. Section 4 explains the background to this exercise and how we set about the work. Section 5 describes the data and the
issues we faced. Section 6 illustrates in a table the key features of the methods we tested and Appendix A gives a summary of the methods. Section 7 contains a numerical summary of the results of the working group testing and Section 8 summarises the qualitative comments about using the different methods, such as ease of use. Appendix F contains the complete Insureware submission.

It is interesting to note that there are areas of significant overlap between the distributions of outcomes produced by the various methods - and there are also important and noteworthy areas of non-overlap. Further, as the results in Section 7 show, it is not unusual for there to be significant differences between indications from various stochastic reserving techniques and not every method will give a reasonable result for every set of data.

It is also interesting to look at the full set of results, say for the Employers' Liability class (Graph 1 in Section 7). Some actuaries may wonder whether many or any of the method generate a $95^{\text {th }}$ centile value which is as high in relation to the mean as seasoned actuaries might expect, based on their experience of reserving such classes.

Having discussed these findings within the working party and with the Reserving Oversight Committee, we believe that - at least in the short-term - there is a clear need to apply a variety of methods rather than relying heavily upon the output of one, or even two methods. There also continues to be a need for careful professional judgement. In the longer term, as a profession, we feel there is an imperative to have a deeper understanding of the causes of variation between the results produced by the methods, and to add an intuitive understanding of the appropriateness of applying particular stochastic methods to the actuary's tool-kit. Since quantifications of reserving uncertainty will be becoming part of the statutory returns in the near future, the results of these methods are likely to attract greater scrutiny from outside the profession.

### 1.2.4 Numerical Simulation Testing of Stochastic Methods

This area of work had originally been intended as somewhat of an aside. We intended to use artificial simulated data to test how well the various methods handle unusual circumstances. We decided to begin by creating a baseline - testing the performance of the stochastic methods when all of their conditions are met. In theory, the simulations should show that $25 \%$ of outcomes fall above the method's $75^{\text {th }}$ centile result, and that $1 \%$ of outcomes fall above the method's $99^{\text {th }}$ centile result. In fact, we found some interesting results, with the methods most commonly used, applied in the way they are generally applied, tending to under-estimate the tails of the distribution. Considerably more than $1 \%$ of observations were falling above the $99^{\text {th }}$ centile produced by the methods in our testing. The work is described in Section 9 and the detailed calculations can be found in Appendix B.

On the one hand these could indicate that we have an important and urgent issue to deal with as a profession; on the other hand, there could be aspects of our analysis which are partly or wholly responsible for generating these anomalous results.

We have thought of three issues in relation to our testing work which could be contributing to the anomalous results at the extremes of the distributions and these are shown below.

- We may have made a calculation error in our testing. To seek to reduce this risk we have had a number of working party members reproducing the calculations independently.
- There may be a fundamental flaw in our approach to testing stochastic methods (for example, the method for generation of the artificial data may have introduced anomalies or inaccuracies at the extremes of the distribution). To reduce this risk we need careful consideration from the wider profession.
- When testing methods that produce only the first two moments of the predictive distribution for reserves (i.e. mean and standard error) we obtained percentiles by fitting particular analytic distributions (Log-Normal and Inverse Gauss) to the mean and standard error. We found that the Log-Normal and Inverse Gauss results did not differ materially, so we have presented detailed results only for the Log-Normal. However, it is still possible that using other analytic distributions would have led to different conclusions. (The use of the Log-Normal for this purposes was recommended by Thomas Mack in his 1993 paper, and we believe is quite common practice).

An alternative explanation for some or all of the anomalies is:

- The stochastic methods tested (including over-dispersed Poisson and Mack) may genuinely not work well at the tails of the distributions (for example, because functions of unbiased estimators, which are themselves approximately unbiased, may not be sufficiently unbiased in these extremes).

If this alternative is a true explanation for some, or all, of these anomalies then such would be a serious finding, because these methods are in widespread use across the profession, and in particular, are used for capital modelling where the extremes tails of the distributions are potentially in play.

We have set out our work here so that more members of the profession can consider the way we have approached it, and the potential implications, in the hopes that together we can better understand the true position and work together towards resolving any issues which may arise.

### 1.3 Judgement as a Method

### 1.3.1 Is Judgement a Method?

One of the methods we looked at and thought about was "Judgement". We had some interesting discussions about what this meant, and whether it was a meaningful method, both within the working party and with the Reserving Oversight Committee. As you will see in Section 8 the working groups had differing views on this.

We thought it would be helpful to add some words about what we mean by judgement as a method, how it can work in combination with other methods, and why it is important.

### 1.3.2 What is the Judgement Method?

Actuaries apply professional judgement in most of their work, from choosing development factors to exclude from averages, to interpreting data and softer information. That is not what we have in mind as the judgement method, but rather the following.

The judgement method, in the context of estimating quantifications of reserving uncertainty, is a use of the actuary's own experience as an implicit benchmark for expected results. In fact this definition may work in wider contexts.

The actuary may have implicit experience-based benchmarks for many possible values, including for example, tail factor extremes, the ratio of tail centiles to means, "worst case" loss ratios for certain classes of business, exposure based method parameters.

There are two methods related to the judgement method, being benchmarking and scenario testing.

Benchmarking is the use of explicit data to create benchmarks which will guide the selected estimates, or parameters involved in their derivation. Judgement is similar because the actuary is essentially using his or her own, potentially complex and interacting, set of benchmarks based on personal experience. When an actuary has become sufficiently confident in their implicit benchmarks they may become "rules of thumb" and be adopted by other actuaries and spread through the profession’s consciousness.

Scenario testing is the creation of possibly hypothetical scenarios to be used as deterministic parameterisations for certain calculations. The choice of scenarios to consider may be based on the actuary's own experience and hence a manifestation of the judgement method.

### 1.3.3 How does Judgement Work in Combination with other Methods?

Actuaries may carry out a number of methods in estimating the distribution of possible ultimate claims outcomes. In so doing they may combine these methods with the judgement method. For example, the actuary may believe that the $95^{\text {th }}$ centile of the distribution of outcomes is being underestimated by his or her chosen method. This could be because in his or her experience they may feel they have seen sufficiently many examples of adverse development to suggest a higher level of deterioration for a 1 in 20 event. They therefore adjust the results of their method to align the aberrant parts of the distribution more in line with their own experience.
1.3.4 Why is the Judgement Method important?

The judgement method is important because it is so often used in practice, whether or not this is stated explicitly. Given that different actuaries have different past experiences it seems likely that they may at times have different implicit benchmarks which they have developed from their respective experiences. It is arguably therefore all the more important that we disclose when we are using this method, and seek to articulate the implicit benchmarks we are using.

There are clear risks with the judgement method, especially if it is not used in a transparent manner. Some actuaries may have limited experience, or their experience may not be typical, and so their personal suite of benchmarks may be unreliable. The ability to assimilate experiences in such as way as to differentiate between different types of circumstances to create sufficiently intricate implicit benchmarks may vary from person to person, so that some develop more robust implicit benchmarks than others. Peer review can clearly reduce these risks.

### 1.3.5 Development of the Profession’s Judgement on Reserving Uncertainty

One of the objectives of this working party is to add some material to assist in developing the profession’s judgement on estimating reserving uncertainty.

We are acutely aware of the need to resolve the issues raised in 1.2 .4 so that we can be sure that the profession's growing experience and rapidly forming implicit personal benchmarks, especially for the tails of the distributions, are based on a robust foundation.

### 1.4 Next Year

While we have carried out a great deal of work this year there remains a lot still to do. In Section 10 we set out some objectives for next year's working party to carry on the work.

## 2. Best Estimates

### 2.1 Definitions of "Best Estimate"

What is an actuary's "best estimate" of an outcome, for example, of ultimate losses?
It is her/his subjective derivation of the mean of all possible outcomes, taking into account all available information about the business being analysed.

This definition allows for the subjective interpretation by the actuary of the available data and the choice of models and methods used. As a default it also excludes allowance for events not reflected in the data such as unanticipated major new types of latent claims, although in some cases it could be argued that all available information includes knowledge about the risk of such new claims potentially arising.

Subjectivity is included in the definition because it emphasises that even with a standard definition of "best estimate" one would still expect different parties to produce different estimates. The uncertainty about the random process being estimated and the data provided gives grounds for different judgements to be made about how the future will unfold. Also, different actuaries may understand the "available information" to include or exclude different matters. Therefore a variety of reasonable best estimates is possible, even with a common definition.

However, if the actuary diverges from the above definition then we would recommend that a detailed description be given, defining exactly what is being estimated. This should include comment on:

- Why the above standard definition is not being used in a particular instance and why an alternative definition is thought to be a more appropriate single point estimate, "best" summarising the true underlying random probability distribution. Other summary statistics which might be favoured include the mode and the median
- Whilst the actuary's view of the mean of the whole distribution of all possible outcomes should not vary depending on the purpose of his or her work, the purpose may influence which areas of divergence from this complete mean may be acceptable or desirable. The actuary should therefore comment on the purpose of the best estimate and how this influences the choices made about what, if any, contingencies to exclude from the set of possible outcomes.
- How prudent, optimistic or pessimistic the estimate is intended to be. In this case it is preferable to define any deviation from the standard definition explicitly; for example "The best estimate is the mean of the underlying distribution of the claims allowing for all possible outcomes, plus a margin of prudence." (Giving reasons for the prudent stance)
- Whether unlikely outcomes are included or excluded (if excluded explain why). It is worth differentiating between remote events that are significant to a particular insurer and those that could cause significant proportions of the industry to collapse
- Whether an attempt has been made to provide for latent events


### 2.2 Mean versus Median

There is sometimes a debate about whether best estimates should be means or medians.

The mean definition of a best estimate accords with UK professional guidance in GN20 and it corresponds to equivalent definitions used internationally - albeit using different terminology (i.e. "expected value estimate" of a reasonable provision in the US - ASOP 36 and "central estimate" in Australia - PS300). It accords with the anticipated reserving requirements under Solvency 2 and IFRS.

The mean is arguably a more intuitive choice and it has the benefit that the expected average surplus equals the expected average deficit. That said some of our stakeholders may assume the best estimate is "the midpoint", or that it is "equally likely to be too high as too low", which indicates a median rather than a mean.

The US system, albeit using an "expected value" terminology, does not necessarily require the actuary to anticipate future contingencies and so, if it is a mean, it is perhaps a mean of a distribution with some of the extreme tail removed. Given the typically skewed nature of distributions of ultimate claims outcomes, such an estimate may be nearer to the median of the whole distribution than the mean.

This working party has come down in favour of the mean in this debate.

## 3. Survey and Results

### 3.1 Survey of GI Actuaries

Given the ROC brief to investigate reserving uncertainty methodology, it seemed appropriate to include practical evidence of the methods and views of the industry towards reserving uncertainty.

A questionnaire was written - see Appendix E1, and distributed to members on the GIRO mailing list by Peter Stirling on $17^{\text {th }}$ May.

The accompanying text follows:
"To General Insurance Actuaries on behalf of GI ROC
As you probably know, four work streams of GI ROC (General Insurance Reserving Oversight Committee) were announced at last year's GIRO convention. The 'Best Estimates and estimating uncertainty' and 'Effectiveness of reserving methods' working parties would like your help in completing a brief on-line survey as follows:

The GI ROC working parties will be reporting on their investigations into the assessment of best estimates and reserving uncertainty at this year's GIRO conference. A survey has been developed to put the results into a practical perspective and to assess where the industry is up to in terms of the methods currently in use. Your participation in this survey would be very much appreciated. All responses will be anonymised and only the combined results in summary form will be publicly disclosed following analysis of the raw data. Please access the survey by clicking on the link below. The survey is open for responses until Friday 25th May.
http://wam.actuaries.org.uk/eforms/eForm.aspx?TAG=ROC
Many thanks in anticipation of your help."

### 3.2 A Previous Survey in 1993

A 1993 GIRO working party produced a paper entitled "Variance in Claims Reserving" that discussed the principles of stochastic claims reserving and tested several methods current at the time on actual datasets. That paper is available under "general insurance convention papers archive" on the Institute of Actuaries website.

An informal survey on use of stochastic reserving methods was also carried out at the 1993 GIRO. $40 \%$ of those present who were involved in claims reserving said they had used a stochastic method at one time or another. The table below shows numbers who had used various classes of stochastic method.

Table 3.2.1 - Numbers of attendees at 1993 GIRO who had previously used stochastic reserving methods in practice

|  | Log incremental <br> regression (static) | Log incremental <br> regression <br> (dynamic) | Bootstrap | Operational <br> time |
| :--- | :--- | :--- | :--- | :--- |
| Consultants | 3 | 3 | 1 | 7 |
| Lloyds/RI | 2 | 0 | 0 | 0 |
| Insurers | 7 | 1 | 2 | 3 |
| Total | 12 | 4 | 3 | 10 |

Log incremental regression (static) refers to regression methods where the dependent variable is the log of the aggregate incremental paid amounts and the underlying runoff pattern is assumed to be the same for all origin years. Log-incremental regression (dynamic) refers to similar methods in which run-off pattern is allowed to vary across origin years using the Kalman filter (this method was advocated by Insureware at the time). Mack's method was not considered by the 1993 working party as it had only just been published.

### 3.3 The Results of our 2007 Survey

Please see Appendix E2 for the detailed results.
There were 47 respondents, mostly based in the UK, from a mixture of Insurance companies, Reinsurance companies, Lloyd's syndicates and consultancies. Just over half represented insurance companies. Please see Chart 1 and Chart 2 of Appendix E2.

Respondents were invited to report on the classes of business for which their organisation performed reserving work. Reinsurance companies covered most of the classes of businesses - see Chart 3. The largest number of responses related to Personal lines - property, Personal lines - motor, Commercial lines - property and Commercial lines - liability, although for Lloyd's syndicates Marine was also significant.

For the 3 largest classes, respondents were asked which methods and models they used to calculate best estimate values and to determine uncertainty.

The results from question 3a (best estimates) indicate that the Basic Chain ladder incurred is most often the key method, followed by Bornhuetter Fergusson, paid chain ladder and then judgement. Alternative approaches included using loss ratios from pricing work.

The results from question 3 b (reserving uncertainty) indicate that the key methods were: Overdispersed Poisson (ODP) stochastic chain ladder/ bootstrap, judgement, scenarios and then basic chain ladder paid and incurred methods. Alternative methods included using benchmark information from the market.

Question 4 asked about the key requirements for determining uncertainty. They were:

1) to identify variability around the best estimate and,
2) to identify the tail of the distribution (above $95^{\text {th }}$ percentile).

Question 5 asked what were the factors influencing a choice of method/model. The choice was decided based upon

1) Practical issues relating to the data
2) The quality of results that the model/method will give.

Interestingly, "knowledge of the actuarial staff" was listed as the least important factor. This indicates a confidence within the industry that actuarial staff understand the relevant models being used.

Question 6 asked what systems/software were used to model uncertainty. When modelling uncertainty organisations were evenly split between using in-house models and a combination of in-house and off-the-shelf packages. The packages mentioned included ResQ, Igloo, RMS Risklink and @Risk. ResQ was most frequently mentioned.

Question 7 asked about the methods used to communicate uncertainty to senior management/ executives.
The most frequent methods used were:

1) Quantitative using percentiles.
2) Quantitative using everyday English.

A satisfyingly low number didn't communicate uncertainty at all.
Question 8 asked about what aspects of uncertainty were covered in communications. Communication of uncertainty most commonly covered process and parameter uncertainty, but less often covered model uncertainty.

Questions 9 to 13 asked for additional information was collected relating to the groupings by time periods commonly used in reserving to calculate best estimates and for calculating uncertainty.

For calculating best estimates, the majority grouped data using annual origin periods. The most common grouping by development period was quarterly.

For calculating uncertainty, the majority grouped data using annual periods and annual development periods.

Reserving, in most instances was done quarterly using a full analysis. Where reserving was done monthly, a reduced analysis was more often performed.

### 3.4 Our Thoughts and Conclusions

When drafting the survey, the extent to which different institutions would use different methods was not known. Since the GRIT paper was issued in 2006, it was
hoped that its key recommendations had been adopted. This survey indicates that most respondents do communicate and quantify uncertainty.

The survey shows reliance on a small number of standard methods for producing estimates of reserve uncertainty, and the key requirements of such methods are to identify variability around the best estimate and identify the tail of the distribution. Other parts of the ROC paper give an independent assessment of several of the methods that can be used and illustrate their limitations.

The answers to question 5 reveal a confidence from respondents that actuarial staff members are equipped with sufficient knowledge to deal with whichever method/model is chosen to measure uncertainty. This may be owing to high quality actuarial training, or may indicate a degree of misplaced complacency - particularly if the limitations associated with the methods used are not fully understood.

The above results are representative of only a small proportion of the industry, and represent the views at a particular point in time. It is hoped that this survey (or a similar one) could be re-issued on an annual basis to detect changes that the industry are making in reserving methodologies.

Ideally the results would have been split and analysed by line of business. This has been done, but not published in Appendix E2, owing to the small sample sizes and similarities between the responses. Where the responses were not very similar, they are detailed below.

Judgement and scenarios were more used commonly to calculate best estimates and reserving uncertainty used in reinsurance lines than was the case for direct writers.

Reinsurers also were more interested than primary insurers in identifying the complete distribution of reserve variability than just the tail of a distribution. The key requirement though was variability around the best estimate for both types.

Personal lines insurers chose a method to identify reserve uncertainty based more upon the quality of results than the quality of the data. Commercial lines and reinsurers chose a method based more upon the available data.

The survey questions about the groupings of data suggest that different groupings by origin/development year are used when determining best estimates compared to reserve uncertainty. Perhaps further studies could be performed to indicate whether this impairs the accuracy of the best estimates/reserve uncertainty results.

## 4. Method Testing: Our Scope and Approach

### 4.1 Background

There have been many developments in recent years that have generated increased interest in the assessment of uncertainty within the reserving process. These have largely been driven by the Morris review, which identified the lack of a consistent approach or set of approaches relating to both the estimation and communication of uncertainty within point estimates for claims reserves within formal actuarial reports.

The work of GRIT, which was intended to pre-empt and also respond to the Morris report investigated the calculation of reserve uncertainty to a limited extent, focussing largely on the way such concepts are understood within the profession, as well as how they are communicated outside, and how such information is received and interpreted.

Further impetus has been provided by the re-interpretation of the GN12 guidance on formal actuarial reports, which has underlined the need for actuaries to indicate both the source and amount of uncertainty within their work, where practicable.

Additionally various authors continue to investigate the statistical credibility of various models, and develop increasingly sophisticated methods to derive more reliable and consistent results to compare to best estimates and standard actuarial reserving tools.

Separately from reserving uncertainty alone, the introduction of the Individual Capital Assessment by the FSA, and to a lesser extent the discussions relating to IFRS, have brought the attention of parties both inside and outside the profession to bear on uncertainty in general, with reserving uncertainty being a strong element of these discussions.

These discussions have focussed on the definition of adequacy around particular centiles of probability, implicitly implying that a full distribution of outcomes is required to derive the associated processes.

The above sources of interest in this area are relatively well reported, however there are other, less public, areas that have increased the need for research and discussion in this field. These are through the proliferation of more advanced computing power and the increasing use of off-the-shelf reserving packages that include "standard" reserving uncertainty methods, which can be run almost as an afterthought to the reserving process.

This has increased the ease of obtaining reserve uncertainty information, but also has opened the question of how much the explicit and implicit assumptions relating to the models are understood by the reserving actuary, let alone the user of the report.

The above issues have raised the level of awareness of reserving uncertainty as a topic within the profession, and we believe there is a need to broaden the general knowledge base within the profession on the methods available, their strengths and
limitations and to generally demystify the area for the general insurance population at large.

### 4.2 GRIT Conclusions

One of the key recommendations from GRIT was: "Providing more information on uncertainty in our reserve estimates. In particular, we recommend that actuaries provide a quantitative indication of the range of outcomes for future claim payments...".

The 2005/6 Estimating Reserving Uncertainty Working Party following GRIT also showed that only a limited number of actuaries were using defined methods within the reserving process or assessing reserve uncertainty at all, and that those that did tended to rely on a single method.

This paper is intended to help members of the Profession understand more about the tools at their disposal, as well as giving some information about the relative merits and limitations of such methods. It is hoped that this may help in the fulfilment of the GRIT recommendation, and also to ensure that a follow-up to the Working Party next year would indicate better understanding and use of such methods.

### 4.3 What the paper is attempting to do

The key objectives of this paper are:

- to educate the majority of general insurance actuaries so that they are aware of the wide variety of methods available;
- to educate the majority of general insurance actuaries so that they are able to understand at a basic level what the differences are between these models are;
- to give a high-level review of the practical and theoretical aspects of a few of the more common models based on application to real data sets;
- to provide resources of data, models and references with which to accustom actuaries to these models.


### 4.4 What the paper is not attempting to do

We are definitely not attempting to suggest that a particular model is somehow "better" than any other, nor are we trying to review all possible methods. In particular we are not attempting to bring the reader up to date with the latest developments in the field, as we are concentrating on educating from the most basic level.

We expect future working parties to extend the work presented here in response to feedback from the profession on this paper, and this is discussed in more detail in Section 10.

Before we get into more detail about the work we have done, we would like to draw the reader's attention to the work of the ROC best estimate working party, which is reviewing more traditional actuarial models, and assessing their effectiveness in various circumstances. We have shared data with them, and intend eventually to present combined results to compare best estimates and ranges. This will follow next year.

### 4.5 Our Approach

Briefly, our approach centred on selecting a few methods we could look at for a number of datasets, and could compare the observations of a number of sub-groups of the working party for each method employed. This process would not only look at the technical results, but also assess each method for ease of use, understanding and interpretation. The results of these analyses have been combined to present an assessment of the methods under various categories to help identify those with strengths and weaknesses in particular areas.

The following paragraphs set out this process in more detail.
We selected a number of methods for assessment. The selection was largely based on the knowledge within the working party itself as to the most common methods within the profession, as well as through literature searches and discussions with colleagues.

This "long" list was then reduced to a selection of eight, including purely subjective approaches and methods that required a level of detail within the data that is not usually available, particularly for consultancy work or for older data sets. We also provided our data to consultancies to allow them to provide results on their own models for comparison, we received a completed response from one consultancy, which is included in full in Appendix F. These methods are briefly discussed in Appendix A.

The data we applied these methods to was sourced anonymously from a number of companies. We attempted to cover "normal" lines of business, as well as some more unusual classes such as Marine. The classes included long and short tail, and had data of different periodicities and levels of detail. In particular we unfortunately were not able to obtain transaction level information, which reduced the number of available methods by one.

The data was adjusted such that the source was not easily identifiable, however this process was not well defined, and hence the consistency of some data sets was questionable. However, as the key objective was to test the methods against each other, such inconsistencies did not necessarily result in a data set being rejected.

To concentrate the results of the models we selected three core data sets for analysis. These represented an employers' liability class, a commercial property class and a motor class. Other classes were modelled for some methods, but they did not form the core of our work. The data is discussed in more detail in Section 5.

The working party was split into a number of sub-groups, each of which was given two methods to investigate, as well as considering the methods of scenario testing and judgement.

Each method was assessed using a standard questionnaire, which investigated various aspects of the method, both numerical and practical. Although we tried to avoid using off-the-shelf packages for all groups working on a particular method, where they were available use was made of them to investigate the effects of implicit assumptions relating to ensuring that the method could operate on the data sets provided.

It was therefore intended that each of the methods was investigated by a number of groups independently. However, there were fewer responses than originally envisioned, as discussed in Section 7. The results of the investigations for each were compared to identify any sources of user error or where off-the-shelf packages had used a more advanced version of the model than from the primary reference. In addition the practical aspects of the models were discussed to assess the less technical areas for investigation. Sections 7 and 8 discuss the numerical and qualitative comparisons respectively, with Section 6 giving a high-level comparison of the results.

We also asked for input from Professional Service firms, and Insureware made a submission using some of their own methods applied to our datasets. In particular, in Section 7 we show the results for the Employers' Liability class which compares the methods tested by the Working Party subgroups with modelling frameworks utilised by Insureware. The full submission from Insureware can be found in Appendix F.

A separate strand of our work focussed on assessing the theoretical accuracy of a number of models under ideal conditions. This was achieved using stochastic simulation whereby the models were applied to a large number of simulated triangles, each with a known ultimate. These data sets have been derived such that the assumptions required by those models are fulfilled. The aim of these tests was to calculate the accuracy of the model in estimating the probability of events at particular centiles. This investigation is set out in Section 9.

## 5 The Data

### 5.1 Introduction

This section considers the data requirements for the application of the various methods and the limitations imposed by the data available to the working party.

### 5.2 Required data characteristics

### 5.2.1 Real or artificial data

The working party set out to test a range of methods for estimating uncertainty by considering the merits and disadvantages associated with their practical application.. This "real world" testing suggested that the methods should be tested using real data. This presented a number of problems. Firstly there is no central repository of insurance data in the UK and secondly distinguishing how different features of the data are impacting different methods can be extremely challenging.

The Institute of Actuaries has arranged the collection and rescaling of real data from companies who were willing to contribute. The rescaling methodology has not been disclosed to the members of the working party but it is understood that the approach used is sufficient to ensure that individual data is not recognisable without distorting the variability within that data. The working party recommends that the resulting data is not appropriate for reaching conclusions around the relative performance of any sectors of the insurance industry as premium income and loss information from the same entity may have been rescaled on different bases.

The amount, timing and uncertainty of emerging cash flows is affected by a wide range of factors arising from policyholder attitudes and behaviour, the companies’ own management strategies and practices and the wider economic and legislative environment. It is often difficult to interpret the results of analyses when using familiar, well understood actuarial techniques such as the chain ladder or BornhuetterFerguson technique. Understanding the results is an even greater challenge when the techniques being applied are not widely used and understood.

Artificial data can be designed to avoid the distorting features of actual live insurance data. An alternative approach to testing using real data is to develop artificial data tailored to the characteristics of the method being tested to produce a "control" set of results. Understanding of the method's response to some of the features of general insurance can be achieved by adding new features to the artificial data and comparing the results with the control.

The working party has tested the Mack and ODP methods (both analytic and bootstrap versions) using artificial data. An overview is given in Section 9, and details in Appendix B.

### 5.2.2 Aggregate or transactional data

Ideally the working party would have preferred to work with transaction level data and aggregated it according to the requirements of each method. Although this approach is time consuming it means that data can be aggregated in a variety of different ways and a deeper level of investigation and understanding is possible. In practice aggregated data is more readily available and, as described above, easier to alter to ensure the source of the information is not recognisable.

The working party, in drawing up a list of methods for measuring uncertainty, identified techniques that required transactional data and could not be run using aggregated triangulation data. Although we understand that access to transaction data is still being sought, this data has not been available to the working party and consequently the techniques that rely on such data have not been tested.

Aggregate data has been made available to the working party. Paid and incurred claims triangulations have been provided for all except one data set which excluded paid data. A reasonable number of years of exposure, varying from 8-20 years, have been available for most classes. For longer tail liability risks this volume of data is not sufficient to show fully mature historical loss development and it is necessary to make assumptions regarding future development to ultimate. Development intervals varied by data set between monthly, quarterly and annual. This has not presented a problem for testing conducted to date but does prevent consideration of how the techniques would respond to quarterly analyses for many of the data sets provided. Exposure information is available for 11 out of the 30 data sets and separate large loss information has been provided for 13 of the data sets. Claim counts are available for half the data sets although for two of these claims counts are only in relation to large losses.

Some very limited qualitative information has also been supplied in respect of each data set namely:

- Class and sub-class of business;
- Gross or net of reinsurance (not known for all data sets).

The sub-class information is useful in separating out some specific risks from the underlying data. For example subsidence and weather losses are separately identified from other property losses. Separate data for bodily injury and non-bodily injury motor claims are also available.

The table attached in Section 5.3 below summarises the aggregate data available to the working party.

Whilst the working party acknowledged from the outset that access to high quality data was likely to be a problem, the information available is not ideal. We anticipate most actuaries would be able to access better quality information in the normal course of their work. It is noted, however, that in the testing of the methods it is useful to know which methods can be applied when there is restricted or less than ideal
information. The practical constraints of applying each method, including data requirements, are considered in Section 8.

### 5.3 Classes of Business

The working party has been relying on information donated to the institute so has not been able to specify its own data requirement. A range of classes of business have been provided including:

- Property - personal lines and commercial lines
- Motor - personal lines and commercial lines
- Marine
- Construction
- Liability - employers' and public

In order to ensure maximum comparability of the results it was decided to test each of the selected methods on a limited number of classes of data in the first instance. A review of the data resulted in the recommendation to use the following three classes:

- Commercial Property
- Personal Motor Non-comprehensive
- Employers’ Liability

The Commercial Property and Personal Motor Non-Comprehensive data sets both provide triangulations by accident quarter and quarterly development period for payments, incurred losses and reported claim counts with corresponding earned premiums and earned exposure. The data is gross of reinsurance. The underlying business is UK risks sourced through intermediaries, direct sales and corporate partners. Large claims are included in the underlying loss experience and are not available separately.

Apart from one anomalous figure, assumed to be a data error, the loss development for the Commercial Property business is very stable. In contrast the Personal Motor Non-comprehensive data demonstrates a greater degree of variation around the average development as would be expected for a liability class of business. Comparing the results for these two classes under a number of different methods will assist in considering the extent to which the various methods respond to the variability in the underlying loss data.

The Employers' Liability data set provides annual development triangulations by annual exposure period for payments, incurred losses, settled claim counts and incurred claim counts. No premium or exposure data is provided. The data is provided for three sub-classes and is further sub-divided between large losses and attritional claims. No qualitative information is provided on the nature of the underlying sub-classes (hence these are not identified separately in the table below).

The sub-classes demonstrate very different features. One subclass appears fully developed after 10 or 11 years whilst the other two are still developing after twenty years (the limit of the data), one of which is demonstrating latent development after 12 years and relatively high variability compared with the other two subsets. This data is particularly interesting in terms of testing the models but the eventual results will be highly dependent on individual assumptions in relation to the tail development beyond the 20 years development history provided.

The following table provides a summary of all the data sets available to the working party.

| Class | Sub Class | Paid / Incurred? | Origin Period | Development Period | Large Losses? | Exposure? | Net of Reinsurance? |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Construction | Employers' Liability | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | ? |
| Construction | Public Liability | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | ? |
| Construction | Damage | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | ? |
| Employers' Liability | Employers' Liability | P \& I | 1991-2005 | Annual | By Class. Not by sub Class | No | ? |
| Public Liability | Public Liability | P \& I | 1991-2005 | Annual | By Class. Not by sub Class | No | ? |
| Employers' Liability | Employers' Liability | P \& I | 1985-2005 | Annual | Yes | No | No |
|  |  |  |  |  |  |  |  |
| Home | Direct Subsidence | P \& I | 1995-2005 | Monthly | No | No | ? |
| Home | Int Subsidence | P \& I | 1995-2005 | Monthly | No | No | ? |
| Home | Direct Non subsidence | P \& I | 1998-2005 | Monthly | No | No | ? |
| Home | Int non Subsidence | P \& I | 1998-2005 | Monthly | No | No | ? |
| Comm Property | Comm Property | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Household Weather | Household Weather | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Property | Weather | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | No |
| Property | Ex Weather | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | No |
|  |  |  |  |  |  |  |  |
| Marine | Cargo | P \& I | 1988-2005 | Annual | Yes (for Gross) | Yes | Yes |
| Marine | Liability | P \& I | 1988-2005 | Annual | Yes (for Gross) | Yes | Yes |
| Marine | Hull | P \& I | 1988-2005 | Annual | Yes (for Gross) | Yes | Yes |
|  |  |  |  |  |  |  |  |
| Motorcycle | Motorcycle | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Motor Fleet | Bodily Injury | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Motor | Direct injury | P \& I | 1990-2005 | Monthly | No | No | ? |
| Motor | Int injury | P \& I | 1998-2005 | Monthly | No | No | ? |
| Motor | Direct Non injury | P \& I | 1990-2005 | Monthly | No | No | ? |
| Motor | Int non injury | P \& I | 1998-2005 | Monthly | No | No | ? |
| Pers Motor Non Comp | Pers Motor Non Comp | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Pers Motor Non Comp | Pers Motor Non Comp | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
|  |  |  |  |  |  |  |  |
| Business Interuption | Weather | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | ? |
| Business Interuption | Ex Weather | P \& I | 1985-2005 | Annual | By Class. Not by sub Class | No | ? |
|  |  |  |  |  |  |  |  |
| Personal Creditor | Personal Creditor | P \& I | 1997-2006 Q3 | Quarterly | No | $>2003$ | No |
|  |  |  |  |  |  |  |  |
| Prof Indemnity | Prof Indemnity | P \& I | 1997-2006 Q3 | Quarterly | No | Yes | No |
| Prof Indemnity | Prof Indemnity | I | 1989-2005 | Annual | No | No | No |

### 5.4 Items Beyond the Data

Assessments of uncertainty need to also consider the potential for differences in future claims experience when compared with past experience as depicted by the historical loss data. The working party considered that appropriate methods for evaluating uncertainty would need to be able to take into account anticipated future changes relative to past experience.

The working party considered the potential sources of uncertainty which may not be reflected or may only be partially reflected in the data sets used and came up with the following non-exhaustive list:

- Inflation
- Policy data e.g. lack of historical split by risk factors
- Rating indices
- Latent claims
- Case estimation practice
- Data quality
- Circumstances of business
- Staff turnover
o Claims handlers
o Underwriters
o Actuaries
- Product bundles
- Aggregate policies - combined data
- Legal changes e.g. Ogden tables
- Pricing change data
- Underwriting cycle
- Terms and conditions
- Period claims order
- Known events not in data
- Social and economic factors
- Knock-for-knock
- Contract certainty
- Large catastrophes (extreme events or just catastrophes?)
o Impact on claims philosophy
o Demand surge
- Climate change
- Return period of large claims
o What should be stripped out
o Potential impact on UPR
- Distribution methods
- Geographical exposure changes
- Changes in reinsurance
- Influence of rating agencies
- Influence of market views
- Reserving process
o Anchored knowledge
o Personal independence
Each of these factors not only has the potential to change the expected amount and timing of future cash flows but the uncertainty around the amount and timing of those cash flows. A challenge for measuring uncertainty will be how actuaries can reflect these and other factors affecting the future performance of the business within the methods available.


## 6 Key Features of the Methods Tested

|  | ODP/Bootstrap | Mack | Bayesian/BF Method | Judgement | Scenarios | Regression/Curve Fitting |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Description | Most common bootstrap model. Potential to use different distribution for the residuals | Calculation of standard error with and without tail factors. | Uses ODP model with a series of prior ULR estimates defined by a distribution | Based on professional experience | Can include any variation such as changing development patterns or single events | Fits Craighead curve to each origin year to derive initial estimate of ULR, then smoothes across origin years using regression |
| Data required | Cumulative claims triangles (paid or incurred) | Cumulative claim triangles (paid or incurred) | Cumulative claim triangles (paid or incurred) | Any | Any | Premium and claim amounts triangles |
| Is the method acceptable to the Profession? | Yes | Yes | Yes | Yes | Yes | Depends on purpose |
| Is the method easy to use and is it practical? | Yes | Yes | No | Yes | Yes | Yes |
| Can judgement or amendments be applied? | Yes | Amendments needed where gaps in published method | Requires prior distribution of ultimate position of each origin year | Yes - essential | Yes via choice of scenarios and manual adjustments or tweaks | Yes, perhaps too easily |
| Is the method easy to explain? | Principles easy to explain | No | Very difficult | Yes | Yes | Yes |
| When is method good? (Or not?) | Good if little negative <br> development and residuals are iid and run-off pattern is same for all years. | Good only if run-off pattern is same for all years | Good if little negative development and residuals are iid and run-off pattern is same for all years. | Good if actuary has additional knowledge; bad if not experienced | Not good if volatile datasets or inexperienced actuary | Good if run-off pattern varies across origin years. Not good if there is much negative development |


|  | ODP/Bootstrap | Mack | Bayesian/BF Method | Judgement | Scenarios | Regression/Curve Fitting |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Are extreme events included? | Only if in data | Only if in data | Yes, if in data and/or in prior distributions | Yes if desired | Yes if desired | Yes, if in data but can exclude if desired |
| Produce complete distribution of outcomes? | Yes if process error is simulated in addition to bootstrapping for parameter error | Produces mean and standard error only | Yes | Yes as any required percentile can be estimated using judgement | No - produces a few possible outcomes to which probabilities can be judgementally applied | No, just an approximate range |
| Type of uncertainty measured | Bootstrap method gives parameter uncertainty, process uncertainty can be simulated in addition | Process and parameter uncertainty | Process and parameter uncertainty | Potentially model error as well as parameter and process error | Usually just parameter uncertainty | Parameter uncertainty only (dependent variable in regression is expected ULR) |
| Time to program and complete | Easy to program in Excel though long time to run | Easy to program and quick to run | Specialist software required and very slow to run |  |  | Easy to do in Excel |
| Comparison of class results to aggregated | Automatic consistency between origin year and aggregate results | Automatic consistency between origin year and aggregate results | Automatic consistency between origin year and aggregate results | Should be consistent given enough care, but this not guaranteed | Does not produce separate assessment of aggregate uncertainty | Does not produce separate assessment of aggregate uncertainty |

## 7 Numerical Summary of Results

### 7.1 Summary of Approach

The Working Party has tested and compared a variety of methods to quantify reserving uncertainty. The working party members were split into groups with each group applying a method on 3 datasets. These datasets were chosen for their dissimilarity - Employers' Liability, Personal Motor and Commercial Property. More than one group attempted some of the methods.

### 7.2 Comparison of Methods

A key aim of this analysis was to assess how comparable the results were from different method frameworks. The following Sections show a graphical comparison of the mean, inter-quartile range and $5^{\text {th }}-95^{\text {th }}$ percentile range of reserve estimates for each of the methods by class.

Possible causes of the variation in estimates produced by the methods could include differences in the:

- theoretical framework of the methods
- practical application of the methods
- subjective choice of underlying parameters/data used

Although the data being used in the trial was real insurance company data, anonymity was given to the source of the data. Therefore, the groups applying the methods did not have a detailed knowledge of the company supplying the data, or have the facility to investigate the causes of any apparent anomalies in the development. A key message from the GRIT report was the importance of understanding the data when applying any actuarial method. Some of the results of the analysis described below highlight this issue, as different treatment of anomalies in the data produced a wider range of results than that seen from applying different methods.

A further difficulty in drawing conclusions from the analysis was the relatively low number of groups involved. There were few cases where more than one group applied the same method to the same data set, making it impossible to reliability adjust for the variation in estimates arising because different people were applying the methods. However, from the few observations possible, it appears that the variation produced by different groups was very significant compared to the variation between methods. This may have arisen from some groups applying the methods automatically, with little judgemental adjustment, whilst others may have incorporated market knowledge into the analysis.

Given all the difficulties outlined above, any conclusions from this analysis about the relative results from the different methods could be spurious.

## Employers' Liability Results Comparison



## Graph 1: Employers' Liability Method Comparison

The graph above shows the results produced for the reserves at the various centue points we discussed above, for a number of methods and modelling frameworks. The majority of results shown are those produced by the working groups. The two right most results shown, (PTF/MPTF/ELRF, PTF/MPTF/ELRF - Inflation trend) are from the submission from Insureware. This is the only class for which we received a distribution from Insureware.

We have discussed the results produced for this class with the Reserving Oversight Committee. We would observe that some actuaries may be concerned that the $95^{\text {th }}$ centiles produced by most of the methods in Graph 1 appear surprisingly low for this class of business when benchmarked against their personal reserving experience. It would be interesting to hear from the profession whether this is so. We refer the reader to the discussion about the judgement method in Section 1.

Some of the methods and modelling frameworks produce fairly narrow ranges in this example - such as Mack Bootstrap Incurred, Mack Variability Incurred, ODP BF Targets, ODP BF Targets (Varying Scale), Bayesian - Group 1, Bayesian Incurred, PTF/MPTF/ELRF, PTF/MPTF/ELRF - Inflation trend.

Some of the methods produce fairly wide ranges in this example, such as ODP Group 1, ODP - BCL Targets, ODP - Large Losses Separate and the four Regression/Curve fitting methods.

There is a wide variety of means produced across the various methods, the highest being around three times the lowest.

One cause of the difference in the results appears to arise from the choice of paid versus incurred triangles. The majority of groups used paid triangles but for the Mack
and Bayesian approaches incurred data was also used. The relationship between Paid and Incurred data changed in more recent origin periods and gave rise to noticeably different results with the same method. The graph below shows the development of paid claims as a proportion of incurred claims.


Graph 2: Employers’ Liability Paid to incurred Data Comparison 1997 to 2005
It can be seen that the paid claims as a proportion of incurred claims increased for more recent origin periods, possibly indicating either a relative weakening in case reserves or a speeding up of payment patterns, or a combination of both. The assumption made about which of these explanations underlies the change in development significantly affects the results produced.

When applying the methods to incurred claims, it is implicitly assumed that the development pattern of incurred claims is stable. As the graph below shows, the older years had shown a substantial proportion of redundancy within outstanding claims case estimates being released after about development year 4.


## Graph 3: Employers' Liability Incurred Claims Development - All Years

When applying the methods to incurred claims, the level of this release was implicitly assumed to continue. The mean estimates produced by the incurred methods appear unreasonably low, as in some cases they are below paid claims to date, with little historical evidence for recoveries on paid. It is notable that the automatic application of these methods produced estimates where even the $95^{\text {th }}$ percentile was significantly below the mean produced by methods applied to paid claims. However, the systematic decrease in incurred claims invalidates the assumptions for some of the methods used.

The decision to analyse large losses separately for the ODP method also gave rise to very different results.

The choice of approach within each method was also influential. For example the choice of informed prior and scaling approach in the Bayesian method led to quite different results. Group 1 applied the deterministic basic chain ladder method to both Paid and Incurred data and then based the informed prior on the average of the ultimates by origin year. Group 2 did not use an informed prior. Neither group incorporated their own external judgement to influence the informed prior.

The Insureware team did not believe in a method as such but rather a modelling framework to identify the structure in the data. The modelling frameworks used were the Probabilistic Trend Family (PTF), Multiple Probabilistic Trend Family (MPTF) and Extended Link Ratio Family(ELRF). Please note that the graph for this approach does not show the $5^{\text {th }}$ and $25^{\text {th }}$ percentiles.

Comparing the $95^{\text {th }}$ percentile to the mean the Regression/Curve Fitting method gave the highest value at $138 \%$ compared to the Insureware approach giving the lowest value at $107 \%$. This is quite a significant variation in the estimated uncertainty.

## Personal Motor Results Comparison



## Graph 4: Personal Motor Method Comparison

The majority of methods gave consistent results for this dataset. The ratio of the $95^{\text {th }}$ percentile to the mean mostly varied between $110 \%$ and $120 \%$. The exceptions were the Regression / Curve fitting methods (varying between $186 \%$ and $243 \%$ ) and the Judgement method on Paid Data from Group 2 giving a ratio of $133 \%$. It is noticeable that these exceptions were the methods where intermediate percentiles were not calculated.

The Mack Variability Incurred projection was significantly skewed due to the high coefficient of variation and the decision to use a lognormal distribution.

## Commercial Property Comparison



## Graph 5: Commercial Property Method Comparison

The estimates seemed to be split into two main groups with the results being quite consistent within these groupings. This was partly explained by the choice of Paid or Incurred data. As for the Employers' Liability class, there was a change in the relationship between paid and incurred claims, albeit in this case paid claims reduced as a proportion of incurred claims for recent origin periods leading to estimates from incurred claims exceeding those from methods applied to paid claims.

The Regression / Curve Fitting results on paid data were noticeably different to the remaining methods.

## 8. Summary of Qualitative Comments

### 8.1 Introduction

In applying the selected methods to the real datasets, sub-groups of the working party considered the following qualitative questions:

- Would the method be acceptable to the Profession?
- Ease of use and practicality of method
- How difficult is it to apply judgement and / or amendments to the results
- How easily would you be able to explain the method to non technicians?
- Does the method include extreme events? (By this we mean can you allow for the sudden emergence of large individual losses, late tail kicks in incurred, surprising developments on known large losses, etc)
- When is the method good, when is it not good, and when does the method fail?

Comments from the sub-groups are summarised in the remainder of this section for each method in turn. In some cases, these comments were collected from more than one-subgroup and the sub-groups may have had differing views: these have been retained.

### 8.2 ODP Bootstrap Method

8.2.1 Would the method be acceptable to the Profession?

Three sub-groups looked at this method. The consensus was that this method would be acceptable (and our survey results, Section 3, show that it is in fact one of the most widely used methods). However, the new simulation results presented in Section 9 and Appendix B of the present paper suggest that a renewed debate on this would be appropriate: published versions of the ODP bootstrap method seem to be inaccurate in the extremes of the predictive distribution.

### 8.2.2 Ease of use and practicality of method.

A group that used an existing proprietary implementation commented that the bootstrap ODP method is easy to use once set up, and as the algorithms are relatively simple, the program runs quickly. This group added that the method is relatively straightforward and can be run in an Excel spreadsheet; but that it is better suited to specialist software without the limitations of Excel.

Another group who used Excel commented that it is easy to programme, taking 1-2 days to programme for a generic application including tests. However, this group found it quite slow to run in Excel (about 30mins for 200 simulations). This group commented that a significant number of simulations is required to generate distributions without "roughness" in the tail, as use of scale parameter and integer

Poisson results can give clumpy results on an individual underwriting year basis, particularly where minimal reserves are expected.
A third group programmed the method in C++ and found that 1,000 simulations took a fraction of a second, so running enough simulations to obtain a smooth predictive distribution was not a problem.

### 8.2.3 How difficult is it to apply judgement and / or amendments to the results

Judgement is required on details of the method where the published literature is silent or inconclusive. If using a proprietary implementation, some of these judgements have been made by the software providers and users should be aware of this.
One group commented that checks are needed to ensure that individual development factors are greater than one. The group that used a C++ implementation did not impose this constraint but instead imposed the weaker constraint that all cumulative paid amounts should be positive in pseudo-data (as described in Section B2.4.1 of the present paper).

The group that used the proprietary application commented that the ODP Bootstrap results can be scaled to achieve any desired mean or coefficient of variation. Judgement is required on whether or not to do this, and in the choice of target mean and/or variance. The mean could be chosen to match results attained by a deterministic statistical method, such as a deterministic chain ladder. The variance could be chosen to match that of another, perhaps more complicated, stochastic method. Otherwise the variance could be scaled (either additively or multiplicatively) in proportion to the scaling of the mean.

However, it is unclear what the bootstrap ODP method achieves if the bootstrap predictive distribution is shifted and scaled in this way. This presupposes that the mean and predictive standard error have been obtained by some other method, in which case an analytic distribution (instead of a bootstrap distribution) could be fitted to give the required mean and standard error.

The group that programmed the method in Excel commented that the method has limited usability on a pure basis as it needs all individual development factors (idfs) to be greater than one, adding that judgement can be applied to select a set of idfs, but this will increase the residuals and hence uncertainty. This group also commented that other amendments to the base model are possible, including changing the assumptions on the predictive distribution.
8.2.4 How easily would you be able to explain the method to non technicians?

One group commented that the bootstrapping part of the method (in respect of parameter uncertainty) would be easier to explain to non-technicians than the forecasting part: it would be possible to explain the principles of bootstrapping in layman's terms. Another commented that principles would be easy to explain, but technical details more difficult. This concurs with experience of other members of the working group who have had to explain the method in practice.

### 8.2.5 Does the method include extreme events?

One group commented that the ODP bootstrap can produce some very extreme outliers that may need to be capped in order for the distribution to be used. Judgement is needed on whether/how to include extreme outcomes in the bootstrap results. Another group commented that extreme events are only included as much as they exist in existing data.

### 8.2.6 When is the method good, not good, and when does it fail?

One group commented that the ODP bootstrap is good if there are few negative developments, and it is reasonable to assume that the residuals are identically distributed (subject to scaling) and independent.

The group that used the proprietary implementation commented that the method will not work if the sum of a column in the incremental triangle is negative. Another group commented that for similar reasons, the method may not work well on incurred data. It was generally recognised that the method requires the underlying run-off pattern to be the same for all origin years, and will not work well if this is not the case. This is confirmed by the simulation results presented in Appendix B (Section B.4.5) of the present paper.

### 8.3 Mack's Method

Two sub-groups tested and commented on this method.

### 8.3.1 Would the method be acceptable to the Profession?

As for the bootstrap ODP method, the general view was that the method is currently quite widely accepted by the profession. However, the new simulation results presented in Section 9 and Appendix B of the present paper suggest that a renewed debate on this would be appropriate because the method can be very inaccurate in the extremes of the predictive distribution.

### 8.3.2 Comment on difficulty of method to program and run

One group used an existing implementation and commented that software was readily available for the method and was quick to run. Another group found it straightforward to programme the method in both Excel and C++.

### 8.3.3 How difficult is it to apply judgement?

The group that used an existing implementation said that adjustments could not be made. The group that programmed the method in accordance with Thomas Mack’s original papers found that nearly all details are specified by Mack, but there is some scope for judgement where there are too few data-points to estimate a variance parameter for each development period, and where tail factors are required. Mack also does not fully prescribe how a complete predictive distribution should be obtained from the mean and standard error that are produced by the method.
8.3.4 How easily would you be able to explain the method to non technicians?

One group commented that it would probably be difficult to explain this to non technicians but no more so than most methods. Another group commented that it is a relatively difficult method to explain to non-technicians.

### 8.3.5 Does the method include extreme events?

It was generally recognised that extreme events are only included to the extent that they occur in the past data.

### 8.3.6 When does the method fail or not work well?

It was generally recognised that the method requires the underlying run-off pattern to be the same for all origin years, and will not work well if this is not the case. The method will not fail if the run-off pattern varies, but will give poor results. This is confirmed by the simulation results presented in Appendix B (Section B.4.5) of the present paper.

### 8.4 Bayesian Bornheutter-Ferguson (BBF)

Two subgroups tested this method - both used proprietary specialist software (Igloo).

### 8.4.1 Would the method be acceptable to the Profession?

The method is a formalisation of the widely used Bornheutter-Ferguson method, so in principle should be acceptable to the profession, provided it is correctly implemented and the judgements required understood.

### 8.4.2 Comment on difficulty of method to program and run

The BBF method has some advanced theory behind it and requires some advanced sampling algorithms. The method requires sampling several parameters from a nonstandard multivariate distribution. This has been done using Adaptive Rejective Metropolis Sampling (ARMS). The Gibbs method is used to reduce the multidimensional sampling problem to an iterative one dimensional problem. This would be very difficult to perform without using specialist software.

Both groups commented that the specialist software used is currently slow to run. One group said that running 50,000 simulations on a 10 by 10 triangle takes over one hour using a PC with processor speed 3 GHz and 2GB of RAM. The other said that run time can be decreased by reducing the sampling rate, but that when unadjusted, 10 scenarios for one triangle takes 20 mins to run.

### 8.4.3 How difficult is it to apply judgement?

The model is flexible as it can accept prior information in the form of a distribution. The prior distributions used in the BBF method influence the results. The output distribution can be adjusted to achieve a desired mean and variance, however this would undermine the decision to use the BBF method in the first place.

The method requires a prior distribution for the ultimates separately for each origin period. Judgement is required in choosing these distributions. The levels of certainty of the prior estimates are reflected in the choice of the coefficients of variation for the
prior distributions. If a high CoV is chosen for the prior distributions then the BBF predictive distribution resembles that produced by the ODP Bootstrap method. If the CoV of the prior distributions are small then the ultimates will have a smaller CoV . The mean of the distribution tends towards that of a deterministic BF model as the CoV of the priors tend towards zero.

### 8.4.4 How easily would you be able to explain the method to non technicians?

One group commented that the BBF method is difficult for a technician to understand so it would be very difficult to explain the method to a non-technician. Even to explain the principles of the BBF method in layman's terms would be very challenging. The other group said it would probably be difficult to explain this to non technicians but no more so than most methods.

### 8.3.5 Does the method include extreme events?

It was generally recognised that extreme events are only included to the extent that they occur in the past data.
8.4.6 When is the method good, when is it not good and when does it fail?

The BBF is useful to use when there is a prior assumption of the ultimate values. It is not so good when time or processor speed is an issue. One group said the method will not work if the sum of a column in the incremental triangle is negative. The other said the method does not seem to work if there is no movement in the upper right of the triangle.

### 8.5 Judgement

Two sub-groups considered this approach.

### 8.5.1 Would the method be acceptable to the Profession?

We refer the reader to this discussion of this method in Section 1. The following discussion is from the working groups.

The first sub-group said: Yes. "Actuarial judgement" is probably the most widely used concept when setting reserves. It does rely solely on the judgement of the person setting the reserves, but it allows that person full control and flexibility to adapt to any trends they may see in the data without much effort. This method could be criticised as the accuracy of the best estimates is dependent on the experience of the actuary applying the judgement. There is always a risk that two reasonable actuaries would generate two very different results through taking a different view around emerging trends or allowances for distortions.

The other sub-group said: I don't think so. It is very subjective and difficult to do. However this second group also recognized that this method does have the merit of potentially allowing for model error as well as parameter and process error.
8.5.2 Comment on difficulty of method to program and run

The first group considered the use of the basic chain ladder method as a starting point for the application of judgement, and commented that to generate a range of results
using judgement it can be somewhat laborious tweaking development factor assumptions to allow for distortions within the data. They added that automatically generated development factors can be easily obtained however as a starting point by sampling across varying numbers of accident periods. Generally, initial selection can be made quickly (several minutes). Generating a sensible range of results through adjustment of factors naturally adds to time taken.

They also commented that input from several modellers provided greater variety of answers. In this way, the subjectivity of the approach could be viewed as an advantage when assessing reserve uncertainty. However, they also commented that it can be difficult to avoid being influenced by an expectation of what reserves should be. For example, the actuary could be 'anchored' to previous valuation results. They commented that the method potentially allows for the many complexities of different classes of business, but taking all factors into account by judgement is demanding and time consuming to do well. Additional information (e.g. large claims data) or discussions with claims handlers can provide valuable insight leading to a better application of judgement.

The other group said it is generally very difficult to translate judgement into percentiles.

### 8.5.3 How difficult is it to apply judgement?

See response to 8.5.2 above.
8.5.4 How easily would you be able to explain the method non technicians?

It is easy to explain how certain factors have been taken into account when presenting results to non technicians. It is also relatively easy to explain the basic chain ladder approach to non technicians to expand on where judgement has been applied.

### 8.5.5 Does the method include extreme events?

Extreme events can be suitably allowed for through additional margins within the provisions by making reasonable assumptions about likelihood and severity. However, there may be much subjectivity over what is 'reasonable' here.
8.5.6 When is the method good, when is it not good and when does it fail?

It is good to apply judgement when the actuary believes he or she has relevant knowledge or experience about the expected future development of the claims, and is adept enough to make suitable allowances. The application of judgement is less necessary for standard short-tail classes where past performance remains a good guide to the future.

Lack of experience or inaccurate supporting data/information could lead to unsuitable application of judgement. Peer review can help to mitigate these risks.

### 8.6 Scenarios

There are several ways to employ scenario testing. This method can be applied to many underlying reserving methods including chain ladder, Bornheutter Ferguson and exposure methods.

One sub-group considered the use of scenarios generated from the basic chain ladder method and commented as follows:

### 8.6.1 Would the method be acceptable to the Profession?

Yes. Application of judgement can play a major role in selecting scenarios to obtain a range of results. A more mechanical process of simply selecting the most optimistic and pessimistic development factors is unlikely to be suitable, especially where large distortions exist within the data.
8.6.2 Comment on difficulty of method to program and run

The basic chain ladder method is very simple and easy to use. A simple mechanical process of selecting the most optimistic and pessimistic development factors is easy to apply. Other considerations should be made however to avoid generating meaningless results. Inherent uncertainty / volatility in smaller datasets e.g. from large claims, ideally need to be considered and allowed for in the scenarios selected.

### 8.6.3 How difficult is it to apply judgement?

It is easy to apply amendments to results via manual adjustments and other tweaks when using a simple, well designed model.
8.6.4 How easily would you be able to explain the method to non technicians?

It is relatively easy to explain the basic chain ladder approach to non technicians and describe the choice of scenarios generating the range of results.

### 8.6.5 Does the method include extreme events?

Extreme development factors can only be selected to the extent they are generated from the underlying dataset. A mechanical selection of factors, with no consideration to the distorting effect of extreme events, would not be sensible.
(The working party would add that the use of scenarios can be a helpful approach to considering extreme events, and can be usefully employed, for example, in exposure reserving for asbestos claims).
8.6.6 When is the method good, when is it not good and when does it fail?

A purely mechanical application of selecting best/worst scenarios would not be advised, especially when used for volatile datasets. Conversely, this type of method works well with large stable datasets. Lack of experience could lead to inappropriate selection of scenarios.

### 8.7 Craighead Curve followed by Regression

This was considered by one sub-group who commented as follows:
8.7.1 Would the method be acceptable to the Profession?

It depends on the purpose. It gives a rough indication of uncertainties of reserves based on past development observations. The method would need heavy adaptations and more research to be used for stochastic simulations.
8.7.2 Comment on difficulty of method to program and run

The method is easy to implement using Excel. The regression method does not give estimation for the first two origin periods (because one needs more than two points to perform regression)
Curve fitting could be an issue if there is much negative developments: the Craighead curve is increasing, and other curves could be vulnerable to over-fitting.

### 8.7.3 How difficult is it to apply judgement

Judgments are required at several points in the methods (e.g. the weight to give in the curve fitting for each data point, the curve to use, the weight to give in the regression for each data point, etc.) But (at least for the curve fitting), it seems that there isn't that much more judgment required than in, say, the chain ladder. The paper (by Benjamin and Eagles) proposed several modifications to the method - very easy to "break into the method" and make modifications
8.7.4 How easily would you be able to explain the method to non technicians?

Should be easily explained to colleagues in other professions.
8.7.5 Does the method include extreme events?

The mechanical application of the model does include them - but one may want to exclude (or put less weight) on them for curve fitting or regression.
8.7.6 When is the method good, when is it not good and when does it fail?

No comment provided.

## 8.8 (Multiple) Probabilistic Trend Family (M)PTF

The comments in this subsection have been taken from the submission to the working party provided by Insureware Pty Ltd. Insureware produces software (ICRFS-plus) for carrying out these methods. Their complete submission is given in Appendix F.
8.8.1 Would the method be acceptable to the Profession?

Yes.
8.8.2 Comment on difficulty of method to program and run

It is extremely difficult to program (but theoretically possible) in a spreadsheet form, but is easy to run in ICRFS-Plus as all the programming has been pre-specified. ICRFS-Plus is a point and click system with extremely fast algorithms. Individual tests do not need to be programmed. It takes seconds for individual tests to be completed.
8.8.3 How difficult is it to apply judgement within the program

It is not difficult. Judgement is made on the basis of accurate information about the volatility in the business.
8.8.4 How easily would you be able to explain the method to non technicians?

No comment provided.
8.8.5 Does the method include extreme events?

No comment provided.
8.8.6 When is the method good, when is it not good and when does it fail?

No comment provided.

## 9. Assessing Performance of Stochastic Reserving Methods by Numerical Simulation

### 9.1 Introduction

The work described in previous sections has given us an initial impression of the main features of several stochastic methods. Although we have given an initial assessment of these methods (as described in previous sections) we cannot formulate definitive conclusions on the performance of these methods from the experience of applying them to a handful of data-sets.

Towards this end, we have also used numerical simulation methods to assess the performance of various stochastic methods. The basic idea is to apply the methods to a large number of simulated run-off triangles for which the 'true' ultimate is known (by simulating it), and to compare the predictions produced by the various methods to the true ultimate.

While this approach can be criticised on the grounds that simulated data may not exhibit all the complications and variety or real-world data, it does have merits:

- It enables methods to be tested on very large numbers of triangles.
- Predictions can be compared to 'true' ultimates without waiting years for the true ultimate position to be reached.
- Triangles can be constructed such that they exactly satisfy the assumptions underlying a method: the performance of a method in this ideal situation indicates the limits of the method's potential performance on real data. (If a method does not perform well on ideal data, it is unlikely to perform well on real data.)
- The robustness of a method to violations of its underlying assumptions (which may be met in practice) can be tested in a controlled way.

We have made a start with this but have not yet done enough to formulate definitive conclusions. We propose that the working party should continue pursuing this approach next year.

In practice, the assumptions of a stochastic method are never perfectly satisfied, and (even if the stochastic assumptions approximate the past run-off data reasonably closely) there is the added risk of shocks and changes in the future unlike anything observed in the past.

### 9.2 What constitutes "good performance" for a stochastic reserving method?

### 9.2.1 Three main types of uncertainty

In the following discussion, it is worth bearing in mind the three main categories of uncertainty when forecasting outstanding claims liabilities:

- model uncertainty
- parameter uncertainty
- process uncertainty


### 9.2.2 Why different stochastic methods give different results

Some stochastic methods tend to give wider confidence intervals for the eventual outcome than others. Some possible reasons are listed below (examples follow this list):
a) Inappropriate stochastic model: if a method is used where the underlying stochastic assumptions are not appropriate for the data, then the results - both best estimates and uncertainty assessments - are unlikely to be good.
b) Failure of some methods to take account of both parameter and process uncertainty.
c) In assessing parameter uncertainty, failure of some methods to take account of uncertainty in all parameters.
d) Failure to make use of all available data.
e) Failure to take account of correlations between different components of the overall forecast.
f) Use of inadequate mathematical approximations in the formulation of a stochastic method (so that even if the method is applied where its stated assumptions are perfectly satisfied, it may not give reliable results).
An example of (a) is misuse of any method based on an assumption that the underlying run-off pattern across development time is the same for all origin years. Where this assumption is a reasonably good approximation to reality, such methods may give realistic results. But if applied in situations where this assumption is unreasonably wide of the mark, such methods may tend to understate uncertainty, that is, give predictive ranges that are unrealistically narrow.

While there will always be some model uncertainty, we should attempt to minimise this by being aware of all model assumptions and of the claim settlement processes involved in the lines of business concerned, so we are in a position to judge whether the model assumptions are reasonably realistic in each application of any stochastic method. The run-off data itself can often be used to check for gross violations of the model assumptions.

An example of (b) is where bootstrapping is used to determine a 'range of best estimates' but no attempt is made to include future process uncertainty.

An example of (c) occurs where run-off data is pre-adjusted for inflation as if the rate of claims inflation were known precisely.

An example of (d): given a triangle of aggregate amounts paid and a corresponding triangle of counts of claims closed, some methods may not make use of the counts triangle. It is tempting to think that using more relevant information necessarily reduces uncertainty, leading to narrower predictive ranges. However this is not necessarily the case: the additional information might indicate that there is more apparently random variation than previously thought. For example, suppose we are using a stochastic chain ladder method and all available diagnostics indicate that the model is appropriate. That is, past data shows no statistically significant departures from a constant run-off pattern and all residual variation apparently satisfies the stochastic assumptions. The quantum of parameter and process variation is assessed from this residual variation (e.g. using Bootstrap, or analytic methods). Then, two diagonals later, there is no material change in the run-off pattern, but residual variation is much greater. This new information indicates that process variation is potentially much greater than it previously appeared: as a result future process uncertainty may increase, despite there being a reduced period of future development.

An example of (e): a method may produce forecasts and standard errors for each origin year separately and in combining these, fail to take proper account of the possibility that over-estimation of a single parameter of the model may cause forecasts for all origin years to be simultaneously overstated.

An example of (f) occurs where a stochastic method relies on asymptotic unbiasedness and efficiency of maximum likelihood estimates, but the number of estimated parameters is so high compared to the number of data-points that these asymptotic approximations are poor.

### 9.2.3 Choice of stochastic method

It is hopefully clear from the above discussion that, in choosing between stochastic methods, the width of predictive ranges produced by the methods should not be the primary consideration.

Users of our results from outside the profession may be forgiven for judging methods in this way, because they may make the implicit assumption that all methods used by actuaries will be 'correct' in the sense of not omitting any major source of uncertainty (such as those listed above). However, this will only be true if we within the profession ensure this is the case by critically examining all stochastic methods we use. For us, the primary considerations should be that stochastic methods:

- attempt to assess all sources of uncertainty,
- are technically accurate in this assessment given the assumptions of the method,
- are applied only where the assumption of the method have a reasonable chance of being approximately true.

If these points are all satisfied, then we have to accept the resulting ranges regardless of their width. (Of course, this may be an iterative process: if ranges come out surprisingly narrow or wide, the above points should be critically examined and the method adjusted and refined if necessary. But when we are eventually satisfied that a
valid method has been validly applied, the results should not be rejected merely on the basis that the width of predictive ranges was initially surprising, or may be surprising or inconvenient to others.)

With reference to model assumptions being satisfied, a related and more difficult question concerns the robustness of a method to violations of its assumptions. That is, how reliant is a method on the correctness of its assumptions? This is an important question as it is rare for model assumptions to be perfectly satisfied in reality.

For example, one key assumption underlying all stochastic chain-ladder methods is (as in the basic chain-ladder) that the underlying run-off pattern is the same for all origin years. If a stochastic chain ladder method is used where this assumption is false, are the results likely to be materially misleading? Perhaps not, because if a method involves assessing the quantum of process variation from residual variation in the data, then systematic departures from this key assumption will be partly taken into account through an increase in the residual variation.

### 9.3 Outline of numerical simulation approach to testing stochastic methods

To test stochastic methods, we can use the well known fact that if X is a random variable, and $\mathrm{F}(\mathrm{x})$ is its cumulative distribution function, then the random variable $F(X)$ has a uniform distribution on the unit interval $[0,1]$.

In the context of stochastic reserving, X represents the total of future claim payments (i.e. ultimate less amount paid-to-date). A stochastic method produces a function F (x) that purports to be the distribution function of X. (Many stochastic methods produce only a best estimate reserve and a root-mean-square predictive error, but these can be extended by using some class of analytic distributions, e.g. Log-Normal, to produce a complete distribution $\mathrm{F}(\mathrm{x})$.)

If a stochastic method is reasonably good, $\mathrm{F}(\mathrm{X})$ should therefore have approximately a uniform distribution. To test this, we need a number of independent instances from this supposedly uniform distribution to see if they are indeed uniformly distributed. The larger the number of instances, the more powerful will be the test.

Having carried out the stochastic method on a particular triangle, we can obtain one instance of the random variable $\mathrm{F}(\mathrm{X})$ by waiting for the triangle to reach its ultimate position: this gives us one instance ( $\mathrm{x}_{0}$ say) of the random variable X , hence one instance $\mathrm{F}\left(\mathrm{x}_{0}\right)$ purportedly from the uniform $[0,1]$ distribution.

By applying the same stochastic method many times to independent triangles, and then waiting for each to develop to ultimate, we could gradually accumulate an increasingly large sample that can be tested for uniformity.

Clearly, if we use only real-world data, it is likely to take a very long time to obtain a large enough sample for a reasonably powerful test of uniformity. The entire process can be accelerated by using a large number of artificial triangles. This approach also has the advantage of allowing us to test the robustness of a method to violations of its underlying assumptions.

So, for each stochastic method we can do two things:
a) Check the performance of the method in situations where its underlying assumptions are perfectly satisfied.
b) Test the robustness of the method to violations of its assumptions.

The main steps of (a) are:
(i) Generate a large number I of artificial datasets (i = 1...I) that follow the assumptions of the method: each dataset to contain run-off arrays (on which to apply the stochastic method) and 'true' reserves (denoted $\mathrm{r}_{\mathrm{i}}$ ).
(ii) For each artificial dataset, apply the stochastic method to obtain the predictive distribution function $\mathrm{F}_{\mathrm{i}}(\mathrm{x})$ (where x is any possible value of aggregate future payments).
(iii) For each artificial dataset, calculate $\mathrm{u}_{\mathrm{i}}=\mathrm{F}_{\mathrm{i}}\left(\mathrm{r}_{\mathrm{i}}\right)$ (where $\mathrm{r}_{\mathrm{i}}$ is the simulated 'true' total of future payments).
(iv) Test the dataset $\left\{\mathrm{u}_{\mathrm{i}}: \mathrm{i}=1 \ldots \mathrm{I}\right\}$ for uniformity. If this is not significantly different from uniform, then the method is 'stochastically correct' (in other words, it appears to correctly assess parameter and process error, given the underlying model assumptions).

For (b) (the assessment of robustness) the steps are as for (a) except that at step (i) the artificial datasets would be generated to have some feature that departs from the assumptions underlying the stochastic method. For example, if the chosen stochastic method assumes that settlement delay is unrelated to the size of a claim, we could generate datasets in which larger claims tend to take longer to settle than smaller claims. If the distribution obtained at step (iv) still does not depart significantly from a uniform distribution, then we can conclude that the method is robust to the particular aspect modelled. In other words, no material additional uncertainty arises from this particular aspect of model uncertainty.

### 9.4 Overview of simulation analysis carried out so far and provisional conclusions

We have done some initial work applying the approach described in the previous section to:

- Thomas Mack's method as described in his 1993 paper.
- The over-dispersed Poisson (ODP) method as described by Arthur Renshaw and Richard Verrall (1998), Peter England and Richard Verrall (1999) and Peter England (2001).
- Tom Wright’s operational time average-cost-per-claim method as described in his 1992 paper.

Appendix B gives a detailed description of the simulations carried out and the results obtained. Here we give just a brief overview and summarise the main findings.

For each of these stochastic methods, we simulated at least 30,000 triangles using algorithms that perfectly satisfy the stochastic assumptions underlying each method. We then applied the stochastic method to each simulated triangle to obtain a predictive probability distribution function $\mathrm{F}(\mathrm{x})$ (where x the total of future payments). We then calculated $\mathrm{u}=\mathrm{F}(\mathrm{r})$ where r is the 'true' outcome (simulated along with each triangle). As discussed, in the previous section, this quantity should be uniformly distributed in the range 0 to 1 . However, for all the stochastic methods tested, too many value of $\mathrm{F}(\mathrm{r})$ were very close to 1 . The table below shows the proportion of simulations in which $\mathrm{F}(\mathrm{r})$ exceeded 0.99 . In other words, this is the proportion of simulations in which the true outcome exceeded the $99^{\text {th }}$ percentile of the predictive distribution obtained from the triangle. Clearly, if the predictive distribution is accurate, this should occur in $1 \%$ of simulations.

Table 9.4.1 Proportion of simulations in which 'true' outcome exceeded $99^{\text {th }}$ percentile

| Mack 1993 (with Log-Normal) | $8 \%$ to $13 \%$ |
| :--- | :--- |
| Analytic ODP (Renshaw \& Verrall, 1998), Pearson dispersion | $2.6 \%$ |
| Analytic ODP (Renshaw \& Verrall, 1998), deviance dispersion | $2.7 \%$ |
| Bootstrap ODP (England \& Verrall, 1999) | $3.1 \%$ |
| Bootstrap ODP (England 2001) | $2.6 \%$ |
| Operational time (Wright 1992), Pearson dispersion | $4.0 \%$ |

Bootstrap ODP (England 2001) is the only one of these methods that produces a full predictive distribution $\mathrm{F}(\mathrm{x})$ : the others produce just a best estimate and a standard error (or "root-mean-square predictive error"). For these methods we used a LogNormal predictive distribution F(x) (as suggested by Mack 1993) in order to calculate $\mathrm{u}=\mathrm{F}(\mathrm{r})$. (We also tried using the Inverse Gauss distribution for $\mathrm{F}(\mathrm{x})$ but found this made no material difference to the main results summarised above.)

It is emphasised that for each of the methods, the above results are based on at least 30,000 triangles generated in a way that perfectly satisfies the stochastic assumptions of the respective method.

For most methods we have so far used only a single set of parameters to generate triangles that perfectly satisfy the assumptions. Other parameters would no doubt lead to different results. Only in the case of Mack's method have we tried more than one set of parameters for generating the triangles, which is why the table above shows a range of results (from 8\% to 13\%). For all stochastic methods the simulated triangles had 10 origin years and annual development. We would expect better performance on larger triangles but have not yet tested this.

We should generally expect stochastic methods to perform worse than this in practice because their assumptions will never be perfectly satisfied. We have done some testing of robustness to violations of assumptions only for Mack's method and the ODP methods so far. The results for all variants of the ODP method indicate, as expected, a deterioration in performance if applied where the ODP assumptions are violated. Surprisingly, results obtained so far for Mack's method show it performs
better where its assumptions are not satisfied than where they are perfectly satisfied (details and a possible explanation are given in Appendix B).

### 9.5 Conclusions to Date

The main conclusion of the simulation work carried out so far is that all these stochastic reserving methods tend to understate the chance of extreme adverse outcomes, even in situations where their underlying assumptions are perfectly satisfied.

In Appendix B we outline some possible explanations of why these stochastic methods do not work well at the extremes. To be fair to the developers of these methods, the goal-posts have moved. The aim at the time these methods were developed was to improve on the usual practice of providing just 'best estimate’ reserves, or perhaps estimates based on just a few scenarios (to give perhaps low, medium and high reserve estimates). It is only in more recent years that regulators have begun specifying particular extreme percentiles for capital requirements. Our results indicate that further work is now needed to develop and refine stochastic reserving methods to better meet this new challenge.

## 10. Objectives for the Working Party for Next Year

There are three main areas of objectives for next year:

### 10.1 Objective 1: Statistical Testing When Conditions Met

We would like to continue the consideration of the performance of statistical methods when all their underlying conditions are met. We would like to better understand to what extent and in what circumstances these methods can be reliable indicators of ultimate claims outcomes at the tails of the distribution. Do we need to alter the way we use such methods and are there any "quick wins"? We see this as a key priority in the context of reserving risk assessment for capital purposes.

Part of the work in this area is likely to include liaison between this working party and the one concentrating on best estimate reserving and the performance of methods in that context.

## 10.2: Objective 2: Test More Methods

We would like to expand the review and testing of methods on "real" data to include additional methods, if possible those which operate on transactional data. This would ideally include further quantitative and qualitative review.

## 10.3: Objective 3: Robustness of Methods in Real Life

We would like to expand the work on simulated data to test the response of methods to circumstances when there underlying conditions are not met - which is the case in most real life scenarios. How do methods respond when there are new trends, changes in claims processing speeds or underwriting cycles, for example. How should the actuary make use of information beyond the triangle? How should an actuary test whether methods are appropriate or not - and how might these tests vary depending on the purpose and the part of the distribution of outcomes which is of particular interest?

Appendix A: Summary of the Methods Tested

## Appendix A - Summary of each of the methods tested

The following methods are described briefly below:

1. Judgement
2. Scenario testing
3. Mack
4. Over-dispersed Poisson stochastic chain ladder
5. Transaction level modelling
6. Operational time
7. Regression/ Curve fitting
8. Bayesian/BF method
9. Probabilistic Trend Family (PTF) - ICRFS

Note that Prof. Richard Verrall has a web-lecture on a number of these methods at: https://talk.city.ac.uk/stochasticreserving. In addition the paper by England and Verrall provides a detailed review of the differences between many methods and approaches (England and Verral, BAJ (2002)).

Additional references to these methods are contained in the bibliography, the references provided here represent the actual methods applied in each case.

## 1. Judgement

Description: This method is described in Section 1 of the report.
Key assumptions: This will vary depending on exactly how the actuary applies judgement. In some cases the only assumptions will be the selected results themselves.

Data required: Any - but relevant experience is also needed.

Reference: n/a

## 2. Scenario testing

Description: although not the only methodology that can be defined under this heading, this is generally perceived as selecting alternative parameters from the best estimate projection that reflect the user's view of the extreme ends of the range to be defined.

Key assumptions: the assumptions under this method are similar to those under the judgement method. However, they are generally more specific and relate to changes in one or more parameters to investigate the effect such changes have on the projected reserves. Note that any assumptions that relate to the underlying model(s) used to calculate the best estimate will still be applicable in this case.

Data required: Any.
Reference: n/a

## 3. Mack

Description: This method calculates the standard error involved in the application of the Chain-ladder method for the reserve estimation for each origin year and for all years combined, allowing for both parameter and process uncertainty. The 1999 paper allows for tail factors within the method.

The method can be applied both as a deterministic calculation, or using a bootstrap approach. It is suggested that standard distributions (LogNormal; Gamma) can be used to derive ranges based on the results of the method.

## Key assumptions:

- The underlying run-off pattern is the same for all origin years (as in the basic chain ladder method)
- Future development factors are independent of past development factors (i.e. the size of the next factor is independent of the size of the last factor).
- Variance of the next cumulative claims amount is proportional to the current cumulative claims amount.

Data required: Cumulative claims triangles (paid or incurred).
Reference: ASTIN Bulletin vol. 23 (1993) (and update in 1999 - ASTIN vol.29)

## 4. Over-dispersed Poisson method

Description: This method calculates the distribution of outcomes for each origin year and for all years combined. It models incremental claims using an Over-dispersed Poisson distribution. The calculation of a distribution of outcomes is performed using a Bootstrapping calculation.

Key assumptions:

- The underlying run-off pattern is the same for all origin years (as in the basic chain ladder method)
- Incremental claims amounts are stochastically independent
- The variance of incremental claim amounts is proportional to the mean (process error).
- Incremental claims are positive for all development periods (though there are adjustments to the method that can allow for data where this is not the case).

Data required: Cumulative claims triangles (paid or incurred).
Reference: "Stochastic Claims Reserving", England and Verrall, BAJ (2002) and others

## 5. Transaction level modelling

Description: This method uses Bootstrapping techniques to model individual claims events, based on investigation of individual claim data. GLM methods are used to fit the claims development factors to each claim's data, based on all available
information e.g. credit rating or exposure measure contained within the policy or claims data to derive the best estimate. Similar techniques are applied to derive estimators for reserve ranges.

Key assumptions:

- Individual claim development is a function of characteristics relating to that policy and claim.
- IBNR claims originate in a similar response to the predicted claims development factors.
- A known mix of "good" and "bad" policies is contained within each origin year to allow for changes in business, i.e. the reserving cycle.

Data required: Policy and developmental claim data for each claim.
Reference: "Loss reserving using claim level data", Guszcza and Lommele, CAS Fall Forum (2006)

## 6. Operational time

Description: An average-cost-per-claim method in which the average-cost-per-claim closed is assumed to depend on the point in operational time at which a claim is closed. 'Operational time' is defined as the number of claims closed expressed as a proportion of the ultimate number of claims in an origin year (so it increases from zero at the beginning of each origin year to one when an origin year reaches ultimate). The main advantage of using operational time is that the method does not require the underlying run-off pattern (across real development time) to be the same for all origin years. Generalized linear modelling is used to find the best fitting formula relating mean claim amount to operational time. Best estimate reserves are determine by evaluating this fitted formula for all future operational times. Predictive standard errors are calculated that allow for both process and parameter uncertainty.

Key assumptions:

- The probability distribution for the amount of an individual claim depends on the point in operational time (between 0 and 1 ) when the claim is settled.
- The variance of the probability distribution for individual claim amounts is proportional to the mean raised to some power (determined from residual analysis).
- Aggregate incremental claim amounts are stochastically independent (given the number of claims closed).


## Data required:

The method requires the usual run-off triangle of aggregate paid amounts, and also a triangle that counts the number of claims closed. Optionally, the method can also make use of counts of reported claim numbers (used to improve estimates of ultimate claim numbers).

Reference: "Stochastic reserving when past claim numbers are known", Wright T. S, Proceedings of Casualty Actuarial Society 1992

## 7. Regression/ Curve fitting

Description: Estimates ultimate claims by fitting a Craighead (or similar) curve to paid and incurred development data (either as ratios of ultimate premium or absolute values) using a weighted least squares fitting algorithm. The range of reserve estimates is derived by plotting IBNR vs. incurred claims for each development [period within a given origin year.

This plot is examined for potential uncertainty for fitting a regression curve to the data for each origin year separately. This method does not therefore explicitly generate a particular statistical property of the reserve distribution; the extent of the range is defined by the fitting process and the implicit assumptions used therein.

This method does not, therefore, produce reserve ranges for all origin years combined.
Key assumptions:
The fitting curve is a good description of the claim development
Data required: Ultimate premium, paid and incurred cumulative claims.
Reference: Benjamin, S. \& Eagles, L. (1997). A curve fitting method and a regression method. Claims Reserving Manual volume 2. London: Institute of Actuaries.

## 8. Bayesian/BF method

Description: Bayesian method in which prior probability distributions are specified for the ultimates of each origin year. Using very vague priors yields the same results as the over-dispersed-Poisson chain-ladder method: using exact priors produces the same results as the Bornheutter-Ferguson method. So the method allows for a complete spectrum of results between these two extremes. The appropriate point on this spectrum is determined by the prior distributions, which must be set by the user on the basis of judgement and/or information from other sources (eg industry data).

Key assumptions:
Same as for over-dispersed Poisson model:

- The underlying run-off pattern is the same for all origin years (as in the basic chain ladder method)
- Incremental claims amounts are stochastically independent
- The variance of incremental claim amounts is proportional to the mean
- Incremental claims are positive for all development periods (though negatives can be removed or adjusted if not too many).
And in addition:
- Assumptions for mean and variance of prior distributions (by judgement, or from industry data, or both).


## Data required:

Usual run-off triangle of aggregate amounts (paid or incurred). However, as nonpositive increments must be removed or adjusted, the method may break-down if there are too many negative increments, so it is often unsuitable for use with incurred data.

In addition, information (or prior experience with other datasets) is needed to inform the judgemental selection of prior means and variances.

Reference: Verrall R. J (2001): A Bayesian generalized linear model for the Bornhuetter-Ferguson method of claims reserving. Actuarial Research Paper No. 139, Department of Actuarial Science and Statistics, City University.

## 9. Probabilistic Trend Family (PTF) (as implemented in ICRFS-plus)

Description: Normal theory linear regression models for the logarithm of aggregate increments. Regression parameters represent origin year effects, and differences ('trends') between successive development years and successive calendar years.

Key assumptions:

- Same development pattern for all origin years after allowing for calendar year effects.
- Aggregate incremental paid data are approximately log-normally distributed. (Note this can never be exact because the sum of log-normals is not lognormal, but it might sometimes be a reasonable approximation: residual analysis can be used to check.)
- Incremental paid amounts are stochastically independent.

Data required:
Aggregate development data with few negative increments. (Because the dependent variable in the regression is the log of the incremental data, negative or zero increments have to be removed, so the method is often unsuitable for use with incurred data.)

If available, the method can also make use of some measure of exposure (eg premium, or number of claims reported in first development period).

Reference: Best Estimates for Reserves, Glen Barnett and Ben Zehnwirth, PCAS 2000 Volume LXXXVII Part 2

## Appendix B: Numerical Simulation - Detailed Results

## B Assessing performance of stochastic reserving methods by numerical simulation

## B.1 Assessment of Mack's 1993 method using simulation

## B.1.1 Mack's assumptions

As discussed in Section 9, we aim to use simulation to:
a) Determine whether Mack's method gives correct results when its assumptions are perfectly satisfied, and
b) Determine the robustness of the method to its assumptions not being perfectly satisfied. In other words, to answer the question: does the method continue to give reasonable uncertainty assessments when the claim payment process deviates (in ways that may be met in reality) from the assumptions underlying the method?

In Mack's 1993 paper he considers the basic chain ladder method (BCL) and sets out to find stochastic assumptions under which the BCL should give good point estimates, and then to find formulas (derived from these stochastic assumptions) for the root-mean-square predictive error of the BCL. (We will use the term 'standard error' to mean the root-mean-square predictive error.) The assumptions identified by Mack are as follows (where $\mathrm{C}_{\mathrm{jk}}$ denotes the cumulative amount paid in origin year j by the end of development period k ):

1. There exist parameters $f_{k}$ such that: $E\left(C_{j, k+1} \mid C_{j 1}, \ldots C_{j k}\right)=f_{k} \cdot C_{j k}$
2. There exist parameters $\alpha_{k}$ such that: $\operatorname{Var}\left(\mathrm{C}_{\mathrm{j}, \mathrm{k}+1} \mid \mathrm{C}_{\mathrm{j} 1}, \ldots \mathrm{C}_{\mathrm{jk}}\right)=\alpha_{\mathrm{k}}{ }^{2} \cdot \mathrm{C}_{\mathrm{jk}}$
3. Accident years are stochastically independent.

## B.1.2 Artificial data satisfying Mack's assumptions (Algorithm A)

To test whether Mack's formulas give a correct assessment of uncertainty when these assumptions are perfectly satisfied, we generate artificial run-off triangles satisfying these assumptions. Since the assumptions concern only the conditional mean and variance of the aggregate run-off data (saying nothing about higher moments), we are free to use any probability distribution in generating the artificial data (provided the above mean and variance assumptions are satisfied). Because of its widespread acceptance in modelling loss data and its ease of use, we have used the Log-Normal distribution initially.

The steps used to generate an artificial run-off triangle (and corresponding 'true' ultimates) satisfying Mack's assumptions are as follows:

## Algorithm A for artificial run-off data:

1. Decide on the dimensions of the triangle. We have used 10 origin years with annual development: the number of development years reducing from 10 in the first origin year to 1 in the last.
2. Decide on values of the parameters $f_{k}$ and $\alpha_{k}$ : the values we used are given in Table B-1 below.
3. For each origin year j , generate a value for $\mathrm{C}_{\mathrm{j} 1}$ (representing the amount paid in the first development year). Mack's assumptions say nothing about how these values are generated so we are free to use any method. We used random sampling from a Log-Normal distribution: the same Log-Normal distribution for all origin years (mean $=$ variance $=1.0$ ) but independent random sampling for each origin year.
4. For each origin year, generate $\mathrm{C}_{\mathrm{jk}}$ (for $\mathrm{k}>1$ ) recursively using Mack's assumptions. We generated $\mathrm{C}_{\mathrm{j} 2}$ by random sampling from the shifted Log-Normal distribution that gives values greater than $\mathrm{C}_{\mathrm{j} 1}$, with mean equal to $\mathrm{f}_{1} \cdot \mathrm{C}_{\mathrm{j} 1}$ and variance equal to $\alpha_{1}{ }^{2} \cdot \mathrm{C}_{\mathrm{j} 1}$. We then generated $\mathrm{C}_{\mathrm{j} 3}$ by random sampling from the shifted Log-Normal that gives values greater than $\mathrm{C}_{\mathrm{j} 2}$, with mean equal to $\mathrm{f}_{2} \cdot \mathrm{C}_{\mathrm{j} 2}$ and variance equal to $\alpha_{2}{ }^{2} . \mathrm{C}_{\mathrm{j} 2}$. We continued recursively in this way, using independent random sampling at each stage, until we obtained a value for $\mathrm{C}_{\mathrm{j}, 10}$, which is the 'true' ultimate figure for origin year j . This was repeated for each origin year, using independent random sampling for each one. The triangle was then constructed by discarding the lower right part of the development array (except the $\mathrm{C}_{\mathrm{j}, 10}$ values which were kept as the 'true' ultimates for comparison with estimates produced by applying Mack's method to the upper left triangle). Note that we assumed complete development after 10 years, so no tailfactors were necessary when producing forecasts by Mack's method.

Table B-1 - Parameter values used at Step 2 of Algorithm A

| Dev-yr (k) | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| $\mathrm{f}_{\mathrm{k}}$ | 4.289 | 2.064 | 1.502 | 1.268 | 1.150 | 1.085 | 1.048 | 1.027 | 1.015 |
| $\alpha_{\mathrm{k}}$ | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |

## B.1.3 Performance of Mack's method where its assumptions are true

## B.1.3.1 Details of simulation method

We created 10,000 artificial triangles (and corresponding 'true' ultimates) by Algorithm A, and applied Mack's method to each triangle. For each triangle, we considered only the reserves for the entire triangle (all origin years combined) and we compared the forecast produced by Mack's method to the 'true' reserve.

Mack's method produces a best estimate (equal to the basic chain ladder or BCL estimate) and an estimated root-mean-square prediction error (or just 'standard error'). We used these to calculate the 'standardised predictive error' as [BCL estimate - true reserve] / (Mack's standard error). If Mack's standard errors are correct, then this should have a mean close to zero and mean square close to one.
We also followed Mack's recommendation of using a Log-Normal distribution for the reserve, with mean equal to the BCL estimate and standard deviation as given by Mack's formula. From this, we found the probability (based on knowledge of the upper left triangle only) that the ultimate outcome would be less than what it turned out to be. In other words, if $\mathrm{F}(\mathrm{x})$ denotes the

Log-Normal cumulative distribution function (with mean equal to the BCL estimate and standard deviation as given by Mack), we calculated F (true reserve). If, in addition to Mack's standard errors being correct, the Log-Normal is a reasonable distribution for the predictive error, then the quantity F (true reserve) should be uniformly distributed on the unit interval. Mack also suggests using a Normal distribution (instead of the Log-Normal) when the coefficient of variation is less than $50 \%$, so we have also investigated this.

## B.1.3.2 Results for first $\mathbf{1 0 , 0 0 0}$ simulations

Table B-2 shows results from the first 10,000 simulations carried out. Each value in the second column is the mean (over all 10,000 simulations) of the quantity indicated in the first column.

Table B-2 - Results of applying Mack's method to data generated using Algorithm A

| Simulated quantity | Mean from 10,000 simulations |
| :--- | :--- |
| BCL estimated reserve | 77.74 |
| Indicator that (BCL > True) | $47.25 \%$ |
| BCL estimate - True reserve | 0.68 |
| Mack standard error | 27.88 |
| (BCL - True) / (Mack std error) | -0.60 |
| Square of the above | 6.6 |
| Log-Normal F(true reserve) | 0.571 |

The graph below shows the cumulative distribution function for the quantity (BCL estimate True reserve).


The figure $47.25 \%$ means that in 4,725 out of the 10,000 simulations, the BCL estimate was greater than the 'true' reserve (in the other $52.75 \%$ of simulations the BCL estimate was less than the true reserve). So BCL reserves are smaller than 'median' reserves in this case.
The fact that the mean predictive error $(0.68)$ is such a small proportion $(0.9 \%)$ of the mean estimated reserve (77.74) supports the belief that the basic chain ladder method is unbiased when the assumptions given by Mack hold true (ie, that BCL reserves are 'mean' reserves).
However, the fact that the mean standardised predictive error appears (at -0.60 ) to be significantly below zero, and its mean square (6.6) significantly greater than one, suggests that when the BCL gives an underestimate, the Mack standard error also tends to be understated. So it seems that Mack's method tends to understate the chance of extremely high outcomes. This is confirmed by analysis of the quantity F (true reserve) based on the Log-Normal predictive distribution.

The mean value of F (true reserve) being substantially higher than 0.50 indicates that this quantity is not uniformly distributed on the unit interval as it should be, but tends to be too high. In other words, it seems that actual outcomes (total future claim payments) tend to be higher than indicated by Mack's method.
We have investigated this further by looking at the empirical distribution function of this quantity which is shown below. The vertical axis shows the proportion of simulations in which F (true reserve) was less than the value shown on the horizontal axis. If this quantity were uniformly distributed (as it should be), then this graph should not differ significantly from a straight line from $(0,0)$ to $(1,1)$. Clearly it does: the shape confirms that Mack's method tends to understate reserve uncertainty, particularly the chance of extremely high outcomes (right end of graph below).


In 3,467 simulation out of the first 10,000 simulations ( $34.67 \%$ ), the quantity F (true reserve) exceeds 0.80 , and in 2,455 of these ( $24.55 \%$ ) exceeds 0.90 . If the Log-Normal distribution based on Mack's standard error were correct, these events should occur in only $20 \%$ and $10 \%$ of simulations respectively. Mack's method clearly tends to understate the chance of very high outcomes even when its underlying assumptions hold precisely (at least, this is so for the particular set of parameters selected to generate the artificial data).
At the other extreme (left hand end of the above graph): F (true reserve) is less than 0.20 in $19.58 \%$ of simulations and less than 0.10 in $12.08 \%$ of simulations, so Mack's method seems better at assessing the chances of very low outcomes than the chances of very high outcomes.
These and similar results are summarised below:
Table B-3 - Further results of applying Mack's method to data generated using Algorithm A

| p | Chance that true out-turn exceeds $\mathrm{F}^{-1}(1-\mathrm{p})$ |
| :--- | :--- |
| $1 \%$ | $10.1 \%$ |
| $5 \%$ | $18.02 \%$ |
| $10 \%$ | $24.55 \%$ |
| $20 \%$ | $34.67 \%$ |
| $30 \%$ | $42.66 \%$ |
| $50 \%$ | $58.30 \%$ |
| $70 \%$ | $72.84 \%$ |
| $80 \%$ | $80.42 \%$ |
| $90 \%$ | $87.92 \%$ |
| $95 \%$ | $92.18 \%$ |
| $99 \%$ | $96.48 \%$ |

To make absolutely clear what these results mean: the first row of the table shows there is approximately a $10 \%$ chance that the outcome will exceed the value that Mack's method (with a Log-Normal distribution for the reserve) indicates is the $99^{\text {th }}$ percentile. In other words: while Mack's method indicates there is only a $1 \%$ chance that the outcome will exceed a certain value, there is actually a $10 \%$ chance.

## B.1.3.3 Results for further sets of $\mathbf{1 0 , 0 0 0}$ simulations

To check that the above results are not unduly affected by sampling error, we repeated the calculations for several additional sets of 10,000 simulations, each time using the same parameters for the artificial data but a different seed for the random number generator. The results are summarised below: the first seven rows give mean values (over 10,000 simulations), the remaining rows give the actual proportion of simulations for which the 'true' ultimate exceeded the Mack percentile indicated.
Table B-4 - Results of applying Mack's method to data generated using Algorithm A

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | Set 4 |
| BCL estimate | 77.74 | 76.78 | 77.55 | 78.18 |
| Indicator that (BCL > True) | $47.25 \%$ | $45.82 \%$ | $46.44 \%$ | $46.52 \%$ |
| (BCL estimate - True reserve) | 0.68 | -1.08 | 0.28 | 0.55 |
| Mack standard error | 27.88 | 27.41 | 27.88 | 28.46 |
| (BCL - True) / (Mack std error) | -0.60 | -0.66 | -0.62 | -0.63 |
| Square of the above | 6.6 | 7.7 | 5.2 | 5.7 |
| Log-Normal F(true liability) | 0.571 | 0.582 | 0.577 | 0.577 |
| $1 \%$ | $10.1 \%$ | $10.3 \%$ | $10.5 \%$ | $10.4 \%$ |
| $5 \%$ | $18.0 \%$ | $18.8 \%$ | $18.0 \%$ | $18.9 \%$ |
| $10 \%$ | $24.6 \%$ | $26.0 \%$ | $25.6 \%$ | $25.2 \%$ |
| $20 \%$ | $34.7 \%$ | $35.9 \%$ | $35.6 \%$ | $35.0 \%$ |
| $30 \%$ | $42.7 \%$ | $44.5 \%$ | $43.6 \%$ | $43.8 \%$ |
| $50 \%$ | $58.3 \%$ | $59.7 \%$ | $58.6 \%$ | $59.2 \%$ |
| $70 \%$ | $72.8 \%$ | $73.9 \%$ | $73.5 \%$ | $73.4 \%$ |
| $80 \%$ | $80.4 \%$ | $81.0 \%$ | $80.1 \%$ | $80.8 \%$ |
| $90 \%$ | $87.9 \%$ | $88.2 \%$ | $87.7 \%$ | $87.9 \%$ |
| $95 \%$ | $92.2 \%$ | $92.1 \%$ | $92.0 \%$ | $92.0 \%$ |
| $99 \%$ | $96.5 \%$ | $96.2 \%$ | $96.2 \%$ | $96.2 \%$ |

These results show that the apparent anomalies in the initial 10,000 simulations do not result from sampling error: there is little difference between results from the four independent sets of 10,000 simulations.

## B.1.3.4 Results based on alternative parameters in Algorithm A

The particular parameters that gave the above results are the first set of parameters that we tried for constructing artificial triangles satisfying Mack's assumptions. The use of a Log-Normal distribution with mean $=$ variance $=1.0$ to generate $\mathrm{C}_{\mathrm{j} 1}$ (Step 3 of the data generation algorithm) tends to produce greater variation between origin years than in a typical real-world triangle. To test the sensitivity of the results to this aspect, we repeated the simulations, but this time with the variance of the Log-Normal distribution used at Step 3 reduced to 0.01 . This produces more realistic looking triangles. Results are summarised in Table B- 5 below.
In this set of results, we have also calculated percentiles using a Normal predictive distribution (instead of a Log-Normal) in those simulations for which Mack's standard error is less than $50 \%$ of the BCL reserve (as suggested by Mack).

Table B-5 - Further results of applying Mack's method to data generated using Algorithm A

| Variation coefficient of $\mathrm{C}_{\mathrm{j} 1}$ <br> reduced from $100 \%$ to $10 \%$. | Results from 10,000 simulations |  |
| :--- | :--- | :--- |
| BCL estimate | 78.0 |  |
| Indicator that (BCL > True) | $48.1 \%$ |  |
| (BCL estimate - True reserve) | 0.93 |  |
| Mack standard error | 29.52 |  |
| (BCL - True) / (Mack std error) | -0.51 |  |
| Square of the above | 4.2 |  |
| F(true liability) | Log-Normal <br> 0.569 | Normal when (Std Error / BCL) <br> 0.555 |
| $1 \%$ | $8.4 \%$ | $12.0 \%$ |
| $5 \%$ | $16.3 \%$ | $18.2 \%$ |
| $10 \%$ | $22.5 \%$ | $23.0 \%$ |
| $20 \%$ | $32.6 \%$ | $30.7 \%$ |
| $30 \%$ | $41.2 \%$ | $37.9 \%$ |
| $50 \%$ | $57.8 \%$ | $53.0 \%$ |
| $70 \%$ | $73.9 \%$ | $71.7 \%$ |
| $80 \%$ | $81.9 \%$ | $82.2 \%$ |
| $90 \%$ | $90.0 \%$ | $92.6 \%$ |
| $95 \%$ | $93.8 \%$ | $97.0 \%$ |
| $99 \%$ | $97.8 \%$ | $99.2 \%$ |

This set of results (using the Log-Normal predictive distribution) shows a slight improvement compared to the previous results. Using the Normal distribution for predictions when the predictive coefficient of variation is less than $50 \%$ produces a less accurate assessment of the chances of extreme adverse outcomes: the proportion of simulations in which the true outcome exceeds the $99^{\text {th }}$ percentile increases from $8.4 \%$ when the Log-Normal is always used to calculated the $99^{\text {th }}$ percentile, to $12.0 \%$ when the $99^{\text {th }}$ percentile is sometimes calculated using a Normal distribution. The main conclusions are unchanged: even where its assumptions are perfectly satisfied, Mack's method does not necessarily give a correct assessment of reserve uncertainty. For triangles constructed with the particular parameter values that we have used, Mack's method significantly understates the chance of very adverse outcomes.

## B.1.3.5 Independent checking of the above results

Our main finding - that Mack's method can substantially understate the chance of extreme adverse outcomes even if its assumptions are perfectly satisfied - is clearly important given the quite widespread use of Mack's method and (with the advent of ICA and Solvency II) the increasing importance of accurately assessing the chances of extreme adverse outcomes. For this reason we carried out thorough checking of these results to ensure that they are genuine and not the result of errors in our implementation of Mack's method or in our simulation of triangles that satisfy Mack's assumptions.

To provide a final comprehensive check, another member of the working party, who was not involved in the work described above, carried out a completely independent simulation exercise (working only from Mack's 1993 paper and preceding sections of the present paper). All the simulation results described in preceding and later sections of the present paper were obtained using C++ programs (written by Tom Wright) to generate the artificial triangles and to apply Mack's method (and other stochastic methods) to those triangles. For the independent exercise, Mack's method and Algorithm A (as described above) were implemented by Gary Dunne in Miscrosoft Excel using VBA macros. This independent implementation was used to carry out a further 10,000 simulations with the same parameter values as those used (in the $\mathrm{C}++$ programs) to produce the results in Table B-5 above. A Log-Normal predictive distribution was used in every simulation. The results are summarised below.

Table B-6 -Results obtained from an independent simulation exercise (Mack's method applied to data generated using Algorithm A, with same parameters as in Table B-5)

|  | Results from 10,000 simulations |
| :--- | :--- |
| BCL estimate | 76.8 |
| Indicator that (BCL > True) | $47.3 \%$ |
| (BCL estimate - True reserve) | -0.34 |
| Mack standard error | 29.12 |
| (BCL - True) / (Mack std error) | -0.53 |
| Square of the above | 3.7 |
| $1 \%$ | $8.4 \%$ |
| $5 \%$ | $16.7 \%$ |
| $10 \%$ | $23.5 \%$ |
| $20 \%$ | $33.1 \%$ |
| $30 \%$ | $41.7 \%$ |
| $50 \%$ | $58.1 \%$ |
| $70 \%$ | $74.2 \%$ |


| $80 \%$ | $82.6 \%$ |
| :--- | :--- |
| $90 \%$ | $90.3 \%$ |
| $95 \%$ | $94.8 \%$ |
| $99 \%$ | $98.0 \%$ |

Differences between the results in the above two tables are not statistically significant: they are consistent with sampling variation resulting from the two sets of results being based on independent sets of 10,000 artificial triangles. (This is clear from the fact that the differences are of the same order of magnitude as differences in results from different sets of 10,000 triangles shown in Table B-4).

Having obtained essentially the same results in two quite separate and independent simulation exercises, we are confident that these results are genuine.

## B.1.3.6 Results for triangles where not all increments are positive

At Step 4 of Algorithm A, we chose initially to use shifted Log-Normals (such that the mean and variance after shifting were as required by Mack's assumptions) so that the artificial data had only positive increments. To test the sensitivity of the results to this aspect, we have also generated artificial triangles using Log-Normal distributions directly for the cumulative amounts $\mathrm{C}_{\mathrm{jk}}$ at Step 4 (so the increments will sometimes be negative, as might occur in incurred data). Results for 10,000 simulations are shown below. Note that the triangles generated in this way still perfectly satisfy Mack's assumptions (as do all triangles analysed in Section 2.4.3).

Table B-7 - Further results of applying Mack's method to data generated using Algorithm A

| Simulated data with negative increments. | Results from 10,000 simulations |
| :--- | :--- |
| BCL estimate | 77.8 |
| Indicator that (BCL > True) | $47.2 \%$ |
| (BCL estimate - True reserve) | 0.76 |
| Mack standard error | 55.4 |
| (BCL - True) / (Mack std error) | -0.47 |
| Square of the above | 4.3 |
| Log-Normal F(true liability) | 0.582 |
| $1 \%$ | $13.4 \%$ |
| $5 \%$ | $23.2 \%$ |
| $10 \%$ | $30.6 \%$ |
| $20 \%$ | $41.0 \%$ |
| $30 \%$ | $48.3 \%$ |


| $50 \%$ | $60.8 \%$ |
| :--- | :--- |
| $70 \%$ | $70.8 \%$ |
| $80 \%$ | $76.0 \%$ |
| $90 \%$ | $81.9 \%$ |
| $95 \%$ | $85.6 \%$ |
| $99 \%$ | $90.7 \%$ |

These results show the performance is worse here than when Mack's method is applied to triangles that only have positive increments. Here, (with the possibility of negative increments) the method significantly understates the chances of extremely low outcomes as well as understating the chance of very high outcomes: in $9.3 \%$ of simulations the eventual outcome was lower than what was supposed to be the $1^{\text {st }}$ percentile, and in $13.4 \%$ the eventual outcome was higher than what was supposed to be the $99^{\text {th }}$ percentile.

## B.1.3.7 Possible explanation of our findings

It is beyond the scope of this paper to definitively explain where Mack's method goes wrong. One possibility is that significant bias is introduced by taking non-linear functions of unbiased parameter estimates. Mack's formulas for the standard error have terms with $\mathrm{f}_{\mathrm{k}}{ }^{2}$ in the denominator. Mack uses unbiased (chain ladder) estimators for $f_{k}$, but an unbiased estimate of $f_{k}$ does not give an unbiased estimate of $1 / \mathrm{f}_{\mathrm{k}}{ }^{2}$. Given the relatively large number of parameters estimated from the data ( 9 f-parameters and $9 \alpha$-parameters from 55 data points), the reliability of some of these will be relatively low, which could result in substantial bias in non-linear functions (such as $1 / \mathrm{f}_{\mathrm{k}}{ }^{2}$ ). In addition, the square root of an unbiased estimate of the mean-square prediction error is not an unbiased estimate of the root-mean-square prediction error (or 'standard error') because square root is not a linear function. Further, even if the method did produce an unbiased estimate of the standard error, it would not necessarily give correct percentiles because the predictive distribution function is not linear (ie the predictive distribution is not uniform).
We expect to find that other stochastic methods suffer from similar problems, but this should be less so for Bayesian methods because these do not focus on unbiased estimation of parameters of the predictive distribution: instead they aim to calculate the predictive distribution function directly.
It would be interesting to investigate how Mack's method performs on triangles constructed (by Algorithm A) using a wider range of parameter values (ie different values of $f_{k}, \alpha_{k}$, and parameters of the loss distribution).

## B.1.4 Performance of Mack's method where its assumptions are false

Given the findings of the previous section, we should not expect that Mack's method will perform well where its underlying assumptions are false. In his 1993 paper Mack advises against applying the method where the assumptions are believed to be false, and gives suggestions for checking the assumptions using the data. However, these diagnostic checks may not always be carried out, and even where they are, they may show no clear evidence that the assumptions are
false where this is in fact the case: this is quite possible given the small volume of data in many run-off triangles. Therefore, it is of interest to explore how the method performs when its assumptions are false.
Testing of Mack's method when its assumptions are false is carried out in later sections by applying Mack's method to artificial triangles generated to satisfy the assumptions of other stochastic methods.

## B. 2 Assessment of over-dispersed Poisson (ODP) methods by simulation

## B.2.1 Assumptions of the ODP model

The over-dispersed Poisson (ODP) model (as described, for example by Renshaw and Verrall 1998) is based on the following three assumptions (we use $\mathrm{Y}_{\mathrm{jk}}$ to denote the aggregate incremental amount paid in development period k of origin year j ):

1. $E\left(Y_{j k}\right)=x_{j} \cdot p_{k}$
2. $\operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}}\right)=\varphi \cdot \mathrm{x}_{\mathrm{j}} \cdot \mathrm{p}_{\mathrm{k}}$
3. The $\mathrm{Y}_{\mathrm{jk}}$ are stochastically independent (across all j and all k ).

If the parameters $p_{k}$ are normalised so that $\Sigma_{k} p_{k}=1$, then they represent the proportion of the ultimate paid amount expected to fall in each development year $k$, and $x_{j}$ represents the expected ultimate amount for origin year $j$ (where all expectations here are those applying before any $\mathrm{Y}_{\mathrm{jk}}$ have been observed).

Renshaw and Verrall's version of the ODP method uses quasi-likelihood maximisation to estimate the parameters of the model and produces 'best estimate' reserves that are identical (provided the run-off array has no missing values) to those produced by the basic chain ladder method. Renshaw and Verrall give formulas for predictive standard errors of the BCL reserves incorporating parameter and process error. England and Verrall (1999) and England (2001) also consider the ODP model, and they describe a bootstrap procedure that is also intended to give predictive standard errors incorporating parameter and process error. Here, we consider both these methods which we call respectively the 'analytic ODP method' and the 'bootstrap ODP method'. Both these methods are based on the same stochastic model, which is fully described by the above three assumptions.

## B.2.2 Artificial data satisfying ODP assumptions (Algorithm B)

To test the ODP methods when the underlying assumptions are true, we generate artificial data that satisfy the above three assumptions. Since the assumptions say nothing about the distribution of the $Y_{j k}$ other than the constraints on the first two moments represented by assumptions 1 and 2, we are free to use any distribution for this purpose. We have chosen to use a compound Poisson/Log-Normal distribution for each $\mathrm{Y}_{\mathrm{jk}}$ because it is possible that actual paid run-off data may approximate this. In other words, each $\mathrm{Y}_{\mathrm{jk}}$ is constructed as the sum of a number $\left(\mathrm{N}_{\mathrm{jk}}\right.$ say) of independent Log-Normal amounts, where the number $\mathrm{N}_{\mathrm{jk}}$ is generated from a Poisson distribution. The parameters of the Log-Normal distribution used for the individual claim amounts are constant across all cells ( $\mathrm{j}, \mathrm{k}$ ) of the run-off triangle: this ensures that the quantity $\operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}}\right) / \mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}}\right)$ is the same in all cells (as required by assumptions 1 and 2, where this ratio is denoted $\varphi$ ).

The algorithm we have used to generate the artificial data is as follows:

## Algorithm B for artificial run-off data:

1. The ultimate number of claims in an origin year is generated by random sampling from a Poisson distribution (same parameters for each origin year, but independent sampling).
2. Each claim is assumed to be settled by a single payment, and the development year of the payment determined by independent random sampling from a Multinomial distribution (same parameters for each origin year: these are the parameters denoted $\mathrm{p}_{\mathrm{k}}$ in the above assumptions of the ODP method).
3. The amount of each individual claim payment is determined by independent random sampling from a Log-Normal distribution (same parameters in every cell of the triangle).
4. The amounts of claims settling in the upper left triangle of the run-off array are accumulated to create this run-off triangle, and all claim amounts (regardless of the development year when settled) are accumulated to obtain the 'true' ultimate position for each origin year.

It is perhaps not immediately obvious that Steps 1 and 2 produce stochastically independent Poisson numbers $\left(\mathrm{N}_{\mathrm{jk}}\right)$ across all development years of a given origin year (which is required by Assumption 3 of the ODP method). However, this is quite a well known result: a proof is given in the last section of this appendix, which also shows that the constant $\varphi$ of Assumption 2 is related to the mean $\mu$ and variance $\sigma^{2}$ of the loss distribution used at Step 3 by $\varphi=\left(\mu^{2}+\sigma^{2}\right) / \mu$.

Algorithm B is clearly simpler in many ways than what typically occurs in reality, but is nevertheless more realistic than Algorithm A (used to generate data strictly in accordance with Mack's assumptions). For this reason, as well as using data generated by Algorithm B to test the performance of the ODP methods where their assumptions are true, we have used the same artificial data to test Mack's method where its assumptions are false. (It is easily proved that triangles generated using Algorithm B do not satisfy Mack's assumptions - see the last section of this appendix.)

## B.2.3 Performance of analytic ODP method when its assumptions are true

We carried out three sets of 10,000 independent simulations. In each simulation, a run-off triangle was generated using Algorithm B, and the analytic ODP method applied to produce reserve estimates (equal to BCL estimates) and root-mean-square prediction errors (called 'standard errors' below) calculated as described by Renshaw and Verrall 1998. Again, we looked at only the total reserve for all years combined (rather than the estimates for separate origin years). In quasi-likelihood modelling, there are two commonly used methods for estimating the dispersion parameter (the parameter $\varphi$ of ODP Assumption 2): using the sum of squared Pearson residuals, or using the deviance. As there is little theoretical basis for favouring one over the other, we have tested both methods: in each set of results below, the first column relates to the Pearson residual method, the second column to the deviance method. Since the ODP method does not produce a full predictive distribution, we have used a Log-Normal distribution fitted to the best estimate and standard error to obtain percentiles (in the same way as described earlier for Mack's method). Results are summarised below.

Table B-8 - Results of applying the analytic ODP method to data generated using Algorithm B

|  | Results from different sets of 10,000 simulations |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 |  |  |  |
|  |  |  |  |  |  |  |
| Best estimate reserve | 3,668 |  | 3,654 |  | 3,632 |  |
| Best estimate > True | $46.9 \%$ |  | $46.8 \%$ |  | $46.4 \%$ |  |
| (Best estimate - True) | 303.7 |  | 292.7 |  | 268.4 |  |
| ODP standard error | 1,511 | 1,491 | 1,514 | 1,494 | 1,502 | 1,482 |
| (Best - True) / (std error) | -0.271 | -0.276 | -0.263 | -0.267 | -0.280 | -0.285 |
| Square of the above | 1.48 | 1.51 | 1.46 | 1.50 | 1.45 | 1.49 |
| Log-Normal F(True) | 0.571 | 0.570 | 0.568 | 0.568 | 0.574 | 0.573 |
| $1 \%$ | $2.4 \%$ | $2.6 \%$ | $2.7 \%$ | $2.7 \%$ | $2.6 \%$ | $2.7 \%$ |
| $5 \%$ | $9.5 \%$ | $9.8 \%$ | $9.1 \%$ | $9.3 \%$ | $9.9 \%$ | $10.2 \%$ |
| $10 \%$ | $17.1 \%$ | $17.4 \%$ | $16.4 \%$ | $16.7 \%$ | $17.2 \%$ | $17.4 \%$ |
| $20 \%$ | $29.8 \%$ | $30.0 \%$ | $29.0 \%$ | $29.2 \%$ | $29.7 \%$ | $29.8 \%$ |
| $30 \%$ | $41.0 \%$ | $41.1 \%$ | $40.3 \%$ | $40.4 \%$ | $41.1 \%$ | $41.1 \%$ |
| $50 \%$ | $60.4 \%$ | $60.2 \%$ | $60.3 \%$ | $60.1 \%$ | $60.9 \%$ | $60.8 \%$ |
| $70 \%$ | $77.3 \%$ | $77.1 \%$ | $76.6 \%$ | $76.3 \%$ | $77.6 \%$ | $77.3 \%$ |
| $80 \%$ | $84.1 \%$ | $83.8 \%$ | $84.1 \%$ | $83.8 \%$ | $84.8 \%$ | $84.5 \%$ |
| $90 \%$ | $90.8 \%$ | $90.7 \%$ | $91.0 \%$ | $90.8 \%$ | $91.4 \%$ | $91.2 \%$ |
| $95 \%$ | $94.4 \%$ | $94.4 \%$ | $94.8 \%$ | $94.7 \%$ | $94.7 \%$ | $94.6 \%$ |
| $99 \%$ | $97.9 \%$ | $97.8 \%$ | $98.1 \%$ | $98.1 \%$ | $98.2 \%$ | $98.2 \%$ |

These results show that the analytic ODP method is biased upwards when applied to triangles generated using Algorithm B. Over all 30,000 simulations the ODP reserve overstates the true reserve by about $7.9 \%$ of the mean true reserve, and this occurs fairly consistently across the three sets of 10,000 simulations. Since the best estimate reserves produced by the ODP method are identical to those produced using the basic chain ladder (BCL) method, this means that the BCL method is biased upwards when applied to data of this type. The results in Table B-8 also show that, although the BCL estimate is positively biased, it exceeds the true reserve in only about $47 \%$ of simulations. It would be interesting to investigate to what extent this positive bias of the BCL depends on the particular parameter values used in generating the triangles.

The fact that there is bias should not be surprising given that the assumptions Mack showed to be necessary for unbiasedness of the BCL are known to be violated here. The bias of the BCL method when applied to triangles with independent increments has previously been noted and
studied by Stanard (1985) (who used a numerical simulation method similar to that described here) and by Taylor (2001 and 2002) who derived approximate formulas for the bias.
The positive bias of the BCL (and the less than $50 \%$ chance that the BCL estimate exceeds the true reserve) is evident from the following graph, which is the cumulative distribution function of the quantity (BCL estimate - True reserve):


The negative mean value of the 'standardised predictive error' implies (given that the 'best estimate' is positively biased) that when the method understates the true outcome, the analytic ODP standard error also tends to be understated.

The distribution of the quantity F (true reserve) (where $\mathrm{F}(\mathrm{x})$ is the Log-Normal predictive distribution) shows that this method understates the chance of very adverse outcomes. For example, in about $2.6 \%$ of simulations the true outcome exceeded what was supposed to be the $99^{\text {th }}$ percentile. The full empirical distribution function of this quantity for the first set of 10,000 simulations is shown below: it clearly differs significantly from a uniform distribution (which would be a straight line from $(0,0)$ to $(1,1))$.


The analytic ODP method appears to perform slightly better when the dispersion parameter is estimated from the Pearson residuals (first column for each set of 10,000 simulations) than when it is estimated from the deviance (second column).

## B.2.4 Performance of bootstrap ODP method when its assumptions are true

## B.2.4.1 Two variants of bootstrap ODP method

We have tested the performance of two variants of the bootstrap ODP method as described by England \& Verrall (1999) and England (2001).
In their 1999 paper England \& Verrall describe a bootstrap procedure that is intended to give a probability distribution in respect of BCL parameter uncertainty only (ie not including future process uncertainty). They point out that the predictive standard error (including process as well as parameter uncertainty) can be obtained by adding the estimated process variance to the variance of the bootstrap distribution (representing parameter uncertainty only) then taking the square root. They propose estimating the process variance as $\varphi$ times the BCL reserve, where $\varphi$ is the estimated dispersion parameter of the ODP assumption. They found that, in the case of one particular triangle, this method gave a predictive standard error close to that given by the analytic ODP method. The 1999 paper does not describe a procedure for obtaining a full predictive probability distribution incorporating process and parameter uncertainty.
In his 2001 paper, England outlines an ODP bootstrap procedure for obtaining a full predictive probability distribution (incorporating both parameter and process uncertainty). In this procedure, the allowance for future process variation is based on BCL projections from bootstrap pseudo-data (not on BCL projections of the original data as in the 1999 paper) and is simulated by adding process variation to each future cell of the pseudo-data triangle. In their 2002 paper, England and Verrall applied this method to one particular triangle and obtained a predictive
standard error somewhat higher than given by the analytic ODP method: the values were 19,267 and 18,193 respectively (which they described as close). They also found that the mean of the bootstrap predictive distribution was higher than the BCL reserve: the values were 53,210 and 52,135 (described as reassuringly close).
We have tested both these variants of the bootstrap ODP method: referred to below as the " 1999 bootstrap ODP method" and " 2001 bootstrap ODP method".

Our implementations of these methods are exactly as described in the papers by England and Verrall except in one respect: we ensure that all pseudo data triangles have no negative cumulative paid amounts. This is done by applying the algorithm to produce each pseudo triangle exactly as described by England and Verrall, but then rejecting any triangle that has any negative cumulatives. Each time a triangle is rejected another one is created and the final bootstrap predictive distribution is based on the outcomes from 1,000 pseudo-triangles in which all cumulative amounts are positive ( 1,000 is the number of bootstrap simulations suggested by England and Verrall). The reason we have made this modification is that we found the performance of the bootstrap method is sometimes extremely poor if pseudo-data triangles with negative cumulatives are allowed.
For example, consider the following cumulative run-off triangle. This is an artificial triangle generated using Algorithm B. The final column gives basic chain ladder (BCL) ultimates.

| 11.1 | 122.1 | 206.5 | 353.6 | 534.7 | 547.0 | 581.9 | 621.4 | 802.3 | 807.8 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 22.1 | 79.0 | 902.4 | $1,037.5$ | $1,450.9$ | $1,521.3$ | $1,556.5$ | $1,604.3$ | $1,604.8$ | $1,615.9$ |
| 22.7 | 115.1 | 233.4 | 408.6 | 546.0 | 616.8 | 692.8 | 712.0 |  | 775.3 |
| 115.5 | 340.3 | 491.4 | 622.6 | 716.5 | 870.9 | 954.0 |  |  | $1,077.9$ |
| 0.0 | 324.2 | 488.3 | 810.7 | 996.7 | $1,073.5$ |  |  |  | $1,291.2$ |
| 105.1 | 222.8 | 392.5 | 629.1 | 743.9 |  |  |  |  | 975.8 |
| 18.2 | 120.2 | 365.3 | 452.1 |  |  |  |  |  | 766.0 |
| 34.3 | 386.7 | 461.5 |  |  |  |  |  |  | $1,095.3$ |
| 130.7 | 249.4 |  |  |  |  |  |  |  | $1,225.5$ |
| 90.4 |  |  |  |  |  |  |  |  | $1,892.9$ |

The triangle below is a pseudo-data set produced by applying the bootstrap algorithm described by England and Verrall (1999) to the above triangle. The final column gives the BCL ultimates for this pseudo triangle.

| 32.248 | 242.0 | 481.5 | 666.0 | 893.3 | 937.1 | $1,029.6$ | $1,116.9$ | $1,117.2$ | $1,115.2$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2.194 | 323.8 | 607.6 | 790.2 | $1,136.9$ | $1,166.7$ | $1,542.3$ | $1,646.7$ | $2,109.3$ | $2,105.5$ |
| -4.485 | -50.4 | 231.9 | 462.7 | 554.2 | 584.4 | 654.3 | 661.6 |  | 771.0 |
| -42.552 | 85.5 | 215.7 | 498.7 | 505.8 | 728.7 | 820.0 |  |  | $1,014.5$ |
| 32.145 | 95.6 | 232.1 | 381.9 | 528.5 | 617.7 |  |  |  | 905.1 |
| -45.767 | 142.1 | 204.4 | 299.0 | 520.3 |  |  |  |  | 850.0 |
| -10.300 | 97.9 | 220.4 | 297.1 |  |  |  |  |  | 648.4 |
| 27.423 | 176.9 | 517.4 |  |  |  |  |  |  | $1,747.8$ |
| 9.118 | 133.5 |  |  |  |  |  |  |  | $1,098.4$ |
| 192.355 |  |  |  |  |  |  |  |  | $77,700,000.0$ |

There is nothing to prevent negative increments in pseudo data-triangles generated as described by England and Verrall, and negative increments may produce negative cumulatives. In this example, the pseudo data has several negatives in the first development period. As a result, the first BCL development factor is extremely large: $(1246.9 / 0.025=49,097.8)$ which produces a very high BCL ultimate for the last origin year ( 77.7 million). If something like this occurs in one of 1,000 bootstrap simulations, it can have a massive impact on both the mean and the variance of the bootstrap distribution. Given that the original triangle has only positive increments, it would in most circumstances be unrealistic to allow for such extreme possibilities in bootstrap simulations. For this reason, in our implementation of the bootstrap methods, we reject any pseudo triangles in which the cumulative amounts are not all positive. Pseudo triangles with negative increments are accepted provided all cumulative figures are positive.
To check that our bootstrap procedures are correct implementations of the methods described by England and Verrall we replicated their results as given in the table on page 293 of their 1999 paper and in Appendix A of England 2001. Our bootstrap results differed from theirs by an amount consistent with bootstrap sampling error. The modification described above (rejection of pseudo-triangles with negative cumulatives) makes no difference to the results obtained for the particular two triangles analysed in the papers by England and Verrall.

For each of the two variants of the bootstrap ODP method (1999 and 2001 variants) we have carried out three sets of 10,000 independent simulations, each based on a run-off triangle generated using Algorithm B. For each of the 30,000 artificial run-off triangles, we used 1,000 bootstrap simulations (as suggested by England \& Verrall).

## B.2.4.2 Performance of 1999 bootstrap ODP method when its assumptions are true

Table B-9 below gives results of applying the 1999 bootstrap ODP method to data generated using Algorithm B. The 1999 bootstrap ODP method does not give a full predictive distribution so we have used a Log-Normal with mean equal to the BCL reserve and variance equal to the predictive variance (calculated as the bootstrap variance representing parameter uncertainty, plus $\varphi$ times the BCL reserve). Note that the mean of the Log-Normal predictive distribution was set equal to the BCL estimate (that is, the ODP best estimate), not to the mean of the bootstrap distribution, because our purpose here is to assess the performance of published methods. Although England \& Verrall do not suggest the use of a Log-Normal predictive distribution, they make it clear that the predictive variance calculated as described above is supposed to be the predictive variance of the ODP best estimate reserve (ie the BCL reserve): they ignore the difference between the BCL reserve and the mean of the bootstrap distribution on the grounds that these two quantities are close.

Table B-9 - Results of applying 1999 bootstrap ODP method to data generated using Algorithm B

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | All combined |
| BCL estimate | 3,668 | 3,654 | 3,632 | 3,651 |
| Best estimate > True | $46.9 \%$ | $46.8 \%$ | $46.4 \%$ | $46.7 \%$ |
| (BCL estimate - True reserve) | 303.7 | 292.7 | 268.4 | 288.3 |
| Bootstrap mean | 3,727 | 3,719 | 3,695 | 3,714 |
| Bootstrap standard error | 1,304 | 1,302 | 1,295 | 1,300 |
| (BCL - True) / (BS std error) | -0.251 | -0.242 | -0.261 | -0.251 |
| Square of the above | 1.727 | 1.722 | 1.703 | 1.717 |
| Log-Normal F(true liability) | 0.559 | 0.556 | 0.562 | 0.559 |
| $1 \%$ | $3.1 \%$ | $3.1 \%$ | $3.0 \%$ | $3.1 \%$ |
| $5 \%$ | $10.6 \%$ | $10.2 \%$ | $11.1 \%$ | $10.6 \%$ |
| $10 \%$ | $18.4 \%$ | $17.8 \%$ | $18.6 \%$ | $18.3 \%$ |
| $20 \%$ | $31.1 \%$ | $30.0 \%$ | $30.8 \%$ | $30.6 \%$ |
| $30 \%$ | $41.4 \%$ | $40.8 \%$ | $41.4 \%$ | $41.2 \%$ |
| $50 \%$ | $58.8 \%$ | $58.6 \%$ | $59.3 \%$ | $58.9 \%$ |
| $70 \%$ | $74.3 \%$ | $73.6 \%$ | $74.4 \%$ | $74.1 \%$ |
| $80 \%$ | $80.9 \%$ | $80.6 \%$ | $81.5 \%$ | $81.0 \%$ |
| $90 \%$ | $87.7 \%$ | $87.8 \%$ | $88.2 \%$ | $87.9 \%$ |
| $95 \%$ | $91.8 \%$ | $91.7 \%$ | $92.1 \%$ | $91.9 \%$ |
| $99 \%$ | $96.1 \%$ | $96.4 \%$ | $96.4 \%$ | $96.3 \%$ |

These results show that there is in fact a statistically significant difference between the BCL reserve and the mean of the bootstrap distribution in this case. Over all 30,000 artificial datasets, the mean of the bootstrap distribution exceeds the mean BCL reserve by about $1.7 \%$. This difference occurs consistently across all three sets of 10,000 triangles, showing that it is statistically significant (ie is not just caused by bootstrap sampling error). The graph below shows the cumulative probability distribution (over all 30,000 triangles) for the ratio of bootstrap mean to BCL reserve. The bootstrap mean exceeds the BCL reserve for approximately $79 \%$ of the triangles we generated using Algorithm B.


Comparing Tables B-8 and B-9 we see that the 1999 bootstrap standard error tends to be lower than the analytic standard error, and percentiles based on the bootstrap standard error are less accurate. The graph below shows the cumulative probability distribution (over all 30,000 triangles) for the ratio of bootstrap standard error to analytic ODP standard error. The bootstrap standard error is smaller than the analytic standard error for approximately $86 \%$ of the triangles generated using Algorithm $B$.


In summary it seems that the bootstrap ODP method (as described by England \& Verrall 1999) does not perform quite as well as the analytic ODP method when the ODP assumptions are perfectly satisfied. (At least, this is the case for triangles generated using the particular set of parameters in Algorithm B that we used here.)

## B.2.4.3 Performance of 2001 bootstrap ODP method when its assumptions are true

The table below gives results of applying the 2001 bootstrap method to data generated using Algorithm B. The 2001 bootstrap method does give a full predictive distribution so we have not used a Log-Normal distribution for percentiles: we have calculated percentiles directly from the full predictive distribution given by the bootstrap method.
Table B-10 - Results of applying 2001 bootstrap ODP method to data generated using Algorithm B

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | All combined |
| BCL estimate | 3,668 | 3,654 | 3,632 | 3,651 |
| Best estimate > True | $46.9 \%$ | $46.8 \%$ | $46.4 \%$ | $46.7 \%$ |
| (BCL estimate - True reserve) | 303.7 | 292.7 | 268.4 | 288.3 |
| Bootstrap mean | 3,727 | 3,719 | 3,695 | 3,714 |
| Bootstrap standard error | 1,307 | 1,309 | 1,299 | 1,305 |
| (BCL - True) / (BS std error) | -0.245 | -0.237 | -0.255 | -0.246 |
| Square of the above | 1.701 | 1.692 | 1.670 | 1.688 |
| F(true liability) | 0.527 | 0.523 | 0.530 | 0.527 |
| $1 \%$ | $2.6 \%$ | $2.8 \%$ | $2.5 \%$ | $2.6 \%$ |
| $5 \%$ | $8.3 \%$ | $8.0 \%$ | $8.4 \%$ | $8.2 \%$ |
| $10 \%$ | $14.3 \%$ | $13.7 \%$ | $14.5 \%$ | $14.2 \%$ |
| $20 \%$ | $24.8 \%$ | $24.4 \%$ | $25.3 \%$ | $24.8 \%$ |
| $30 \%$ | $34.8 \%$ | $34.2 \%$ | $35.3 \%$ | $34.8 \%$ |
| $50 \%$ | $53.9 \%$ | $53.4 \%$ | $54.4 \%$ | $53.9 \%$ |
| $70 \%$ | $71.8 \%$ | $71.2 \%$ | $71.9 \%$ | $71.6 \%$ |
| $80 \%$ | $80.1 \%$ | $79.8 \%$ | $80.8 \%$ | $80.2 \%$ |
| $90 \%$ | $88.6 \%$ | $88.4 \%$ | $89.0 \%$ | $88.7 \%$ |
| $95 \%$ | $93.0 \%$ | $93.1 \%$ | $93.2 \%$ | $93.1 \%$ |
| $99 \%$ | $97.3 \%$ | $97.5 \%$ | $97.6 \%$ | $97.5 \%$ |

The bootstrap 2001 standard error is smaller than the analytic ODP standard error in $82 \%$ of simulations. However, comparing the above results with Table B-8 we see that the assessment of
uncertainty for adverse outcomes, while not perfect, is about as good as provided by the analytic standard error. (Using either method the true outcome exceeds what is supposed to be the $99^{\text {th }}$ percentile in approximately $2.6 \%$ of cases.) Since the standard error tends to be smaller but does not understate the chance of extreme adverse events any more than the analytic standard error, the 2001 bootstrap method appears to be the better method for assessing the reliability of BCL reserves when the ODP assumptions are perfectly satisfied. This may be because the method gives a full predictive distribution so it is not necessary to use an analytic predictive distribution (such as the Log-Normal) which may not have an appropriate shape at the extremes.

## B.2.5 Performance of Mack's method when its assumptions are false

We applied Mack's method to the same artificial data (generated using Algorithm B) as used to test the performance of the ODP method when its assumptions are true. It is easily shown that data generate in this way do not satisfy Mack's assumptions (see the last section of this appendix). Results are summarised below. The final column gives results for all 30,000 simulations combined.
Table B-11 - Results of applying Mack's method to data generated using Algorithm B

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | All combined |
| BCL estimate | 3,668 | 3,654 | 3,632 | 3,651 |
| Best estimate > True | $46.9 \%$ | $46.8 \%$ | $46.4 \%$ | $46.70 \%$ |
| (BCL estimate - True reserve) | 303.7 | 292.7 | 268.4 | 288.3 |
| Mack standard error | 1,961 | 1,960 | 1,951 | 1,957 |
| (BCL - True) / (Mack std error) | -0.222 | -0.210 | -0.227 | -0.220 |
| Square of the above | 1.11 | 1.10 | 1.08 | 1.10 |
| Log-Normal F(true liability) | 0.586 | 0.583 | 0.589 | 0.586 |
| $1 \%$ | $1.4 \%$ | $1.5 \%$ | $1.4 \%$ | $1.4 \%$ |
| $5 \%$ | $6.8 \%$ | $6.4 \%$ | $6.7 \%$ | $6.6 \%$ |
| $10 \%$ | $13.1 \%$ | $12.8 \%$ | $13.5 \%$ | $13.1 \%$ |
| $20 \%$ | $27.2 \%$ | $26.5 \%$ | $27.4 \%$ | $27.0 \%$ |
| $30 \%$ | $41.2 \%$ | $40.5 \%$ | $41.2 \%$ | $41.0 \%$ |
| $50 \%$ | $64.7 \%$ | $64.4 \%$ | $64.5 \%$ | $64.5 \%$ |
| $70 \%$ | $81.4 \%$ | $81.0 \%$ | $82.1 \%$ | $81.5 \%$ |
| $80 \%$ | $87.8 \%$ | $87.6 \%$ | $88.5 \%$ | $88.0 \%$ |
| $90 \%$ | $93.4 \%$ | $93.5 \%$ | $93.7 \%$ | $93.5 \%$ |
| $95 \%$ | $96.0 \%$ | $96.2 \%$ | $96.3 \%$ | $96.2 \%$ |
| $99 \%$ | $98.3 \%$ | $98.5 \%$ | $98.4 \%$ | $98.4 \%$ |

The best estimate reserve (and the difference between the best estimate and true reserve) is the same as shown in the previous Table B-10: this is because the ODP method and Mack's method give identical best estimates when fitted to triangles with no missing values (both give the BCL estimates).
The negative mean value of the 'standardised predictive error' ( -0.220 ) implies (given the positive bias) that when the BCL understates the true outcome, the Mack standard error also tends to be understated. The same occurred when Mack's method was run on triangles that satisfy its assumptions.
Despite the positive bias of the BCL estimates (or perhaps because of this bias), Mack's standard errors appear to give a better assessment of overall predictive error here than in our simulations on triangles that satisfy Mack's assumptions: the mean square of the 'standardised predictive error' is not much greater (at 1.10 over all 30,000 simulations) than the ideal value of one.
The distribution of the quantity F (true reserve) (where $\mathrm{F}(\mathrm{x})$ is the Log-Normal predictive distribution) shows that the method slightly understates the chance of very adverse outcomes. For example, in $1.4 \%$ of simulations the true outcome exceeded the $99^{\text {th }}$ percentile as given by Mack's method. Clearly the performance of Mack's method here is much better than on triangles constructed (using Algorithm A) to perfectly satisfy Mack's assumptions!

## B.2.6 Performance of ODP method when its assumptions are false

For completeness, we have also tested the ODP methods (analytic and bootstrap) on data produced using Algorithm A (satisfying the assumptions of Mack's method). It is easily proven that data generated in this way do not satisfy the assumptions of the ODP method (see the last section of this appendix). Results (based on the same 10,000 data-sets as used for the results in Table B-4) are given below. The results obtained using Mack's method are repeated here for comparison. As before, the results in the first column for the analytic ODP method are based on estimating the dispersion parameter using the Pearson residuals, and the second column using the deviance. For the bootstrap ODP results, the first column relates to the 1999 method, the second column to the 2001 method.
Table B-12 - Results of applying Mack and ODP methods to data generated using Algorithm A

|  | Mack | analytic ODP |  | bootstrap ODP |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| BCL estimate | 78.0 | 78.0 |  | 78.0 |  |
| BCL estimate > True reserve | $48.1 \%$ | $48.1 \%$ |  | $48.1 \%$ |  |
| (BCL estimate - True reserve) | 0.93 | 0.93 |  |  |  |
| Standard error | 29.52 | 30.8 | 31.02 | 26.81 | 27.07 |
| (BCL - True) / (Std error) | -0.51 | -0.40 | -0.40 | -0.41 | -0.41 |
| Square of the above | 4.2 | 2.9 | 2.8 | 3.5 | 3.5 |
| F(true liability) | 0.569 | 0.569 | 0.570 | 0.558 | 0.526 |
| $1 \%$ | $8.4 \%$ | $5.4 \%$ | $5.4 \%$ | $6.5 \%$ | $5.9 \%$ |
| $5 \%$ | $16.3 \%$ | $12.8 \%$ | $12.7 \%$ | $14.7 \%$ | $13.7 \%$ |


| $10 \%$ | $22.5 \%$ | $19.1 \%$ | $19.0 \%$ | $21.2 \%$ | $20.1 \%$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $20 \%$ | $32.6 \%$ | $30.7 \%$ | $30.6 \%$ | $32.0 \%$ | $29.9 \%$ |
| $30 \%$ | $41.2 \%$ | $40.5 \%$ | $40.5 \%$ | $40.9 \%$ | $38.1 \%$ |
| $50 \%$ | $57.8 \%$ | $59.1 \%$ | $59.1 \%$ | $57.4 \%$ | $52.9 \%$ |
| $70 \%$ | $73.9 \%$ | $75.7 \%$ | $75.9 \%$ | $72.7 \%$ | $68.3 \%$ |
| $80 \%$ | $81.9 \%$ | $83.6 \%$ | $84.0 \%$ | $79.9 \%$ | $75.6 \%$ |
| $90 \%$ | $90.0 \%$ | $91.1 \%$ | $91.4 \%$ | $87.8 \%$ | $84.5 \%$ |
| $95 \%$ | $93.8 \%$ | $94.5 \%$ | $94.8 \%$ | $92.0 \%$ | $89.8 \%$ |
| $99 \%$ | $97.8 \%$ | $97.8 \%$ | $97.9 \%$ | $96.3 \%$ | $95.4 \%$ |

Although these results indicate that the analytic ODP methods perform better than Mack's method when applied to triangles that satisfy Mack's assumptions (and the analytic ODP method performs better than the bootstrap ODP methods), the results also show that the ODP methods do not perform particularly well on this type of data. However, it is difficult to imagine any real-life claims settlement process that would approximate Algorithm A: we think it is of greater value to test the ODP method on data that violate the assumptions of the method in more realistic ways. For this purpose, we have generated data by Algorithm C below. This is just like algorithm B except that the Negative Binomial distribution is used at Step 1 instead of the Poisson. It is easily shown (see the last section of this appendix) that data generated in this way do not have independent increments ( $\mathrm{Y}_{\mathrm{jk}}$ ) so Assumption 3 of the ODP method is violated. (In fact, the $\mathrm{Y}_{\mathrm{jk}}$ values generated in this way are positively correlated within any origin year j ).

## Algorithm C for artificial run-off data:

1. The ultimate number of claims in an origin year is generated by random sampling from a Negative Binomial distribution (same parameters for each origin year, but independent sampling).
2. Each claim is assumed to be settled by a single payment, and the development year of the payment determined by independent random sampling from a Multinomial distribution (same parameters for each origin year: these are the parameters denoted $\mathrm{p}_{\mathrm{k}}$ in the above assumptions of the ODP method).
3. The amount of each individual claim payment is determined by independent random sampling from a Log-Normal distribution (same parameters in every cell of the triangle).
4. The amounts of claims settling in the upper left triangle of the run-off array are accumulated to create this run-off triangle, and all claim amounts (regardless of the development year when settled) are accumulated to obtain the 'true' ultimate position for each origin year.
The table below gives results for the analytic ODP method. As before, for each set of 10,000 simulations, the first column gives results based on estimating the dispersion parameter using the Pearson residuals, and the second column using the deviance (this has been done for the first set
of 10,000 simulated triangles only). For the bootstrap ODP results, the first column relates to the 1999 method, the second column to the 2001 method.

Table B-13 - Results of applying analytic ODP method to data generated using Algorithm C

|  | Results from different sets of 10,000 simulations |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  | Set 1 |  | Set 2 | Set 3 |  |
|  |  |  |  |  |  |
| Best estimate reserve | 3,612 | 3,600 | 3,574 |  |  |
| Best estimate > True | $45.5 \%$ | $45.8 \%$ | $44.8 \%$ |  |  |
| (Best estimate - True) | 244.1 | 229.0 | 211.8 |  |  |
| ODP standard error | 1,487 | 1,467 | 1,489 | 1,483 |  |
| (Best - True) / (std error) | -0.31 | -0.31 | -0.31 | -0.32 |  |
| Square of the above | 1.53 | 1.56 | 1.51 | 1.48 |  |
| Log-Normal F(True) | 0.579 | 0.579 | 0.580 | 0.584 |  |
| $1 \%$ | $2.6 \%$ | $2.8 \%$ | $2.8 \%$ | $2.5 \%$ |  |
| $5 \%$ | $10.2 \%$ | $10.3 \%$ | $9.8 \%$ | $9.7 \%$ |  |
| $10 \%$ | $17.7 \%$ | $17.9 \%$ | $17.3 \%$ | $17.7 \%$ |  |
| $20 \%$ | $30.8 \%$ | $31.0 \%$ | $30.7 \%$ | $30.7 \%$ |  |
| $30 \%$ | $42.4 \%$ | $42.3 \%$ | $42.4 \%$ | $42.8 \%$ |  |
| $50 \%$ | $61.8 \%$ | $61.6 \%$ | $61.6 \%$ | $62.3 \%$ |  |
| $70 \%$ | $77.6 \%$ | $77.4 \%$ | $78.2 \%$ | $78.3 \%$ |  |
| $80 \%$ | $84.9 \%$ | $84.6 \%$ | $85.0 \%$ | $85.4 \%$ |  |
| $90 \%$ | $91.5 \%$ | $91.4 \%$ | $91.9 \%$ | $92.0 \%$ |  |
| $95 \%$ | $94.9 \%$ | $94.8 \%$ | $95.0 \%$ | $95.3 \%$ |  |
| $99 \%$ | $98.1 \%$ | $98.1 \%$ | $98.2 \%$ | $98.3 \%$ |  |

Table B-14 below gives results for the bootstrap ODP method. As before, for each set of 10,000 simulations, the first column relates to the 1999 method, the second column to the 2001 method.

Table B-14 - Results of applying bootstrap ODP methods to data generated using Algorithm C

|  | Results from different sets of 10,000 simulations |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 |  |  | Set 3 |  |
| Best estimate reserve | 3,612 | 3,600 | 3,574 |  |  |  |
| Best estimate > True | $45.5 \%$ | $45.8 \%$ | $44.8 \%$ |  |  |  |
| (Best estimate - True) | 244.1 | 229.0 | 211.8 |  |  |  |
| Bootstrap mean | 3,675 | 3,676 | 3,668 | 3,668 |  |  |
| ODP standard error | 1,284 | 1,289 | 1,284 | 1,290 |  |  |
| (Best - True) / (std error) | -0.287 | -0.281 | -0.289 | -0.284 |  |  |
| Square of the above | 1.77 | 1.75 | 1.76 | 1.73 |  |  |
| F(True) | 0.568 | 0.535 | 0.549 | 0.535 |  |  |
| $1 \%$ | $3.2 \%$ | $2.8 \%$ | $3.3 \%$ | $2.9 \%$ |  |  |
| $5 \%$ | $11.3 \%$ | $8.8 \%$ | $11.1 \%$ | $8.7 \%$ |  |  |
| $10 \%$ | $18.8 \%$ | $14.7 \%$ | $18.8 \%$ | $14.7 \%$ |  |  |
| $20 \%$ | $31.9 \%$ | $25.6 \%$ | $31.8 \%$ | $25.8 \%$ |  |  |
| $30 \%$ | $42.7 \%$ | $36.2 \%$ | $42.8 \%$ | $36.1 \%$ |  |  |
| $50 \%$ | $60.4 \%$ | $55.2 \%$ | $60.1 \%$ | $55.4 \%$ |  |  |
| $70 \%$ | $74.8 \%$ | $72.6 \%$ | $75.1 \%$ | $72.4 \%$ |  |  |
| $80 \%$ | $81.5 \%$ | $80.9 \%$ | $81.7 \%$ | $80.7 \%$ |  |  |
| $90 \%$ | $88.6 \%$ | $89.4 \%$ | $88.5 \%$ | $89.4 \%$ |  |  |
| $95 \%$ | $92.2 \%$ | $93.5 \%$ | $92.3 \%$ | $93.7 \%$ |  |  |
| $99 \%$ | $96.3 \%$ | $97.6 \%$ | $96.5 \%$ | $97.7 \%$ |  |  |

These results show that the ODP methods performs nearly as well here as where the underlying assumptions are satisfied (ie when applied to data generated using Algorithm B). The results for the analytic ODP method again show slightly better performance when the dispersion parameter is estimated from Pearson residuals (first column) than when it is estimated from the deviance (second column). These results also show that the 1999 bootstrap ODP method performs worse than the analytic method, but the 2001 bootstrap method performs slightly better than the analytic method (which is the same ranking as when the ODP assumptions are satisfied). Clearly more testing is needed to see whether the performance of the analytic ODP method is generally insensitive to the ultimate claim number distribution.

## B.2.7 Further results for Mack's method when its assumptions are false

We have also tested Mack's method on the same data (generated using Algorithm C) and the results (below) again show that it performs surprisingly well given its poor performance on data that satisfies its assumptions precisely.

Table B-15 - Results of applying Mack's method to data generated using Algorithm C

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | ---: |
|  | Set 1 | Set 2 | Set 3 | All combined |
| BCL estimate | 3,612 | 3,600 | 3,574 | 3,595 |
| BCL estimate > True reserve | $45.5 \%$ | $45.8 \%$ | $44.8 \%$ | $45.4 \%$ |
| BCL estimate - True reserve | 244.1 | 229.0 | 211.8 | 228.3 |
| Mack standard error | 1,917 | 1,923 | 1,903 | 1,914 |
| (BCL - True) / (Mack std error) | -0.248 | -0.249 | -0.256 | -0.251 |
| Square of the above | 1.15 | 1.14 | 1.12 | 1.14 |
| Log-Normal F(true liability) | 0.593 | 0.594 | 0.597 | 0.595 |
| $1 \%$ | $1.7 \%$ | $1.7 \%$ | $1.5 \%$ | $1.6 \%$ |
| $5 \%$ | $6.8 \%$ | $7.1 \%$ | $6.8 \%$ | $6.9 \%$ |
| $10 \%$ | $13.8 \%$ | $13.8 \%$ | $13.9 \%$ | $13.8 \%$ |
| $20 \%$ | $27.9 \%$ | $28.1 \%$ | $28.0 \%$ | $28.0 \%$ |
| $30 \%$ | $42.4 \%$ | $42.4 \%$ | $43.1 \%$ | $42.6 \%$ |
| $50 \%$ | $65.8 \%$ | $65.4 \%$ | $65.8 \%$ | $65.7 \%$ |
| $70 \%$ | $82.0 \%$ | $82.2 \%$ | $82.4 \%$ | $82.2 \%$ |
| $80 \%$ | $88.5 \%$ | $88.4 \%$ | $88.9 \%$ | $88.6 \%$ |
| $90 \%$ | $93.7 \%$ | $93.9 \%$ | $94.0 \%$ | $93.9 \%$ |
| $95 \%$ | $96.4 \%$ | $96.4 \%$ | $96.6 \%$ | $96.5 \%$ |
| $99 \%$ | $98.5 \%$ | $98.6 \%$ | $98.6 \%$ | $98.6 \%$ |

## B. 3 Summary of results so far for Mack and ODP methods

Both Mack's method and the ODP methods can be applied to any triangle and give exactly the same reserve estimates as the basic chain ladder (BCL) method. However, the two sets of assumptions (those underlying Mack's method and those underlying the ODP methods) cannot both hold true. If Mack's assumptions hold true, then the BCL estimates are unbiased, but Mack's standard errors tend to understate the chance of extreme outcomes: substantially in the simulations we have carried out. If the ODP assumptions hold true, then the BCL estimates usually understate the true ultimates but when they overstate they do so on average by much
more than the average understatement, so that overall the BCL estimates are positively biased (by about $8 \%$ in simulations carried out so far). ODP standard errors (whether calculated analytically or by bootstrapping) understate the chances of extreme outcomes when the ODP assumptions are perfectly satisfied. Although the analytic and bootstrap methods of calculating ODP standard errors sometimes give very different results for the same triangle, there is little difference in the overall performance of the two methods in assessing reliability of BCL estimates when the ODP assumptions hold true.

Given its poor performance when its assumptions are true, Mack's method performs surprisingly well in the scenarios we have tested where its assumptions are false. Curiously, in the simulations carried out so far: Mack's method performs better than the ODP methods when applied to triangles that satisfy the ODP assumptions (Algorithm B), and the ODP methods perform better than Mack's method when applied to triangles that satisfy Mack's assumptions (Algorithm A).
Further results for Mack's method and the ODP methods applied where their assumptions are false are given later (by applying these methods to triangles that satisfy the assumptions of Wright's method).

## B. 4 Assessment of Wright's 1992 (operational time) method

## B.4.1 Assumptions of Wright's method

The original paper (Wright 1992) describes several related stochastic methods for use in various different circumstances. What they all have in common is that they are average-cost-per-claim methods in which the mean and variance of individual claim payments are assumed to be functions of 'operational time' (the same functions of operational time across all origin years). Operational time is defined as the proportion of the ultimate number of claims that have been closed, so it increases from zero at the start of an origin year to one when a year reaches full development.

Given a development triangle of aggregate claim amounts, and a corresponding triangle of the numbers of claims closed, the original paper describes how generalized linear modelling can be used to determine the relationship between mean claim amount and operational time, and how this can be used to forecast the total of future payments and its root-mean-square predictive error (allowing for both future process variation and parameter estimation uncertainty).

Some of the different circumstances covered by the original paper are:

- The individual payment amount distribution can be assumed to have the same coefficient of variation (ratio of standard deviation to mean) at all operational times, or this can be assumed to vary across operational times in a defined way.
- Claims inflation can be assumed to be absent (on the grounds that the claim amounts triangle has been pre-adjusted perhaps) or the average rate of claims inflation can be estimated as one of the parameters of the generalized linear model.
- The ultimate number of claims can be assumed known with certainty (as would be the case with reporting year cohorts or if there were no reporting delays in accident year cohorts) or uncertainty in ultimate numbers of claims can be taken into account.
- Each claim payment in the amounts triangle can assumed to be separately counted in the claim numbers triangle (as when every claim is settled with a single payment) or there may be partial payments in the amounts triangle that are not separately counted in the claim numbers triangle.
To date, we have tested the method by numerical simulation in only the simplest of these situations:
- The coefficient of variation of individual claim payments is the same at all operational times.
- There is no claims inflation.
- The ultimate number of claims is fully known for each origin year.
- All claims are settled by a single payment (so each payment contributing to the aggregate claim amounts triangle is counted in the triangle of numbers of claims closed).


## B.4.2 Artificial data satisfying Wright's assumptions (Algorithms D and E)

To test the performance of the method when these assumptions are true, we generated artificial run-off datasets as follows:

## Algorithm D for artificial run-off data:

1. The ultimate number of claims in an origin year is generated by random sampling from a Negative Binomial distribution (same parameters for each origin year but independent sampling).
2. Each claim is assumed to be settled by a single payment, and the development year of the payment determined by independent random sampling from a multinomial distribution (same parameters for each origin year).
3. The amount of each individual claim payment was determined by independent random sampling from a Log-Normal distribution. The mean of the Log-Normal distribution was determined, for each claim, from the operational time at which the claim settles. We used the formula: mean $=\exp \left(b_{0}+b_{1} . t+b_{2} \cdot t^{2}\right)$ where $t$ is operational time (for example, the $86^{\text {th }}$ claim to settle in an origin year in which the ultimate number of claim is 112 has $t=85.5 / 112$ ). For the parameters, we used the values $b_{0}=b_{1}=b_{2}=1.0$ : the mean amount of individual payments then rises from 2.718 at $t=0$ to 20.09 at $t=1$. The shape of the Log-Normal distribution was kept the same for all claims: the dependence of the mean on operational time was achieved by scaling the entire distribution. We used Log-Normal distributions all with coefficient of variation equal to 2.528 and skewness coefficient equal to 23.7.
4. The amounts of claims settling in the upper left triangle of the run-off array were accumulated to create this run-off triangle, and all claim amounts (regardless of the development year when settled) were accumulated to obtain the 'true' ultimate position for each origin year.

Note that it is only at Step 3 that this differs from Algorithm C (used in the previous section for testing the Mack and ODP methods). In Algorithm C the same probability distribution for individual payments is used at all stages of development: in Algorithm D this distribution is multiplied by a scaling factor that is a function of operational time. We used the same parameters in Steps 1 and 2 of Algorithm D as we used when generating data by Algorithm C to test the ODP and Mack methods.
Because Wright's method does not assume that the underlying run-off pattern (over real development time) is the same for all origin years, we have also generated data (by Algorithm E) in which the mean delay to settlement decreases across the origin years:

## Algorithm E for artificial run-off data:

Algorithm E is the same as Algorithm D except that the probabilities of the multinomial distribution used at Step 2 are not the same for all origin years. The probabilities we used for each origin year are shown below (one row for each origin year: each row adds to $100 \%$ ).

Table B-16 - Multinomial probabilities of claim settlement delay used in Algorithm E

|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| OYr 1 | $4.3 \%$ | $14.3 \%$ | $19.8 \%$ | $19.3 \%$ | $15.5 \%$ | $11.0 \%$ | $7.2 \%$ | $4.4 \%$ | $2.6 \%$ | $1.5 \%$ |
| 2 | $4.6 \%$ | $15.0 \%$ | $20.4 \%$ | $19.5 \%$ | $15.3 \%$ | $10.7 \%$ | $6.8 \%$ | $4.1 \%$ | $2.4 \%$ | $1.3 \%$ |
| 3 | $5.0 \%$ | $15.7 \%$ | $20.9 \%$ | $19.6 \%$ | $15.1 \%$ | $10.3 \%$ | $6.4 \%$ | $3.8 \%$ | $2.1 \%$ | $1.2 \%$ |
| 4 | $5.3 \%$ | $16.5 \%$ | $21.5 \%$ | $19.6 \%$ | $14.8 \%$ | $9.9 \%$ | $6.0 \%$ | $3.5 \%$ | $1.9 \%$ | $1.0 \%$ |
| 5 | $5.7 \%$ | $17.3 \%$ | $22.0 \%$ | $19.7 \%$ | $14.5 \%$ | $9.4 \%$ | $5.6 \%$ | $3.2 \%$ | $1.7 \%$ | $0.9 \%$ |
| 6 | $6.2 \%$ | $18.2 \%$ | $22.5 \%$ | $19.7 \%$ | $14.1 \%$ | $9.0 \%$ | $5.2 \%$ | $2.9 \%$ | $1.5 \%$ | $0.8 \%$ |
| 7 | $6.6 \%$ | $19.1 \%$ | $23.1 \%$ | $19.6 \%$ | $13.7 \%$ | $8.5 \%$ | $4.8 \%$ | $2.6 \%$ | $1.3 \%$ | $0.7 \%$ |
| 8 | $7.2 \%$ | $20.0 \%$ | $23.6 \%$ | $19.5 \%$ | $13.3 \%$ | $8.0 \%$ | $4.4 \%$ | $2.3 \%$ | $1.2 \%$ | $0.6 \%$ |
| 9 | $7.7 \%$ | $21.0 \%$ | $24.1 \%$ | $19.3 \%$ | $12.8 \%$ | $7.5 \%$ | $4.1 \%$ | $2.1 \%$ | $1.0 \%$ | $0.5 \%$ |
| 10 | $8.4 \%$ | $22.1 \%$ | $24.5 \%$ | $19.1 \%$ | $12.3 \%$ | $7.0 \%$ | $3.7 \%$ | $1.8 \%$ | $0.8 \%$ | $0.4 \%$ |

Assuming origin years are accident years, this implies that mean delay between loss event and claim settlement steadily reduces from about 3.4 years in the first origin year to 2.6 years in the latest origin year.

## B.4.3 Performance of Wright's method when its assumptions hold

For testing Wright's method, each dataset generated by Algorithm D or Algorithm E comprises:

- Ultimate number of claims in each origin year (from Step 1)
- Upper left triangle of numbers of claims closed (from Step 2)
- Upper left triangle of aggregate claim amounts (from Step 4)
- Ultimate aggregate claim amount for each origin year (from Step 4).

The first three of these items were treated as if known, and were used to estimate the ultimate aggregate amount (and its standard error) using Wright's operational time method. The last item is the 'true' ultimate aggregate amount which was compared to the estimate produced by Wright's method. We compared only the estimated and true totals for the entire triangle, not the results for individual origin years. As in Mack's method, we used a Log-Normal distribution (with mean equal to the best estimate reserve and standard deviation equal to the estimated predictive standard error) to determine the values F (true reserve) for comparison with a uniform distribution.
The original paper (Wright, 1992, pages 270 to 272) describes two methods for estimating the dispersion parameter (the squared coefficient of variation of individual payments, denoted $\varphi^{2}$ in the original paper). This can be estimated using the mean-squared Pearson residual or the meansquared deviance residual. As neither has a clear theoretical advantage, we have used both methods in these simulations to try to establish which performs better.
Results are summarised below for three independent sets of 10,000 simulations. For each set of 10,000 simulations, the first column gives results obtained using Pearson residuals to estimate the dispersion parameter; the second column gives results based on deviance residuals.

Table B-17 - Results of applying Wright's method to data from Algorithm $\mathbf{D}$ (constant run-off pattern)

|  | Results from different sets of 10,000 simulations |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Set 1 |  | Set 2 |  | Set 3 |  |
|  |  |  |  |  |  |  |
| Best estimate reserve | 3,434 |  | 3,436 |  | 3,440 |  |
| Best estimate > True | $49.6 \%$ |  | $49.2 \%$ |  | $49.5 \%$ |  |
| Best estimate - True | 37.4 |  | 21.7 |  | 39.4 |  |
| Wright standard error | 763.7 | 673.2 | 764.8 | 673.0 | 765.8 | 674.1 |
| (Best - True) / (std error) | -0.222 | -0.227 | -0.242 | -0.248 | -0.214 | -0.218 |
| Square of the above | 1.58 | 1.86 | 1.59 | 1.85 | 1.53 | 1.78 |
| Log-Normal F(True) | 0.540 | 0.533 | 0.545 | 0.538 | 0.540 | 0.533 |
| $1 \%$ | $4.1 \%$ | $5.0 \%$ | $3.7 \%$ | $5.1 \%$ | $3.7 \%$ | $4.5 \%$ |
| $5 \%$ | $10.3 \%$ | $11.7 \%$ | $9.4 \%$ | $12.2 \%$ | $9.9 \%$ | $11.3 \%$ |
| $10 \%$ | $15.8 \%$ | $17.5 \%$ | $14.6 \%$ | $18.3 \%$ | $15.8 \%$ | $17.7 \%$ |
| $20 \%$ | $25.9 \%$ | $27.7 \%$ | $24.8 \%$ | $28.6 \%$ | $26.1 \%$ | $27.7 \%$ |
| $30 \%$ | $35.7 \%$ | $36.7 \%$ | $33.8 \%$ | $37.6 \%$ | $35.7 \%$ | $36.6 \%$ |
| $50 \%$ | $54.4 \%$ | $53.5 \%$ | $52.1 \%$ | $54.0 \%$ | $54.7 \%$ | $53.8 \%$ |
| $70 \%$ | $73.2 \%$ | $70.6 \%$ | $71.4 \%$ | $70.8 \%$ | $73.1 \%$ | $70.5 \%$ |
| $80 \%$ | $82.4 \%$ | $79.3 \%$ | $81.1 \%$ | $79.6 \%$ | $82.6 \%$ | $79.2 \%$ |
| $90 \%$ | $91.5 \%$ | $88.6 \%$ | $90.9 \%$ | $89.0 \%$ | $91.8 \%$ | $88.8 \%$ |
| $95 \%$ | $96.3 \%$ | $94.2 \%$ | $95.7 \%$ | $94.3 \%$ | $95.9 \%$ | $94.0 \%$ |
| $99 \%$ | $99.3 \%$ | $98.5 \%$ | $99.1 \%$ | $98.7 \%$ | $99.3 \%$ | $98.6 \%$ |

These results suggest there may be a small positive bias in the best estimate of the reserve produced by Wright's method in this case: over all 30,000 simulations, the mean estimation error is about $0.96 \%$ of the mean true reserve. This is not very consistent across all three sets of 10,000 simulations (the mean error in Set 2 is about half the value in Set 1 or 3 ) so it might be caused by sampling error rather than being a bias. As well as being close to unbiased, estimates exceeded the true reserve in close to half (49.4\%) of simulations.

As with Mack's method and the ODP method, there is clear evidence that Wright's method tends to understate the predictive standard error when it understates the true reserve: the mean value of the 'standardised predictive error' is consistently negative (and its mean square significantly greater than one). The performance is slightly better when the dispersion parameter is estimated
from Pearson residuals (first column for each set 10,000 simulations) than when it is estimated from deviance residuals: Pearson residuals tend to give a higher estimate of the dispersion parameter, hence a higher predictive standard error (about 764 instead of 673 in this example) which more closely approximates the reliability of the best estimate reserve.
The tendency of the method to understate the predictive standard error when it understates reserves is also apparent in the results based on a Log-Normal predictive distribution: in $3.8 \%$ of the simulations the true reserve exceeded what was supposed to be the $99^{\text {th }}$ percentile (based on dispersion parameter estimated from Pearson residuals: this rises to $4.9 \%$ if deviance residuals are used).
One possible explanation of why Wright's method seems to understate the chance of adverse outcomes is that the method relies on asymptotic results for unbiasedness and standard errors of the parameter estimates of generalized linear models, and the volume of data may not be sufficiently large for the asymptotic results to work well. To investigate this further, we looked at the mean estimated values of the parameters $\left(b_{0}, b_{1}, b_{2}\right)$ across all simulations. It seems that the estimates are biased: the true values of these parameters were all exactly one, but the mean estimated values were $(0.975,1.113,0.873)$ in the first set of 10,000 simulations, $(0.977,1.118$, $0.864)$ in the second set and $(0.979,1.107,0.876)$ in the third: the differences between the estimated values and true values appear to be consistent and not just caused by sampling error.
The table below only shows results with the dispersion parameter estimated from Pearson residuals (because the earlier results indicate this is slightly better than using the deviance).

Table B-18 - Results of applying Wright's method to data from Algorithm E (varying run-off pattern)

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | All combined |
| Best estimate | 3,048 | 3,057 | 3,061 | 3,055 |
| Best estimate > True reserve | $50.4 \%$ | $49.6 \%$ | $50.2 \%$ | $50.1 \%$ |
| Best estimate - True reserve | 29.7 | 24.9 | 38.9 | 31.2 |
| Wright standard error | 697.5 | 700.7 | 704.7 | 701 |
| (Best - True) / (Std error) | -0.213 | -0.222 | -0.203 | -0.213 |
| Square of the above | 1.58 | 1.57 | 1.50 | 1.55 |
| Log-Normal F(true liability) | 0.539 | 0.541 | 0.537 | 0.539 |
| $1 \%$ | $4.0 \%$ | $4.0 \%$ | $3.9 \%$ | $4.0 \%$ |
| $5 \%$ | $9.9 \%$ | $10.2 \%$ | $9.6 \%$ | $9.9 \%$ |
| $10 \%$ | $15.7 \%$ | $16.5 \%$ | $15.3 \%$ | $15.8 \%$ |
| $20 \%$ | $25.7 \%$ | $26.8 \%$ | $25.8 \%$ | $26.1 \%$ |
| $30 \%$ | $35.2 \%$ | $36.2 \%$ | $35.2 \%$ | $35.5 \%$ |


| $50 \%$ | $53.9 \%$ | $54.3 \%$ | $54.0 \%$ | $54.1 \%$ |
| :--- | :--- | :--- | :--- | :--- |
| $70 \%$ | $73.6 \%$ | $72.8 \%$ | $72.9 \%$ | $73.1 \%$ |
| $80 \%$ | $82.7 \%$ | $81.9 \%$ | $82.2 \%$ | $82.3 \%$ |
| $90 \%$ | $92.0 \%$ | $91.4 \%$ | $91.8 \%$ | $91.7 \%$ |
| $95 \%$ | $96.1 \%$ | $96.0 \%$ | $96.1 \%$ | $96.1 \%$ |
| $99 \%$ | $99.1 \%$ | $99.3 \%$ | $99.2 \%$ | $99.2 \%$ |

These results are very similar to those based on data from Algorithm D, and the same comments apply. This confirms that Wright's method is not sensitive to whether or not the underlying runoff pattern (across real development time) is the same for all origin years.

Further work is needed to investigate:

- The extent to which these conclusions depend on the particular set of parameters chosen in simulating the artificial data by Algorithms D and E .
- How more general versions of Wright's method perform in situations where the underlying assumptions are perfectly satisfied (for example, in situations of non-zero claims inflation, or where not all claims are settled by a single payment).
- The performance of the method if applied where its underlying assumptions are not true.


## B.4.4 Performance of Mack's method on data from Algorithms D and E

As a further test of Mack's method in situations where its underlying assumptions are false, we applied Mack's method to the same claim amounts triangles (generated using Algorithms D and E) as using in testing Wright's method when its assumption hold true. Results are summarised below:

Table B-19 - Results of applying Mack's method to data generated using Algorithm D

|  | Results from different sets of 10,000 simulations |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
|  | Set 1 | Set 2 | Set 3 | All combined |
| BCL estimate | 3,654 | 3,659 | 3,666 | 3,660 |
| BCL estimate > True reserve | $47.7 \%$ | $46.4 \%$ | $47.6 \%$ | $47.2 \%$ |
| BCL estimate - True reserve | 257 | 245 | 265 | 256 |
| Mack standard error | 1,856 | 1,855 | 1,865 | 1,859 |
| (BCL - True) / (Mack std error) | -0.227 | -0.241 | -0.231 | -0.233 |
| Square of the above | 1.17 | 1.18 | 1.15 | 1.17 |
| Log-Normal F(true liability) | 0.580 | 0.586 | 0.581 | 0.582 |
| $1 \%$ | $1.6 \%$ | $1.8 \%$ | $1.7 \%$ | $1.7 \%$ |


| $5 \%$ | $6.8 \%$ | $7.1 \%$ | $7.3 \%$ | $7.1 \%$ |
| :--- | :--- | :--- | :--- | :--- |
| $10 \%$ | $13.8 \%$ | $13.5 \%$ | $13.7 \%$ | $13.7 \%$ |
| $20 \%$ | $27.5 \%$ | $27.2 \%$ | $27.0 \%$ | $27.2 \%$ |
| $30 \%$ | $40.3 \%$ | $41.0 \%$ | $40.1 \%$ | $40.5 \%$ |
| $50 \%$ | $62.4 \%$ | $63.8 \%$ | $62.9 \%$ | $63.0 \%$ |
| $70 \%$ | $80.5 \%$ | $81.5 \%$ | $80.8 \%$ | $80.9 \%$ |
| $80 \%$ | $87.7 \%$ | $88.1 \%$ | $88.2 \%$ | $88.0 \%$ |
| $90 \%$ | $93.5 \%$ | $93.7 \%$ | $93.9 \%$ | $93.7 \%$ |
| $95 \%$ | $96.4 \%$ | $96.3 \%$ | $96.5 \%$ | $96.4 \%$ |
| $99 \%$ | $98.7 \%$ | $98.6 \%$ | $98.7 \%$ | $98.7 \%$ |

These results are very similar to those obtained by applying Mack's method to data generated using Algorithm C. Compared to the results of applying Wright's method to data generated using Algorithm D, we see that Mack's method gives much higher predictive standard errors (mean of around 1,860 for Mack, compared to around 765 for Wright). This is to be expected because Mack's method makes use of the claim amounts triangle only, whereas Wright's method makes use of the claim amounts and claim numbers triangles. Clearly, in situations where claim numbers data are available, more reliable reserve estimates will usually be obtained by making use of this information rather than ignoring it.
The following results show that Mack's method (and the BCL method) breaks down badly if applied where the underlying run-off pattern is not the same for all origin years. The data simulated using Algorithm E has an accelerating development pattern. In this case, the BCL estimates exceed the true outcome in $84.5 \%$ of simulations, and on average overestimate the true reserves by $61 \%$. In only $0.1 \%$ of simulations does the true ultimate exceed what Mack's method indicates is the $99^{\text {th }}$ percentile. None of this is at all surprising - a key assumption of all chainladder methods (including Mack's) is that the underlying run-off pattern is the same for all origin years. These results confirm that such methods should not be used where this is not the case.

Table B-20 - Results of applying Mack's method to data generated using Algorithm E

|  | Results from 10,000 simulations <br> (Set 1) |
| :--- | :--- |
| BCL estimate | 4,873 |
| BCL estimate > True reserve | $84.5 \%$ |
| BCL estimate - True reserve | 1,854 |
| Mack standard error | 2,257 |
| (BCL - True) / (Mack std error) | 0.665 |


| Square of the above | 0.985 |
| :--- | :--- |
| Log-Normal F(true liability) | 0.295 |
| $1 \%$ | $0.1 \%$ |
| $5 \%$ | $0.5 \%$ |
| $10 \%$ | $1.4 \%$ |
| $20 \%$ | $4.3 \%$ |
| $30 \%$ | $8.8 \%$ |
| $50 \%$ | $22.3 \%$ |
| $70 \%$ | $42.1 \%$ |
| $80 \%$ | $54.9 \%$ |
| $90 \%$ | $70.9 \%$ |
| $95 \%$ | $80.6 \%$ |
| $99 \%$ | $92.1 \%$ |

## B.4.5 Performance of ODP methods on data generated using Algorithms D and E

As a further test of the ODP methods (analytic and bootstrap) in situations where the underlying assumptions are false, we have applied these methods to the same claim amounts triangles (generated using Algorithms D and E) as used in testing Wright's method when its assumptions hold. We have done this for only the first set of 10,000 triangles produced by each algorithm. Results are summarised below. For the analytic ODP method, since earlier results show that estimating the dispersion parameter from Pearson residuals is generally better than estimating it from the deviance, results are given below for the Pearson residual method only.

Table B-21 - Results of applying ODP methods to data generated using Algorithm D

|  | Analytic ODP | Bootstrap 99 | Bootstrap 01 |
| :--- | :--- | :--- | :--- |
| BCL estimate | 3,654 | 3,654 | 3,654 |
| BCL estimate > True reserve | $47.7 \%$ | $47.7 \%$ | $47.7 \%$ |
| BCL estimate - True reserve | 257 | 257 | 257 |
| Bootstrap mean |  | 3,723 | 3,724 |
| ODP standard error | 1,723 | 1,259 | 1,261 |
| (BCL - True) / (Std error) | -0.278 | -0.270 | -0.270 |
| Square of the above | 1.38 | 1.95 | 1.93 |
| F(true liability) | 0.578 | 0.553 | 0.522 |
| $1 \%$ | $2.3 \%$ | $3.6 \%$ | $2.8 \%$ |
| $5 \%$ | $8.5 \%$ | $11.8 \%$ | $9.2 \%$ |
| $10 \%$ | $16.0 \%$ | $19.7 \%$ | $15.9 \%$ |
| $20 \%$ | $28.6 \%$ | $31.0 \%$ | $26.4 \%$ |
| $30 \%$ | $40.1 \%$ | $41.0 \%$ | $35.9 \%$ |
| $50 \%$ | $61.2 \%$ | $57.6 \%$ | $52.9 \%$ |
| $70 \%$ | $79.7 \%$ | $72.4 \%$ | $69.3 \%$ |
| $80 \%$ | $87.0 \%$ | $79.7 \%$ | $78.4 \%$ |
| $90 \%$ | $93.6 \%$ | $87.3 \%$ | $87.3 \%$ |
| $95 \%$ | $96.8 \%$ | $91.2 \%$ | $92.2 \%$ |
| $99 \%$ | $99.0 \%$ | $96.1 \%$ | $96.9 \%$ |

These results show that, while it does not perform as well as Mack's method on triangles generated by Algorithm D, the analytic ODP method does perform almost as well here as where its assumptions are perfectly satisfied (Algorithm B, Table B-8). The parameters we have used in Algorithm D produce individual claim amounts with a mean value that increases by a factor of about seven between development periods 1 and 10 , which implies that the parameter $\varphi$ (defined as ratio of variance to mean of aggregate increments) increases by a similar factor. One of the ODP assumptions is that the value of $\varphi$ is constant across the entire triangle. This assumption is far from true here, so it seems that the analytic ODP method is quite robust to violations of this assumption. The bootstrap methods appear to be less robust to violations of this assumption.
In tests of the ODP methods where their assumptions are true (data generated using Algorithm B) we found that the 1999 bootstrap method gave a predictive standard error smaller than the analytic ODP standard error in $86 \%$ of simulations. Here (date generated using Algorithm D), the

1999 bootstrap method gave a smaller standard error than the analytic ODP standard error in $96.2 \%$ of simulations, and the 2001 bootstrap standard error was smaller than the analytic standard error in $95.7 \%$ of simulations. It seems that the bootstrap methods tend to understate the predictive standard error when data are generated by Algorithm D.
Another of the ODP assumptions is that the underlying run-off pattern across development periods is the same for all origin years: this assumption is satisfied in triangles generated by Algorithm D. However, the following results show that the analytic ODP method (like Mack's method) are not robust to violations of this assumption: it performs very poorly when applied to data generated by Algorithm E.

Table B-22 - Results of applying analytic ODP method to data generated using Algorithm E

|  | Results from 10,000 simulations <br> (Set 1) |
| :--- | :--- |
| BCL estimate | 4,873 |
| BCL estimate > True reserve | $84.5 \%$ |
| BCL estimate - True reserve | 1,854 |
| ODP standard error | 2,154 |
| (BCL - True) / (ODP std error) | 0.663 |
| Square of the above | 0.990 |
| Log-Normal F(true liability) | 0.287 |
| $1 \%$ | $0.1 \%$ |
| $5 \%$ | $0.7 \%$ |
| $10 \%$ | $1.8 \%$ |
| $20 \%$ | $4.6 \%$ |
| $30 \%$ | $8.8 \%$ |
| $50 \%$ | $21.0 \%$ |
| $70 \%$ | $39.4 \%$ |
| $80 \%$ | $52.7 \%$ |
| $90 \%$ | $69.2 \%$ |
| $95 \%$ | $80.4 \%$ |
| $99 \%$ | $93.0 \%$ |

## B. 5 Provisional conclusions

Much more work is needed before we can formulate definitive conclusions, but based on the simulations carried out so far it seems (provisionally) that:

- When the basic chain ladder method is applied to triangles where the underlying run-off pattern is the same for all origin years, it understates the true reserve more often than not. However, the mean value of BCL reserves exceeds the mean value of true outcomes: in other words, the BCL is positively biased. These results apply when the data are generated using a range of simplified but reasonably realistic algorithms (subject to the constraint that the underlying run-off pattern is the same for all origin years).
- When calculated from triangles where the underlying run-off pattern is the same for all origin years (this being an example of when the conditions are NOT satisfied), Mack's standard errors, used in conjunction with positively biased BCL estimates, provide reasonably accurate assessments of the chances of very adverse outcomes. This applies when the data are generated using a range of increasingly realistic algorithms that do not satisfy Mack's explicit assumption (but do have the same underlying run-off pattern across all origin years).
- Strangely however, Mack's method performs very poorly on data generated in such a way that Mack's explicit assumptions are perfectly satisfied. (This is perhaps because the BCL estimates are not biased in this case, but the Mack standard errors tend to be understated.) As it is difficult to imagine any real-world claim settlement process that would approximate Mack's assumptions, this may not be a major drawback of the method.
- Given the above, there seems to be little value in trying to check (through diagnostics) that Mack's explicit assumptions hold true when applying Mack's method (or perhaps the assumptions should be checked and the method used only where the assumptions appear to be false!) It is however important to check that the run-off pattern does not differ significantly between origin years.
- Standard errors for BCL estimates calculated using the over-dispersed Poisson (ODP) method can understate the chance of adverse outcomes even where the explicit assumptions of the ODP method are perfectly satisfied: in simulations carried out so far, what was supposed to be the $95^{\text {th }}$ percentile was exceeded in nearly $10 \%$ of simulations, and what was supposed to be the $99^{\text {th }}$ percentile was exceeded in $2.6 \%$ of simulations.
- The performance of the analytic OPD method is insensitive to the probability distribution of the ultimate number of claims in an origin year (even if this is such as to violate the assumptions of the method) and insensitive to the relationship between mean size of claim payment and delay to settlement (provided this is the same for all origin years).
- The bootstrap ODP method does not in general give the same predictive standard error as the analytic ODP method, even where the ODP assumptions are perfectly satisfied.
- The bootstrap ODP method is more sensitive than the analytic ODP method to violations in the ODP assumptions: its performance deteriorates if there is a strong relationship between the mean claim size and the delay to settlement (such that the ratio of variance to mean is not the same for all aggregate increments in the paid claims triangle).
- When applied to triangles for which the underlying run-off pattern is not the same for all origin years, the BCL method can give very poor estimates, and both Mack's method and the ODP methods give very unreliable standard errors. (Neither method was ever intended to be used in these circumstances, of course.)
- Wright's operational time method also tends to understate the chances of adverse outcomes: in simulations carried out so far, what was supposed to be the $95^{\text {th }}$ percentile was exceeded in about $10 \%$ of simulations, and what was supposed to be the $99^{\text {th }}$ percentile was exceeded in $3.9 \%$ of simulations.
- Provided the assumptions of the method continue to hold, the performance of Wright's method is not sensitive to variation in the underlying run-off pattern between origin years.
- We have not yet done any testing of the sensitivity of Wright's method to violations of its assumptions.


## B. 6 Technical results for Mack and ODP methods

## B.6.1 Mutually exclusivity of assumptions

Here we show that the assumptions of Mack's (1993) method and the over-dispersed Poisson (ODP) method are mutually exclusive, that is, the two sets of assumptions cannot both be true simultaneously. This explains how it is that - when applied to triangles with no missing data both methods produce 'best estimates' identical to those given by the basic chain ladder (BCL) method, yet the two methods give different root-mean-square predictive errors (or 'standard errors' for short). Since the two methods rely on two mutually exclusive sets of assumptions, both have the potential to give correct standard errors in different circumstances, but they can never both be correct.
Mack's assumptions are:
M1: There exist parameters $f_{k}$ such that: $E\left(C_{j, k+1} \mid C_{j 1}, \ldots C_{j k}\right)=f_{k} \cdot C_{j k}$
M2: There exist parameters $\alpha_{\mathrm{k}}$ such that: $\operatorname{Var}\left(\mathrm{C}_{\mathrm{j}, \mathrm{k}+1} \mid \mathrm{C}_{\mathrm{j} 1}, \ldots \mathrm{C}_{\mathrm{jk}}\right)=\alpha_{\mathrm{k}}{ }^{2} \cdot \mathrm{C}_{\mathrm{jk}}$
M3: Accident years are stochastically independent.
The ODP assumptions are:

$$
\begin{aligned}
& \mathrm{P} 1: \mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}}\right)=\mathrm{x}_{\mathrm{j} \cdot} \cdot \mathrm{p}_{\mathrm{k}} \\
& \mathrm{P} 2: \operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}}\right)=\varphi \cdot \mathrm{x}_{\mathrm{j}} \cdot \mathrm{p}_{\mathrm{k}}
\end{aligned}
$$

P3: The $\mathrm{Y}_{\mathrm{jk}}$ are stochastically independent (across all j and all k ).

First we show that if P1 and P3 hold, then M1 necessarily does not hold.
From the definitions of $\mathrm{C}_{\mathrm{jk}}$ and $\mathrm{Y}_{\mathrm{jk}}$ as respectively cumulative and incremental aggregate paid amounts:

$$
\mathrm{E}\left(\mathrm{C}_{\mathrm{j}, \mathrm{k}} \mid \mathrm{C}_{\mathrm{j}, \mathrm{k}-1}\right)=\mathrm{C}_{\mathrm{j}, \mathrm{k}-1}+\mathrm{E}\left(\mathrm{Y}_{\mathrm{j}, \mathrm{k}} \mid \mathrm{C}_{\mathrm{j}, \mathrm{k}-1}\right)
$$

$$
\begin{aligned}
& =C_{j, k-1}+E\left(Y_{j, k}\right) \text { by assumption P3 (independent increments) } \\
& =C_{j, k-1}+x_{j} \cdot p_{k} \text { by assumption P1 } \\
& =\left(1+x_{j} \cdot p_{k} / C_{j, k-1}\right) \cdot C_{j, k-1}
\end{aligned}
$$

Comparing this to assumption M1 we see that assumption M1 cannot hold because the quantity in parentheses in the last equation is a function of the random variable $\mathrm{C}_{\mathrm{j}, \mathrm{k}-1}$ so cannot be a constant $f_{k}$ (the same for all origin years).

Next we show that if M1 holds, then P3 necessarily does not.
Assumption M 1 implies that $\mathrm{E}\left(\mathrm{Y}_{\mathrm{j}, \mathrm{k}+1} \mid \mathrm{C}_{\mathrm{j} 1} \ldots \mathrm{C}_{\mathrm{j}, \mathrm{k}}\right)=\left(\mathrm{f}_{\mathrm{k}}-1\right) . \mathrm{C}_{\mathrm{jk}}$
Putting $\mathrm{k}=1$, this implies: $\mathrm{E}\left(\mathrm{Y}_{\mathrm{j} 2} \mid \mathrm{Y}_{\mathrm{j} 1}\right)=\left(\mathrm{f}_{1}-1\right) . \mathrm{Y}_{\mathrm{j} 1}$
which implies that $\mathrm{Y}_{\mathrm{j} 2}$ and $\mathrm{Y}_{\mathrm{j} 1}$ are not stochastically independent, contradicting P3.

## B.6.2 Algorithm $B$ satisfies the ODP assumptions

We next show that Algorithm B satisfies the ODP assumptions (and therefore, by the above results, does not satisfy Mack's assumptions).
Algorithm B is:

1. Generate the ultimate number of claims $\mathrm{N}_{\mathrm{j}}$ from a Poisson distribution with parameter $\lambda$ say.
2. Find numbers $\mathrm{N}_{\mathrm{jk}}$ settled in each development year k by splitting $\mathrm{N}_{\mathrm{j}}$ using a multinomial distribution, with probabilities $\mathrm{p}_{\mathrm{k}}$, where sum is 1 .
3. Generate the amount of each claim $Z_{j k i}\left(i=1\right.$ to $\left.N_{j k}\right)$ by independent sampling from a distribution with mean $\mu$ and variance $\sigma^{2}$ say.
4. Calculate $Y_{j k}=\Sigma_{i} Z_{j k i}$

First we show that Steps 1 and 2 produce independent Poisson numbers $\mathrm{N}_{\mathrm{jk}}$. We consider just two development years (the argument easily generalises to more than two). The joint probability density function of $\mathrm{N}_{\mathrm{j} 1}$ and $\mathrm{N}_{\mathrm{j} 2}$ is given by:

$$
\begin{aligned}
\mathrm{f}\left(\mathrm{~N}_{\mathrm{j} 1}, \mathrm{~N}_{\mathrm{j} 2}\right) & =\mathrm{f}\left(\mathrm{~N}_{\mathrm{j} 1}, \mathrm{~N}_{\mathrm{j} 2} \mid \mathrm{N}_{\mathrm{j}}\right) \cdot \mathrm{f}\left(\mathrm{~N}_{\mathrm{j}}\right) \\
& =\text { multinomial } * \text { Poisson } \\
& =\text { product of two Poisson pdfs for } \mathrm{N}_{\mathrm{j} 1} \text { and } \mathrm{N}_{\mathrm{j} 2} .
\end{aligned}
$$

Since their joint pdf factorises, $\mathrm{N}_{\mathrm{j} 1}$ and $\mathrm{N}_{\mathrm{j} 2}$ are stochastically independent.
Clearly $\mathrm{Y}_{\mathrm{j} 1}$ and $\mathrm{Y}_{\mathrm{j} 2}$ generated by Steps 3 and 4 of Algorithm B are therefore also independent, so assumption P3 is satisfied.

To show that P 1 and P 2 are also satisfied we must consider $\mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}}\right)$ and $\operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}}\right)$ :

$$
\mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}}\right) \quad=\mathrm{E}\left(\mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}} \mid \mathrm{N}_{\mathrm{jk}}\right)\right)
$$

$$
=\mathrm{E}\left(\mu \cdot \mathrm{~N}_{\mathrm{jk}}\right) \quad \text { from Step } 4 \text { of Algorithm } \mathrm{B}
$$

$$
=\mu \cdot \lambda \cdot \mathrm{p}_{\mathrm{k}} \quad \text { because } \mathrm{N}_{\mathrm{jk}} \text { is Poisson }\left(\lambda \cdot \mathrm{p}_{\mathrm{k}}\right)
$$

from which we see that P1 holds, with $\mathrm{x}_{\mathrm{j}}=\mu . \lambda$ (ie expected ultimate is expected number of claims $\lambda$ multiplied by expected amount $\mu$ of each one).

$$
\begin{aligned}
\operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}}\right) & =\mathrm{E}\left(\operatorname{Var}\left(\mathrm{Y}_{\mathrm{jk}} \mid \mathrm{N}_{\mathrm{jk}}\right)\right)+\operatorname{Var}\left(\mathrm{E}\left(\mathrm{Y}_{\mathrm{jk}} \mid \mathrm{N}_{\mathrm{jk}}\right)\right) & & \\
& =\mathrm{E}\left(\sigma^{2} \cdot \mathrm{~N}_{\mathrm{jk}}\right)+\operatorname{Var}\left(\mu \cdot \mathrm{N}_{\mathrm{jk}}\right) & & \text { from Step } 4 \text { of Algorithm B } \\
& =\sigma^{2} \cdot \lambda \cdot \mathrm{p}_{\mathrm{k}}+\mu^{2} \cdot \lambda \cdot \mathrm{p}_{\mathrm{k}} & & \text { because } \mathrm{N}_{\mathrm{jk}} \text { is Poisson }\left(\lambda \cdot \mathrm{p}_{\mathrm{k}}\right) \\
& =\left(\sigma^{2}+\mu^{2}\right) \cdot \lambda \cdot \mathrm{p}_{\mathrm{k}} & &
\end{aligned}
$$

from which we see that P 2 holds with $\varphi=\left(\sigma^{2}+\mu^{2}\right) / \mu$.
If algorithm $B$ were generalised to have a different ultimate expected number of claims $\lambda_{j}$ for each origin year, the above argument is unaffected and the ODP assumptions continue to hold (with $\mathrm{x}_{\mathrm{j}}=\mu . \lambda_{\mathrm{j}}$ ).

## B.6.3 Algorithm C does not satisfy ODP or Mack assumptions

Algorithm C is the same as Algorithm B except that the distribution used in Step 1 is Negative Binomial instead of Poisson. It is easily shown (by replacing the Poisson pdf with the Negative Binomial pdf in the above) that in this case, the join distribution $f\left(N_{j 1}, N_{j 2}\right)$ does not factorise, which implies that $\mathrm{N}_{\mathrm{j} 1}$ and $\mathrm{N}_{\mathrm{j} 2}$ are stochastically dependent, which in turn implies that assumption P3 does not hold. It is also easy to show that assumption M1 does not hold in this case.

Appendix C: Long List of Methods

## Appendix C

Long List of Methods

| No. | Category | Paper / Workshop / Reference name | Link |
| :---: | :---: | :---: | :---: |
| 1 | Mack | Mack, T. (1993). Distribution Free Calculation of the Standard Error of Chain Ladder Reserve Estimates. ASTIN Bulletin, 23, pp. 213-225. | $\frac{\text { http://www.casact.org/library/astin/vol2 }}{\text { 3no2/213.pdf }}$ |
| 2 | Mack | Mack, T. (1999). The Standard Error of Chain Ladder Reserve Estimates: Recursive Calculation and Inclusion of a Tail Factor. ASTIN Bulletin, 29, pp. 361-366. | $\frac{\text { http://www.casact.org/library/astin/vol2 }}{\text { 9no2/361.pdf }}$ |
| 3 |  | Over Dispersed Poisson, including GLM |  |
| 4 |  | Benchmarking |  |
| 5 |  | McLennan, A. \& Murphy, K. P. (2006). A Method For Projecting Individual Large Claims. Casualty Actuarial Society, Fall Forum 2006, pp. 205-236. | http://www.casact.org/pubs/forum/06ff orum/209.pdf |
| 6 | Individual Level Claim Reserving | Guszcza, J. C. \& Lommele, J. (2006). Loss Reserving Using Claim-Level Data. Casualty Actuarial Society, Fall Forum 2006, pp. 111140. | $\frac{\text { http://www.casact.org/pubs/forum/06ff }}{\text { orum/115.pdf }}$ |
| 7 |  | Patel, C. C., \& Raws, A. (1998). Statistical Modeling Techniques for Reserve Ranges: A Simulation Approach. Casualty Actuarial Society, Fall Forum 1998, pp. 229-255. | http://www.casact.org/pubs/forum/98ff orum/patel.pdf |
| 8 |  | Extrapolation - Curve Fitting for the Tail |  |
| 9 |  | Taylor (check CAS - Australian) |  |
| 10 |  | Halliwell, L. J. (1996). Loss Prediction by Generalized Least Squares. Proceedings of the Casualty Actuarial Society 83, pp. 436489. | $\frac{\text { http://www.casact.org/pubs/proceed/pr }}{\text { oceed96/96436.pdf }}$ |
| 11 |  | Deterministic - Curve Fit |  |
| 12 |  | Historical Ults - Restating |  |
| 13 |  | Stressing - Best Estimate and Scenario |  |
| 14 |  | McLennan - Fitting Curves to Loss Development Factors |  |
| 5 |  | Panning, W. H.(2006). Measuring Loss Reserve Uncertainty. Casualty Actuarial Society, Fall Forum 2006, pp. 237-267. | http://www.casact.org/pubs/forum/06ff orum/241.pdf |
| 16 |  | Scenarios Using Exposure Based Methods |  |
| 17 |  | Meyers, G. G. (2006). Estimating Predictive Distributions for Loss Reserve Models. Casualty Actuarial Society, Fall Forum 2006, pp. 159-203. | $\frac{\text { http://www.casact.org/pubs/forum/06ff }}{\text { orum/163.pdf }}$ |


| 18 |  | Verrall, R. J. \& England, P. D. (2005). Incorporating Expert Opinion into a Stochastic Model for the Chain-Ladder Technique. Insurance: Mathematics and Economics Vol. 37, pp. 355-370. |  |
| :---: | :---: | :---: | :---: |
| 19 |  | Verall/England - New Paper to be released Annals of Actuarial Science |  |
| 20 |  | Renshaw, A. E. \& Verrall, R. J. (1994). The Stochastic Model Underlying the ChainLadder Technique. Proceedings of the ASTIN Colloquium. |  |
| 21 |  | Murphy, D. M. (1994). Unbiased Loss Development Factors. Proceedings of the Casualty Actuarial Society 81, pp. 154-222. | http://www.casact.org/pubs/proceed/pr oceed94/94154.pdf |
| 22 | Bootstrap / Overdisperse d Poisson | England, P. D. \& Verrall, R. J. (2002). Stochastic Claims Reserving in General Insurance. British Actuarial Journal Vol. 8, pp. 443-544. |  |
| 23 | Operational Time | Wright, T. S. (1992). Stochastic Claims Reserving When Past Claim Numbers Are Known. Proceedings of the Casualty Actuarial Society 79, pp. 255-361. | http://www.casact.org/pubs/proceed/pr oceed92/92255.pdf |
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| 25 |  | Reid, D. H. (1997). Operational Time and a Fundamental Problem of Insurance in a DataRich Environment: Claims Reserving Manual Vol. 2. London: Institute of Actuaries. |  |
| 26 | Regression /Curve Fitting | Benjamin, S. \& Eagles, L. (1997). A Curve Fitting Method and a Regression Method. Claims Reserving Manual Vol. 2. London: Institute of Actuaries. |  |
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| :---: | :---: | :---: | :---: |
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# Appendix D: Bibliography 

Basic Chain Ladder

Paper/Link

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## Mack

## Paper/Link

Mack, T. (1993). Distribution free calculation of the standard error of chain ladder reserve estimates. ASTIN Bulletin, 23, 213-225. http://www.casact.org/library/astin/vol23no2/213.pdf

Mack, T. (1994). Which stochastic model is underlying the chain ladder method? Insurance: Mathematics and Economics, 15, 133-138.
Mack T. (1994). Measuring the variability of chain ladder reserve estimates. Casualty Actuarial Society, Spring Forum, 101-182. http://www.casact.org/pubs/forum/94spforum/94spf101.pdf

Mack, T. (1997). Measuring the variability of chain ladder reserve estimates. Claims Reserving Manual volume 2. London: Institute of Actuaries. http://www.actuaries.org.uk/files/pdf/library/crm2-D6.pdf

Mack, T. (1999). The standard error of chain ladder reserve estimates: Recursive calculation and inclusion of a tail factor. ASTIN Bulletin, 29 No. 2, 361-366. http://www.casact.org/library/astin/vol29no2/361.pdf

## Bootstrap/Overdispersed Poisson

## Paper/Link

Renshaw, A. E. \& Verral, R. J. (1998). A stochastic model underlying the chain ladder technique. British Actuarial Journal, 4, 903-923.
England, P. D. \& Verrall, R. J. (1999). Analytic and bootstrap estimates of prediction errors in claims reserving. Insurance: Mathematics and Economics, 25, 281-293.
England, P. D. (2001). Addendum to `Analytic and bootstrap estimates of prediction errors in claims reserving'. Actuarial Research Paper No. 138, Department of Actuarial Science and Statistics, City University, London, EC1V 0HB.
England, P. D. \& Verrall, R. J. (2002). Stochastic claims reserving in general insurance. British Actuarial Journal, 8, 443-544.
Online web lecture for stochastic claim reserving models including the over-dispersed Poisson:
https://talk.city.ac.uk/stochasticreserving

## Individual level claim reserving

## Paper/Link

Guszcza, J. C. \& Lommele, J. (2006). Loss reserving using claim level data. Casualty Actuarial Society, Fall Forum, 111-140
Taylor, G. C., McGuire, G. \& Sullivan, J. (2006). Individual claim loss reserving conditioned by case estimates. Taylor Fry Consulting Actuaries.
Wright T S (1997). "Probability distribution of outstanding claim payments from individual payments data." Claims Reserving Manual volume 2. London: Institute of Actuaries.
http://www.actuaries.org.uk/files/pdf/library/crm2-D7.pdf

## Operational time

## Paper/Link

Wright, T. S. (1992). Stochastic Claims Reserving When Past Claim Numbers Are Known. Casualty Actuarial Society. Proceedings. 150 \& 151, 255-361.
http://www.casact.org/pubs/proceed/proceed92/92255.pdf
Bain, D. (2003). A practical implementation of Wright's Operational Time Model. Society of Actuaries in Ireland.
Lowe, J. (1994). A practical guide to measuring reserve variability using: bootstrapping, operational time and a distribution-free approach. General Insurance Convention.

## Regression/Curve Fitting

## Paper/Link

Benjamin, S. \& Eagles, L. (1997). A curve fitting method and a regression method. Claims Reserving Manual volume 2. London: Institute of Actuaries.
http://www.actuaries.org.uk/files/pdf/library/crm2-D3.pdf

## Bayesian \& Borhuetter/Fergusson

## Paper/Link

England, P. D. \& Verrall, R. J. (2002). Stochastic claims reserving in general insurance. British Actuarial Journal, 8, 443-544.
England, P. D. (2002). A beginner's guide to Bayesian modelling. General Insurance Convention.
Verrall, R. J. (2001). A Bayesian generalised linear model for the Bornhuetter-Ferguson method of claims reserving. Actuarial Research Paper No. 139, Department of Actuarial Science and Statistics, City University, London.

## Appendix E: Detailed Survey Results

This appendix has two components: Appendix E1 and Appendix E2.

## Appendix E1-ROC Working Party Questionnaire

The following survey has been developed to assist the Reserving Oversight Committee (ROC) in the
formulation of its conclusions relating to the working parties 'Best estimates and estimating uncertainty' and 'Effectiveness of reserving methods'.

1) In order to classify responses, please identify the type of organisation you represent:

Insurance company
Reinsurance company
Lloyd's syndicate
Consultancy
1a) Where is your company based?
UK
Overseas
2) What are your key products/classes of business that you perform reserving for? (Please select a maximum of three)

Personal lines - property
Personal lines - motor
Commercial lines - property
Commercial lines - motor
Commercial lines - liability
Health insurance
Credit insurance
Marine/aviation/transport
Reinsurance - Property
Reinsurance - Motor
Reinsurance - Liability
Other - please state
3) For the classes above please check the appropriate box in relation to the following questions:
a) Method not used
b) Method partially used (for background)
c) Method is a key supplementary method
d) Key/principal method

3a) What are the methods and models you use to determine best estimate values?
(In respect of each of the lines of business selected above).
Name of method/model a)
Mack
Overdispersed Poisson stochastic chain ladder
Bootstrap
Transaction level/ individual claim
Operational time
Regression/Curve fitting
Bayesian/BF method
Basic Chain Ladder - paid
Basic Chain Ladder - incurred
Bornhuetter Fergusson
Average cost per claim
Judgement
Scenarios
Other - please state
3b What are the methods and models you use to determine uncertainty?
Name of model
Mack
Overdispersed Poisson stochastic chain ladder/
Bootstrap
Transaction level/ individual claim
Operational time
Regression/Curve fitting
Bayesian/BF method
Basic Chain Ladder - paid
Basic Chain Ladder - incurred
Bornhuetter Fergusson
Average cost per claim
Judgemen
Scenarios
Other - please state
4 When determining uncertainty what are your key requirements - please state in order of importance
Identifying variability around the best estimate
Identifying the complete distribution
Identifying the tail of the distribution (above 95th percentile).
5 When determining which model/method to use in assessing uncertainty, what are the key criteria you use (in order of importance)?
Ease of use
Quality of results
Communicability
Practical issues - relating to available data
Time/ resource constraints
Knowledge of actuarial staff
6 When modelling uncertainty do you use?
In-house models
A combination of in house models/ off the shelf packages
If "off the shelf" packages are used - please list them:

Quantitative using everyday English
Whole account scenarios with indication of likelihood
Whole account scenarios without indication of likelihood
Scenarios relating to parts of the reserves with indication of likelihood
Scenarios relating to parts of the reserves without indication of likelihood
Stress tests
Other (please specify)
Uncertainty only communicated qualitatively
Uncertainty not communicated
8 Does your communication of uncertainty cover? (select all that apply)
Process uncertainty
Parameter uncertainty
Model uncertainty
9a When calculating best estimates what grouping of data do you use by origin period?
Monthly
Quarterly
Ansua

9b When calculating best estimates what grouping of data do you use by development period?
Monthly
Quarterly
Half-yearly
Annual
10a When calculating uncertainty what grouping of data do you use by origin period?
Monthly
Quarterly
Half-yearly
Annual
10b When calculating uncertainty what grouping of data do you use by development period?
Monthly
Quarterly
Half-yearly
Annual
11 How frequently do you reserve each class of business mentioned in 2) above?
Monthly
Quarterly
Half-yearly
Annually
12 If you reserve monthly - do you carry out a full analysis or a reduced analysis each month?
Full analysis
Reduced analysis
13 If you reserve quarterly - do you carry out a full analysis or a reduced analysis each quarter?
Full analysis
Reduced analysis
Thank you for completing this survey. Your results will be kept anonymous and used to assist the ROC working party's work.
This questionnaire may be repeated annually to monitor the take-up of new methodology.


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Analysis of data using all data combined
Data has been combined for all classes

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Key

1) Most important
2) Least important






| $\begin{array}{\|l\|} \hline \text { All lines } \\ \hline \text { Methods used } \\ \hline \end{array}$ | Count of preferences |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | Total |
| Quantitative using percentiles | 21 | 8 | 3 | 2 | 0 | 1 | 0 | 0 | 1 | 36 |
| Quantitative using everyday English | 10 | 15 | 4 | 2 | 1 | 1 | 0 | 2 | 0 | 35 |
| Whole account scenarios with indication of likelihood | 1 | 2 | 6 | 4 | 2 | 4 | 3 | 1 | 0 | 23 |
| Whole account scenarios without indication of likelihood | 1 | 2 | 0 | 7 | 5 | 4 | 3 | 0 | 0 | 22 |
| Scenarios relating to parts of the reserves with indication o likelihood | 3 | 3 | 4 | 2 | 5 | 3 | 2 | 2 | 0 | 24 |
| Scenarios relating to parts of the reserves without indication of likelihood | 4 | 1 | 3 | 4 | 2 | 4 | 4 | 1 | 1 | 24 |
| Stress tests | 2 | 3 | 3 | 2 | 4 | 2 | 6 | 4 | 0 | 26 |
| Uncertainty only communicated qualitatively | 5 | 1 | 2 | 0 | 1 | 1 | 2 | 11 | 1 | 24 |
| Uncertainty not communicated | 3 | 1 | 0 | 0 | 0 | 1 | 0 | 0 | 17 | 22 |
| Other (please specify) | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| Complete distribution of outcomes with reserves for given scenarions and at various exceedence probabilities |  |  |  |  |  |  |  |  |  |  |
| This seems to be the best place to comment on the questions on this page. Q7 depends on the context. I will be looking at the best estimate when recommending a balance sheet but at the tail when looking at a transaction. Q8 and Q10 again depend on |  |  |  |  |  |  |  |  |  |  |



## Appendix F: Insureware Submission

This appendix contains the Insureware submission in the full form received by the working party. It is written by Insureware, not by the working party. The approaches are described in more detail than we have described the methods tested by the working groups, although we have referenced relevant papers in relation to those methods. Some of the material is beyond the scope of the working party.

This appendix has six components as follows:
F1. Uncertainty Working Group Test Template.doc
F2. Variability and Uncertainty.doc
F3. Employers Liability_UC_O110.xls
F4. ELRF study.xls
F5. EmployersLiabilityUC01100012.xls
F6. MarineCargoGrossvNet.xls

## Appendix F

# Best Estimate \& Uncertainty Working Group 

## Method Evaluation Template

## Name of Method:

We do not advise a method. We use modelling frameworks to identify the structure in the data. In actuarial parlance, we design a method that is appropriate for the data.

There are three modelling frameworks we typically use, Probabilistic Trend Family (PTF), Multiple Probabilistic Trend Family (MPTF), and the Extended Link Ratio Family (ELRF).

The PTF and ELRF frameworks are discussed in the paper 'Best Estimates for Reserves' (available on the CAS website at: www.casact.org/pubs/proceed/proceed00/00245.pdf). The paper has been on the CAS syllabus since January 2005.

## PTF Modelling Framework

In the PTF modelling framework an optimal model is identified, equivalently, built or designed that captures the variability (volatility) in the incremental loss development array. The variability is described using four components of interest. Namely, trends in the three directions: development period, accident period and calendar period, and the variability of the data about the trend structure. The (process) variability is an integral part of the model.

A PTF model is succinctly described by four graphs; three graphs describe the trend structure in the three directions, the fourth graph depicts the process variability. The identified model is tested to ensure that the model assumptions are consistent with the data; including validation testing (by removal of years). The triangle is regarded as a sample path from the fitted model. Thus data simulated from the model should not be distinguishable from the original data in respect of trend structure and volatility about the trend structure.

## MPTF Modelling Framework

The MPTF modelling framework is used to design (build) an optimal composite model for multiple incremental loss development arrays. The identified composite model captures (describes) the variability in each loss development array (a la PTF) and the relationships between them.

This has applications to modelling multiple lines of business, multiple segments, multiple layers and credibility modelling.

Relationships between lines of business, for example, involve two types of correlations; process correlation and parameter correlation.

## ELRF Modelling Framework

The ELRF modelling framework in the first instance formalizes average link ratios (Age-to-age development factors) as regression estimators through the origin. It is also extended to include intercepts and constant trends for each development period across the accident periods. The framework provides statistical tests of link ratio methods including Mack, Murphy, and many extensions thereof.

The ELRF regression framework provides the ability to determine statistically whether the selected link ratios quantify salient features of the data and whether they have any predictive power.

It is very important to recognize that apart from the many failings of link ratio methods the actuary has no control over assumptions going forward in respect of projections. Indeed, there are no simple descriptors of the volatility in the data including how much inflation has been captured by the method!

By contrast, a PTF or MPTF has simple descriptors of the volatility in the data and assumptions made in deriving reserve distributions by accident year, calendar year and total are explicit and can be controlled (chosen).

## Datasets used:

We have used the Employers' Liability and Marine Cargo data sets.
The Employers' Liability data is sub-divided into three data sets. As noted in Section 5 , no information is available on the nature of these data sets. We have used two of these datasets to demonstrate the application of the modelling frameworks.

We have also tested a composite model created by combining the two subclasses of Employers' Liability data already used.

Marine Cargo paid loss data was available gross and net of reinsurance.
We have only modelled a small number of datasets to illustrate the rich statistical ideas incorporated in the above mentioned modelling frameworks.

## General Comments on the Method

## Does the Method produce a complete distribution of Outcomes?

Yes. The PTF and MPTF identified models forecast distributions for every cell in the future and their correlations, conditional on an explicit set of easily interpretable assumptions. These distributions include both process variability and parameter
uncertainty. Distributions of aggregates across accident periods and calendar periods can also be obtained.

## Would the Method be acceptable to the Profession?

Yes.
What Uncertainty has been measured by the method?
PTF and MPTF models include both process and parameter uncertainty.
There is an important distinction between variability and uncertainty and the two should not be used interchangeably.
"Variability is a phenomenon in the physical world to be measured, analyzed and where appropriate explained. By contrast uncertainty is an aspect of knowledge."

Sir David Cox.
For further details see:
http://www.insureware.com/Library/Technical/VariabilityandUncertainty.doc

## Comment on difficulty of method to program \& run

It is extremely difficult to program (but theoretically possible) in a spreadsheet form, but is easy to run in ICRFS-Plus as all the programming has been pre-specified.
ICRFS-Plus is a point and click system with extremely fast algorithms.
How long did the tests take to program and complete.
Individual tests do not need to be programmed. It takes seconds for individual tests to be completed.

How difficult is it to apply judgement and / or amendments to the results within the program

It is not difficult. Judgement is made on the basis of accurate information about the volatility in the business.

## How do the Class by Class results compare to the Aggregated results.

We do not believe in aggregating data across classes (and usually not even segments). We design a composite model that captures the variability and trend structure in each class and the relationships between classes. For example, see the composite model: Appendix F5

However, if two segments exhibit the same trend structure then it might make sense to aggregate the data.

## Qualitative Results:

Qualitative results are provided for each dataset as follows.
Appendix F3: Employers Liability subclass "a".
Appendix F4: Employers Liability subclass "b".
Appendix F5: Employers Liability Composite Model.
Appendix F6: Marine Cargo Gross and Net.
The Employers' Liability data used was provided sub-divided between three subclasses. Two of those subclasses have been used in developing the results shown in this appendix. For the purpose of this appendix only these two sub-classes have been referred to as " a " and " b ". No information has been available regarding the nature of these Employers' liability sub-classes as explained in Section 5.

## Quantitative Results

All quantitative results are included in the same workbooks as the qualitative results. Please see corresponding Quantile and model diagnostic spreadsheets.

## Tests to establish if model is reasonable

Comment on validity of underlying model on data set.
All the models we have produced for the datasets on the accompanying spreadsheets are statistically valid. The model assumptions have been tested and met. In one case we simulated three triangles from the (PTF identified) model for the real data. It is almost impossible to distinguish the simulated triangles and the real data in respect of salient features. See: Appendix F3

## Comments on results by accident year

The forecast means and standard deviations of distributions by accident year have been included. They are based on explicit easily interpretable assumptions that can be related to the historical experience.

## Comments on Coefficient of Variations (COV) \& standard deviation measured by accident year

The coefficient of variations decrease as the accident year increases for all our PTF and MPTF based models as should often be expected of a good model. Similarly, the standard deviation of the reserves also increases. HOWEVER, this is not true for an Extended Link Ratio Family (ELRF) based models that also include average link ratios!

## Comments on correlations of standard error \& COV when aggregating classes together compared to class by class results.

As mentioned above, we do not aggregate the data. We design a composite model for the individual classes that captures the volatility in each class and their inter relationships (two types of correlations). For an example of a composite model: See: Appendix F5

## Comment on how stable results are when development factors or residuals are changed

A PTF or an MPTF based model captures the volatility in the data in respect of trend structure and volatility about the trend structure, called process variability. Assumptions in respect of future trend (and uncertainty thereof) and process variability are explicit and need to be argued for in respect of past volatility. To illustrate this very important point, suppose the calendar year trend has been relatively stable in the last seven years and its estimate is $10 \%+\_3 \%$. Typically when doing a projection of reserve distributions you would assume that the true trend is a random number from a normal distribution with mean .10 and std. dev. 0.03 . That is, all the reserve distributions by Calendar year, accident year and Total are based on this explicit assumption. Suppose we now change the assumption to $15 \%+\ldots 3 \%$, all the resulting distributions will be different.

Incidentally, if the trend has been relatively stable for the last seven years (as in this example) then removal of years (diagonals) should yield statistically stable reserve distributions beyond the last year and the volatility of the numbers in the cells removed should have been predicted years ago by the then-estimated model.
See: Appendix F3

## Tests for Goodness of Fit and Prediction Error.

Comment on tests of Standardised Residuals
Normality tests were applied to all residuals including during validation analysis where normality tests were applied to the residuals not used during model estimation.

## Comment on tests of residual patterns

In model identification various discriminatory statistical criteria are used to ensure parsimony. All remaining patterns in the residuals are random.

We also perform validation analysis and simulate triangles from fitted models to ensure that the real data is indistinguishable from the simulated data.

## Test the Amended Model Fit by Removing 2 diagonals

Removal of calendar years (diagonals) is called validation analysis, and is part of extracting information from the data. In order to observe stability in the reserve distributions the most recent trends must be stable. If you do have stability of most recent trends, then the most important test is whether the model years ago predicted

## Appendix F1.6

the volatility of the numbers in the years that have been removed. Results of these kinds of critical tests are done shown in Appendix F3

## Variability and Uncertainty

There is an important distinction between variability and uncertainty and the two should not be used interchangeably.
"Variability is a phenomenon in the physical world to be measured, analyzed and where appropriate explained. By contrast uncertainty is an aspect of knowledge."

Sir David Cox.
Uncertainty and variability are philosophically very different and it is common for them to be kept separate in risk analyses modeling.

Variability is the effect of chance and a function of the system. It is not reducible through either study or further measurement (may be reduced through changing the system).

Uncertainty is the assessor's lack of knowledge (level of ignorance) about the parameters that characterize the physical system that is being modeled. It is sometimes reducible through further measurement or study.

Uncertainly has also been called "fundamental uncertainly" or "degree of belief".

## Simple examples

Suppose a symmetric coin is tossed 100 times and X denotes the no. of heads.
The mean number of heads (the mean of X ) is 50 . The SD of X is 5 . The Binomial probability of each possible outcome of $\mathrm{X}(0,1,2, . .100)$ is known precisely. There is no uncertainty about the coin's variability.

A $100 \%$ confidence interval for the mean is [50, 50]. There is no uncertainty in the mean and indeed in any of the probabilities of the outcome X . The probability that $\mathrm{X}=50$ (the mean) is approximately 0.08 . A $95 \%$ prediction interval for the outcome X is $[40,60]$. This 95\% prediction interval cannot be shortened.

Suppose we do not know the true probability of a head, p , because the coin is mutilated. Suppose also that before the coin is tossed 100 times, it is tossed 10 times to get an estimate of the probability p . Suppose for the sake of argument 5 heads are observed. Now the estimate of .5 of the probability of a head (in one toss) is uncertain.

We can create a confidence interval (CI) for p , and also for the mean 100p of the number of heads in 100 tosses. The CI is an interval around the estimated mean, namely, $50=100 * .5$. The CI is not the same as a prediction interval for the outcome in 100 tosses.

A prediction interval includes process variability and is wider than a confidence interval. A confidence interval is for a parameter, which is a constant. A prediction interval is for a random variable. [The loss reserve is a random variable].

Now, a $95 \%$ prediction interval for the number of heads in 100 tosses will be wider than [40, 60], say, [35, 65]. This interval can only be reduced to at best [40, 60] by reducing the parameter uncertainty through more (prior) sampling. But you cannot make a $95 \%$ prediction interval narrower than [40, 60].

Consider another example with the same mean. A symmetric roulette wheel, numbered, $0,1,2,3, \ldots . ., 100$ that is turned only once, and let X be the random variable that represents the outcome. The mean of X is 50 , SD is 29 . There is no uncertainty about the variability in the outcome $X$. The probability that $X=50$ is $1 / 101$.

A $100 \%$ confidence interval for the mean is [50, 50] (just like the coin). A $95 \%$ prediction interval for the outcome X is [2,97], for example.

Each process (symmetric coin and symmetric roulette wheel) has the same mean, or if you like the same "best estimate". Which one requires more capital?

So "best estimate" is pretty useless and a "range of estimates", also pretty useless, notwithstanding the fact that how do you know if an estimate is "best"?

In general, only in the presence of a probabilistic framework can you assess "best". Indeed it is only in a probabilistic framework that you identify (build) a model that represents the variability in the data.

## Probabilistic Trend Family (PTF) of models described in the paper "Best Estimates for Reserves"

In the PTF modeling framework a model is identified (built) that quantifies the variability in the data.

Variability is decomposed into Trends + Process Variability. You cannot reduce process variability. You can reduce parameter uncertainty (for the past) by having a larger triangle or a related triangle so that you can do some credibility modeling. Future variability on a $\underline{\log \text { scale }}=$ parameter uncertainty + process variability.

In general, parameter uncertainty increases the length of a prediction interval, alternatively, the skewness (both mean and coefficient of variation of the loss reserve distribution). In the case of no parameter uncertainty the prediction distribution's skewness is determined by only the process variability inherent in the data that cannot be reduced.

## Summary of Appendix F3

Appendix F3 is outputs from the model for Paid Losses adjusted by number of claims reported in first Development year for Employers' Liability subclass "a". (Note Appendix F4 is based on a different subclass of Employers' Liability data referred to as subclass "b".)

Appendix F3.2 provides the model displays
Appendices F3.3 and F3.4 give forecast distributions under two different future scenarios, based on the volatility found in the data.

Appendix F3.5 summarises validation analysis based on removal of the last three Calendar years. Whether a model validates well or not depends on the trend structure in the data.

Appendix F3.6 demonstrates that the fitted model captures accurately the volatility in the data. Three triangles are simulated from the fitted model and are seen to be indistinguishable from the real data. Moreover forecast distributions are the same.

Appendix F3.7 shows the relationship between the Paid Losses and the Case Reserve Estimates.


The model display depicts the trends in the three directions plus the process variability about the trend structure.
Note that Weighted Std Residuals are normally distributed - see Normality display below


The volatility of the residuals (process variability) is an integral part of the model.

Note that in respect of Calendar year trends, there is a huge drop in respect of 2001/2002 and the more recent trend is zero.

To assume that the trend will be zero for the next 20 years is very optimistic. We call this scenario 1.

Scenario 2 assumes that the trend for the next six years is 7.16\%+-0.78\% and thereafter reduces to zero.

Note that all forecasts distributions are conditional on an explicit, easily interpretable set of assumptions. Two key assumptions involve the future Calendar year trend volatility and the process variability in the past data depicted by the bottom right graph in the model display.

Another scenario that can easily be considered is the 17.5\%+$2.3 \%$ at least for the next two years.

More importantly it would be important to investigate what might have caused the huge drop from 2001-2002 and the relatively constant trend from 1995-2001

## Reserve Forecast Table

Accident Period vs Development Period

| Accident Period vs Development Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cal. Per. Total | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 | Reserve | Ultimate |
| 1985 | 327,998 | 327,998 | 2,822,314 | 4,286,674 | 5,005,747 | 4,429,542 | 2,695,789 | 1,994,968 | 1,447,601 | 902,569 | 572,356 | 298,319 | 167,149 | 118,820 | 66,699 | 37,476 | 21,076 | 11,864 | 2,506 | 1,315 | 691 | 363 | 0 | 25,580,529 |
|  | 363,421 | 363,421 | 2,823,353 | 3,740,983 | 4,851,694 | 4,302,127 | 3,337,231 | 2,176,450 | 1,561,371 | 1,390,771 | 628,642 | 87,377 | 200,919 | 39,315 | 23,958 | 29,482 | 9,494 | 11,340 | - | 3,845 | 0 | -1,245 | 0 |  |
| 1986 | 3,137,155 | 314,841 | 2,709,078 | 4,114,650 | 4,804,828 | 4,251,711 | 3,082,343 | 2,282,197 | 1,389,616 | 866,408 | 549,417 | 307,566 | 172,337 | 122,512 | 68,775 | 38,644 | 21,734 | 4,588 | 2,405 | 1,262 | 663 | 348 | 348 | 23,570,729 |
|  | 3,165,803 | 342,451 | 2,787,749 | 3,771,543 | 4,460,001 | 4,101,358 | 2,822,690 | 2,145,689 | 1,196,979 | 827,302 | 537,642 | 372,893 | 48,850 | -7,546 | 123,409 | 51,765 | -13,445 | 215 | 0 | 237 | 600 | 487 | 487 | 487 |
| 1987 | 7,323,749 | 327,998 | 2,822,314 | 4,286,674 | 5,005,747 | 5,276,625 | 3,826,934 | 2,377,407 | 1,447,601 | 902,569 | 614,748 | 344,160 | 192,853 | 137,106 | 76,972 | 43,252 | 9,126 | 4,780 | 06 | 1,315 | 691 | 363 | 1,054 | 26,164,111 |
|  | 6,900,955 | 372,223 | 2,597,547 | 4,597,842 | 4,439,529 | 4,171,869 | 4,073,713 | 2,327,095 | 1,894,814 | 807,301 | 428,835 | 165,943 | 100,930 | 94,769 | 1,811 | 88,567 | 269 | 0 |  | 0 | 951 | 507 | 1,103 | 1,10 |
| 1988 | 12,304,967 | 362,256 | 3,117,370 | 4,735,250 | 6,584,540 | 6,941,315 | 4,224,345 | 2,624,591 | 1,598,294 | 1,070,395 | 729,103 | 408,206 | 228,757 | 162,641 | 91,313 | 19,252 | 10,076 | 5,278 | 2,767 | 1,452 | 763 | 401 | 2,616 | 32,732,389 |
|  | 11,547,134 | 326,349 | 3,172,571 | 5,176,011 | 5,379,296 | 6,387,889 | 4,521,772 | 2,811,563 | 1,323,411 | 1,445,179 | 311,985 | 562,611 | 405,192 | 31,637 | 434,136 | 433,892 | 7,580 | 2,729 | -4,029 | 1,973 | 1,050 | 560 | 2,388 | 2,388 |
| 1989 | 16,990,391 | 351,978 | 3,028,922 | 5,478,566 | 7,621,744 | 6,744,373 | 4,104,489 | 2,550,125 | 1,667,891 | 1,117,061 | 760,927 | 426,045 | 238,766 | 169,766 | 35,773 | 18,706 | 9,790 | 5,128 | 2,689 | 1,411 | 741 | 390 | 5,230 | ,488,631 |
|  | 16,930,957 | 398,416 | 3,384,938 | 6,261,944 | 7,232,344 | 6,056,747 | 4,172,737 | 2,638,001 | 1,999,745 | 1,280,583 | 518,592 | 432,369 | 47,274 | 15,294 | 16,795 | 199 | 4,548 | 22,876 | 3,611 | 1,917 | 1,020 | 544 | 4,416 | 4,416 |
| 1990 | 20,097,157 | 379,738 | 3,890,728 | 7,039,194 | 8,219,945 | 7,273,483 | 4,426,533 | 2,953,994 | 1,932,258 | 1,294,265 | 881,725 | 493,729 | 276,725 | 73,872 | 38,592 | 20,180 | 10,561 | 5,532 | 2,900 | 1,522 | 799 | 420 | 11,174 | 9,101,530 |
|  | 20,777,343 | 338,276 | 3,981,902 | 7,269,814 | 7,273,202 | 6,764,962 | 4,663,626 | 3,778,364 | 1,786,086 | 2,035,300 | 966,986 | 143,383 | 21,026 | -117 | 20,750 | 28,458 | 18,336 | 7,353 | 3,896 | 2,069 | 1,101 | 587 | 8,988 | 8,988 |
| 1991 | 26,697,014 | 389,243 | 3,989,913 | 6,059,969 | 7,076,465 | 6,261,666 | 4,093,134 | 2,731,640 | 1,786,901 | 1,196,962 | 815,477 | 456,656 | 96,124 | 63,595 | 33,224 | 17,372 | 9,092 | 4,763 | 2,497 | 1,310 | 688 | 362 | 18,712 | 33,795,441 |
|  | 25,151,751 | 357,601 | 3,930,777 | 6,254,476 | 7,270,856 | 5,543,963 | 4,385,023 | 2,271,060 | 2,163,980 | 1,131,551 | 314,255 | -113,095 | -18,502 | 267,189 | 15,518 | 2,076 | 11,973 | 6,330 | 3,354 | 1,781 | 948 | 50 | 14,566 | 14,566 |
| 1992 | 33,503,376 | 354,477 | 3,050,078 | 4,632,526 | 5,409,584 | 5,141,467 | 3,361,048 | 2,243,179 | 1,467,447 | 983,024 | 669,757 | 140,906 | 73,482 | 48,615 | 25,398 | 13,280 | 6,950 | 3,641 | 1,909 | 1,002 | 526 | 277 | 27,585 | ,187,447 |
|  | 33,007,909 | 406,312 | 3,485,198 | 4,787,544 | 6,016,700 | 4,772,276 | 3,259,134 | 1,978,110 | 1,388,877 | 645,655 | 305,919 | 90,952 | 22,476 | 12,200 | -11,490 | 17,351 | 9,153 | 4,839 | 2,564 | 1,361 | 724 | 386 | 20,975 | 20,975 |
| 1993 | 33,263,505 | 295,203 | 2,540,061 | 3,857,901 | 4,838,880 | 4,599,349 | 3,006,808 | 2,006,857 | 1,312,915 | 879,548 | 225,216 | 117,345 | 61,195 | 40,486 | 21,151 | 11,060 | 5,788 | 3,032 | 1,590 | 834 | 438 | 230 | 44,123 | 24,957,383 |
|  | 32,793,140 | 286,900 | 2,500,351 | 4,252,026 | 5,428,687 | 4,340,034 | 3,008,771 | 2,731,563 | 1,181,851 | 728,782 | 371,273 | 68,942 | 10,783 | 3,295 | 27,456 | 14,450 | 7,622 | 4,030 | 2,135 | 1,134 | 603 | 32 | 32,974 | 32,974 |
| 1994 | 31,473,817 | 335,836 | 2,889,680 | 4,714,267 | 5,913,229 | 5,620,882 | 3,674,814 | 2,452,831 | 1,604,758 | 404,177 | 256,216 | 133,496 | 69,618 | 46,059 | 24,062 | 12,582 | 6,585 | 3,449 | 1,808 | 949 | 498 | 262 | 96,255 | 27,930,657 |
|  | 32,018,107 | 359,335 | 2,940,485 | 4,749,103 | 5,360,802 | 5,962,373 | 4,131,260 | 2,780,882 | 1,056,343 | 386,307 | 226,250 | -15,475 | -103,263 | 59,485 | 31,235 | 16,439 | 8,671 | 4,584 | 2,429 | 1,290 | 68 | 366 | 71,008 | 71,008 |
| 1995 | 29,114,365 | 370,278 | 3,422,255 | 5,583,414 | 7,003,692 | 6,657,867 | 4,352,989 | 2,905,638 | 715,087 | 445,629 | 282,493 | 147,187 | 76,757 | 50,782 | 26,530 | 13,872 | 7,260 | 3,803 | 1,994 | 1,046 | 550 | 28 | 182,88 | ,632,347 |
|  | 29,082,604 | 272,072 | 3,351,803 | 5,088,547 | 7,048,157 | 6,656,030 | 3,895,374 | 1,595,575 | 794,846 | 513,843 | 47,316 | 185,902 | 62,026 | 65,586 | 34,439 | 18,125 | 9,561 | 5,054 | 2,678 | 1,422 | 757 | 404 | 101,046 | 101,046 |
| 1996 | 29,365,438 | 373,692 | 3,453,624 | 5,634,678 | 7,067,685 | 6,718,980 | 4,392,885 | 1,103,615 | 671,648 | 418,554 | 265,331 | 138,246 | 72,095 | 47,698 | 24,919 | 13,030 | 6,819 | 3,572 | 1,873 | 98 | 516 | 27 | 310,02 | 32,887,118 |
|  | 31,282,166 | 369,338 | 3,373,646 | 5,580,846 | 7,733,248 | 9,576,139 | 4,414,667 | 466,863 | 777,454 | 204,599 | 80,294 | 111,281 | 58,242 | 61,592 | 32,342 | 17,022 | 8,979 | 4,747 | 2,515 | 1,336 | 71 | 379 | 147,76 | 147,766 |
| 1997 | 30,508,236 | 452,194 | 4,179,321 | 6,819,032 | 8,553,572 | 8,132,086 | 2,002,289 | 1,243,249 | 756,628 | 471,512 | 298,902 | 155,738 | 81,217 | 53,733 | 28,072 | 14,679 | 7,682 | 4,024 | 2,110 | 1,107 | 582 | 306 | 648,151 | 32,400,477 |
|  | 27,665,2 | 339,282 | 3,732,991 | 5,227,094 | 9,228,797 | 8,435,768 | 2,888,546 | 1,135,434 | 547,640 | 216,773 | 240,279 | 125,361 | 65,611 | 69,385 | 36,434 | 19,175 | 10,115 | 5,348 | 2,833 | 1,505 | 80 | 42 | 294,613 | 294,613 |
| 1998 | 32,778,418 | 457,495 | 4,228,506 | 6,899,649 | 8,655,029 | 3,100,593 | 1,885,782 | 1,170,908 | 712,602 | 444,076 | 281,510 | 146,676 | 76,491 | 50,607 | 26,438 | 13,824 | 7,235 | 3,790 | 1,98 | 1,043 | 548 | 288 | 1,054,514 | 9,515,120 |
|  | 32,987,794 | 278,722 | 3,494,212 | 5,806,965 | 11,247,156 | 10,692,710 | 4,906,059 | 1,673,282 | 361,501 | 147,905 | 226,298 | 118,066 | 61,793 | 65,347 | 34,314 | 18,060 | 9,526 | 5,036 | 2,669 | 1,417 | 754 | 40 | 318,921 | 318,921 |
| 1999 | 36,465,760 | 1,566,708 | 14,482,146 | 23,629,812 | 11,083,082 | 9,807,436 | 5,968,885 | 3,708,648 | 2,258,555 | 1,408,419 | 893,178 | 465,557 | 242,881 | 160,753 | 84,015 | 43,948 | 23,009 | 12,058 | 6,324 | 3,320 | 1,744 | 917 | 5,604,68 | 46,760,396 |
|  | 34,807,630 | 261,587 | 3,279,725 | 6,885,099 | 10,432,638 | 11,101,285 | 6,269,954 | 2,925,429 | 532,779 | 463,153 | 716,414 | 374,293 | 196,167 | 207,678 | 109,162 | 57,510 | 30,367 | 16,070 | 8,523 | 4,530 | 2,413 | 1,288 | 1,155,713 | 1,155,713 |
| 2000 | 49,001,211 | 1,397,800 | 12,921,401 | 7,882,100 | 9,204,080 | 8,144,705 | 4,956,933 | 3,079,892 | 1,875,645 | 1,169,639 | 741,750 | 386,627 | 201,704 | 133,499 | 69,771 | 36,497 | 19,109 | 10,013 | 5,252 | 2,757 | 1,449 | 76 | 7,734,366 | 41,191,749 |
|  | 37,162,551 | 167,544 | 3,509,353 | 6,962,743 | 9,221,279 | 9,569,081 | 4,027,383 | 718,487 | 442,452 | 384,631 | 594,954 | 310,836 | 162,909 | 172,468 | 90,655 | 47,760 | 25,218 | 13,346 | 7,078 | 3,762 | 2,004 | ,07 | 1,223,561 | 1,223,561 |
| 2001 | 66,251,939 | 1,389,714 | 4,802,771 | 7,293,884 | 8,517,209 | 7,536,891 | 4,587,013 | 2,850,049 | 1,735,671 | 1,082,352 | 686,396 | 357,774 | 186,651 | 123,537 | 64,564 | 33,773 | 17,683 | 9,266 | 4,860 | 2,551 | 1,341 | 705 | 11,744,187 | 37,391,773 |
|  | 38,780,673 | 158,189 | 2,663,632 | 6,601,908 | 8,424,381 | 7,799,475 | 498,312 | 664,868 | 409,434 | 355,928 | 550,555 | 287,639 | 150,752 | 159,598 | 83,890 | 44,196 | 23,336 | 12,350 | 6,550 | 3,48 | 1,854 | 99 | 1,268,375 | 1,268,375 |
| 2002 | 32,258,341 | 557,264 | 4,794,893 | 7,281,920 | $8,503,238$ | 7,524,529 | 4,579,489 | 2,845,374 | 1,732,824 | 1,080,577 | 685,270 | 357,188 | 186,345 | 123,334 | 64,458 | 33,718 | 17,654 | 9,251 | 4,852 | 2,547 | 1,338 | 70 | 19,249,453 | 38,546,594 |
|  | 36,338,872 | 155,809 | 3,402,607 | 7,405,743 | 8,332,982 | 821,511 | 497,495 | 663,778 | 408,762 | 355,344 | 549,652 | 287,168 | 150,504 | 159,336 | 83,752 | 44,123 | 23,298 | 12,329 | 6,539 | 3,476 | 1,851 | 988 | 1,548,518 | 1,548,518 |
| 003 | 36,334,581 | 401,468 | 3,454,371 | 5,246,093 | 6,125,964 | 5,420,875 | 3,299,188 | 2,049,885 | 1,248,374 | 778,477 | 493,687 | 257,328 | 134,248 | 88,853 | 46,438 | 24,291 | 12,718 | 6,665 | 3,496 | 1,835 | 964 | 507 | 19,993,792 | 29,515,09 |
|  | 38,753,674 | 476,574 | 3,196,497 | 5,848,228 | 669,867 | 591,839 | 358,409 | 478,204 | 294,483 | 255,999 | 395,984 | 206,883 | 108,427 | 114,790 | 60,337 | 31,788 | 16,785 | 8,882 | 4,711 | 2,504 | 1,334 | 712 | 1,339,824 | 1,339,824 |
| 2004 | 36,625,332 | 313,423 | 2,696,799 | 4,095,582 | 4,782,489 | 4,232,032 | 2,575,649 | 1,600,328 | 974,595 | 607,751 | 385,418 | 200,894 | 104,806 | 69,367 | 36,253 | 18,964 | 9,929 | 5,203 | 2,729 | 1,433 | 753 | 396 | 19,704,570 | 22,920,58 |
|  | 37,600,773 | 203,281 | 3,012,732 | 448,295 | 522,960 | 462,044 | 279,807 | 373,330 | 229,901 | 199,857 | 309,141 | 161,512 | 84,648 | 89,616 | 47,105 | 24,816 | 13,104 | 6,934 | 3,678 | 1,955 | 1,041 | 556 | 1,168,228 | 1,168,22 |
| 2005 | 34,703,794 | 284,888 | 2,451,278 | 3,722,712 | 4,347,083 | 3,846,740 | 2,341,157 | 1,454,631 | 885,866 | 552,420 | 350,328 | 182,604 | 95,265 | 63,052 | 32,953 | 17,238 | 9,025 | 4,729 | 2,481 | 1,302 | 684 | 360 | 20,361,907 | 20,635,105 |
|  | 32,991,052 | 273,198 | 270,806 | 407,481 | 475,349 | 419,978 | 254,333 | 339,341 | 208,970 | 181,661 | 280,997 | 146,808 | 76,942 | 81,457 | 42,816 | 22,557 | 11,911 | 6,303 | 3,343 | 1,777 | 946 | 505 | 1,119,926 | 1,119,926 |
|  | Total Fitted/Paid |  | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | 2023 | 2024 | 202 | Total Reserve | Total Ulitimate |
| Cal. Per. Total | 598,526,543 |  | 31,181,087 | 25,252,076 | 18,851,459 | 12,677,799 | 7,692,458 | 4,668,475 | 2,777,393 | 1,633,673 | 940,569 | 509,561 | 280,703 | 159,942 | 83,535 | 43,601 | 22,373 | 11,337 | 5,581 | 2,562 | 1,080 | 360 | 06,795,624 | 666,905,202 |
|  | 560,109,578 |  | 1,751,257 | 1,534,726 | 1,457,655 | 1,198,589 | 976,256 | 859,996 | 641,822 | 492,459 | 397,906 | 233,529 | 146,915 | 105,717 | 56,116 | 29,878 | 15,888 | 8,435 | 4,441 | 2,275 | 1,127 | 505 | 4,607,172 | 4,607,1 |
| 1 Unit $=\$ 1$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

Inside the table the blues are the observed, the black are means of log-normal distributions and the reds are the corresponding Std Devs of the log-normals.

The predictive log-normal distributions include parameter uncertainty and process variability.

## Summaries by Year

| Accident Yr Summary |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Acc. Yr | Mean |  |  | $\begin{aligned} & \text { Standard } \\ & \text { Dev. } \end{aligned}$ | cV |  |
|  | Reserve | Ultimate |  |  | Reserve | Ultimate |
| 1985 | 0 | 25,580,529 |  | 0 | **** | 0 |
| 1986 | 348 | 23,570,729 |  | 487 | 1.4 | 0 |
| 1987 | 1,054 | 26,164,111 |  | 1,103 | 1.05 | 0 |
| 1988 | 2,616 | 32,732,389 |  | 2,388 | 0.91 | 0 |
| 1989 | 5,230 | 34,488,631 |  | 4,416 | 0.84 | 0 |
| 1990 | 11,174 | 39,101,530 |  | 8,988 | 0.8 | 0 |
| 1991 | 18,712 | 33,795,441 |  | 14,566 | 0.78 | 0 |
| 1992 | 27,585 | 27,187,447 |  | 20,975 | 0.76 | 0 |
| 1993 | 44,123 | 24,957,383 |  | 32,974 | 0.75 | 0 |
| 1994 | 96,255 | 27,930,657 |  | 71,008 | 0.74 | 0 |
| 1995 | 182,884 | 29,632,347 |  | 101,046 | 0.55 | 0 |
| 1996 | 310,023 | 32,887,118 |  | 147,766 | 0.48 | 0 |
| 1997 | 648,151 | 32,400,477 |  | 294,613 | 0.45 | 0.01 |
| 1998 | 1,054,514 | 39,515,120 |  | 318,921 | 0.3 | 0.01 |
| 1999 | 5,604,680 | 46,760,396 |  | 1,155,713 | 0.21 | 0.02 |
| 2000 | 7,734,366 | 41,191,749 |  | 1,223,561 | 0.16 | 0.03 |
| 2001 | 11,744,187 | 37,391,773 |  | 1,268,375 | 0.11 | 0.03 |
| 2002 | 19,249,453 | 38,546,594 |  | 1,548,518 | 0.08 | 0.04 |
| 2003 | 19,993,792 | 29,515,092 |  | 1,339,824 | 0.07 | 0.05 |
| 2004 | 19,704,570 | 22,920,584 |  | 1,168,228 | 0.06 | 0.05 |
| 2005 | 20,361,907 | 20,635,105 |  | 1,119,926 | 0.06 | 0.05 |
| Total | 106,795,624 | 666,905,202 |  | 4,607,172 | 0.04 | 0.01 |
|  |  |  | 1 Unit $=$ \$ |  |  |  |


| Calendar Yr Summary |  |  |  |
| :---: | :---: | :---: | :---: |
| Calendar <br> Yr | Mean Reserve | Standard Dev. | CV Reserve |
| 2006 | 31,181,087 | 1,751,257 | 0.06 |
| 2007 | 25,252,076 | 1,534,726 | 0.06 |
| 2008 | 18,851,459 | 1,457,655 | 0.08 |
| 2009 | 12,677,799 | 1,198,589 | 0.09 |
| 2010 | 7,692,458 | 976,256 | 0.13 |
| 2011 | 4,668,475 | 859,996 | 0.18 |
| 2012 | 2,777,393 | 641,822 | 0.23 |
| 2013 | 1,633,673 | 492,459 | 0.3 |
| 2014 | 940,569 | 397,906 | 0.42 |
| 2015 | 509,561 | 233,529 | 0.46 |
| 2016 | 280,703 | 146,915 | 0.52 |
| 2017 | 159,942 | 105,717 | 0.66 |
| 2018 | 83,535 | 56,116 | 0.67 |
| 2019 | 43,601 | 29,878 | 0.69 |
| 2020 | 22,373 | 15,888 | 0.71 |
| 2021 | 11,337 | 8,435 | 0.74 |
| 2022 | 5,581 | 4,441 | 0.8 |
| 2023 | 2,562 | 2,275 | 0.89 |
| 2024 | 1,080 | 1,127 | 1.04 |
| 2025 | 360 | 505 | 1.4 |
| Total | 106,795,624 | 4,607,172 | 0.04 |
| 1 Unit $=$ \$1 |  |  |  |

In order to find distributions of aggregates by accident year, calendar year and total of all years, we simulate from the predictive log-normal distributions of each cell, incorporating their correlations.
We do this because there is no analytical form for the sum of log-normals.

## However, note that the means and Std Devs of the distributions of aggregates are computed via the model using second order moments,

## The sample quaniles (perceniles) below are based on 10,000 simulations of the forecast trangle.

The Kernel is fitted to the sample frequency plot whereas the log-normal and gamma are just based on the means and std. devs. of the actua
forecast distributions.
Note that the distributions are slightly skewed.

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \#S.D.'s | V-a-R | Quantile | \#S.D.'s | V -a-R |
| 99.995 | 124.733 | 3.893 | 17.937 | 125.664 | 4.095 | 18.869 | 126.185 | 4.209 | 19.39 | 125.664 | 4.095 | 18.868 |
| 99.99 | 124.519 | 3.847 | 17.723 | 125.259 | 4.007 | 18.463 | 125.256 | 4.007 | 18.46 | 124.786 | 3.905 | 17.99 |
| 99.98 | 124.349 | 3.81 | 17.554 | 124.752 | 3.898 | 17.957 | 124.293 | 3.798 | 17.498 | 123.874 | 3.707 | 17.078 |
| 99.97 | 124.185 | 3.774 | 17.389 | 124.386 | 3.818 | 17.59 | 123.713 | 3.672 | 16.918 | 123.323 | 3.587 | 16.528 |
| 99.96 | 124.104 | 3.757 | 17.309 | 124.083 | 3.752 | 17.287 | 123.294 | 3.581 | 16.498 | 122.924 | 3.501 | 16.129 |
| 99.95 | 123.393 | 3.603 | 16.598 | 123.815 | 3.694 | 17.02 | 122.963 | 3.509 | 16.168 | 122.61 | 3.432 | 15.814 |
| 99.94 | 123.114 | 3.542 | 16.319 | 123.58 | 3.643 | 16.784 | 122.69 | 3.45 | 15.894 | 122.349 | 3.376 | 15.553 |
| 99.93 | 123.108 | 3.541 | 16.313 | 123.366 | 3.597 | 16.571 | 122.456 | 3.399 | 15.66 | 122.126 | 3.328 | 15.33 |
| 99.92 | 122.869 | 3.489 | 16.073 | 123.172 | 3.555 | 16.377 | 122.252 | 3.355 | 15.456 | 121.931 | 3.285 | 15.135 |
| 99.91 | 122.833 | 3.481 | 16.037 | 122.995 | 3.516 | 16.199 | 122.07 | 3.315 | 15.274 | 121.758 | 3.248 | 14.962 |
| 99.9 | 122.422 | 3.392 | 15.626 | 122.831 | 3.481 | 16.036 | 121.906 | 3.28 | 15.11 | 121.601 | 3.214 | 14.805 |
| 99.8 | 121.548 | 3.202 | 14.753 | 121.605 | 3.214 | 14.809 | 120.797 | 3.039 | 14.001 | 120.539 | 2.983 | 13.744 |
| 99.7 | 120.51 | 2.977 | 13.714 | 120.732 | 3.025 | 13.937 | 120.12 | 2.892 | 13.324 | 119.889 | 2.842 | 13.094 |
| 99.6 | 119.768 | 2.816 | 12.972 | 120.095 | 2.887 | 13.299 | 119.625 | 2.785 | 12.829 | 119.414 | 2.739 | 12.618 |
| 99.5 | 119.334 | 2.722 | 12.539 | 119.619 | 2.783 | 12.824 | 119.232 | 2.699 | 12.437 | 119.036 | 2.657 | 12.24 |
| 99.4 | 119.034 | 2.656 | 12.239 | 119.238 | 2.701 | 12.443 | 118.905 | 2.628 | 12.11 | 118.721 | 2.588 | 11.925 |
| 99.3 | 118.738 | 2.592 | 11.943 | 118.92 | 2.632 | 12.124 | 118.624 | 2.567 | 11.829 | 118.45 | 2.53 | 11.654 |
| 99.2 | 118.52 | 2.545 | 11.725 | 118.644 | 2.572 | 11.848 | 118.377 | 2.514 | 11.582 | 118.211 | 2.478 | 11.416 |
| 99.1 | 118.24 | 2.484 | 11.445 | 118.401 | 2.519 | 11.606 | 118.157 | 2.466 | 11.361 | 117.998 | 2.432 | 11.202 |
| 99 | 117.946 | 2.42 | 11.15 | 118.184 | 2.472 | 11.388 | 117.957 | 2.423 | 11.161 | 117.805 | 2.39 | 11.009 |
| 98 | 116.635 | 2.136 | 9.839 | 116.785 | 2.168 | 9.989 | 116.578 | 2.123 | 9.783 | 116.469 | 2.1 | 9.674 |
| 97 | 115.812 | 1.957 | 9.017 | 115.93 | 1.983 | 9.134 | 115.712 | 1.935 | 8.916 | 115.627 | 1.917 | 8.832 |
| 96 | 115.125 | 1.808 | 8.33 | 115.271 | 1.84 | 8.475 | 115.065 | 1.795 | 8.269 | 114.996 | 1.78 | 8.201 |
| 95 | 114.638 | 1.702 | 7.842 | 114.728 | 1.722 | 7.933 | 114.541 | 1.681 | 7.745 | 114.485 | 1.669 | 7.689 |
| 94 | 114.156 | 1.598 | 7.361 | 114.266 | 1.621 | 7.47 | 114.096 | 1.585 | 7.301 | 114.051 | 1.575 | 7.255 |
| 93 | 113.735 | 1.506 | 6.94 | 113.867 | 1.535 | 7.071 | 113.708 | 1.5 | 6.913 | 113.671 | 1.492 | 6.876 |
| 92 | 113.411 | 1.436 | 6.615 | 113.513 | 1.458 | 6.718 | 113.362 | 1.425 | 6.566 | 113.332 | 1.419 | 6.536 |
| 91 | 113.101 | 1.369 | 6.305 | 113.196 | 1.389 | 6.401 | 113.048 | 1.357 | 6.252 | 113.024 | 1.352 | 6.228 |
| 90 | 112.795 | 1.302 | 5.999 | 112.907 | 1.326 | 6.111 | 112.759 | 1.294 | 5.964 | 112.741 | 1.29 | 5.945 |
| 89 | 112.542 | 1.247 | 5.747 | 112.64 | 1.268 | 5.844 | 112.492 | 1.236 | 5.696 | 112.478 | 1.233 | 5.683 |
| 88 | 112.319 | 1.199 | 5.524 | 112.39 | 1.214 | 5.594 | 112.242 | 1.182 | 5.446 | 112.233 | 1.18 | 5.437 |
| 87 | 112.097 | 1.151 | 5.302 | 112.154 | 1.163 | 5.358 | 112.007 | 1.131 | 5.211 | 112.001 | 1.13 | 5.206 |
| 86 | 111.863 | 1.1 | 5.068 | 111.929 | 1.114 | 5.134 | 111.784 | 1.083 | 4.989 | 111.782 | 1.082 | 4.987 |
| 85 | 111.637 | 1.051 | 4.841 | 111.716 | 1.068 | 4.92 | 111.573 | 1.037 | 4.777 | 111.574 | 1.037 | 4.778 |
| 84 | 111.445 | 1.009 | 4.65 | 111.51 | 1.023 | 4.714 | 111.371 | 0.993 | 4.575 | 111.375 | 0.994 | 4.579 |
| 83 | 111.278 | 0.973 | 4.483 | 111.311 | 0.98 | 4.516 | 111.177 | 0.951 | 4.382 | 111.184 | 0.953 | 4.389 |
| 82 | 111.095 | 0.933 | 4.3 | 111.119 | 0.938 | 4.324 | 110.991 | 0.911 | 4.196 | 111.001 | 0.913 | 4.205 |
| 81 | 110.88 | 0.887 | 4.084 | 110.934 | 0.898 | 4.138 | 110.812 | 0.872 | 4.017 | 110.824 | 0.874 | 4.028 |
| 80 | 110.705 | 0.849 | 3.909 | 110.754 | 0.859 | 3.958 | 110.639 | 0.834 | 3.843 | 110.653 | 0.837 | 3.857 |
| 79 | 110.503 | 0.805 | 3.707 | 110.579 | 0.821 | 3.783 | 110.471 | 0.798 | 3.675 | 110.487 | 0.801 | 3.691 |
| 78 | 110.351 | 0.772 | 3.556 | 110.408 | 0.784 | 3.612 | 110.308 | 0.762 | 3.512 | 110.325 | 0.766 | 3.53 |
| 77 | 110.182 | 0.735 | 3.386 | 110.242 | 0.748 | 3.446 | 110.149 | 0.728 | 3.354 | 110.168 | 0.732 | 3.373 |
| 76 | 110.051 | 0.707 | 3.255 | 110.079 | 0.713 | 3.284 | 109.995 | 0.694 | 3.199 | 110.015 | 0.699 | 3.22 |
| 75 | 109.894 | 0.673 | 3.099 | 109.921 | 0.678 | 3.125 | 109.844 | 0.662 | 3.048 | 109.866 | 0.666 | 3.07 |
| 74 | 109.728 | 0.636 | 2.932 | 109.765 | 0.645 | 2.97 | 109.696 | 0.63 | 2.901 | 109.72 | 0.635 | 2.924 |
| 73 | 109.573 | 0.603 | 2.777 | 109.613 | 0.612 | 2.818 | 109.552 | 0.598 | 2.756 | 109.577 | 0.604 | 2.781 |
| 72 | 109.435 | 0.573 | 2.639 | 109.464 | 0.579 | 2.669 | 109.41 | 0.567 | 2.615 | 109.436 | 0.573 | 2.641 |
| 71 | 109.285 | 0.54 | 2.49 | 109.318 | 0.548 | 2.523 | 109.271 | 0.537 | 2.476 | 109.298 | 0.543 | 2.503 |
| 70 | 109.153 | 0.512 | 2.357 | 109.175 | 0.516 | 2.379 | 109.135 | 0.508 | 2.339 | 109.163 | 0.514 | 2.367 |

## Quantiles and VaR Mean $=106.796$, S.D. $=4.607$, Provision $=106.796,1$ Unit $=\$ 1,000,000$

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \#S.D.'s | V-a-R | Quantile | \#S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \#S.D.'s | V-a-R |
| 69 | 108.999 | 0.478 | 2.203 | 109.035 | 0.486 | 2.239 | 109 | 0.479 | 2.205 | 109.029 | 0.485 | 2.234 |
| 68 | 108.823 | 0.44 | 2.028 | 108.897 | 0.456 | 2.101 | 108.868 | 0.45 | 2.072 | 108.898 | 0.456 | 2.102 |
| 67 | 108.698 | 0.413 | 1.903 | 108.762 | 0.427 | 1.966 | 108.738 | 0.422 | 1.942 | 108.768 | 0.428 | 1.973 |
| 66 | 108.565 | 0.384 | 1.769 | 108.629 | 0.398 | 1.834 | 108.609 | 0.394 | 1.813 | 108.64 | 0.4 | 1.845 |
| 65 | 108.454 | 0.36 | 1.658 | 108.499 | 0.37 | 1.703 | 108.482 | 0.366 | 1.686 | 108.514 | 0.373 | 1.718 |
| 64 | 108.346 | 0.336 | 1.55 | 108.37 | 0.342 | 1.574 | 108.356 | 0.339 | 1.561 | 108.389 | 0.346 | 1.593 |
| 63 | 108.209 | 0.307 | 1.414 | 108.243 | 0.314 | 1.447 | 108.232 | 0.312 | 1.436 | 108.265 | 0.319 | 1.469 |
| 62 | 108.094 | 0.282 | 1.299 | 108.118 | 0.287 | 1.323 | 108.109 | 0.285 | 1.313 | 108.142 | 0.292 | 1.347 |
| 61 | 107.976 | 0.256 | 1.18 | 107.995 | 0.26 | 1.199 | 107.987 | 0.259 | 1.192 | 108.021 | 0.266 | 1.225 |
| 60 | 107.861 | 0.231 | 1.066 | 107.872 | 0.234 | 1.077 | 107.866 | 0.232 | 1.071 | 107.9 | 0.24 | 1.105 |
| 59 | 107.734 | 0.204 | 0.938 | 107.751 | 0.207 | 0.955 | 107.747 | 0.206 | 0.951 | 107.781 | 0.214 | 0.985 |
| 58 | 107.611 | 0.177 | 0.815 | 107.631 | 0.181 | 0.835 | 107.628 | 0.181 | 0.832 | 107.662 | 0.188 | 0.866 |
| 57 | 107.5 | 0.153 | 0.705 | 107.511 | 0.155 | 0.716 | 107.509 | 0.155 | 0.714 | 107.544 | 0.162 | 0.748 |
| 56 | 107.389 | 0.129 | 0.594 | 107.392 | 0.129 | 0.597 | 107.392 | 0.129 | 0.596 | 107.426 | 0.137 | 0.631 |
| 55 | 107.265 | 0.102 | 0.47 | 107.274 | 0.104 | 0.478 | 107.275 | 0.104 | 0.479 | 107.309 | 0.111 | 0.514 |
| 54 | 107.151 | 0.077 | 0.356 | 107.156 | 0.078 | 0.36 | 107.158 | 0.079 | 0.363 | 107.193 | 0.086 | 0.397 |
| 53 | 107.045 | 0.054 | 0.249 | 107.038 | 0.053 | 0.242 | 107.042 | 0.054 | 0.247 | 107.076 | 0.061 | 0.281 |
| 52 | 106.915 | 0.026 | 0.119 | 106.92 | 0.027 | 0.124 | 106.927 | 0.028 | 0.131 | 106.961 | 0.036 | 0.165 |
| 51 | 106.799 | 0.001 | 0.004 | 106.802 | 0.001 | 0.007 | 106.811 | 0.003 | 0.016 | 106.845 | 0.011 | 0.049 |
| 50 | 106.684 | -0.024 | -0.112 | 106.684 | -0.024 | -0.111 | 106.696 | -0.022 | -0.099 | 106.729 | -0.014 | -0.066 |

## Below are qunatiles and VaR statistics for the next calendar year 2006

Note that the distributions here are more skewed than for the aggregate reserves.
Quantile Statistics and Value at Risk (Cal. Yr: 2006)

| Quantile Statistics and Value at Risk (Cal. Yr: 2006) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \#S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \#S.D.'s | V-a-R | Quantile | \#S.D.'s | V-a-R |
| 99.95 | 37.521 | 3.62 | 6.34 | 37.663 | 3.701 | 6.482 | 37.447 | 3.578 | 6.265 | 37.268 | 3.476 | 6.087 |
| 99.9 | 37.109 | 3.385 | 5.928 | 37.222 | 3.449 | 6.041 | 37.028 | 3.339 | 5.847 | 36.874 | 3.251 | 5.693 |
| 99.8 | 36.577 | 3.081 | 5.396 | 36.739 | 3.174 | 5.558 | 36.59 | 3.089 | 5.409 | 36.461 | 3.015 | 5.28 |
| 99.7 | 36.371 | 2.964 | 5.19 | 36.464 | 3.017 | 5.283 | 36.323 | 2.936 | 5.142 | 36.208 | 2.871 | 5.027 |
| 99.6 | 36.185 | 2.857 | 5.004 | 36.264 | 2.902 | 5.083 | 36.129 | 2.825 | 4.948 | 36.023 | 2.765 | 4.842 |
| 99.5 | 36.114 | 2.817 | 4.933 | 36.102 | 2.81 | 4.921 | 35.975 | 2.737 | 4.794 | 35.877 | 2.681 | 4.695 |
| 99.4 | 35.846 | 2.664 | 4.665 | 35.967 | 2.733 | 4.785 | 35.846 | 2.664 | 4.665 | 35.754 | 2.611 | 4.573 |
| 99.3 | 35.771 | 2.621 | 4.59 | 35.851 | 2.666 | 4.67 | 35.736 | 2.601 | 4.555 | 35.649 | 2.551 | 4.468 |
| 99.2 | 35.678 | 2.568 | 4.497 | 35.749 | 2.608 | 4.567 | 35.639 | 2.546 | 4.458 | 35.557 | 2.499 | 4.376 |
| 99.1 | 35.523 | 2.48 | 4.342 | 35.659 | 2.557 | 4.478 | 35.553 | 2.496 | 4.372 | 35.474 | 2.451 | 4.293 |
| 99 | 35.492 | 2.462 | 4.311 | 35.579 | 2.511 | 4.397 | 35.475 | 2.452 | 4.293 | 35.399 | 2.409 | 4.218 |
| 98 | 34.968 | 2.163 | 3.787 | 35.015 | 2.189 | 3.834 | 34.936 | 2.144 | 3.755 | 34.882 | 2.113 | 3.701 |
| 97 | 34.596 | 1.95 | 3.415 | 34.653 | 1.982 | 3.472 | 34.599 | 1.951 | 3.417 | 34.557 | 1.928 | 3.376 |
| 96 | 34.338 | 1.803 | 3.157 | 34.39 | 1.832 | 3.209 | 34.347 | 1.808 | 3.166 | 34.314 | 1.789 | 3.133 |
| 95 | 34.143 | 1.691 | 2.962 | 34.182 | 1.714 | 3.001 | 34.143 | 1.691 | 2.962 | 34.116 | 1.676 | 2.935 |
| 94 | 33.964 | 1.589 | 2.783 | 34.01 | 1.615 | 2.829 | 33.971 | 1.593 | 2.79 | 33.949 | 1.581 | 2.768 |
| 93 | 33.82 | 1.507 | 2.639 | 33.859 | 1.529 | 2.678 | 33.821 | 1.507 | 2.64 | 33.803 | 1.497 | 2.622 |
| 92 | 33.683 | 1.428 | 2.501 | 33.724 | 1.452 | 2.543 | 33.687 | 1.431 | 2.506 | 33.673 | 1.423 | 2.491 |
| 91 | 33.578 | 1.368 | 2.396 | 33.601 | 1.382 | 2.42 | 33.565 | 1.361 | 2.384 | 33.554 | 1.355 | 2.373 |
| 90 | 33.464 | 1.304 | 2.283 | 33.486 | 1.316 | 2.305 | 33.454 | 1.298 | 2.273 | 33.445 | 1.293 | 2.264 |
| 89 | 33.366 | 1.247 | 2.185 | 33.378 | 1.254 | 2.197 | 33.351 | 1.239 | 2.17 | 33.345 | 1.235 | 2.163 |
| 88 | 33.256 | 1.185 | 2.074 | 33.277 | 1.197 | 2.096 | 33.254 | 1.184 | 2.073 | 33.25 | 1.182 | 2.069 |
| 87 | 33.134 | 1.115 | 1.953 | 33.181 | 1.142 |  | 33.164 | 1.132 | 1.983 | 33.162 | 1.131 | 1.98 |
| 86 | 33.052 | 1.068 | 1.871 | 33.092 | 1.091 | 1.911 | 33.078 | 1.083 | 1.897 | 33.078 | 1.083 | 1.896 |
| 85 | 32.975 | 1.024 | 1.794 | 33.007 | 1.042 | 1.826 | 32.996 | 1.037 | 1.815 | 32.998 | 1.037 | 1.817 |
| 84 | 32.888 | 0.975 | 1.707 | 32.926 | 0.996 | 1.745 | 32.919 | 0.992 | 1.738 | 32.921 | 0.994 | 1.74 |
| 83 | 32.819 | 0.935 | 1.638 | 32.85 | 0.953 | 1.669 | 32.844 | 0.95 | 1.663 | 32.848 | 0.952 | 1.667 |
| 82 | 32.749 | 0.895 | 1.568 | 32.777 | 0.911 | 1.596 | 32.773 | 0.909 | 1.592 | 32.778 | 0.912 | 1.597 |
| 81 | 32.681 | 0.857 | 1.5 | 32.707 | 0.871 | 1.526 | 32.704 | 0.87 | 1.523 | 32.71 | 0.873 | 1.529 |
| 80 | 32.619 | 0.821 | 1.438 | 32.639 | 0.833 | 1.458 | 32.637 | 0.832 | 1.456 | 32.645 | 0.836 | 1.464 |
| 79 | 32.562 | 0.789 | 1.381 | 32.574 | 0.796 | 1.393 | 32.573 | 0.795 | 1.392 | 32.581 | 0.799 | 1.4 |
| 78 | 32.485 | 0.745 | 1.304 | 32.511 | 0.76 | 1.33 | 32.51 | 0.759 | 1.329 | 32.519 | 0.764 | 1.3388 |
| 77 | 32.423 | 0.709 | 1.242 | 32.45 | 0.725 | 1.269 | 32.45 | 0.724 | 1.268 | 32.459 | 0.73 | 1.278 |
| 76 | 32.367 | 0.677 | 1.186 | 32.391 | 0.691 | 1.21 | 32.39 | 0.69 | 1.209 | 32.401 | 0.697 | 1.22 |
| 75 | 32.308 | 0.643 | 1.127 | 32.333 | 0.658 | 1.152 | 32.332 | 0.657 | 1.151 | 32.344 | 0.664 | 1.163 |
| 74 | 32.255 | 0.613 | 1.074 | 32.277 | 0.626 | 1.096 | 32.276 | 0.625 | 1.095 | 32.288 | 0.632 | 1.107 |
| 73 | 32.204 | 0.584 | 1.023 | 32.221 | 0.594 | 1.04 | 32.221 | 0.594 | 1.04 | 32.233 | 0.601 | 1.052 |
| 72 | 32.149 | 0.553 | 0.968 | 32.167 | 0.563 | 0.986 | 32.166 | 0.563 | 0.985 | 32.18 | 0.57 | 0.998 |
| 71 | 32.106 | 0.528 | 0.925 | 32.114 | 0.533 | 0.933 | 32.113 | 0.532 | 0.932 | 32.127 | 0.54 | 0.946 |
| 70 | 32.056 | 0.5 | 0.875 | 32.062 | 0.503 | 0.88 | 32.061 | 0.503 | 0.88 | 32.075 | 0.511 | 0.894 |
| 69 | 32.004 | 0.47 | 0.823 | 32.01 | 0.473 | 0.829 | 32.01 | 0.473 | 0.829 | 32.024 | 0.481 | 0.843 |
| 68 | 31.958 | 0.443 | 0.777 | 31.959 | 0.444 | 0.778 | 31.959 | 0.444 | 0.778 | 31.974 | 0.453 | 0.793 |
| 67 | 31.91 | 0.416 | 0.729 | 31.909 | 0.415 | 0.728 | 31.909 | 0.416 | 0.728 | 31.925 | 0.425 | 0.744 |
| 66 | 31.85 | 0.382 | 0.669 | 31.859 | 0.387 | 0.678 | 31.86 | 0.388 | 0.679 | 31.876 | 0.397 | 0.695 |
| 65 | 31.8 | 0.354 | 0.619 | 31.81 | 0.359 | 0.629 | 31.812 | 0.36 | 0.631 | 31.828 | 0.369 | 0.646 |
| 64 | 31.747 | 0.323 | 0.566 | 31.762 | 0.332 | 0.581 | 31.764 | 0.333 | 0.583 | 31.78 | 0.342 | 0.599 |
| 63 | 31.701 | 0.297 | 0.52 | 31.714 | 0.304 | 0.533 | 31.716 | 0.306 | 0.535 | 31.733 | 0.315 | 0.552 |
| 62 | 31.648 | 0.267 | 0.467 | 31.666 | 0.277 | 0.485 | 31.67 | 0.279 | 0.488 | 31.686 | 0.288 | 0.505 |
| 61 | 31.607 | 0.243 | 0.426 | 31.62 | 0.25 | 0.439 | 31.623 | 0.252 | 0.442 | 31.64 | 0.262 | 0.459 |
| 60 | 31.558 | 0.215 | 0.377 | 31.574 | 0.224 | 0.392 | 31.577 | 0.226 | 0.396 | 31.594 | 0.236 | 0.413 |
| 59 | 31.508 | 0.187 | 0.327 | 31.528 | 0.198 | 0.347 | 31.531 | 0.2 | 0.35 | 31.548 | 0.21 | 0.367 |
| 58 | 31.473 | 0.167 | 0.292 | 31.482 | 0.172 | 0.301 | 31.486 | 0.174 | 0.305 | 31.503 | 0.184 | 0.322 |
| 57 | 31.43 | 0.142 | 0.249 | 31.437 | 0.146 | 0.256 | 31.441 | 0.148 | 0.26 | 31.458 | 0.158 | 0.277 |
| 56 | 31.39 | 0.119 | 0.209 | 31.393 | 0.121 | 0.212 | 31.396 | 0.123 | 0.215 | 31.413 | 0.133 | 0.232 |
| 55 | 31.351 | 0.097 | 0.17 | 31.348 | 0.096 | 0.167 | 31.352 | 0.098 | 0.171 | 31.369 | 0.107 | 0.188 |
| 54 | 31.313 | 0.075 | 0.132 | 31.304 | 0.07 | 0.123 | 31.308 | 0.072 | 0.126 | 31.324 | 0.082 | 0.143 |
| 53 | 31.263 | 0.047 | 0.082 | 31.26 | 0.045 | 0.079 | 31.263 | 0.047 | 0.082 | 31.28 | 0.057 | 0.099 |
| 52 | 31.213 | 0.019 | 0.032 | 31.216 | 0.02 | 0.035 | 31.22 | 0.022 | 0.038 | 31.236 | 0.031 | 0.055 |
| 51 | 31.174 | -0.004 | -0.007 | 31.172 | -0.005 | -0.009 | 31.176 | -0.003 | -0.005 | 31.192 | 0.006 | 0.011 |
| 50 | 31.128 | -0.03 | -0.053 | 31.128 | -0.03 | -0.053 | 31.132 | -0.028 | -0.049 | 31.148 | -0.019 | -0.033 |
|  |  |  |  | an $=31.181$, S.D | . $=1.751$, Pro | vision $=31$ | = \$1,000,000 |  |  |  |  |  |

## Dataset: Employers Liability: Forecast

Reserve Forecast Table
Model: PTF-g

Accident Period vs Development Period

|  |  |  |  |  |  |  |  | Acc | 7 | vs | elop | 10 |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cal. Per. Total | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |  | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 | Reserve | Ultimate |
| 1985 | 327,998 | 327,998 | 2,822,314 | 4,286,674 | 5,005,747 | 4,429,542 | 2,695,789 | 1,994,968 | 1,447,601 | 902,569 | 572,356 | 298,319 | 167,149 | 118,820 | 66,699 | 37,476 | 21,076 | 11,864 | 2,506 | 1,315 | 691 | 363 | 0 | 25,580,529 |
|  | 363,421 | 363,421 | 2,823,353 | 3,740,983 | 4,851,694 | 4,302,127 | 3,337,231 | 2,176,450 | 1,561,371 | 1,390,771 | 628,642 | 87,377 | 200,919 | 39,315 | 23,958 | 29,482 | 9,494 | 11,340 | 0 | 3,845 | 0 | 1,245 | 0 |  |
| 1986 | 3,137,155 | 314,841 | 2,709,078 | 4,14,650 | 4,804,828 | 4,251,711 | 3,082,343 | 2,882,197 | 1,389,616 | 866,408 | 549,417 | 307,566 | 172,337 | 122,512 | 68,775 | 38,644 | 21,734 | 4,588 | 2,405 | 1,262 | 663 | 374 | 4 | 23,570,755 |
|  | 3,165,803 | 342,451 | 2,787,749 | 3,771,543 | 4,460,001 | 4,101,358 | 2,822,690 | 2,145,689 | 1,196,979 | 827,302 | 537,642 | 372,893 | 48,850 | -7,546 | 123,409 | 51,765 | -13,445 | 215 | 0 | 237 | 600 | 523 | 523 | 523 |
| 1987 | 7,323,749 | 327,998 | 2,822,314 | 4,286,674 | 5,005,747 | 5,276,625 | 3,826,934 | 2,377,407 | 1,447,601 | 902,569 | 614,748 | 344,160 | 192,853 | 137,106 | 76,972 | 43,252 | 9,126 | 4,780 | 2,506 | 1,315 | 742 | 419 | 1,161 | 26,164,218 |
|  | 6,900,955 | 372,223 | 2,597,547 | 4,597,842 | 4,439,529 | 4,171,869 | 4,073,713 | 2,327,095 | 1,894,814 | 807,301 | 428,835 | 165,943 | 100,930 | 94,769 | 1,811 | 88,567 | 269 | 0 | 0 | 0 | 1,021 | 585 | 1,206 | 1,206 |
| 1988 | 12,304,967 | 362,256 | 3,117,370 | 4,735,250 | 6,584,540 | 6,941,315 | 4,224,345 | 2,624,591 | 1,598,294 | 1,070,395 | 729,103 | 408,206 | 228,757 | 162,641 | 91,313 | 19,252 | 10,076 | 5,278 | 2,767 | 1,560 | 880 | 497 | 2,937 | 2,732,711 |
|  | 11,547,134 | 326,349 | 3,172,571 | 5,176,011 | 5,379,296 | 6,387,889 | 4,521,772 | 2,811,563 | 1,323,411 | 1,445,179 | 311,985 | 562,611 | 405,192 | 31,637 | 434,136 | 433,892 | 7,580 | 2,729 | -4,029 | 2,120 | ,212 | 695 | 2,637 | 2,637 |
| 1989 | 16,990,391 | 351,978 | 3,028,922 | 5,478,566 | 7,621,744 | 6,744,373 | 4,104,489 | 2,550,125 | 1,667,891 | 1,117,061 | 760,927 | 426,045 | 238,766 | 169,766 | 35,773 | 18,706 | 9,790 | 5,128 | 2,888 | 1,628 | 919 | 519 | 5,954 | 89,355 |
|  | 16,930,957 | 398,416 | 3,384,938 | 6,261,944 | 7,232,344 | 6,056,747 | 4,172,737 | 2,638,001 | 1,999,745 | 1,280,583 | 518,592 | 432,369 | 47,274 | 15,294 | 16,795 | 199 | 4,548 | 22,876 | 3,879 | 2,213 | ,266 | 725 | 4,903 | 4,903 |
| 1990 | 20,097,157 | 379,738 | 3,890,728 | 7,039,194 | 8,219,945 | 7,273,483 | 4,426,533 | 2,953,994 | 1,932,258 | 1,294,265 | 881,725 | 493,729 | 276,725 | 73,872 | 38,592 | 20,180 | 10,561 | 5,943 | 3,347 | 1,887 | 1,065 | 602 | 12,844 | 103,200 |
|  | 20,777,343 | 338,276 | 3,981,902 | 7,269,814 | 7,273,202 | 6,764,962 | 4,663,626 | 3,778,364 | 1,786,086 | 2,035,300 | 966,986 | 143,383 | 21,026 | -117 | 20,750 | 28,458 | 18,336 | 7,899 | 4,497 | 2,566 | 1,468 | 84 | 10,003 | 10,003 |
| 1991 | 26,697,014 | 389,243 | 3,989,913 | 6,059,969 | 7,076,465 | 6,261,666 | 4,093,134 | 2,731,640 | 1,786,901 | 1,196,962 | 815,477 | 456,656 | 96,124 | 63,595 | 33,224 | 17,372 | 9,767 | 5,497 | 3,096 | 1,746 | 985 | 55 | 21,647 | 3,798,376 |
|  | 25,151,751 | 357,601 | 3,930,777 | 6,254,476 | 7,270,856 | 5,543,963 | 4,385,023 | 2,271,060 | 2,163,980 | 1,131,551 | 314,255 | -113,095 | -18,502 | 267,189 | 15,518 | 2,076 | 12,863 | 7,307 | 4,160 | 2,374 | 1,358 | 779 | 16,221 | 16,221 |
| 1992 | 33,503,376 | 354,477 | 3,050,078 | 4,632,526 | 5,409,584 | 5,141,467 | 3,361,048 | 2,243,179 | 1,467,447 | 983,024 | 669,757 | 140,906 | 73,482 | 48,615 | 25,398 | 14,266 | 8,021 | 4,514 | 2,543 | 1,434 | 809 | 425 | 32,014 | 27,191,877 |
|  | 33,007,909 | 406,312 | 3,485,198 | 4,787,544 | 6,016,700 | 4,772,276 | 3,259,134 | 1,978,110 | 1,388,877 | 645,655 | 305,919 | 90,952 | 22,476 | 12,200 | -11,490 | 18,641 | 10,565 | 6,002 | 3,418 | 1,951 | 1,116 | 59 | 23,349 | 23,349 |
| 1993 | 33,263,505 | 295,203 | 2,540,061 | 3,857,901 | 4,838,880 | 4,599,349 | 3,006,808 | 2,006,857 | 1,312,915 | 879,548 | 225,216 | 117,345 | 61,195 | 40,486 | 22,722 | 12,764 | 7,177 | 4,039 | 2,276 | 1,283 | 674 | 354 | 51,289 | 24,964,549 |
|  | 32,793,140 | 286,900 | 2,500,351 | 4,252,026 | 5,428,687 | 4,340,034 | 3,008,771 | 2,731,563 | 1,181,851 | 728,782 | 371,273 | 68,942 | 10,783 | 3,295 | 29,496 | 16,680 | 9,455 | 5,373 | 3,060 | 1,747 | 930 | 49 | 36,675 | 36,67 |
| 1994 | 31,473,817 | 335,836 | 2,889,680 | 4,714,267 | 5,913,229 | 5,620,882 | 3,674,814 | 2,452,831 | 1,604,758 | 404,177 | 256,216 | 133,496 | 69,618 | 49,479 | 27,770 | 15,601 | 8,773 | 4,938 | 2,782 | 1,460 | 767 | 403 | 111,97 | 27,946,374 |
|  | 32,018,107 | 359,335 | 2,940,485 | 4,749,103 | 5,360,802 | 5,962,373 | 4,131,260 | 2,780,882 | 1,056,343 | 386,307 | 226,250 | -15,475 | -103,263 | 63,905 | 36,056 | 20,392 | 11,561 | 6,571 | 3,743 | 1,987 | 1,058 | 56 | 78,900 | 78,90 |
| 1995 | 29,114,365 | 370,278 | 3,422,255 | 5,583,414 | 7,003,692 | 6,657,867 | 4,352,989 | 2,905,638 | 715,087 | 445,629 | 282,493 | 147,187 | 82,457 | 58,608 | 32,896 | 18,482 | 10,393 | 5,850 | 3,067 | 1,610 | 845 | 444 | 214,653 | 29,664,117 |
|  | 29,082,604 | 272,072 | 3,351,803 | 5,088,547 | 7,048,157 | 6,656,030 | 3,895,374 | 1,595,575 | 794,846 | 513,843 | 47,316 | 185,902 | 66,637 | 75,707 | 42,721 | 24,166 | 13,703 | 7,789 | 4,127 | 2,191 | 1,166 | 622 | 116,174 | 116,174 |
| 1996 | 29,365,438 | 373,692 | 3,453,624 | 5,634,678 | 7,067,685 | 6,718,980 | 4,392,885 | 1,103,615 | 671,648 | 418,554 | 265,331 | 148,512 | 83,205 | 59,144 | 33,199 | 18,653 | 10,490 | 5,495 | 2,881 | 1,512 | 794 | 417 | 364,303 | 2,941,398 |
|  | 31,282,166 | 369,338 | 3,373,646 | 5,580,846 | 7,733,248 | 9,576,139 | 4,414,667 | 466,863 | 777,454 | 204,599 | 80,294 | 119,554 | 67,238 | 76,404 | 43,122 | 24,397 | 13,836 | 7,315 | 3,876 | 2,058 | 1,095 | 584 | 169,269 | 169,26 |
| 1997 | 30,508,236 | 452,194 | 4,179,321 | 6,819,032 | 8,553,572 | 8,132,086 | 2,002,289 | 1,243,249 | 756,628 | 471,512 | 321,098 | 179,737 | 100,705 | 71,587 | 40,186 | 22,580 | 11,818 | 6,190 | 3,246 | 1,703 | 895 | 470 | 760,216 | 2,512,542 |
|  | 27,665,268 | 339,282 | 3,732,991 | 5,227,094 | 9,228,797 | 8,435,768 | 2,888,546 | 1,135,434 | 547,640 | 216,773 | 258,142 | 144,724 | 81,411 | 92,511 | 52,220 | 29,549 | 15,587 | 8,240 | 4,366 | 2,318 | 1,234 | 658 | 332,39 | 332,39 |
| 1998 | 32,778,418 | 457,495 | 4,228,506 | 6,899,649 | 8,655,029 | 3,100,593 | 1,885,782 | 1,170,908 | 712,602 | 477,052 | 324,891 | 181,872 | 101,907 | 72,446 | 40,671 | 21,266 | 11,130 | 5,830 | 3,057 | 1,604 | 842 | 443 | 1,243,012 | 9,703,619 |
|  | 32,987,794 | 278,722 | 3,494,212 | 5,806,965 | 11,247,156 | 10,692,710 | 4,906,059 | 1,673,282 | 361,501 | 158,936 | 261,252 | 146,499 | 82,427 | 93,662 | 52,878 | 27,830 | 14,680 | 7,761 | 4,112 | 2,184 | 1,162 | 620 | 377,154 | 377,154 |
| 1999 | 36,465,760 | 1,566,708 | 14,482,146 | 23,629,812 | 11,083,082 | 9,807,436 | 5,968,885 | 3,708,648 | 2,426,272 | 1,625,458 | 1,107,499 | 620,249 | 347,698 | 247,291 | 129,242 | 67,606 | 35,396 | 18,549 | 9,729 | 5,107 | 2,684 | 1,41 | 6,644,190 | 47,799,907 |
|  | 34,807,630 | 261,587 | 3,279,725 | 6,885,999 | 10,432,638 | 11,101,285 | 6,269,954 | 2,925,429 | 572,672 | 535,192 | 888,941 | 499,279 | 281,364 | 320,036 | 168,220 | 88,623 | 46,794 | 24,763 | 13,133 | 6,981 | 3,718 | 1,98 | 1,420,868 | 1,42,868 |
| 2000 | 49,001,211 | 1,397,800 | 12,921,401 | 7,882,100 | 9,204,080 | 8,144,705 | 4,956,933 | 3,308,599 | 2,164,683 | 1,450,298 | 988,214 | 553,478 | 310,286 | 205,366 | 107,331 | 56,144 | 29,395 | 15,404 | 8,079 | 4,241 | 2,229 | 1,172 | 9,204,920 | 42,662,303 |
|  | 37,162,551 | 167,544 | 3,509,353 | 6,962,743 | 9,221,279 | 9,569,081 | 4,027,383 | 772,295 | 511,812 | 478,261 | 793,627 | 445,841 | 251,302 | 265,778 | 139,700 | 73,598 | 38,861 | 20,565 | 10,907 | 5,797 | 3,088 | 1,648 | 1,512,375 | 1,512,375 |
| 2001 | 66,251,939 | 1,389,714 | 4,802,771 | 7,293,884 | 8,517,209 | 7,536,891 | 4,927,636 | 3,289,244 | 2,152,151 | 1,441,989 | 982,613 | 550,374 | 287,131 | 190,040 | 99,321 | 51,954 | 27,201 | 14,254 | 7,476 | 3,925 | 2,062 | 1,085 | 14,028,458 | 39,676,043 |
|  | 38,780,673 | 158,189 | 2,663,632 | 6,601,908 | 8,424,381 | 7,799,475 | 536,710 | 769,132 | 510,309 | 476,553 | 789,680 | 443,717 | 232,548 | 245,943 | 129,275 | 68,105 | 35,961 | 19,030 | 10,093 | 5,365 | 2,857 | 1,525 | 1,636,505 | 1,636,50 |
|  | 32,258,341 | 557,264 | 4,794,893 | 7,281,920 | 8,503,238 | 8,083,287 | 5,285,193 | 3,528,131 | 2,308,596 | 1,546,904 | 1,054,169 | 549,471 | 286,660 | 189,728 | 99,158 | 51,869 | 27,157 | 14,231 | 7,464 | 3,918 | 2,059 | 1,083 | 23,039,079 | 42,336,220 |
|  | 36,338,872 | 155,809 | 3,402,607 | 7,405,743 | 8,332,982 | 884,791 | 580,119 | 827,409 | 549,591 | 512,646 | 847,908 | 442,990 | 232,167 | 245,540 | 129,063 | 67,994 | 35,902 | 18,999 | 10,076 | 5,356 | 2,853 | 1,523 | 2,022,154 | 2,022,15 |
| 2003 | 36,334,581 | 401,468 | 3,454,371 | 5,246,093 | 6,580,867 | 6,256,237 | 4,090,840 | 2,731,007 | 1,787,114 | 1,197,552 | 759,452 | 395,854 | 206,517 | 136,685 | 71,436 | 37,368 | 19,565 | 10,252 | 5,377 | 2,823 | 1,483 | 780 | 24,291,210 | 33,812,510 |
|  | 38,753,674 | 476,574 | 3,196,497 | 5,848,228 | 721,461 | 690,062 | 454,725 | 643,081 | 427,612 | 398,210 | 610,856 | 319,142 | 167,259 | 176,894 | 92,980 | 48,984 | 25,865 | 13,687 | 7,259 | 3,858 | 2,055 | 1,097 | 1,780,645 | 1,780,64 |
| 2004 | 36,625,332 | 313,423 | 2,696,799 | 4,399,712 | 5,519,476 | 5,247,523 | 3,431,468 | 2,290,956 | 1,499,245 | 934,919 | 592,898 | 309,040 | 161,226 | 106,709 | 55,770 | 29,173 | 15,274 | 8,004 | 4,198 | 2,204 | 1,158 | 609 | 24,609,562 | 27,825,576 |
|  | 37,600,773 | 203,281 | 3,012,732 | 482,820 | 609,733 | 586,081 | 388,030 | 542,265 | 360,943 | 310,879 | 476,890 | 249,151 | 130,578 | 138,099 | 72,589 | 38,242 | 20,192 | 10,686 | 5,667 | 3,012 | 1,604 | 85 | 1,631,864 | 1,631,864 |
| 2005 | 34,703,794 | 284,888 | 2,633,305 | 4,296,386 | 5,390,180 | 5,124,909 | 3,351,493 | 2,237,698 | 1,362,751 | 849,802 | 538,919 | 280,904 | 146,548 | 96,994 | 50,692 | 26,517 | 13,883 | 7,275 | 3,816 | 2,003 | 1,053 | 554 | 26,415,683 | 26,688,881 |
|  | 32,991,052 | 273,198 | 291,649 | 475,084 | 602,916 | 582,197 | 387,120 | 532,990 | 328,082 | 282,576 | 433,473 | 226,468 | 118,690 | 125,527 | 65,980 | 34,760 | 18,354 | 9,713 | 5,151 | 2,738 | 1,458 | 77 | 1,687,389 | 1,687,389 |
| $\begin{aligned} & \text { Cal. Per. } \\ & \text { Total } \end{aligned}$ | Total Fitted/Paid |  | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | 2023 | 2024 | 2025 | Total Reserve | Total Ultimate |
|  | 598,526,543 |  | 33,496,540 | 29,143,446 | 23,374,930 | 16,890,292 | 11,012,169 | 7,181,639 | 4,272,538 | 2,513,123 | 1,446,903 | 783,871 | 431,813 | 246,042 | 128,505 | 67,072 | 34,417 | 17,440 | 8,586 | 3,941 | 1,662 | 554 | 131,055,481 | 691,165,059 |
|  | 560,109,578 |  | 1,899,416 | 1,828,862 | 1,888,870 | 1,682,334 | 1,463,129 | 1,366,430 | 1,008,458 | 767,469 | 616,515 | 361,506 | 227,152 | 163,213 | 86,629 | 46,120 | 24,521 | 13,015 | 6,851 | 3,508 | 1,737 | 778 | 6,580,520 | 6,580,520 |

Inside the table the blues are the observed, the black are means of log-normal distributions and the reds are the corresponding Std Devs of the log
normals.
Given that on the log-scale we have normal distributions in every cell about the trend structure, the corresponding distributions on the dollar scale
The predictive log-normal distributions include parameter uncertainty and process variability.


In order to find distributions of aggregates by accident year, calendar year and total of all years, we simulate from the predictive log-normal
distributions of each cell, incorporating their correlations.
We do this because there is no analytical form for the sum of log-normals.
However, note that the means and Std Devs of the distributions of aggregates are computed via the model using second order moments, including
covariances.
The sample quantiles (percentiles) below are based on 10,000 simulations of the forecast triangle.
The Kernel is fitted to the sample frequency plot whereas the log-normal and gamma are just based on the means and std. devs. of the actual
forecast distributions.
Note that the distributions are slightly skewed.

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \#S.D.'s | V-a-R | Quantile | \#S.D.'s | V-a-R |
| 99.995 | 159.428 | 4.312 | 28.372 | 161.013 | 4.552 | 29.957 | 159.11 | 4.263 | 28.054 | 158.228 | 4.129 | 27.173 |
| 99.99 | 157.901 | 4.08 | 26.846 | 159.712 | 4.355 | 28.657 | 157.746 | 4.056 | 26.691 | 156.953 | 3.935 | 25.897 |
| 99.98 | 157.823 | 4.068 | 26.768 | 158.143 | 4.116 | 27.087 | 156.337 | 3.842 | 25.281 | 155.63 | 3.734 | 24.574 |
| 99.97 | 156.612 | 3.884 | 25.557 | 157.161 | 3.967 | 26.106 | 155.488 | 3.713 | 24.433 | 154.831 | 3.613 | 23.776 |
| 99.96 | 155.929 | 3.78 | 24.873 | 156.376 | 3.848 | 25.32 | 154.875 | 3.62 | 23.819 | 154.253 | 3.525 | 23.197 |
| 99.95 | 154.8 | 3.608 | 23.745 | 155.73 | 3.75 | 24.674 | 154.392 | 3.546 | 23.336 | 153.797 | 3.456 | 22.741 |
| 99.94 | 154.564 | 3.572 | 23.508 | 155.205 | 3.67 | 24.149 | 153.992 | 3.486 | 22.937 | 153.419 | 3.398 | 22.364 |
| 99.93 | 154.201 | 3.517 | 23.146 | 154.774 | 3.604 | 23.719 | 153.651 | 3.434 | 22.595 | 153.096 | 3.349 | 22.041 |
| 99.92 | 153.971 | 3.482 | 22.915 | 154.424 | 3.551 | 23.368 | 153.352 | 3.388 | 22.297 | 152.814 | 3.306 | 21.758 |
| 99.91 | 153.801 | 3.456 | 22.745 | 154.124 | 3.506 | 23.068 | 153.087 | 3.348 | 22.032 | 152.562 | 3.268 | 21.507 |
| 99.9 | 153.21 | 3.367 | 22.155 | 153.864 | 3.466 | 22.808 | 152.848 | 3.312 | 21.793 | 152.336 | 3.234 | 21.28 |
| 99.8 | 151.838 | 3.158 | 20.783 | 152.276 | 3.225 | 21.22 | 151.231 | 3.066 | 20.175 | 150.799 | 3 | 19.743 |
| 99.7 | 151.185 | 3.059 | 20.13 | 151.336 | 3.082 | 20.28 | 150.245 | 2.916 | 19.189 | 149.859 | 2.857 | 18.804 |
| 99.6 | 150.313 | 2.926 | 19.258 | 150.614 | 2.972 | 19.558 | 149.525 | 2.807 | 18.469 | 149.172 | 2.753 | 18.116 |
| 99.5 | 149.797 | 2.848 | 18.742 | 150.014 | 2.881 | 18.959 | 148.954 | 2.72 | 17.899 | 148.626 | 2.67 | 17.57 |
| 99.4 | 149.328 | 2.777 | 18.272 | 149.494 | 2.802 | 18.439 | 148.479 | 2.648 | 17.423 | 148.171 | 2.601 | 17.115 |
| 99.3 | 148.859 | 2.706 | 17.804 | 149.035 | 2.732 | 17.98 | 148.071 | 2.586 | 17.015 | 147.779 | 2.541 | 16.724 |
| 99.2 | 148.435 | 2.641 | 17.379 | 148.624 | 2.67 | 17.568 | 147.712 | 2.531 | 16.656 | 147.435 | 2.489 | 16.379 |
| 99.1 | 148.155 | 2.598 | 17.099 | 148.256 | 2.614 | 17.201 | 147.391 | 2.482 | 16.336 | 147.127 | 2.442 | 16.072 |
| 99 | 147.576 | 2.511 | 16.521 | 147.92 | 2.563 | 16.864 | 147.101 | 2.438 | 16.046 | 146.848 | 2.4 | 15.793 |
| 98 | 145.333 | 2.17 | 14.277 | 145.675 | 2.222 | 14.619 | 145.103 | 2.135 | 14.047 | 144.922 | 2.107 | 13.866 |
| 97 | 144.227 | 2.002 | 13.171 | 144.376 | 2.024 | 13.321 | 143.849 | 1.944 | 12.793 | 143.708 | 1.923 | 12.653 |
| 96 | 143.289 | 1.859 | 12.234 | 143.407 | 1.877 | 12.351 | 142.912 | 1.802 | 11.857 | 142.8 | 1.785 | 11.745 |
| 95 | 142.514 | 1.741 | 11.459 | 142.587 | 1.752 | 11.532 | 142.155 | 1.687 | 11.1 | 142.064 | 1.673 | 11.009 |
| 94 | 141.726 | 1.621 | 10.67 | 141.869 | 1.643 | 10.814 | 141.514 | 1.589 | 10.458 | 141.44 | 1.578 | 10.384 |
| 93 | 141.096 | 1.526 | 10.04 | 141.237 | 1.547 | 10.182 | 140.954 | 1.504 | 9.898 | 140.893 | 1.495 | 9.838 |
| 92 | 140.536 | 1.441 | 9.481 | 140.682 | 1.463 | 9.626 | 140.454 | 1.428 | 9.399 | 140.406 | 1.421 | 9.35 |
| 91 | 139.982 | 1.356 | 8.926 | 140.186 | 1.387 | 9.13 | 140.001 | 1.359 | 8.946 | 139.963 | 1.354 | 8.907 |
| 90 | 139.6 | 1.299 | 8.545 | 139.739 | 1.32 | 8.683 | 139.586 | 1.296 | 8.53 | 139.556 | 1.292 | 8.501 |
| 89 | 139.137 | 1.228 | 8.082 | 139.33 | 1.257 | 8.275 | 139.201 | 1.238 | 8.145 | 139.179 | 1.234 | 8.12 |
| 88 | 138.795 | 1.176 | 7.74 | 138.954 | 1.2 | 7.899 | 138.841 | 1.183 | 7.785 | 138.826 | 1.181 | 7.771 |
| 87 | 138.494 | 1.13 | 7.438 | 138.602 | 1.147 | 7.547 | 138.502 | 1.132 | 7.447 | 138.494 | 1.13 | 7.439 |
| 86 | 138.193 | 1.085 | 7.138 | 138.271 | 1.096 | 7.215 | 138.182 | 1.083 | 7.127 | 138.18 | 1.083 | 7.125 |
| 85 | 137.906 | 1.041 | 6.85 | 137.956 | 1.049 | 6.9 | 137.878 | 1.037 | 6.823 | 137.881 | 1.037 | 6.826 |
| 84 | 137.608 | 0.996 | 6.552 | 137.656 | 1.003 | 6.6 | 137.58 | 0.993 | 6.532 | 137.596 | 0.994 | 6.5 |
| 83 | 137.279 | 0.946 | 6.223 | 137.369 | 0.959 | 6.314 | 137.31 | 0.95 | 6.254 | 137.322 | 0.952 | 6.266 |
| 82 | 136.964 | 0.898 | 5.908 | 137.094 | 0.918 | 6.039 | 137.042 | 0.91 | 5.987 | 137.059 | 0.912 | 6.003 |
| 81 | 136.682 | 0.855 | 5.626 | 136.83 | 0.877 | 5.774 | 136.785 | 0.871 | 5.729 | 136.805 | 0.874 | 5.749 |
| 80 | 136.44 | 0.818 | 5.385 | 136.575 | 0.839 | 5.52 | 136.536 | 0.833 | 5.48 | 136.559 | 0.836 | 5.50 |
| 79 | 136.213 | 0.784 | 5.158 | 136.329 | 0.801 | 5.274 | 136.295 | 0.796 | 5.239 | 136.321 | 0.8 | 5.266 |
| 78 | 135.997 | 0.751 | 4.942 | 136.091 | 0.765 | 5.036 | 136.06 | 0.761 | 5.005 | 136.09 | 0.765 | 5.035 |
| 77 | 135.809 | 0.722 | 4.754 | 135.86 | 0.73 | 4.804 | 135.833 | 0.726 | 4.777 | 135.865 | 0.731 | 4.81 |
| 76 | 135.6 | 0.691 | 4.544 | 135.635 | 0.696 | 4.579 | 135.611 | 0.692 | 4.555 | 135.646 | 0.698 | 4.591 |
| 75 | 135.392 | 0.659 | 4.337 | 135.415 | 0.662 | 4.359 | 135.394 | 0.659 | 4.339 | 135.432 | 0.665 | 4.376 |
| 74 | 135.172 | 0.626 | 4.117 | 135.199 | 0.63 | 4.144 | 135.183 | 0.627 | 4.127 | 135.223 | 0.633 | 4.167 |
| 73 | 134.958 | 0.593 | 3.903 | 134.988 | 0.598 | 3.933 | 134.976 | 0.596 | 3.92 | 135.017 | 0.602 | 3.962 |
| 72 | 134.722 | 0.557 | 3.666 | 134.781 | 0.566 | 3.726 | 134.773 | 0.565 | 3.717 | 134.816 | 0.572 | 3.761 |
| 71 | 134.52 | ${ }_{0}^{0.526}$ | 3.464 | 134.577 | 0.535 | 3.522 | 134.574 <br> 13478 | 0.535 | 3.518 | 134.619 | 0.542 | 3.563 3.369 |
| 70 | 134.325 | 0.497 | 3.269 | 134.377 | 0.505 | 3.322 | 134.378 | 0.505 | 3.322 | 134.425 | 0.512 | 3.369 |

## Quantiles and VaR Mean $=131.055$, S.D. $=6.581$, Provision $=131.055,1$ Unit $=\$ 1,000,000$

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R |
| 69 | 134.136 | 0.468 | 3.081 | 134.18 | 0.475 | 3.125 | 134.185 | 0.476 | 3.13 | 134.234 | 0.483 | 3.178 |
| 68 | 133.936 | 0.438 | 2.881 | 133.986 | 0.445 | 2.931 | 133.996 | 0.447 | 2.94 | 134.046 | 0.454 | 2.99 |
| 67 | 133.714 | 0.404 | 2.659 | 133.795 | 0.416 | 2.739 | 133.809 | 0.418 | 2.754 | 133.86 | 0.426 | 2.805 |
| 66 | 133.545 | 0.378 | 2.49 | 133.607 | 0.388 | 2.551 | 133.625 | 0.39 | 2.569 | 133.677 | 0.398 | 2.622 |
| 65 | 133.378 | 0.353 | 2.323 | 133.421 | 0.36 | 2.366 | 133.443 | 0.363 | 2.388 | 133.496 | 0.371 | 2.441 |
| 64 | 133.19 | 0.324 | 2.134 | 133.238 | 0.332 | 2.183 | 133.263 | 0.336 | 2.208 | 133.317 | 0.344 | 2.262 |
| 63 | 133.005 | 0.296 | 1.949 | 133.057 | 0.304 | 2.001 | 133.086 | 0.308 | 2.03 | 133.14 | 0.317 | 2.085 |
| 62 | 132.847 | 0.272 | 1.792 | 132.877 | 0.277 | 1.822 | 132.91 | 0.282 | 1.854 | 132.965 | 0.29 | 1.909 |
| 61 | 132.69 | 0.248 | 1.635 | 132.7 | 0.25 | 1.645 | 132.735 | 0.255 | 1.68 | 132.791 | 0.264 | 1.736 |
| 60 | 132.544 | 0.226 | 1.488 | 132.524 | 0.223 | 1.469 | 132.562 | 0.229 | 1.507 | 132.619 | 0.238 | 1.563 |
| 59 | 132.366 | 0.199 | 1.311 | 132.35 | 0.197 | 1.294 | 132.391 | 0.203 | 1.336 | 132.448 | 0.212 | 1.392 |
| 58 | 132.187 | 0.172 | 1.132 | 132.176 | 0.17 | 1.121 | 132.221 | 0.177 | 1.165 | 132.278 | 0.186 | 1.222 |
| 57 | 131.993 | 0.142 | 0.938 | 132.004 | 0.144 | 0.949 | 132.052 | 0.151 | 0.996 | 132.109 | 0.16 | 1.053 |
| 56 | 131.796 | 0.112 | 0.74 | 131.833 | 0.118 | 0.778 | 131.884 | 0.126 | 0.828 | 131.941 | 0.135 | 0.885 |
| 55 | 131.635 | 0.088 | 0.58 | 131.663 | 0.092 | 0.608 | 131.717 | 0.1 | 0.661 | 131.774 | 0.109 | 0.718 |
| 54 | 131.447 | 0.059 | 0.391 | 131.495 | 0.067 | 0.439 | 131.55 | 0.075 | 0.495 | 131.607 | 0.084 | 0.552 |
| 53 | 131.296 | 0.037 | 0.241 | 131.327 | 0.041 | 0.271 | 131.385 | 0.05 | 0.329 | 131.441 | 0.059 | 0.386 |
| 52 | 131.136 | 0.012 | 0.08 | 131.16 | 0.016 | 0.105 | 131.22 | 0.025 | 0.164 | 131.276 | 0.033 | 0.22 |
| 51 | 130.99 | -0.01 | -0.065 | 130.995 | -0.009 | -0.061 | 131.055 | 0 | -0.001 | 131.11 | 0.008 | 0.055 |
| 50 | 130.836 | -0.033 | -0.219 | 130.83 | -0.034 | -0.226 | 130.891 | -0.025 | -0.165 | 130.945 | -0.017 | -0.11 |


| Quantile Statistics and Value at Risk (Cal. Yr: 2006) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \#S.D.'s | V-a-R | Quantile |  | $V$-a-R | Quantile | \#S.D.'s | V-a-R | Quantile | \#S.D.'s | $V-a-R$ |
| 99.95 | 40.089 | 3.471 | 6.592 | 40.21 | 3.535 | 6.714 | 40.297 | 3.581 | 6.801 | 40.101 | ${ }^{3.477}$ | 6.605 |
| 99.9 | 39.58 | 3.203 | 6.084 | 39.814 | 3.326 | 6.317 | 39.843 | 3.341 | 6.346 | 39.674 | 3.252 | 6.178 |
| 99.8 | 39.291 | 3.05 | 5.794 | 39.437 | 3.128 | 5.941 | 39.367 | 3.091 | 5.871 | 39.225 | 3.016 | 5.729 |
| 99.7 | 39.12 | 2.96 | 5.623 | 39.207 | 3.007 | 5.711 | 39.077 | 2.938 | 5.581 | 38.951 | 2.872 | 5.455 |
| 99.6 | 38.984 | 2.889 | 5.488 | 39.029 | 2.913 | 5.533 | 38.866 | 2.827 | 5.37 | 38.75 | 2.766 | 5.254 |
| 99.5 | 38.816 | 2.801 | 5.32 | 38.879 | 2.834 | 5.383 | 38.699 | 2.739 | 5.202 | 38.591 | 2.682 | 5.095 |
| 99.4 | 38.683 | 2.731 | 5.186 | 38.747 | 2.764 | 5.251 | 38.559 | 2.665 | 5.063 | 38.458 | 2.612 | 4.962 |
| 99.3 | 38.567 | 2.67 | 5.071 | 38.629 | 2.702 | 5.133 | 38.44 | 2.602 | 4.943 | 38.344 | 2.552 | 4.848 |
| 99.2 | 38.445 | 2.605 | 4.949 | 38.522 | 2.646 | 5.026 | 38.334 | 2.547 | 4.838 | 38.244 | 2.499 | 4.747 |
| 99.1 | 38.346 | 2.553 | 4.849 | 38.424 | 2.594 | 4.928 | 38.241 | 2.498 | 4.744 | 38.154 | 2.452 | 4.658 |
| 99 | 38.276 | 2.516 | 4.78 | 38.335 | 2.548 | 4.839 | 38.156 | 2.453 | 4.659 | 38.073 | 2.409 | 4.576 |
| 98 | 37.607 | 2.164 | 4.111 | 37.697 | 2.212 | 4.201 | 37.571 | 2.145 | 4.074 | 37.512 | 2.114 | 4.015 |
| 97 | 37.234 | 1.967 | 3.737 | 37.284 | 1.994 | 3.788 | 37.204 | 1.952 | 3.708 | 37.159 | 1.928 | 3.662 |
| 96 | 36.924 | 1.804 | 3.427 | 36.978 | 1.833 | 3.482 | 36.931 | 1.808 | 3.434 | 36.895 | 1.789 | 3.398 |
| 95 | 36.672 | 1.672 | 3.175 | 36.739 | 1.707 | 3.242 | 36.71 | 1.692 | 3.214 | 36.681 | 1.676 | 3.184 |
| 94 | 36.486 | 1.574 | 2.99 | 36.543 | 1.604 | 3.046 | 36.523 | 1.593 | 3.027 | 36.499 | 1.581 | 3.003 |
| 93 | 36.312 | 1.482 | 2.815 | 36.376 | 1.516 | 2.879 | 36.36 | 1.508 | 2.864 | 36.341 | 1.497 | 2.844 |
| 92 | 36.182 | 1.414 | 2.686 | 36.229 | 1.439 | 2.733 | 36.215 | 1.431 | 2.718 | 36.199 | 1.423 | 2.703 |
| 91 | 36.069 | 1.355 | 2.573 | 36.098 | 1.369 | 2.601 | 36.083 | 1.362 | 2.586 | 36.071 | 1.355 | 2.574 |
| 90 | 35.933 | 1.283 | 2.436 | 35.978 | 1.306 | 2.481 | 35.962 | 1.298 | 2.465 | 35.953 | 1.293 | 2.456 |
| 89 | 35.824 | 1.225 | 2.328 | 35.866 | 1.247 | 2.369 | 35.85 | 1.239 | 2.353 | 35.843 | 1.235 | 2.347 |
| 88 | 35.728 | 1.175 | 2.231 | 35.762 | 1.193 | 2.265 | 35.745 | 1.184 | 2.249 | 35.741 | 1.182 | 2.244 |
| 87 | 35.633 | 1.125 | 2.137 | 35.663 | 1.141 | 2.167 | 35.647 | 1.132 | 2.15 | 35.645 | 1.131 | 2.148 |
| 86 | 35.539 | 1.076 | 2.043 | 35.57 | 1.092 | 2.074 | 35.554 | 1.083 | 2.057 | 35.553 | 1.083 | 2.057 |
| 85 | 35.461 | 1.034 | 1.964 | 35.481 | 1.045 | 1.985 | 35.466 | 1.037 | 1.969 | 35.467 | 1.037 | 1.97 |
| 84 | 35.372 | 0.987 | 1.875 | 35.396 | 1 | 1.9 | 35.381 | 0.992 | 1.885 | 35.384 | 0.994 | 1.887 |
| 83 | 35.289 | 0.944 | 1.793 | 35.314 | 0.957 | 1.818 | 35.3 | 0.95 | 1.804 | 35.305 | 0.952 | 1.808 |
| 82 | 35.205 | 0.899 | 1.708 | 35.236 | 0.916 | 1.739 | 35.223 | 0.909 | 1.726 | 35.228 | 0.912 | 1.732 |
| 81 | 35.139 | 0.865 | 1.642 | 35.16 | 0.876 | 1.663 | 35.148 | 0.87 | 1.652 | 35.155 | 0.873 | 1.658 |
| 80 | 35.067 | 0.827 | 1.57 | 35.086 | 0.837 | 1.59 | 35.076 | 0.831 | 1.579 | 35.084 | 0.836 | 1.587 |
| 79 | 34.989 | 0.786 | 1.492 | 35.015 | 0.799 | 1.519 | 35.006 | 0.795 | 1.509 | 35.015 | 0.799 | 1.518 |
| 78 | 34.92 | 0.749 | 1.424 | 34.946 | 0.763 | 1.45 | 34.938 | 0.759 | 1.442 | 34.948 | 0.764 | 1.451 |
| 77 | 34.867 | 0.721 | 1.37 | 34.879 | 0.728 | 1.382 | 34.872 | 0.724 | 1.376 | 34.883 | 0.73 | 1.386 |
| 76 | 34.804 | 0.688 | 1.307 | 34.814 | 0.693 | 1.317 | 34.808 | 0.69 | 1.311 | 34.819 | 0.696 | 1.323 |
| 75 | 34.722 | 0.645 | 1.225 | 34.75 | 0.66 | 1.254 | 34.745 | 0.657 | 1.248 | 34.757 | 0.664 | 1.261 |
| 74 | 34.663 | 0.614 | 1.166 | 34.688 | 0.627 | 1.192 | 34.684 | 0.625 | 1.187 | 34.697 | 0.632 | 1.2 |
| 73 | 34.603 | 0.582 | 1.106 | 34.627 | 0.595 | 1.131 | 34.624 | 0.593 | 1.127 | 34.637 | 0.601 | 1.141 |
| 72 | 34.557 | 0.558 | 1.06 | 34.568 | 0.564 | 1.072 | 34.565 | 0.562 | 1.068 | 34.579 | 0.57 | 1.083 |
| 71 | 34.501 | 0.529 | 1.004 | 34.51 | 0.534 | 1.013 | 34.507 | 0.532 | 1.011 | 34.522 | 0.54 | 1.026 |
| 70 | 34.438 | 0.496 | 0.942 | 34.453 | 0.503 | 0.956 | 34.451 | 0.502 | 0.954 | 34.466 | 0.51 | 0.969 |
| 69 | 34.376 | 0.463 | 0.879 | 34.397 | 0.474 | 0.9 | 34.395 | 0.473 | 0.898 | 34.411 | 0.481 | 0.914 |
| 68 | ${ }^{34.322}$ | 0.434 | 0.825 | ${ }^{34.341}$ | 0.445 | 0.845 | 34.34 | 0.444 | 0.844 | ${ }^{34.356}$ | 0.453 |  |
| 67 66 | 34.274 34.222 | 0.409 0.382 | 0.777 | 34.287 34.234 | 0.416 0.388 | 0.791 0.737 | 34.286 34.233 | 0.416 0.388 | 0.789 0.736 | $\begin{array}{r}34.303 \\ 34.25 \\ \hline\end{array}$ | 0.424 0 | ${ }^{0.806}$ |
| 66 65 | 34.222 34.173 | 0.382 0.356 | 0.725 0.676 | 34.234 34.181 | $\begin{array}{r}0.388 \\ 0.36 \\ \hline\end{array}$ | 0.684 | 34.233 34.18 | 0.388 <br> 0.36 | 0.736 0.684 | 34.25 34.197 | 0.397 0.369 | 0.753 |
| 64 | 34.125 | ${ }_{0}^{0.331}$ | 0.628 | 34.129 | 0.333 | 0.632 | 34.128 | 0.333 | 0.632 | ${ }_{34.146}$ | 0.342 | 0.649 |
| 63 | 34.07 | 0.302 | 0.574 | 34.077 | 0.306 | 0.58 | 34.077 | 0.305 | 0.58 | 34.095 | 0.315 | 0.598 |
| 62 | 34.02 | 0.275 | 0.523 | 34.026 | 0.279 | 0.529 | 34.026 | 0.279 | 0.529 | 34.044 | 0.288 | 0.547 |
| 61 | 33.963 | 0.245 | 0.466 | 33.975 | 0.252 | 0.479 | 33.975 | 0.252 | 0.479 | 33.994 | 0.262 | 0.497 |
| 60 | 33.907 | 0.216 | 0.41 | 33.925 | 0.226 | 0.428 | 33.926 | 0.226 | 0.429 | 33.944 | 0.236 | 0.447 |
| 59 | 33.868 | 0.195 | 0.371 | 33.875 | 0.199 | 0.379 | 33.876 | 0.2 | 0.379 | 33.894 | 0.209 | 0.398 |
| 58 | 33.827 | 0.174 | 0.33 | ${ }^{33.826}$ | 0.173 | 0.329 | 33.827 | 0.174 | 0.33 | 33.845 | 0.184 | 0.349 |
| 57 | 33.78 | 0.149 | 0.284 | 33.776 | 0.147 | 0.28 | 33.778 | 0.148 | 0.281 | 33.797 | 0.158 | 0.3 |
| 56 | 33.722 | 0.119 | 0.225 | 33.727 | 0.122 | 0.231 | 33.73 | 0.123 | 0.233 | 33.748 | 0.132 | 0.251 |
| 55 | ${ }^{33.678}$ | 0.095 | 0.181 | 33.679 | 0.096 | 0.182 | 33.681 | 0.097 | 0.185 | 33.7 | 0.107 | 0.203 |
| 54 | 33.633 | 0.072 | 0.137 | 33.63 | 0.07 | 0.134 | 33.633 | 0.072 | 0.137 | 33.652 | 0.082 | 0.155 |
| 53 | 33.586 | 0.047 | 0.089 | 33.582 | 0.045 | 0.085 | 33.585 | 0.047 | 0.089 | 33.604 | 0.056 | 0.107 |
| 52 | 33.531 | 0.018 | 0.035 | ${ }^{33.533}$ | 0.019 | 0.037 | 33.538 | 0.022 | 0.041 | 33.556 | 0.031 | 0.059 |
| 51 | 33.484 | -0.007 | -0.013 | 33.485 33.437 | -0.006 | -0.011 | 33.49 | -0.003 | -0.006 | 33.508 | 0.006 | 0.012 |
| 50 | 33.429 | -0.035 | -0.067 | 33 $=33.4977$, S | $\xrightarrow{-0.031}$ | $\stackrel{-0.06}{ }$ | 000.443 | -0.028 | -0.054 | 33.461 | -0.019 | -0.036 |

## Dataset: Employers Liability: Forecast

Model: PTF-good1
Forecasts from model omitting data from one to three calendar years.

The graph below depicts the how the means and std. devs. of the reserve distributions (beyond 2005) change as years are removed, assuming that Calendar year trend of zero continues.

The pair of numbers for any particular year represent the mean and std. dev. of the reserve distribution as predicted at that year's end. For example this would mean that at year end 2003 the data for 2004 and 2005 are not used in estimating the parameters of the model, including process variability.

The prinicipal reasons for stability are that 1. the Calendar year trend since 2002 is zero 2. the decay parameter has been stable for many years and 3. the recent accident years have been flat.

Note that the slightly lower mean value at year end 2002 does not represent significant instability if you also consider the std. devs. of the distributions.


Prediction Errors vs Cal. Yr



Quantiles after removing two years then forecasting using scenario 2 (all quantiles are stable compared with the Quantile results obtained on 4 Forecast Sc.2)

Quantile Statistics and Value at Risk (Acc Year: Total)

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 99.995 | 159.34 | 3.844 | 29.615 | 160.836 | 4.038 | 31.111 | 163.128 | 4.335 | 33.404 | 161.881 | 4.173 | 32.156 |
| 99.99 | 158.399 | 3.721 | 28.674 | 159.985 | 3.927 | 30.26 | 161.476 | 4.121 | 31.752 | 160.355 | 3.975 | 30.63 |
| 99.98 | 157.578 | 3.615 | 27.853 | 159.025 | 3.803 | 29.3 | 159.771 | 3.9 | 30.046 | 158.774 | 3.77 | 29.049 |
| 99.97 | 157.464 | 3.6 | 27.739 | 158.408 | 3.723 | 28.683 | 158.747 | 3.767 | 29.022 | 157.821 | 3.646 | 28.096 |
| 99.96 | 157.101 | 3.553 | 27.376 | 157.934 | 3.661 | 28.209 | 158.006 | 3.67 | 28.281 | 157.131 | 3.557 | 27.406 |
| 99.95 | 157.056 | 3.547 | 27.331 | 157.539 | 3.61 | 27.814 | 157.423 | 3.595 | 27.698 | 156.586 | 3.486 | 26.862 |
| 99.94 | 156.937 | 3.532 | 27.212 | 157.192 | 3.565 | 27.468 | 156.942 | 3.532 | 27.217 | 156.136 | 3.428 | 26.411 |
| 99.93 | 156.921 | 3.53 | 27.196 | 156.881 | 3.525 | 27.157 | 156.531 | 3.479 | 26.806 | 155.751 | 3.378 | 26.026 |
| 99.92 | 156.032 | 3.414 | 26.307 | 156.598 | 3.488 | 26.873 | 156.171 | 3.432 | 26.446 | 155.414 | 3.334 | 25.69 |
| 99.91 | 155.991 | 3.409 | 26.266 | 156.335 | 3.454 | 26.61 | 155.852 | 3.391 | 26.127 | 155.115 | 3.295 | 25.39 |
| 99.9 | 155.864 | 3.392 | 26.139 | 156.089 | 3.422 | 26.364 | 155.564 | 3.354 | 25.839 | 154.845 | 3.26 | 25.12 |
| 99.8 | 153.791 | 3.123 | 24.066 | 154.261 | 3.184 | 24.536 | 153.619 | 3.101 | 23.894 | 153.015 | 3.023 | 23.29 |
| 99.7 | 152.622 | 2.972 | 22.897 | 153.067 | 3.029 | 23.342 | 152.435 | 2.947 | 22.71 | 151.897 | 2.878 | 22.172 |
| 99.6 | 151.91 | 2.879 | 22.185 | 152.191 | 2.916 | 22.466 | 151.572 | 2.835 | 21.847 | 151.079 | 2.771 | 21.354 |
| 99.5 | 151.118 | 2.777 | 21.393 | 151.51 | 2.827 | 21.785 | 150.888 | 2.747 | 21.163 | 150.43 | 2.687 | 20.705 |
| 99.4 | 150.385 | 2.681 | 20.66 | 150.962 | 2.756 | 21.237 | 150.319 | 2.673 | 20.594 | 149.89 | 2.617 | 20.165 |
| 99.3 | 149.963 | 2.627 | 20.238 | 150.503 | 2.697 | 20.778 | 149.83 | 2.609 | 20.105 | 149.425 | 2.557 | 19.7 |
| 99.2 | 149.744 | 2.598 | 20.019 | 150.099 | 2.644 | 20.374 | 149.401 | 2.554 | 19.676 | 149.016 | 2.504 | 19.291 |
| 99.1 | 149.367 | 2.549 | 19.642 | 149.74 | 2.598 | 20.015 | 149.018 | 2.504 | 19.293 | 148.651 | 2.456 | 18.926 |
| 99 | 149.089 | 2.513 | 19.364 | 149.413 | 2.555 | 19.688 | 148.671 | 2.459 | 18.946 | 148.32 | 2.413 | 18.595 |
| 98 | 146.744 | 2.209 | 17.019 | 146.967 | 2.238 | 17.242 | 146.285 | 2.149 | 16.56 | 146.036 | 2.117 | 16.311 |
| 97 | 145.092 | 1.994 | 15.367 | 145.341 | 2.027 | 15.616 | 144.791 | 1.955 | 15.066 | 144.599 | 1.93 | 14.874 |
| 96 | 143.919 | 1.842 | 14.194 | 144.164 | 1.874 | 14.439 | 143.677 | 1.811 | 13.952 | 143.524 | 1.791 | 13.799 |
| 95 | 143.086 | 1.734 | 13.361 | 143.229 | 1.753 | 13.504 | 142.777 | 1.694 | 13.053 | 142.653 | 1.678 | 12.928 |
| 94 | 142.253 | 1.626 | 12.528 | 142.441 | 1.65 | 12.716 | 142.016 | 1.595 | 12.291 | 141.915 | 1.582 | 12.19 |
| 93 | 141.521 | 1.531 | 11.796 | 141.759 | 1.562 | 12.034 | 141.352 | 1.509 | 11.627 | 141.27 | 1.498 | 11.545 |
| 92 | 140.928 | 1.454 | 11.203 | 141.155 | 1.483 | 11.43 | 140.759 | 1.432 | 11.034 | 140.694 | 1.424 | 10.969 |
| 91 | 140.431 | 1.389 | 10.706 | 140.612 | 1.413 | 10.887 | 140.223 | 1.362 | 10.498 | 140.172 | 1.356 | 10.447 |
| 90 | 140.004 | 1.334 | 10.279 | 140.113 | 1.348 | 10.388 | 139.731 | 1.299 | 10.006 | 139.692 | 1.294 | 9.967 |
| 89 | 139.513 | 1.27 | 9.788 | 139.647 | 1.288 | 9.922 | 139.275 | 1.239 | 9.55 | 139.247 | 1.236 | 9.522 |
| 88 | 139.147 | 1.223 | 9.422 | 139.205 | 1.23 | 9.48 | 138.849 | 1.184 | 9.124 | 138.831 | 1.182 | 9.106 |
| 87 | 138.656 | 1.159 | 8.931 | 138.783 | 1.176 | 9.058 | 138.449 | 1.132 | 8.724 | 138.44 | 1.131 | 8.715 |
| 86 | 138.353 | 1.12 | 8.628 | 138.379 | 1.123 | 8.654 | 138.071 | 1.083 | 8.346 | 138.07 | 1.083 | 8.345 |
| 85 | 137.874 | 1.058 | 8.149 | 137.99 | 1.073 | 8.265 | 137.711 | 1.037 | 7.987 | 137.717 | 1.037 | 7.992 |
| 84 | 137.517 | 1.011 | 7.792 | 137.615 | 1.024 | 7.891 | 137.369 | 0.992 | 7.644 | 137.381 | 0.994 | 7.656 |
| 83 | 137.15 | 0.964 | 7.425 | 137.255 | 0.977 | 7.53 | 137.04 | 0.949 | 7.315 | 137.059 | 0.952 | 7.334 |
| 82 | 136.811 | 0.92 | 7.086 | 136.908 | 0.932 | 7.183 | 136.725 | 0.908 | 7 | 136.749 | 0.912 | 7.024 |
| 81 | 136.476 | 0.876 | 6.752 | 136.575 | 0.889 | 6.85 | 136.421 | 0.869 | 6.696 | 136.45 | 0.873 | 6.725 |
| 80 | 136.16 | 0.835 | 6.436 | 136.254 | 0.847 | 6.529 | 136.127 | 0.831 | 6.402 | 136.161 | 0.835 | 6.436 |
| 79 | 135.788 | 0.787 | 6.063 | 135.945 | 0.807 | 6.22 | 135.843 | 0.794 | 6.118 | 135.881 | 0.799 | 6.156 |
| 78 | 135.491 | 0.748 | 5.766 | 135.648 | 0.769 | 5.923 | 135.567 | 0.758 | 5.842 | 135.609 | 0.764 | 5.884 |
| 77 | 135.239 | 0.716 | 5.514 | 135.361 | 0.731 | 5.636 | 135.299 | 0.723 | 5.574 | 135.345 | 0.729 | 5.62 |

TABLE CONTINUES ON NEXT PAGE

Quantile Statistics and Value at Risk (Acc Year: Total)

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R |
| 76 | 134.973 | 0.681 | 5.248 | 135.084 | 0.696 | 5.36 | 135.038 | 0.69 | 5.313 | 135.087 | 0.696 | 5.362 |
| 75 | 134.714 | 0.648 | 4.989 | 134.817 | 0.661 | 5.092 | 134.783 | 0.656 | 5.058 | 134.835 | 0.663 | 5.111 |
| 74 | 134.488 | 0.618 | 4.763 | 134.558 | 0.627 | 4.833 | 134.534 | 0.624 | 4.809 | 134.589 | 0.631 | 4.864 |
| 73 | 134.218 | 0.583 | 4.493 | 134.306 | 0.595 | 4.581 | 134.29 | 0.592 | 4.565 | 134.348 | 0.6 | 4.623 |
| 72 | 133.965 | 0.55 | 4.24 | 134.061 | 0.563 | 4.336 | 134.051 | 0.561 | 4.326 | 134.112 | 0.569 | 4.387 |
| 71 | 133.716 | 0.518 | 3.991 | 133.822 | 0.532 | 4.097 | 133.817 | 0.531 | 4.092 | 133.88 | 0.539 | 4.155 |
| 70 | 133.516 | 0.492 | 3.791 | 133.588 | 0.501 | 3.863 | 133.587 | 0.501 | 3.862 | 133.652 | 0.51 | 3.927 |
| 69 | 133.313 | 0.466 | 3.588 | 133.359 | 0.472 | 3.634 | 133.361 | 0.472 | 3.636 | 133.428 | 0.481 | 3.703 |
| 68 | 133.072 | 0.434 | 3.347 | 133.134 | 0.443 | 3.41 | 133.138 | 0.443 | 3.413 | 133.207 | 0.452 | 3.482 |
| 67 | 132.851 | 0.406 | 3.126 | 132.913 | 0.414 | 3.188 | 132.918 | 0.414 | 3.193 | 132.989 | 0.424 | 3.264 |
| 66 | 132.646 | 0.379 | 2.921 | 132.695 | 0.386 | 2.97 | 132.702 | 0.386 | 2.977 | 132.774 | 0.396 | 3.049 |
| 65 | 132.439 | 0.352 | 2.714 | 132.48 | 0.358 | 2.756 | 132.488 | 0.359 | 2.763 | 132.562 | 0.368 | 2.837 |
| 64 | 132.206 | 0.322 | 2.481 | 132.268 | 0.33 | 2.543 | 132.277 | 0.331 | 2.552 | 132.352 | 0.341 | 2.627 |
| 63 | 131.999 | 0.295 | 2.274 | 132.059 | 0.303 | 2.334 | 132.069 | 0.304 | 2.344 | 132.144 | 0.314 | 2.42 |
| 62 | 131.848 | 0.276 | 2.123 | 131.851 | 0.276 | 2.126 | 131.862 | 0.277 | 2.137 | 131.939 | 0.287 | 2.214 |
| 61 | 131.644 | 0.249 | 1.919 | 131.645 | 0.249 | 1.92 | 131.658 | 0.251 | 1.933 | 131.735 | 0.261 | 2.01 |
| 60 | 131.435 | 0.222 | 1.71 | 131.441 | 0.223 | 1.716 | 131.455 | 0.225 | 1.73 | 131.533 | 0.235 | 1.808 |
| 59 | 131.239 | 0.196 | 1.514 | 131.238 | 0.196 | 1.513 | 131.254 | 0.198 | 1.529 | 131.332 | 0.209 | 1.607 |
| 58 | 131.02 | 0.168 | 1.295 | 131.036 | 0.17 | 1.311 | 131.055 | 0.173 | 1.33 | 131.133 | 0.183 | 1.408 |
| 57 | 130.849 | 0.146 | 1.124 | 130.836 | 0.144 | 1.111 | 130.857 | 0.147 | 1.132 | 130.935 | 0.157 | 1.21 |
| 56 | 130.608 | 0.115 | 0.883 | 130.637 | 0.118 | 0.912 | 130.66 | 0.121 | 0.935 | 130.738 | 0.132 | 1.013 |
| 55 | 130.422 | 0.09 | 0.697 | 130.439 | 0.093 | 0.714 | 130.464 | 0.096 | 0.739 | 130.542 | 0.106 | 0.817 |
| 54 | 130.226 | 0.065 | 0.501 | 130.242 | 0.067 | 0.517 | 130.269 | 0.071 | 0.544 | 130.347 | 0.081 | 0.622 |
| 53 | 130.021 | 0.038 | 0.296 | 130.046 | 0.042 | 0.321 | 130.075 | 0.045 | 0.35 | 130.153 | 0.056 | 0.428 |
| 52 | 129.829 | 0.014 | 0.105 | 129.851 | 0.016 | 0.126 | 129.882 | 0.02 | 0.157 | 129.959 | 0.03 | 0.234 |
| 51 | 129.656 | -0.009 | -0.069 | 129.657 | -0.009 | -0.068 | 129.689 | -0.005 | -0.036 | 129.766 | 0.005 | 0.041 |
| 50 | 129.438 | -0.037 | -0.287 | 129.463 | -0.034 | -0.262 | 129.497 | -0.03 | -0.228 | 129.572 | -0.02 | -0.153 |
| 49 | 129.228 | -0.064 | -0.497 | 129.27 | -0.059 | -0.455 | 129.305 | -0.055 | -0.42 | 129.379 | -0.045 | -0.345 |
| 48 | 129.053 | -0.087 | -0.672 | 129.078 | -0.084 | -0.647 | 129.113 | -0.079 | -0.612 | 129.187 | -0.07 | -0.538 |
| 47 | 128.905 | -0.106 | -0.82 | 128.886 | -0.109 | -0.839 | 128.921 | -0.104 | -0.804 | 128.994 | -0.095 | -0.731 |
| 46 | 128.707 | -0.132 | -1.018 | 128.694 | -0.134 | -1.031 | 128.729 | -0.129 | -0.996 | 128.801 | -0.12 | -0.924 |
| 45 | 128.492 | -0.16 | -1.233 | 128.503 | -0.159 | -1.222 | 128.537 | -0.154 | -1.188 | 128.607 | -0.145 | -1.118 |
| 44 | 128.297 | -0.185 | -1.428 | 128.311 | -0.183 | -1.414 | 128.344 | -0.179 | -1.381 | 128.413 | -0.17 | -1.311 |
| 43 | 128.133 | -0.207 | -1.592 | 128.119 | -0.208 | -1.606 | 128.151 | -0.204 | -1.574 | 128.219 | -0.195 | -1.506 |
| 42 | 127.961 | -0.229 | -1.764 | 127.927 | -0.233 | -1.798 | 127.957 | -0.229 | -1.768 | 128.024 | -0.221 | -1.701 |
| 41 | 127.769 | -0.254 | -1.956 | 127.734 | -0.258 | -1.991 | 127.763 | -0.255 | -1.962 | 127.828 | -0.246 | -1.897 |
| 40 | 127.563 | -0.281 | -2.162 | 127.54 | -0.284 | -2.185 | 127.567 | -0.28 | -2.157 | 127.631 | -0.272 | -2.093 |
| 39 | 127.398 | -0.302 | -2.327 | 127.344 | -0.309 | -2.381 | 127.371 | -0.305 | -2.354 | 127.434 | -0.297 | -2.291 |
| 38 | 127.236 | -0.323 | -2.489 | 127.147 | -0.335 | -2.578 | 127.174 | -0.331 | -2.551 | 127.234 | -0.323 | -2.49 |
| 37 | 127.029 | -0.35 | -2.695 | 126.948 | -0.36 | -2.777 | 126.975 | -0.357 | -2.75 | 127.034 | -0.349 | -2.691 |
| 36 | 126.813 | -0.378 | -2.912 | 126.748 | -0.386 | -2.977 | 126.774 | -0.383 | -2.95 | 126.832 | -0.375 | -2.893 |
| 35 | 126.582 | -0.408 | -3.143 | 126.545 | -0.413 | -3.18 | 126.573 | -0.409 | -3.152 | 126.628 | -0.402 | -3.097 |
| 34 | 126.382 | -0.434 | -3.343 | 126.34 | -0.439 | -3.385 | 126.369 | -0.436 | -3.356 | 126.422 | -0.429 | -3.302 |
| 33 | 126.205 | -0.457 | -3.52 | 126.132 | -0.466 | -3.592 | 126.163 | -0.462 | -3.562 | 126.215 | -0.456 | -3.51 |
| 32 | 125.992 | -0.484 | -3.733 | 125.922 | -0.494 | -3.803 | 125.955 | -0.489 | -3.77 | 126.005 | -0.483 | -3.72 |
| 31 | 125.752 | -0.516 | -3.973 | 125.709 | -0.521 | -4.016 | 125.745 | -0.517 | -3.98 | 125.792 | -0.51 | -3.933 |
| 30 | 125.547 | -0.542 | -4.178 | 125.492 | -0.549 | -4.233 | 125.532 | -0.544 | -4.193 | 125.576 | -0.538 | -4.148 |

TABLE CONTINUES ON NEXT PAGE

Quantile Statistics and Value at Risk (Acc Year: Total)

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| \% | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 29 | 125.334 | -0.57 | -4.391 | 125.272 | -0.578 | -4.453 | 125.316 | -0.572 | -4.409 | 125.358 | -0.567 | -4.367 |
| 28 | 125.101 | -0.6 | -4.623 | 125.048 | -0.607 | -4.677 | 125.097 | -0.601 | -4.628 | 125.136 | -0.596 | -4.589 |
| 27 | 124.898 | -0.626 | -4.827 | 124.82 | -0.637 | -4.905 | 124.875 | -0.63 | -4.85 | 124.911 | -0.625 | -4.814 |
| 26 | 124.684 | -0.654 | -5.041 | 124.588 | -0.667 | -5.137 | 124.648 | -0.659 | -5.077 | 124.682 | -0.655 | -5.043 |
| 25 | 124.446 | -0.685 | -5.279 | 124.35 | -0.698 | -5.375 | 124.418 | -0.689 | -5.307 | 124.448 | -0.685 | -5.277 |
| 24 | 124.205 | -0.716 | -5.52 | 124.107 | -0.729 | -5.618 | 124.183 | -0.719 | -5.542 | 124.21 | -0.716 | -5.515 |
| 23 | 123.953 | -0.749 | -5.772 | 123.858 | -0.761 | -5.867 | 123.943 | -0.75 | -5.782 | 123.966 | -0.747 | -5.758 |
| 22 | 123.699 | -0.782 | -6.026 | 123.603 | -0.795 | -6.122 | 123.698 | -0.782 | -6.027 | 123.717 | -0.78 | -6.008 |
| 21 | 123.414 | -0.819 | -6.311 | 123.341 | -0.829 | -6.384 | 123.447 | -0.815 | -6.278 | 123.462 | -0.813 | -6.263 |
| 20 | 123.162 | -0.852 | -6.563 | 123.071 | -0.864 | -6.654 | 123.189 | -0.848 | -6.536 | 123.2 | -0.847 | -6.525 |
| 19 | 122.848 | -0.893 | -6.877 | 122.794 | -0.9 | -6.931 | 122.924 | -0.883 | -6.801 | 122.93 | -0.882 | -6.795 |
| 18 | 122.593 | -0.926 | -7.132 | 122.508 | -0.937 | -7.217 | 122.651 | -0.918 | -7.074 | 122.652 | -0.918 | -7.073 |
| 17 | 122.326 | -0.96 | -7.399 | 122.211 | -0.975 | -7.513 | 122.368 | -0.955 | -7.356 | 122.364 | -0.955 | -7.361 |
| 16 | 122.006 | -1.002 | -7.719 | 121.905 | -1.015 | -7.82 | 122.076 | -0.993 | -7.649 | 122.065 | -0.994 | -7.659 |
| 15 | 121.686 | -1.043 | -8.039 | 121.588 | -1.056 | -8.137 | 121.772 | -1.032 | -7.953 | 121.755 | -1.034 | -7.97 |
| 14 | 121.355 | -1.086 | -8.37 | 121.256 | -1.099 | -8.469 | 121.455 | -1.073 | -8.27 | 121.431 | -1.076 | -8.294 |
| 13 | 120.977 | -1.135 | -8.748 | 120.91 | -1.144 | -8.815 | 121.123 | -1.116 | -8.602 | 121.092 | -1.12 | -8.633 |
| 12 | 120.692 | -1.172 | -9.033 | 120.545 | -1.191 | -9.179 | 120.774 | -1.162 | -8.951 | 120.735 | -1.167 | -8.99 |
| 11 | 120.316 | -1.221 | -9.409 | 120.163 | -1.241 | -9.562 | 120.405 | -1.21 | -9.32 | 120.356 | -1.216 | -9.368 |
| 10 | 119.864 | -1.28 | -9.86 | 119.757 | -1.294 | -9.968 | 120.012 | -1.261 | -9.713 | 119.954 | -1.268 | -9.771 |
| 9 | 119.424 | -1.337 | -10.301 | 119.325 | -1.35 | -10.4 | 119.591 | -1.315 | -10.134 | 119.521 | -1.324 | -10.204 |
| 8 | 118.969 | -1.396 | -10.755 | 118.864 | -1.41 | -10.861 | 119.135 | -1.374 | -10.59 | 119.053 | -1.385 | -10.672 |
| 7 | 118.51 | -1.456 | -11.215 | 118.367 | -1.474 | -11.358 | 118.636 | -1.439 | -11.089 | 118.539 | -1.452 | -11.186 |
| 6 | 117.958 | -1.527 | -11.767 | 117.82 | -1.545 | -11.905 | 118.081 | -1.511 | -11.644 | 117.967 | -1.526 | -11.758 |
| 5 | 117.442 | -1.594 | -12.283 | 117.206 | -1.625 | -12.519 | 117.451 | -1.593 | -12.274 | 117.317 | -1.61 | -12.408 |
| 4 | 116.726 | -1.687 | -12.999 | 116.481 | -1.719 | -13.244 | 116.716 | -1.688 | -13.009 | 116.556 | -1.709 | -13.169 |
| 3 | 115.782 | -1.81 | -13.943 | 115.573 | -1.837 | -14.152 | 115.818 | -1.805 | -13.907 | 115.625 | -1.83 | -14.1 |
| 2 | 114.53 | -1.972 | -15.195 | 114.365 | -1.993 | -15.36 | 114.635 | -1.958 | -15.09 | 114.396 | -1.989 | -15.329 |
| 1 | 112.741 | -2.204 | -16.984 | 112.582 | -2.225 | -17.143 | 112.796 | -2.197 | -16.929 | 112.476 | -2.239 | -17.249 |
| Mean $=129.725$, S.D. $=7.705$, Provision $=129.725,1$ Unit $=\$ 1,000,000$ |  |  |  |  |  |  |  |  |  |  |  |  |

The real paid loss array is regarded as a sample path from the fitted model depicted in sheet $\mathbf{2}$ under Model Displays.
Below are the residual graphs for a model that has only one parameter in each direction for each of four triangles, the real data and three simulated triangles from the fitted model

It is almost impossible to distinguish in patterns between the real data and the three simulated triangles.
This demonstrates that the features in the model, fitted to the real data correspond to the features in the data
Moreover the forecast distributions for the simulated triangles and the real data are statistically the same.

|  | Mean Reserve SD Reserve |  |
| :--- | ---: | ---: |
| Real data | $106,795,624$ | $4,607,172$ |
| Sim 1 | $103,514,833$ | $4,431,096$ |
| Sim 2 | $105,497,719$ | $4,222,907$ |
| Sim 3 | $105,890,574$ | $4,324,997$ |

These features in the real data cannot be replicated by any link-ratio method or any derivative thereof, period.


|  |  |
| :---: | :---: |



Dataset: Employers Liability: Comparison of models for PL(I) and CRE(I)
Final Correlations Final Weighted Residual Correlations Between Datasets
6 iterations were executed

| $\mathrm{PL}(\mathrm{l}) 1$ | 1 | 0.20364 |
| :--- | :--- | :--- |

$\begin{array}{lrr} & 1 & 0.20364 \\ \text { CRE(I) } 1 & 0.20364 & 1\end{array}$
CRE = Case Reserve Estimates
PL = Paid Losses
CRE Model

he following comments are pertinent:

1. Along the calendar periods there some similarity in trend patterns, for example a drop from 2001, an increase around 1990, and both are ero from around 2003. Quite often the PL lag the CRE.

## However the PL are growing from 95-01 and the CRE are not - they are decreasing

2. Along the Accident years the total increase from 97-03 is about the same in both loss development arrays
3. The process correlation between the two loss development arrays is 0.203 . This means that if a random outcome in a cell in one array is above its trend line the probability that it is above the trend line in the other array is no longer 0.5
4. There is no overall indication that the PL have been increasing at a higher rate than the CRE

## Summary of Appendix F4

Here we use the ELRF modelling framework to model the cumulative Paid Losses.
We show that the features in the data cannot be captured by any link-ratio method.
We only illustrate this with volume-weighted averages, but all our arguments apply to any average link-ratios.

The accident years were adjusted by the exposures below.
Appendix F4 is uses data from Employers' Liability subclass "b".

## Subclass of Employers Liability

## Accident Exposures

Period
1985446
1986512
1987883
19881171
19891381
1990962
1991651
1992612
1993443
1994325
1995286
1996336
1997503
1998301
1999240
2000152
$2001 \quad 113$
2002130
200392
$2004 \quad 88$
200566

Note that for any of the models in the ELRF framework the modeller cannot choose nor has control over the salient assumptions in the projections. This is further discussed in Uncertainty Working Group Test Template.doc


An average ratio is a regression through the origin, yet for these data it is obvious that an intercept is needed.


Incrementals in Dev Yr. 1 are not correlated to the cumulatives in Dev. Yr. zero, so any link ratio from zero to one has no predictive power.

The weighted residuals below represent the residuals of fitting volume-weighted averages (as regression estimators).

Note that the trend structure in the data is not captured. Moreover we showed in the previous sheet that for some periods link ratios have no predictive power whatsoever.


## Residuals are not normally distributed.



| Accident Yr Summary |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Acc. Yr | Mean |  | Standard Dev. | CV |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1985 | 0 | 2,681,470 | 0 | **** | 0 |
| 1986 | 207,382 | 4,260,652 | 53,256 | 0.26 | 0.01 |
| 1987 | 377,929 | 4,064,324 | 118,813 | 0.31 | 0.03 |
| 1988 | 739,274 | 5,747,407 | 197,551 | 0.27 | 0.03 |
| 1989 | 1,010,346 | 6,619,931 | 243,124 | 0.24 | 0.04 |
| 1990 | 1,132,616 | 6,379,035 | 217,777 | 0.19 | 0.03 |
| 1991 | 1,403,619 | 7,005,605 | 203,861 | 0.15 | 0.03 |
| 1992 | 1,294,431 | 5,716,867 | 202,417 | 0.16 | 0.04 |
| 1993 | 1,114,836 | 4,532,711 | 160,100 | 0.14 | 0.04 |
| 1994 | 1,156,373 | 4,271,824 | 151,558 | 0.13 | 0.04 |
| 1995 | 1,410,553 | 4,730,236 | 164,061 | 0.12 | 0.03 |
| 1996 | 1,710,512 | 5,241,271 | 208,552 | 0.12 | 0.04 |
| 1997 | 2,438,467 | 6,956,118 | 306,944 | 0.13 | 0.04 |
| 1998 | 2,127,977 | 5,372,060 | 234,747 | 0.11 | 0.04 |
| 1999 | 2,350,370 | 5,136,188 | 273,875 | 0.12 | 0.05 |
| 2000 | 2,969,212 | 5,713,587 | 292,897 | 0.1 | 0.05 |
| 2001 | 3,305,219 | 5,417,264 | 308,274 | 0.09 | 0.06 |
| 2002 | 3,198,321 | 4,320,671 | 450,185 | 0.14 | 0.1 |
| 2003 | 3,430,801 | 3,963,033 | 742,660 | 0.22 | 0.19 |
| 2004 | 1,551,995 | 1,622,149 | 699,885 | 0.45 | 0.43 |
| 2005 | 1,410,124 | 1,421,962 | 1,980,809 | 1.4 | 1.39 |
|  |  |  |  |  |  |
| Total | 34,340,358 | 101,174,366 | 2,989,449 | 0.09 | 0.03 |
| $1 \text { Unit }=\$ 1$ <br> CV of forecast for last accident yr is $140.47 \%$ Model may be inappropriate |  |  |  |  |  |

[^0]All the above answers are meaningless because the features of the model have nothing to do with features in the data.

## Summary of Appendix F5

This appendix provides the forecasted reserve distribution and quantiles for the aggregate of the two subclasses of business Employers' Liability business.

The aggregate distribution of two lines is obtained by designing a composite model for the two lines which includes correlations between the lines of business (the data are not aggregated).

There are two types of correlations between two lines of business; process correlation and parameter correlation. Process correlation is the correlation between two sets of residuals. When you run the individual PTF models in MPTF for the two lines of business, the process correlation is $22 \%$.

This means that if in one line of business an observation in a particular cell is above the trend line, then the likelihood that it is above the trend line in the other line of business is greater than $50 \%$. Process correlation induces parameter correlation because composite model estimates parameters in one line of business also using the data in the other line of business.

Process correlation and parameter correlation induce correlations between all pairs of accident years, calendar years, and aggregates in the two lines of business. The correlation between the reserves is $7.7 \%$. As a result of this low correlation, the required capital for the two lines of business at the 95th percentile, say, is almost the same as if they were independent.

The forecast scenario used included continuing calendar trends as follows.
For Employers Liability subclass "a"

| 2005-2011 | $8.39 \%+-0.76 \%$ | These assumptions are discussed in Appendix F5.3. |
| :--- | :--- | :--- |
| 2011-2025 | $0 \%+-0 \%$ |  |
| For Employers Liability subclass "b" |  |  |
| $2005-2006$ | $24 \%+-2.7 \%$ | These assumptions are discussed in Appendix F5.4. |
| $2006-2025$ | $0 \%+-0 \%$ |  |

The forecasts for the reserve distribution for the subclasses are shown below.
Mean of Reserve Dsn. Std. Dv. Of Reserve Dsn. CoV
Subclass "a"
131,366,189
6,202,006 0.04721159
Subclass "b" 37,154,492 3,429,684 0.09230873

It is important to recognise that all forecast distributions are conditional on an explicit set of assumptions that are easily interpretable. If any of these assumptions are varied then the resulting predictive distributions will be different.

Appendix F5.1 Quantiles and V@R for the aggregate reserve for both LOBs.

Appendix F5.2 Process correlation between the two LOBs

Appendices F5.3 Model displays for the two LOBs.
and F5.4

Appendix F5.5 Forecast table for the aggregate of the two LOBs. Means and SDs of distributions are displayed for each cell and for aggregates across cells.

Appendix F 5.6 Accident year summaries of the aggregate of the two LOBs and for each LOB

Appendix F5.7 Calendar year summaries of the aggregate of the two LOBs and for each LOB

Appendix F5.8 Capital allocation by LOB based on a covariance formula.

| Employers Liability:Composite DS:MPTF[optimal-1]:PALD:Acc. Yr:Total:Quantiles \& VAR |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| Mean $=168.521$, S.D. $=7.314$, Provision $=168.521,1$ Unit $=\$ 1,000,000$ |  |  |  |  |  |  |  |  |  |  |  |  |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 99.995 | 196.92 | 3.883 | 28.399 | 198.44 | 4.091 | 29.92 | 199.318 | 4.211 | 30.797 | 198.485 | 4.097 | 29.964 |
| 99.99 | 195.655 | 3.71 | 27.134 | 197.454 | 3.956 | 28.934 | 197.84 | 4.009 | 29.32 | 197.089 | 3.906 | 28.569 |
| 99.98 | 195.51 | 3.69 | 26.99 | 196.346 | 3.804 | 27.825 | 196.311 | 3.8 | 27.79 | 195.641 | 3.708 | 27.12 |
| 99.97 | 195.25 | 3.654 | 26.729 | 195.65 | 3.709 | 27.129 | 195.39 | 3.674 | 26.869 | 194.766 | 3.588 | 26.246 |
| 99.96 | 194.611 | 3.567 | 26.091 | 195.119 | 3.636 | 26.598 | 194.723 | 3.582 | 26.202 | 194.132 | 3.502 | 25.612 |
| 99.95 | 194.141 | 3.503 | 25.62 | 194.676 | 3.576 | 26.155 | 194.198 | 3.511 | 25.677 | 193.633 | 3.433 | 25.112 |
| 99.94 | 193.931 | 3.474 | 25.41 | 194.291 | 3.523 | 25.77 | 193.763 | 3.451 | 25.242 | 193.219 | 3.377 | 24.698 |
| 99.93 | 193.69 | 3.441 | 25.169 | 193.949 | 3.477 | 25.428 | 193.392 | 3.4 | 24.871 | 192.865 | 3.328 | 24.344 |
| 99.92 | 193.11 | 3.362 | 24.589 | 193.644 | 3.435 | 25.124 | 193.067 | 3.356 | 24.546 | 192.555 | 3.286 | 24.034 |
| 99.91 | 192.773 | 3.316 | 24.252 | 193.369 | 3.397 | 24.848 | 192.778 | 3.317 | 24.258 | 192.279 | 3.248 | 23.759 |
| 99.9 | 192.635 | 3.297 | 24.114 | 193.118 | 3.363 | 24.597 | 192.518 | 3.281 | 23.997 | 192.031 | 3.214 | 23.51 |
| 99.8 | 190.948 | 3.066 | 22.428 | 191.421 | 3.131 | 22.901 | 190.755 | 3.04 | 22.235 | 190.344 | 2.984 | 21.823 |
| 99.7 | 190.165 | 2.959 | 21.644 | 190.372 | 2.987 | 21.851 | 189.68 | 2.893 | 21.159 | 189.312 | 2.843 | 20.791 |
| 99.6 | 189.156 | 2.821 | 20.635 | 189.596 | 2.881 | 21.075 | 188.894 | 2.785 | 20.373 | 188.557 | 2.739 | 20.036 |
| 99.5 | 188.629 | 2.749 | 20.108 | 188.982 | 2.797 | 20.461 | 188.27 | 2.7 | 19.75 | 187.957 | 2.657 | 19.436 |
| 99.4 | 188.236 | 2.696 | 19.715 | 188.472 | 2.728 | 19.951 | 187.751 | 2.629 | 19.23 | 187.456 | 2.589 | 18.935 |
| 99.3 | 187.689 | 2.621 | 19.169 | 188.032 | 2.668 | 19.511 | 187.305 | 2.568 | 18.784 | 187.026 | 2.53 | 18.505 |
| 99.2 | 187.344 | 2.574 | 18.823 | 187.646 | 2.615 | 19.125 | 186.912 | 2.515 | 18.392 | 186.647 | 2.478 | 18.126 |
| 99.1 | 187.101 | 2.54 | 18.581 | 187.296 | 2.567 | 18.775 | 186.562 | 2.467 | 18.041 | 186.308 | 2.432 | 17.788 |
| 99 | 186.728 | 2.489 | 18.207 | 186.979 | 2.524 | 18.458 | 186.244 | 2.423 | 17.723 | 186.001 | 2.39 | 17.481 |
| 98 | 184.469 | 2.18 | 15.948 | 184.667 | 2.208 | 16.146 | 184.054 | 2.124 | 15.534 | 183.88 | 2.1 | 15.36 |
| 97 | 182.91 | 1.967 | 14.39 | 183.08 | 1.99 | 14.559 | 182.679 | 1.936 | 14.158 | 182.543 | 1.917 | 14.022 |
| 96 | 181.586 | 1.786 | 13.065 | 181.865 | 1.824 | 13.345 | 181.65 | 1.795 | 13.129 | 181.541 | 1.78 | 13.021 |
| 95 | 180.689 | 1.664 | 12.168 | 180.915 | 1.695 | 12.394 | 180.818 | 1.681 | 12.297 | 180.729 | 1.669 | 12.208 |
| 94 | 179.837 | 1.547 | 11.316 | 180.14 | 1.589 | 11.619 | 180.112 | 1.585 | 11.592 | 180.04 | 1.575 | 11.519 |
| 93 | 179.269 | 1.469 | 10.748 | 179.486 | 1.499 | 10.965 | 179.496 | 1.501 | 10.975 | 179.437 | 1.492 | 10.916 |
| 92 | 178.74 | 1.397 | 10.219 | 178.916 | 1.421 | 10.395 | 178.946 | 1.425 | 10.425 | 178.898 | 1.419 | 10.377 |
| 91 | 178.227 | 1.327 | 9.707 | 178.408 | 1.352 | 9.887 | 178.447 | 1.357 | 9.926 | 178.409 | 1.352 | 9.888 |
| 90 | 177.775 | 1.265 | 9.254 | 177.947 | 1.289 | 9.427 | 177.989 | 1.295 | 9.468 | 177.96 | 1.29 | 9.439 |
| 89 | 177.395 | 1.213 | 8.875 | 177.524 | 1.231 | 9.003 | 177.564 | 1.236 | 9.044 | 177.543 | 1.233 | 9.022 |
| 88 | 176.97 | 1.155 | 8.449 | 177.132 | 1.177 | 8.611 | 177.167 | 1.182 | 8.647 | 177.152 | 1.18 | 8.632 |
| 87 | 176.665 | 1.113 | 8.144 | 176.764 | 1.127 | 8.244 | 176.794 | 1.131 | 8.273 | 176.785 | 1.13 | 8.265 |
| 86 | 176.278 | 1.061 | 7.758 | 176.418 | 1.08 | 7.897 | 176.441 | 1.083 | 7.92 | 176.438 | 1.082 | 7.917 |
| 85 | 175.97 | 1.019 | 7.45 | 176.088 | 1.035 | 7.568 | 176.105 | 1.037 | 7.584 | 176.107 | 1.037 | 7.586 |
| 84 | 175.673 | 0.978 | 7.153 | 175.775 | 0.992 | 7.255 | 175.784 | 0.993 | 7.264 | 175.791 | 0.994 | 7.27 |
| 83 | 175.376 | 0.937 | 6.855 | 175.475 | 0.951 | 6.954 | 175.477 | 0.951 | 6.956 | 175.488 | 0.953 | 6.967 |
| 82 | 175.059 | 0.894 | 6.538 | 175.185 | 0.911 | 6.664 | 175.182 | 0.911 | 6.661 | 175.197 | 0.913 | 6.676 |
| 81 | 174.83 | 0.863 | 6.309 | 174.905 | 0.873 | 6.384 | 174.897 | 0.872 | 6.376 | 174.915 | 0.874 | 6.395 |
| 80 | 174.555 | 0.825 | 6.034 | 174.633 | 0.836 | 6.112 | 174.622 | 0.834 | 6.101 | 174.644 | 0.837 | 6.123 |
| 79 | 174.275 | 0.787 | 5.754 | 174.369 | 0.8 | 5.848 | 174.355 | 0.798 | 5.834 | 174.38 | 0.801 | 5.859 |
| 78 | 174.035 | 0.754 | 5.515 | 174.112 | 0.764 | 5.591 | 174.096 | 0.762 | 5.575 | 174.124 | 0.766 | 5.603 |
| 77 | 173.775 | 0.718 | 5.254 | 173.86 | 0.73 | 5.339 | 173.844 | 0.728 | 5.323 | 173.875 | 0.732 | 5.354 |
| 76 | 173.576 | 0.691 | 5.055 | 173.614 | 0.696 | 5.093 | 173.599 | 0.694 | 5.078 | 173.632 | 0.699 | 5.111 |
| 75 | 173.362 | 0.662 | 4.841 | 173.372 | 0.663 | 4.852 | 173.359 | 0.662 | 4.838 | 173.395 | 0.666 | 4.874 |
| 74 | 173.128 | 0.63 | 4.607 | 173.136 | 0.631 | 4.615 | 173.125 | 0.629 | 4.604 | 173.163 | 0.635 | 4.642 |
| 73 | 172.849 | 0.592 | 4.329 | 172.903 | 0.599 | 4.382 | 172.896 | 0.598 | 4.375 | 172.935 | 0.604 | 4.415 |
| 72 | 172.594 | 0.557 | 4.074 | 172.675 | 0.568 | 4.154 | 172.671 | 0.567 | 4.15 | 172.712 | 0.573 | 4.192 |
| 71 | 172.377 | 0.527 | 3.856 | 172.451 | 0.537 | 3.931 | 172.45 | 0.537 | 3.929 | 172.493 | 0.543 | 3.973 |
| 70 | 172.149 | 0.496 | 3.628 | 172.232 | 0.507 | 3.711 | 172.233 | 0.508 | 3.713 | 172.278 | 0.514 | 3.757 |
| 69 | 171.915 | 0.464 | 3.394 | 172.016 | 0.478 | 3.495 | 172.02 | 0.478 | 3.499 | 172.066 | 0.485 | 3.546 |
| 68 | 171.717 | 0.437 | 3.196 | 171.804 | 0.449 | 3.283 | 171.81 | 0.45 | 3.289 | 171.858 | 0.456 | 3.337 |
| 67 | 171.519 | 0.41 | 2.998 | 171.596 | 0.42 | 3.075 | 171.603 | 0.421 | 3.082 | 171.652 | 0.428 | 3.131 |
| 66 | 171.305 | 0.381 | 2.785 | 171.391 | 0.392 | 2.87 | 171.399 | 0.393 | 2.878 | 171.449 | 0.4 | 2.928 |
| 65 | 171.132 | 0.357 | 2.611 | 171.19 | 0.365 | 2.669 | 171.197 | 0.366 | 2.676 | 171.248 | 0.373 | 2.727 |
| 64 | 170.927 | 0.329 | 2.407 | 170.991 | 0.338 | 2.47 | 170.997 | 0.339 | 2.477 | 171.049 | 0.346 | 2.529 |
| 63 | 170.762 | 0.306 | 2.241 | 170.795 | 0.311 | 2.275 | 170.8 | 0.312 | 2.28 | 170.853 | 0.319 | 2.332 |
| 62 | 170.577 | 0.281 | 2.056 | 170.602 | 0.285 | 2.082 | 170.605 | 0.285 | 2.084 | 170.658 | 0.292 | 2.138 |
| 61 | 170.398 | 0.257 | 1.878 | 170.412 | 0.259 | 1.891 | 170.412 | 0.259 | 1.891 | 170.465 | 0.266 | 1.945 |
| 60 | 170.213 | 0.231 | 1.692 | 170.223 | 0.233 | 1.702 | 170.22 | 0.232 | 1.699 | 170.274 | 0.24 | 1.753 |
| 59 | 169.979 | 0.199 | 1.459 | 170.035 | 0.207 | 1.515 | 170.029 | 0.206 | 1.509 | 170.084 | 0.214 | 1.563 |

TABLE CONTINUES ON NEXT PAGE

| Mean $=168.521$, S.D. $=7.314$, Provision $=168.521,1$ Unit $=\$ 1,000,000$ |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \#S.D.'s | V -a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 58 | 169.792 | 0.174 | 1.271 | 169.85 | 0.182 | 1.329 | 169.84 | 0.18 | 1.32 | 169.895 | 0.188 | 1.375 |
| 57 | 169.63 | 0.152 | 1.109 | 169.666 | 0.157 | 1.145 | 169.653 | 0.155 | 1.132 | 169.708 | 0.162 | 1.187 |
| 56 | 169.463 | 0.129 | 0.942 | 169.483 | 0.132 | 0.962 | 169.466 | 0.129 | 0.945 | 169.521 | 0.137 | 1 |
| 55 | 169.277 | 0.103 | 0.756 | 169.3 | 0.107 | 0.78 | 169.28 | 0.104 | 0.76 | 169.335 | 0.111 | 0.815 |
| 54 | 169.09 | 0.078 | 0.569 | 169.119 | 0.082 | 0.598 | 169.096 | 0.079 | 0.575 | 169.15 | 0.086 | 0.63 |
| 53 | 168.932 | 0.056 | 0.411 | 168.937 | 0.057 | 0.416 | 168.911 | 0.053 | 0.391 | 168.966 | 0.061 | 0.445 |
| 52 | 168.776 | 0.035 | 0.255 | 168.756 | 0.032 | 0.235 | 168.728 | 0.028 | 0.207 | 168.782 | 0.036 | 0.261 |
| 51 | 168.609 | 0.012 | 0.089 | 168.574 | 0.007 | 0.053 | 168.545 | 0.003 | 0.024 | 168.598 | 0.011 | 0.078 |
| 50 | 168.434 | -0.012 | -0.086 | 168.392 | -0.018 | -0.129 | 168.362 | -0.022 | -0.159 | 168.415 | -0.014 | -0.106 |
| 49 | 168.238 | -0.039 | -0.283 | 168.209 | -0.043 | -0.311 | 168.18 | -0.047 | -0.341 | 168.232 | -0.04 | -0.289 |
| 48 | 168.071 | -0.061 | -0.449 | 168.026 | -0.068 | -0.495 | 167.997 | -0.072 | -0.523 | 168.048 | -0.065 | -0.472 |
| 47 | 167.924 | -0.082 | -0.596 | 167.842 | -0.093 | -0.679 | 167.815 | -0.097 | -0.706 | 167.865 | -0.09 | -0.656 |
| 46 | 167.721 | -0.109 | -0.8 | 167.656 | -0.118 | -0.864 | 167.632 | -0.122 | -0.889 | 167.682 | -0.115 | -0.839 |
| 45 | 167.508 | -0.138 | -1.013 | 167.47 | -0.144 | -1.051 | 167.449 | -0.147 | -1.072 | 167.498 | -0.14 | -1.023 |
| 44 | 167.307 | -0.166 | -1.214 | 167.282 | -0.169 | -1.239 | 167.265 | -0.172 | -1.255 | 167.313 | -0.165 | -1.207 |
| 43 | 167.123 | -0.191 | -1.398 | 167.093 | -0.195 | -1.427 | 167.081 | -0.197 | -1.439 | 167.129 | -0.19 | -1.392 |
| 42 | 166.928 | -0.218 | -1.592 | 166.903 | -0.221 | -1.617 | 166.897 | -0.222 | -1.624 | 166.943 | -0.216 | -1.578 |
| 41 | 166.721 | -0.246 | -1.8 | 166.712 | -0.247 | -1.808 | 166.711 | -0.247 | -1.809 | 166.757 | -0.241 | -1.764 |
| 40 | 166.556 | -0.269 | -1.965 | 166.52 | -0.273 | -2 | 166.525 | -0.273 | -1.996 | 166.569 | -0.267 | -1.951 |
| 39 | 166.338 | -0.298 | -2.183 | 166.327 | -0.3 | -2.194 | 166.337 | -0.298 | -2.183 | 166.381 | -0.293 | -2.14 |
| 38 | 166.099 | -0.331 | -2.422 | 166.133 | -0.326 | -2.388 | 166.149 | -0.324 | -2.372 | 166.191 | -0.318 | -2.329 |
| 37 | 165.945 | -0.352 | -2.576 | 165.937 | -0.353 | -2.584 | 165.959 | -0.35 | -2.562 | 166 | -0.345 | -2.521 |
| 36 | 165.78 | -0.375 | -2.74 | 165.74 | -0.38 | -2.78 | 165.767 | -0.376 | -2.753 | 165.808 | -0.371 | -2.713 |
| 35 | 165.589 | -0.401 | -2.931 | 165.542 | -0.407 | -2.978 | 165.574 | -0.403 | -2.946 | 165.613 | -0.397 | -2.907 |
| 34 | 165.385 | -0.429 | -3.136 | 165.343 | -0.435 | -3.178 | 165.38 | -0.429 | -3.141 | 165.417 | -0.424 | -3.104 |
| 33 | 165.188 | -0.456 | -3.333 | 165.141 | -0.462 | -3.38 | 165.183 | -0.456 | -3.338 | 165.219 | -0.451 | -3.302 |
| 32 | 165.008 | -0.48 | -3.512 | 164.937 | -0.49 | -3.583 | 164.984 | -0.484 | -3.537 | 165.018 | -0.479 | -3.502 |
| 31 | 164.813 | -0.507 | -3.708 | 164.732 | -0.518 | -3.789 | 164.782 | -0.511 | -3.738 | 164.815 | -0.507 | -3.705 |
| 30 | 164.565 | -0.541 | -3.956 | 164.524 | -0.546 | -3.997 | 164.578 | -0.539 | -3.943 | 164.61 | -0.535 | -3.911 |
| 29 | 164.328 | -0.573 | -4.192 | 164.313 | -0.575 | -4.208 | 164.371 | -0.567 | -4.149 | 164.401 | -0.563 | -4.12 |
| 28 | 164.132 | -0.6 | -4.389 | 164.1 | -0.604 | -4.421 | 164.161 | -0.596 | -4.36 | 164.189 | -0.592 | -4.331 |
| 27 | 163.92 | -0.629 | -4.601 | 163.884 | -0.634 | -4.637 | 163.948 | -0.625 | -4.573 | 163.974 | -0.622 | -4.547 |
| 26 | 163.697 | -0.66 | -4.824 | 163.665 | -0.664 | -4.856 | 163.731 | -0.655 | -4.79 | 163.755 | -0.652 | -4.766 |
| 25 | 163.475 | -0.69 | -5.045 | 163.442 | -0.694 | -5.079 | 163.509 | -0.685 | -5.011 | 163.531 | -0.682 | -4.989 |
| 24 | 163.293 | -0.715 | -5.228 | 163.216 | -0.725 | -5.305 | 163.284 | -0.716 | -5.237 | 163.303 | -0.713 | -5.217 |
| 23 | 163.062 | -0.746 | -5.458 | 162.986 | -0.757 | -5.535 | 163.053 | -0.748 | -5.468 | 163.07 | -0.745 | -5.45 |
| 22 | 162.806 | -0.781 | -5.714 | 162.751 | -0.789 | -5.769 | 162.817 | -0.78 | -5.703 | 162.832 | -0.778 | -5.689 |
| 21 | 162.592 | -0.811 | -5.929 | 162.512 | -0.821 | -6.008 | 162.575 | -0.813 | -5.945 | 162.587 | -0.811 | -5.933 |
| 20 | 162.358 | -0.843 | -6.163 | 162.268 | -0.855 | -6.253 | 162.327 | -0.847 | -6.194 | 162.336 | -0.846 | -6.185 |
| 19 | 162.077 | -0.881 | -6.444 | 162.018 | -0.889 | -6.503 | 162.072 | -0.882 | -6.449 | 162.077 | -0.881 | -6.443 |
| 18 | 161.818 | -0.916 | -6.703 | 161.761 | -0.924 | -6.76 | 161.808 | -0.918 | -6.712 | 161.811 | -0.917 | -6.71 |
| 17 | 161.584 | -0.948 | -6.936 | 161.497 | -0.96 | -7.024 | 161.536 | -0.955 | -6.985 | 161.535 | -0.955 | -6.986 |
| 16 | 161.313 | -0.985 | -7.208 | 161.225 | -0.998 | -7.296 | 161.254 | -0.994 | -7.267 | 161.248 | -0.994 | -7.273 |
| 15 | 161.066 | -1.019 | -7.455 | 160.942 | -1.036 | -7.579 | 160.96 | -1.034 | -7.561 | 160.95 | -1.035 | -7.57 |
| 14 | 160.78 | -1.058 | -7.741 | 160.648 | -1.076 | -7.873 | 160.654 | -1.076 | -7.867 | 160.639 | -1.078 | -7.882 |
| 13 | 160.423 | -1.107 | -8.098 | 160.341 | -1.118 | -8.179 | 160.332 | -1.119 | -8.188 | 160.313 | -1.122 | -8.208 |
| 12 | 160.116 | -1.149 | -8.405 | 160.017 | -1.163 | -8.504 | 159.995 | -1.166 | $-8.526$ | 159.969 | -1.169 | -8.551 |
| 11 | 159.833 | -1.188 | -8.687 | 159.674 | -1.209 | -8.846 | 159.637 | -1.215 | -8.884 | 159.606 | -1.219 | -8.915 |
| 10 | 159.457 | -1.239 | -9.064 | 159.304 | -1.26 | -9.217 | 159.256 | -1.267 | -9.265 | 159.218 | -1.272 | -9.303 |
| 9 | 159.103 | -1.288 | -9.417 | 158.901 | -1.315 | -9.619 | 158.847 | -1.323 | -9.673 | 158.801 | -1.329 | -9.719 |
| 8 | 158.696 | -1.343 | -9.824 | 158.455 | -1.376 | -10.066 | 158.405 | -1.383 | -10.116 | 158.35 | -1.391 | -10.171 |
| 7 | 158.169 | -1.415 | -10.351 | 157.953 | -1.445 | -10.568 | 157.919 | -1.449 | -10.602 | 157.854 | -1.458 | -10.667 |
| 6 | 157.498 | -1.507 | -11.022 | 157.382 | -1.523 | -11.138 | 157.379 | -1.523 | -11.142 | 157.301 | -1.534 | -11.219 |
| 5 | 156.853 | -1.595 | -11.667 | 156.733 | -1.612 | -11.788 | 156.765 | -1.607 | -11.756 | 156.673 | -1.62 | -11.848 |
| 4 | 156.136 | -1.693 | -12.385 | 155.986 | -1.714 | -12.534 | 156.046 | -1.706 | -12.474 | 155.937 | -1.72 | -12.584 |
| 3 | 155.289 | -1.809 | -13.232 | 155.087 | -1.837 | -13.433 | 155.168 | -1.826 | -13.353 | 155.035 | -1.844 | -13.485 |
| 2 | 154.183 | -1.96 | -14.338 | 153.861 | -2.004 | -14.66 | 154.008 | -1.984 | -14.513 | 153.842 | -2.007 | -14.679 |
| 1 | 151.941 | -2.267 | -16.579 | 151.874 | -2.276 | -16.647 | 152.197 | -2.232 | -16.323 | 151.974 | -2.262 | -16.547 |
| Mean $=168.521$, S.D. $=7.314$, Provision $=168.521,1$ Unit $=\$ 1,000,000$ |  |  |  |  |  |  |  |  |  |  |  |  |


| Employers Liability:Composite DS:MPTF[optimal-1]:Weighted Residual Covariances Between Datasets:Final Correlations |  |  |
| :---: | :---: | :---: |
| Final Weighted Residual Correlations Between Datasets |  |  |
|  | Employers Liability subclass "a":PL(l)1 | Employers Liability subclass "b":PL(l)1 |
| Employers Liability subclass "a":PL(I)1 | 1 | 0.2245 |
| Employers Liability subclass "b":PL(I)1 | 0.224567 |  |

The process correlation between the two datasets as measured by this model is $\mathbf{2 2 \%}$.



The above model display shows the trends in the three directions along with process variance for the Employers Liability UC 0110 data after incoporating process correlation.

Model comments:
Process variance is high particularly after development period 11.
The development trends show that the company is still paying a significant amount even at development period 6 .
There have been a number of changes in accident levels, but the level has been stable since 1999.
The calendar direction is fairly volatile with a number of calendar year trends. A conservative forecast would incoporate a positive calendar trend (eg: $8 \%+-0.78 \%$ ) in the near future

[^1]
## Model Display for Employers Liability subclass "b"






The model display above shows the trends in the three directions along with process variance for the Employers Liability UC 0012 data after incoporating process
correlation.
Model comments:
There are no changes in process variance by development period. Initially, process variance is higher than for Uc0110.
The development trends show that the company is still paying a significant amount even at development period 5.
There have been a number of changes in accident levels. The current level has been stable since 2002.

The calendar direction is fairly volatile with a number of calendar year trends.
In respect of future calendar year trends we have adopted a consertive view and assume that the recent large positive calendar trend of $24 \%+-2.87 \%$ continues for at one more year before reverting to zero. This assumption is critical to the forecast distributions.

## Forecast Table

| Accident Period vs Development P |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cal. Per. Total | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 | Reserve | Ultimate |
| 1985 | 371,807 | 371,807 | 3,092,22 | 4,802,58 | 5,486,03 | 4,860,18 | 3,028,50 | 2,317,30 | 1,743,011 | 1,110,873 | 616,037 | 3,270 | 186,52 | 154,874 | 25,926 | 106,49 | 93,409 | 55 | 69,794 | 6,02 | 107,834 | 136,389 |  | 28,262,00 |
|  | 478,653 | 478,653 | 3,054,306 | 4,004,796 | 5,044,041 | 46 | 3,513,731 | 2,382,079 | 1,731,117 | 1,526,103 | 724,896 | 257,662 | 23 | 91,136 | 99,462 | 60,147 | 67,813 | 12,023 | 585 | 65,787 | 85,019 | 129,272 |  |  |
| 1986 | 3,44 | 357,574 | 2,942,008 | 4,636,973 | 5,263,752 | 4,689,467 | 3,521,368 | 2,692,0 | ,03 | 1,106,639 | 636,888 | 369,998 | 223,892 | 185,036 | ,16 | 136,783 | 123,726 | 102,204 | 125,964 | 157,876 | 199,651 | 25, 802 | 253,802 | 7,877,45 |
|  | 3,485,668 | 431,362 | 3,040,670 | 4,150,149 | 4,838,383 | 4,541,616 | 3,444,698 | 2,361,280 | 143 | 1,045,055 | 701,428 | 525,607 | 163,379 | 178,214 | 244,901 | 149,780 | 61,427 | 82,034 | 136,531 | 199,783 | 228,210 | 94,140 | 94,140 | 94,12 |
| 987 | 8,122,527 | 377,931 | 3,131,300 | 4,888,345 | 5,571,796 | 5,966,720 | 4,487,202 | 2,844,860 | 1,800,190 | 1,172,113 | 711,454 | 385,059 | 222,201 | 184,457 | 150,234 | 127,263 | 89,821 | 107,147 | 132,018 | 165,527 | 209,637 | 209,037 | 418,674 | 30,268,12 |
|  | 7,533,589 | 488,123 | 2,859,152 | 5,117,642 | 4,826,260 | 4,573,057 | 4,382,958 | 2,558,125 | 2,087,463 | 991, | 617,930 | 278,437 | 236,606 | 202,842 | 61,510 | 180,778 | 85,971 | 69,729 | 101,876 | 172 | 77,841 | 77,688 | 113,942 | 113,942 |
| 1988 | 13,674,163 | 419,850 | 3,462,326 | 5,440,002 | 49,753 | 7,978,614 | 9,044 | 3,201,795 | 2,046,253 | 1,441,388 | 862,979 | 472,775 | 794 | 229,470 | 188,992 | 30 | 144,94 | ,85 | 220,62 | 278,765 | 277,731 | 277,021 | 833,516 | 38,571,423 |
|  | , 488,89 | 499,556 | 3,423,021 | ,871,1 | 6,045,053 | 72,348 | 98,324 | 3,225,429 | 1,593,862 | 1,887,727 | 2,823 | 816,403 | 644 | 53 | 4,119 | 523,950 | 30,285 | 648 | 112,030 | 103,428 | 3,15 | 2,97 | 389 | 191,38 |
| 1989 | 18,892,53 | 417,931 | 27,902 | 9,938 | 3,858 | 7,892,357 | 999,557 | 3,251,600 | 2,249,332 | 1,556,327 | ,279 | 9,228 | 33 | 5,27 | ,28 | 175,488 | 211,379 | 261,934 | 329,943 | ,34 | 327,251 | 326,501 | 1,312,040 | 41,405,02 |
|  | 18,376,437 | 498,745 | 0,019 | 5,137 | ,961 | 6,608,136 | 7,67 | 3,170,227 | 2,305,151 | 1,587,376 | ,373 | 5,909 | 176,879 | 79,590 | 04,940 | 137,114 | 166,351 | 143,40 | 122,300 | 1,86 | 121,580 | 121,39 | 269,010 |  |
| 1990 | 22,557,2 | 399, | 3,988,842 | 594 | 8,581,660 | 7,6 | 4,846,897 | 3,375,477 | 96 | 1,573,737 | 262 | 502 | 287,382 | 145,449 | 944 | 176,785 | 213,385 | 265,442 | 262,937 | 230 | 260,063 | 259,262 | 935 | 5,645,7 |
|  | 30,560 | 407,783 | 4,437,986 | ,963 | 7,855,717 | ,00,485 | 5,323,151 | 043 | 2,218,437 | 2,888,046 | 1,159,598 | 559 | 923 | 16,912 | ,900 | 144,581 | 218,692 | 99,329 | 98,528 | 98,033 | 97,717 | 97,511 | 256,899 | 256, |
| 1991 | 30,153,425 | 419,494 | 4,201,532 | 6,588,480 | 7,535,288 | 6,770,916 | 4,629,556 | 80, | 2,150,602 | 1,484,205 | 883,566 | 476, | 171,186 | 162,782 | 180,486 | 213,4 | 262,901 | 259,44 | 257,100 | 255,502 | 254,409 | 253,659 | 1,543,015 | 40,921,730 |
|  | 28,319,915 | 401,905 | 4,302,360 | 45,7 | 56,8 | 78,9 | 51,4 | ,771,5 | 2,478,978 | 1,376,124 | 9,8 | 5,680 | 68,456 | 392,145 | 186,054 | 352,631 | 97,016 | 95,7 | 95,018 | 9,55 | 94,256 | 94,062 | 270,193 | 27,19 |
| 1992 | 899,32 | 394,513 | 3,261,224 | 5,175,392 | 58,794 |  | 17,5 | 2,694,284 | 1,824,917 | 1,261,044 | 4,440 | 231,113 | 182,025 | 17,106 | 5,232 | 250,2 | 245,961 | 243,088 | 241,138 | 239,809 | 238,901 | 238,16 | 1,697,268 | 33,279,566 |
|  | 36,090, 149 | 445,163 | 3,58 | 6,80 | ,30,624 | 53 | 328,27 | 2,250,251 | 1,526,529 | 769,774 | 14,761 | 287,097 | 40,743 | 166,297 | 68,243 | 2,44 | 0,745 | 89,754 | 39,154 | 8,779 | 88,538 | 88,350 | 81,898 | 281,89 |
| 1993 | 37,212,539 | 324,038 | 9,774 | 4,295,191 | 5,327,322 | 5,104,779 | 3,440,070 | 2,361,211 | ,95,053 | 1,099,858 | 303,451 | 204,788 | 164,315 | 160,655 | ,261 | 546 | 97 | ,95 | 329 | 26 | 153 | 172,538 | 4,746 | 29,75,882 |
|  | 36,085,684 | 320,763 | 2,695,707 | 4,874,905 | 5,969,137 | 4,834,813 | 3,263,004 | 3,062,041 | 776 | 967,392 | 456,236 | ,398 | 205,731 | 1,232 | 71,103 | 68,042 | 66,341 | 65,362 | 64,778 | 64,418 | 64,149 | 63,989 | 269 | 226,2 |
| 1994 | 35,176,920 | 359,142 | 3,015,018 | 5,095,314 | 6,294,645 | 6,019,655 | 4,039,691 | 2,763,058 | 1,857,621 | 515,638 | 320,125 | 207,283 | 160,59 | 158,361 | 147,822 | 140,751 | 135,989 | 132,770 | 130,584 | 128,827 | 127,720 | 127,020 | 1,229,846 | 2,179,6 |
|  | 35,705,302 | 374,943 | 3,134,302 | ,278 | 5,775,133 | 6,563,261 | 4,511,286 | 2,911,676 | 86 | 553,108 | 319,403 | 117,463 | 44,214 | 67,390 | 58,359 | 53,384 | 50,643 | 49,108 | 48,222 | 47,602 | 47,260 | 47,06 | 190,665 | 90,6 |
| 1995 | 32,706,6 | 399,726 | 3,625,022 | 6,076,233 | 7,54,616 | 7,185,281 | 96 | 3,295,943 | 878,795 | 609,108 | 390,946 | 267,6 | 221,045 | 180,198 | 167,562 | 159, | 153,373 | 149,512 | 146,426 | 144,488 | 143, | 142,496 | 1,607,45 | 4,376,597 |
|  | 33,652,902 | 296,004 | 3,57 | 485,848 | 3,465 | 11,298 | 4,415,095 | 1,860,767 | 1,092,444 | 776,126 | 386 | 218,139 | 80,723 | 77,95 | 66,809 | 60,63 | 57,233 | 55,3 | 54,04 | 53,368 | 52,992 | 52,77 | 238,892 | 238,89 |
| 1996 | 33,215,7 | 404,720 | 3,619,707 | 6,088,764 | ,88,161 | 7,183,143 | 4,824,141 | 1,324,922 | 909,371 | 657,34 | 452,65 | 351,337 | 251,039 | 204,368 | 191,896 | 183,527 | 177,890 | 173,410 | 170,608 | 168,850 | 167,74 | 167,0 | 2,207,711 | 38,315,5 |
|  | ,084, | 393,373 | 543,22 | 895,78 | 70,497 | 10,375,471 | 5,192,045 | 624,56 | 1,158,156 | 502,229 | 252,511 | 132,994 | 90,697 | 34,860 | 74,676 | 9,098 | 66,027 | 64,015 | 62,974 | 2,40 | 6,09 | 61,903 | 317,56 | 317,56 |
| 1997 | ,164,46 | 506,838 | 4,505,027 | 7,614,497 | 9,353,865 | 8,956,143 | 2,462,754 | 1,730,480 | 1,246,743 | 948,449 | 724,780 | 499,791 | 363,452 | 295,467 | 280,185 | 269,931 | 261,827 | 256,779 | 253,623 | 251,643 | 250,395 | 249,605 | 3,957,478 | 40,227,455 |
|  | 31,38,779 | 358,375 | 3,938,368 | 5,627,544 | 10,1 | 9,389,914 | 9 | 1,884,588 | 1,068,636 | 2,859 | 290,618 | 184,803 | 130,370 | 118,146 | 106,819 | 100,663 | 96,724 | 94,732 | 93,676 | 93,090 | 92,752 | 92,551 | 20 | 568,5 |
| The above table shows | 36,425,028 | 496,056 | 4,467,052 | 7,476,163 | 9,258,248 | 3,377,814 | 2,244,086 | 1,559,867 | 1,109,268 | 869,312 | 562,343 | 366,965 | 255,456 | 208,424 | 192,769 | 180,466 | 172,833 | 168,079 | 165,107 | 163,241 | 162,066 | 161,323 | 8,38 | 45,33,02 |
|  | 36,5 | 229 | 3,689,1 | 6,253,143 | 11,995,665 | 11,231,031 | 5,317,770 | 22 | 956,095 | 7,470 | 2,127 | 5,941 | ,880 | 92,986 | ,75 | 69,5 | 65,143 | ,93 | 61,805 | 1,19 | 60,854 | 60,65 | 27,34, | 27,3 |
| 1999 | 40,402,067 | 1,738,650 | 15,959,383 | 26,325,986 | 11,449,881 | 10,208,268 | 6,433,608 | 4,182,506 | 2,918,997 | 2,014,392 | 1,198,587 | 645,814 | 370,597 | 307,854 | 239,782 | 197,689 | 171,559 | 155,27 | 145,09 | 138,695 | 134,66 | 132,11 | 8,771,108 | 52,712,64 |
|  | 38,43,626 | 307,131 | 3,464,725 | 7,324,949 | 11,062,092 | 11,742,461 | 6,739,436 | 3,300,742 | 628,979 | 582,395 | 775,116 | 379,199 | 192,815 | 241,502 | 157,640 | 108,366 | 80,241 | 64,880 | 56,8 | 52,74 | 50,649 | 49,548 | 1,346,629 | 1,346,62 |
| 2000 | 54,289,24 | 1,573,041 | 14,432, | 8,342,251 | 9,694,352 | 8,714, | 564,1 | 3,949, | 2,673,431 | 1,846 | 1,102,643 | 599,445 | 347,714 | 275,376 | 218,844 | 183,888 | 162,188 | 148,66 | 140,2 | 134,89 | 131,546 | 129,42 | 2,044 | 48,245,8 |
|  | 166,09 | 307,195 | 3,617,963 | 7,511,621 | 9,917,515 | 10,347,0 | ,500,453 | 844,522 | 577,022 | 532,010 | 703,287 | 344,578 | 177,266 | 203,660 | 134,941 | 95,165 | 72,990 | 61,207 | 55,178 | 52,130 | 50,567 | 49,740 | 1,552,288 | 1,552,2 |
|  | 73,924,017 | 1,575,232 | 4,993,416 | 7,728,899 | 8,982,986 | 79,8 | 562 | 3,814,10 | 2,573,22 | 1,772,609 | 1,046,7 | 52,5 | 293,418 | 231,961 | 179,6 | 147,3 | 127,220 | 114,707 | 106,87 | 101,9 | 98,865 | 96,9 | 16,820,12 | 44,579,75 |
|  | 43,321,2 | 160,458 | 2,7 | 6,929,566 | 08, | 8,712,400 | 621,597 | 825,509 | 561, | 523,2 | 701,1 | 0,99 | 158,885 | 185,195 | 120,6 | 82,7 | 60,9 | 49,1 | 42,8 | 39,715 | 38,095 | 37,248 | 1,677,408 | ,,677,4 |
| 2002 | 35, | 5,947 | 4,967,294 | 66,375 | 8,905,292 | 8,635,328 | 5,792,865 | 3,961,021 | 2,661,936 | 1,827,928 | 1,065,588 | 506,859 | 66,836 | 2,461 | 150,234 | 117,939 | 97,892 | 85,400 | 77,585 | 72,67 | 69,58 | 67,626 | 25,649,757 | 46,069,249 |
|  | 40, | 156,405 | 89,409 | , 88,53 | $9,035,141$ | 933,223 | ,112 | 262 | 2,027 | 556,348 | 4,644 | 334,747 | 152,640 | 105 | 5,63 | 76,3 | 53,041 | 39,599 | 3,10 | 28,227 | 26,188 | 25,131 | 2023,136 | 2,023,13 |
| 2003 | 40,180,074 | 424,976 | 01,831 | 888,35 | 7,032,127 | 6,721,987 | 4,506,889 | 3,0 | 2,068,756 | 1,419,879 | 766,118 | ,898 | 184,019 | 45,042 | 107,416 | 50 | 69,707 | 60,707 | 55,077 | 51,542 | 49,313 | 47,902 | 841 | 36,868,37 |
|  | 43,407,82 | 489 | 3,288,602 | 6,275,73 | ,69 | 729, | 493,040 | 683,205 | 463,952 | 435, | 543,522 | 241,025 | 109, | 130,384 | 83,199 | 54,862 | 38,035 | 28,311 | 22,923 | 20,2 | 18,57 | 17,808 | 1,810,0 | 1,810,06 |
| 2004 | 41,283,942 | 337,98 | 2,862,648 | 4,864,978 | 6,015,218 | 5,752,190 | 3,859,583 | 2,639,606 | 1,774,303 | 1,129,303 | 613,351 | 296,372 | 153,558 | 121,215 | 91,841 | 73,677 | 62,402 | 55,376 | 50,98 | 48,22 | 46,48 | 45,37 | 27,694,034 | 30,980,201 |
|  | 42,771,658 | 221,0 | 3,065,087 | 530, | 654,347 | 632,685 | 429,961 | 585,851 | 398,401 | 342, | 425,764 | 89,540 | 87,143 | 102,693 | 6,082 | 44,269 | 31,509 | 2,30 | 20,417 | 18,39 | 17,355 | 16,815 | 1,704,135 | 1,704,13 |
|  | ,206,282 | 308,042 | 2,825,509 | 4,715,393 | 5,852,836 | 5,593,098 | 3,747,330 | 2,559,669 | 1,588,190 | 1,008,490 | 544,324 | 258,771 | 131,020 | 03,27 | 76,57 | 60,067 | 49,818 | 3,432 | 9,437 | 36,928 | 35,34 | 4,34 | 29,303,85 | ,58 |
|  | 38,641,302 | 285,03 | 306,514 | 512,653 | 640,198 | 62,866 | 422,490 | 576,701 | 359,125 | 309,455 | 385,754 | 171,094 | 77,996 | 92,561 | 59,087 | 38,991 | 27,067 | 20,184 | 16,375 | 14,354 | 13,308 | 2,765 | 1,747,421 | 1,747,4 |
|  | Total FittedPaid |  | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | 2023 | 2024 | 2025 | Total Reserve | Total Ulitimate |
| $\begin{aligned} & \text { Cal. Per. } \\ & \text { Total } \end{aligned}$ | 669,629,300 |  | 40,152,894 | 35,155,581 | 28,479,904 | 21,014,185 | 14,187,514 | 9,614,941 | 6,101,316 | 3,936,035 | 2,623,111 | 1,899,358 | 1,479,245 | 1,182,439 | 843,120 | 626,381 | 461,240 | 312,441 | 204,596 | 131,311 | 30,725 | ,345 | 68,520,681 | 95,464 |
|  | 626,943,587 |  | 2,134,235 | 2,020,694 | 0,152 | 20,152 | 1,458,652 | 1,314,889 | 590 | 706,590 | 558,450 | 349,163 | 25 | 219,099 | 155,839 | 543 | 634 | . 04 | 43,397 | 30,576 | 22,150 | 12,765 | 7,314,170 | 7,314,170 |
| 1 Unit $=\$ 1$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

The above table shows the forecasted reserve distribution for each cell (mean is in black, standard deviation in red) along with the data (mean in black, observed in blue).

## Employers Liability: Composite DS: MPTF[optimal-1]: Forecast:

Aggregate: Reserve Forecast Summaries:Accident Periods

| Aggregate of both LOBs |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Accident Yr Summary |  |  |  |  |  |
| Acc. Yr | Mean |  | StandardDev. | CV |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1985 | 0 | 28,262,000 | 0 | **** | 0 |
| 1986 | 253,802 | 27,877,452 | 94,140 | 0.37 | 0 |
| 1987 | 418,674 | 30,268,126 | 113,942 | 0.27 | 0 |
| 1988 | 833,516 | 38,571,423 | 191,389 | 0.23 | 0 |
| 1989 | 1,312,040 | 41,405,025 | 269,010 | 0.21 | 0.01 |
| 1990 | 1,308,935 | 45,645,710 | 256,899 | 0.2 | 0.01 |
| 1991 | 1,543,015 | 40,921,730 | 270,193 | 0.18 | 0.01 |
| 1992 | 1,697,268 | 33,279,566 | 281,898 | 0.17 | 0.01 |
| 1993 | 1,424,746 | 29,755,882 | 226,269 | 0.16 | 0.01 |
| 1994 | 1,229,846 | 32,179,698 | 190,665 | 0.16 | 0.01 |
| 1995 | 1,607,451 | 34,376,597 | 238,892 | 0.15 | 0.01 |
| 1996 | 2,207,711 | 38,315,565 | 317,569 | 0.14 | 0.01 |
| 1997 | 3,957,478 | 40,227,455 | 568,520 | 0.14 | 0.01 |
| 1998 | 3,628,383 | 45,333,073 | 527,345 | 0.15 | 0.01 |
| 1999 | 8,771,108 | 52,712,643 | 1,346,629 | 0.15 | 0.03 |
| 2000 | 12,044,097 | 48,245,854 | 1,552,288 | 0.13 | 0.03 |
| 2001 | 16,820,122 | 44,579,752 | 1,677,408 | 0.1 | 0.04 |
| 2002 | 25,649,757 | 46,069,249 | 2,023,136 | 0.08 | 0.04 |
| 2003 | 26,814,841 | 36,868,372 | 1,810,063 | 0.07 | 0.05 |
| 2004 | 27,694,034 | 30,980,201 | 1,704,135 | 0.06 | 0.06 |
| 2005 | 29,303,857 | 29,588,893 | 1,747,421 | 0.06 | 0.06 |
|  |  |  |  |  |  |
| Total | \#\#\#\#\#\#\#\#\# | \#\#\#\#\#\#\#\#\# | 7,314,170 | 0.04 | 0.01 |
| 1 Unit = \$1 |  |  |  |  |  |


| LOB UC 0110 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Accident Yr Summary |  |  |  |  |  |
| Acc. Yr | Mean |  | Standard Dev. | CV |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1985 | 0 | 25,580,529 | 0 | **** | 0 |
| 1986 | 1,155 | 23,571,536 | 1,714 | 1.48 | 0 |
| 1987 | 3,341 | 26,166,398 | 3,637 | 1.09 | 0 |
| 1988 | 7,319 | 32,737,092 | 6,679 | 0.91 | 0 |
| 1989 | 12,890 | 34,496,291 | 10,382 | 0.81 | 0 |
| 1990 | 21,775 | 39,112,131 | 15,915 | 0.73 | 0 |
| 1991 | 31,346 | 33,808,075 | 21,187 | 0.68 | 0 |
| 1992 | 39,308 | 27,199,170 | 25,008 | 0.64 | 0 |
| 1993 | 53,175 | 24,966,436 | 32,382 | 0.61 | 0 |
| 1994 | 97,835 | 27,932,237 | 58,018 | 0.59 | 0 |
| 1995 | 162,340 | 29,611,803 | 79,820 | 0.49 | 0 |
| 1996 | 254,902 | 32,831,997 | 111,835 | 0.44 | 0 |
| 1997 | 559,913 | 32,312,238 | 243,491 | 0.43 | 0.01 |
| 1998 | 1,021,125 | 39,481,731 | 293,548 | 0.29 | 0.01 |
| 1999 | 6,238,208 | 47,393,925 | 1,238,792 | 0.2 | 0.03 |
| 2000 | 8,939,135 | 42,396,518 | 1,380,636 | 0.15 | 0.03 |
| 2001 | 13,892,202 | 39,539,787 | 1,524,646 | 0.11 | 0.04 |
| 2002 | 23,055,053 | 42,352,195 | 1,924,838 | 0.08 | 0.05 |
| 2003 | 24,566,249 | 34,087,548 | 1,726,268 | 0.07 | 0.05 |
| 2004 | 25,148,795 | 28,364,809 | 1,605,288 | 0.06 | 0.06 |
| 2005 | 27,260,125 | 27,533,323 | 1,680,000 | 0.06 | 0.06 |
|  |  |  |  |  |  |
| Total | 131,366,189 | 691,475,767 | 6,202,006 | 0.05 | 0.01 |
| 1 Unit = \$1 |  |  |  |  |  |


| LOB UC 0112 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Accident Yr Summary |  |  |  |  |  |
| Acc. Yr | Mean |  | Standard Dev. | CV |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1985 | 0 | 2,681,470 | 0 | **** | 0 |
| 1986 | 252,647 | 4,305,917 | 93,868 | 0.37 | 0.02 |
| 1987 | 415,333 | 4,101,728 | 113,365 | 0.27 | 0.03 |
| 1988 | 826,197 | 5,834,331 | 190,344 | 0.23 | 0.03 |
| 1989 | 1,299,150 | 6,908,735 | 267,402 | 0.21 | 0.04 |
| 1990 | 1,287,160 | 6,533,579 | 254,360 | 0.2 | 0.04 |
| 1991 | 1,511,669 | 7,113,656 | 266,610 | 0.18 | 0.04 |
| 1992 | 1,657,960 | 6,080,396 | 277,661 | 0.17 | 0.05 |
| 1993 | 1,371,571 | 4,789,446 | 220,105 | 0.16 | 0.05 |
| 1994 | 1,132,010 | 4,247,461 | 175,254 | 0.15 | 0.04 |
| 1995 | 1,445,111 | 4,764,794 | 216,004 | 0.15 | 0.05 |
| 1996 | 1,952,809 | 5,483,568 | 284,060 | 0.15 | 0.05 |
| 1997 | 3,397,566 | 7,915,217 | 485,516 | 0.14 | 0.06 |
| 1998 | 2,607,259 | 5,851,342 | 403,170 | 0.15 | 0.07 |
| 1999 | 2,532,900 | 5,318,719 | 393,933 | 0.16 | 0.07 |
| 2000 | 3,104,962 | 5,849,336 | 547,751 | 0.18 | 0.09 |
| 2001 | 2,927,921 | 5,039,965 | 520,202 | 0.18 | 0.1 |
| 2002 | 2,594,704 | 3,717,054 | 401,527 | 0.15 | 0.11 |
| 2003 | 2,248,592 | 2,780,823 | 345,912 | 0.15 | 0.12 |
| 2004 | 2,545,238 | 2,615,392 | 385,782 | 0.15 | 0.15 |
| 2005 | 2,043,732 | 2,055,570 | 302,516 | 0.15 | 0.15 |
|  |  |  |  |  |  |
| Total | 37,154,492 | 103,988,500 | 3,429,684 | 0.09 | 0.03 |
| 1 Unit $=$ \$1 |  |  |  |  |  |

The above table shows the forecasted reserve distribution for each cell (mean is in black, standard deviation in red) along with the data (mean in black, observed in blue).
The aggregate total reserve has a distribution with a mean of 168.5 Million and a standard deviation of 7.3 Million)

Employers Liability: Composite DS: MPTF[optimal-1]:
Forecast:Aggregate:Reserve Forecast Summaries:Calendar Periods

| Aggregate of both LOBs |  |  |  |
| :---: | :---: | :---: | :---: |
| Calendar Yr Summary |  |  |  |
| Calendar Yr | Mean Reserve | Standard Dev. | CV <br> Reserve |
| 2006 | 40,152,894 | 2,134,235 | 0.05 |
| 2007 | 35,155,581 | 2,020,694 | 0.06 |
| 2008 | 28,479,904 | 1,980,152 | 0.07 |
| 2009 | 21,014,185 | 1,720,152 | 0.08 |
| 2010 | 14,187,514 | 1,458,652 | 0.1 |
| 2011 | 9,614,941 | 1,314,889 | 0.14 |
| 2012 | 6,101,316 | 939,590 | 0.15 |
| 2013 | 3,936,035 | 706,590 | 0.18 |
| 2014 | 2,623,111 | 558,450 | 0.21 |
| 2015 | 1,899,358 | 349,163 | 0.18 |
| 2016 | 1,479,245 | 262,725 | 0.18 |
| 2017 | 1,182,439 | 219,099 | 0.19 |
| 2018 | 843,120 | 155,839 | 0.18 |
| 2019 | 626,381 | 117,543 | 0.19 |
| 2020 | 461,240 | 90,634 | 0.2 |
| 2021 | 312,441 | 62,804 | 0.2 |
| 2022 | 204,596 | 43,397 | 0.21 |
| 2023 | 131,311 | 30,576 | 0.23 |
| 2024 | 80,725 | 22,150 | 0.27 |
| 2025 | 34,345 | 12,765 | 0.37 |
|  |  |  |  |
| Total | 168,520,681 | 7,314,170 | 0.04 |
| 1 Unit = \$1 |  |  |  |


| LOB Subclass |  |  |  |
| :---: | :---: | :---: | :---: |
| Calendar Yr Summary |  |  |  |
| Calendar Yr | Mean Reserve | Standard Dev. | CV <br> Reserve |
| 2006 | 33,519,508 | 1,848,318 | 0.06 |
| 2007 | 29,508,011 | 1,802,724 | 0.06 |
| 2008 | 23,768,077 | 1,824,205 | 0.08 |
| 2009 | 17,179,088 | 1,601,101 | 0.09 |
| 2010 | 11,142,452 | 1,370,170 | 0.12 |
| 2011 | 7,132,762 | 1,248,581 | 0.18 |
| 2012 | 4,068,435 | 879,103 | 0.22 |
| 2013 | 2,259,762 | 650,795 | 0.29 |
| 2014 | 1,186,083 | 504,771 | 0.43 |
| 2015 | 624,074 | 276,022 | 0.44 |
| 2016 | 365,224 | 178,727 | 0.49 |
| 2017 | 242,027 | 137,920 | 0.57 |
| 2018 | 150,914 | 88,522 | 0.59 |
| 2019 | 94,180 | 57,741 | 0.61 |
| 2020 | 56,618 | 37,342 | 0.66 |
| 2021 | 33,482 | 24,071 | 0.72 |
| 2022 | 19,058 | 15,215 | 0.8 |
| 2023 | 9,921 | 9,075 | 0.91 |
| 2024 | 4,734 | 5,197 | 1.1 |
| 2025 | 1,777 | 2,658 | 1.5 |
|  |  |  |  |
| Total | 131,366,189 | 6,202,006 | 0.05 |
| 1 Unit = \$1 |  |  |  |


| LOB Subclass |  |  |  |
| :---: | ---: | ---: | ---: |
| Calendar <br> Yr | Calendar Yr Summary |  |  |
| 2006 | $6,633,386$ | 831,656 | 0.13 |
| 2007 | $5,647,570$ | 700,405 | 0.12 |
| 2008 | $4,711,827$ | 584,537 | 0.12 |
| 2009 | $3,835,097$ | 479,536 | 0.13 |
| 2010 | $3,045,061$ | 386,608 | 0.13 |
| 2011 | $2,482,179$ | 323,867 | 0.13 |
| 2012 | $2,032,880$ | 273,862 | 0.13 |
| 2013 | $1,676,274$ | 234,483 | 0.14 |
| 2014 | $1,437,028$ | 211,034 | 0.15 |
| 2015 | $1,275,283$ | 197,173 | 0.15 |
| 2016 | $1,114,021$ | 181,158 | 0.16 |
| 2017 | 940,412 | 161,632 | 0.17 |
| 2018 | 692,206 | 121,301 | 0.18 |
| 2019 | 532,201 | 97,096 | 0.18 |
| 2020 | 404,622 | 78,949 | 0.2 |
| 2021 | 278,959 | 55,320 | 0.2 |
| 2022 | 185,538 | 38,751 | 0.21 |
| 2023 | 121,389 | 28,007 | 0.23 |
| 2024 | 75,992 | 20,828 | 0.27 |
| 2025 | 32,568 | 12,100 | 0.37 |
|  |  |  |  |
| Total | $37,154,492$ | $3,429,684$ | 0.09 |
|  | 10 Unit $=\$ 1$ |  |  |



The reserve forecast correlation of $7.7 \%$ shows a good diversification level between these two lines. That is, the risk capital required at the 95 th percentile, for example is not significantly higher than were

| Capital Allocation (Totals) |  |
| :---: | :---: |
|  | $\%$ |
| Employers Liability subclass a | 74.96 |
| Employers Liability subclass b | 25.04 |

The capital allocation formula is based on the variances of the distributions for each LOB and the covariance between them.

## Summary of Appendix F6

This workbook provides the forecasted reserve distribution and quantiles for the Gross and Net data of the Marine Cargo Line of Business.

Estimates of the Net of reinsurance and Gross reserve distribution for the Marine Cargo Line of business is obtained by designing a composite model for the two lines which includes correlations between the Net and Gross data (the data are not aggregated).

The process correlation between the Net and Gross data is very high: 90\%. The correlation between reserve distributions is also very high at $85 \%$.

The trend structure is approximately the same in each triangle (see model displays). The process variance is higher for the Net data than the Gross data!

The forecast scenario used included continuing calendar trends as follows.
For both Net and Gross data
2005-2022 2.48\% +- 1.72\%

The forecasts for the reserve distribution for the Gross and Net data are shown below.

|  |  | Reserve |  |
| :---: | :---: | :---: | :---: |
|  | Reserve | Standard | Coefficient |
|  | Mean | Deviation | of Variation |
| Gross | 2,267 | 634 | 0.27966476 |
| Net | 2,144 | 613 | 0.28591418 |

The coefficient of variation of the reserves for the Net data is slightly higher than the coefficient of variation for the Gross data. This indicates that the Reinsurance program for the cedant is not optimal in terms of ceding risk. The reason for this phenomenon is that the process variability for the Net data is higher than for the Gross data. See the Model Displays.

| Marine Cargo Gross Data: Quantiles \& VAR Mean $=2.267$, S.D. $=0.634$, Provision $=2.267,1$ Unit $=\$ 1,000$ |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 99.995 | 6.643 | 6.903 | 4.376 | 6.883 | 7.281 | 4.615 | 6.349 | 6.439 | 4.082 | 5.601 | 5.259 | 3.333 |
| 99.99 | 6.299 | 6.36 | 4.031 | 6.556 | 6.767 | 4.289 | 6.057 | 5.979 | 3.79 | 5.408 | 4.955 | 3.141 |
| 99.98 | 5.972 | 5.845 | 3.705 | 6.22 | 6.236 | 3.953 | 5.767 | 5.521 | 3.5 | 5.211 | 4.645 | 2.944 |
| 99.97 | 5.894 | 5.722 | 3.627 | 6.053 | 5.972 | 3.786 | 5.598 | 5.255 | 3.331 | 5.094 | 4.46 | 2.827 |
| 99.96 | 5.854 | 5.659 | 3.587 | 5.933 | 5.783 | 3.666 | 5.478 | 5.066 | 3.211 | 5.011 | 4.328 | 2.743 |
| 99.95 | 5.774 | 5.532 | 3.506 | 5.833 | 5.626 | 3.566 | 5.385 | 4.919 | 3.118 | 4.945 | 4.224 | 2.678 |
| 99.94 | 5.735 | 5.471 | 3.468 | 5.743 | 5.484 | 3.476 | 5.31 | 4.8 | 3.043 | 4.891 | 4.139 | 2.624 |
| 99.93 | 5.733 | 5.468 | 3.466 | 5.658 | 5.35 | 3.391 | 5.246 | 4.699 | 2.978 | 4.845 | 4.066 | 2.578 |
| 99.92 | 5.552 | 5.182 | 3.285 | 5.578 | 5.223 | 3.311 | 5.19 | 4.611 | 2.923 | 4.805 | 4.003 | 2.537 |
| 99.91 | 5.423 | 4.978 | 3.155 | 5.502 | 5.102 | 3.234 | 5.141 | 4.534 | 2.874 | 4.769 | 3.947 | 2.502 |
| 99.9 | 5.341 | 4.85 | 3.074 | 5.431 | 4.991 | 3.164 | 5.098 | 4.465 | 2.83 | 4.737 | 3.897 | 2.47 |
| 99.8 | 4.915 | 4.177 | 2.648 | 5.005 | 4.319 | 2.738 | 4.81 | 4.011 | 2.542 | 4.523 | 3.559 | 2.256 |
| 99.7 | 4.756 | 3.926 | 2.489 | 4.791 | 3.981 | 2.523 | 4.641 | 3.744 | 2.373 | 4.395 | 3.357 | 2.128 |
| 99.6 | 4.575 | 3.641 | 2.308 | 4.648 | 3.756 | 2.381 | 4.52 | 3.555 | 2.253 | 4.302 | 3.21 | 2.035 |
| 99.5 | 4.496 | 3.516 | 2.228 | 4.542 | 3.589 | 2.275 | 4.427 | 3.407 | 2.16 | 4.229 | 3.095 | 1.962 |
| 99.4 | 4.418 | 3.393 | 2.151 | 4.457 | 3.454 | 2.189 | 4.35 | 3.286 | 2.083 | 4.169 | 3 | 1.902 |
| 99.3 | 4.348 | 3.283 | 2.081 | 4.384 | 3.34 | 2.117 | 4.285 | 3.184 | 2.018 | 4.117 | 2.918 | 1.85 |
| 99.2 | 4.279 | 3.174 | 2.012 | 4.321 | 3.24 | 2.054 | 4.229 | 3.094 | 1.961 | 4.072 | 2.847 | 1.805 |
| 99.1 | 4.239 | 3.11 | 1.971 | 4.265 | 3.151 | 1.997 | 4.179 | 3.016 | 1.912 | 4.032 | 2.784 | 1.765 |
| 99 | 4.206 | 3.058 | 1.939 | 4.214 | 3.072 | 1.947 | 4.134 | 2.945 | 1.867 | 3.996 | 2.727 | 1.729 |
| 98 | 3.854 | 2.504 | 1.587 | 3.877 | 2.54 | 1.61 | 3.836 | 2.475 | 1.569 | 3.75 | 2.34 | 1.483 |
| 97 | 3.653 | 2.185 | 1.385 | 3.68 | 2.229 | 1.413 | 3.658 | 2.195 | 1.391 | 3.6 | 2.102 | 1.332 |
| 96 | 3.518 | 1.973 | 1.25 | 3.539 | 2.007 | 1.272 | 3.53 | 1.992 | 1.263 | 3.489 | 1.927 | 1.222 |
| 95 | 3.399 | 1.786 | 1.132 | 3.43 | 1.835 | 1.163 | 3.429 | 1.833 | 1.162 | 3.4 | 1.788 | 1.133 |
| 94 | 3.315 | 1.653 | 1.048 | 3.342 | 1.695 | 1.074 | 3.345 | 1.701 | 1.078 | 3.326 | 1.671 | 1.059 |
| 93 | 3.242 | 1.538 | 0.975 | 3.268 | 1.578 | 1 | 3.274 | 1.588 | 1.006 | 3.262 | 1.57 | 0.995 |
| 92 | 3.186 | 1.449 | 0.919 | 3.203 | 1.477 | 0.936 | 3.211 | 1.488 | 0.943 | 3.206 | 1.48 | 0.938 |
| 91 | 3.133 | 1.366 | 0.866 | 3.146 | 1.386 | 0.879 | 3.154 | 1.4 | 0.887 | 3.154 | 1.4 | 0.887 |
| 90 | 3.083 | 1.287 | 0.816 | 3.094 | 1.304 | 0.826 | 3.104 | 1.319 | 0.836 | 3.108 | 1.326 | 0.841 |
| 89 | 3.04 | 1.219 | 0.773 | 3.045 | 1.228 | 0.778 | 3.057 | 1.246 | 0.79 | 3.065 | 1.259 | 0.798 |
| 88 | 2.992 | 1.143 | 0.725 | 3 | 1.157 | 0.733 | 3.014 | 1.178 | 0.747 | 3.025 | 1.196 | 0.758 |
| 87 | 2.939 | 1.06 | 0.672 | 2.958 | 1.091 | 0.691 | 2.974 | 1.115 | 0.707 | 2.988 | 1.137 | 0.721 |
| 86 | 2.907 | 1.009 | 0.64 | 2.919 | 1.028 | 0.652 | 2.937 | 1.056 | 0.67 | 2.953 | 1.082 | 0.686 |
| 85 | 2.868 | 0.947 | 0.6 | 2.882 | 0.971 | 0.615 | 2.902 | 1.001 | 0.634 | 2.92 | 1.03 | 0.653 |
| 84 | 2.833 | 0.893 | 0.566 | 2.848 | 0.916 | 0.581 | 2.868 | 0.948 | 0.601 | 2.889 | 0.981 | 0.622 |
| 83 | 2.804 | 0.847 | 0.537 | 2.815 | 0.865 | 0.548 | 2.837 | 0.899 | 0.57 | 2.859 | 0.933 | 0.592 |
| 82 | 2.772 | 0.796 | 0.505 | 2.785 | 0.817 | 0.518 | 2.807 | 0.851 | 0.539 | 2.83 | 0.888 | 0.563 |
| 81 | 2.739 | 0.744 | 0.471 | 2.756 | 0.771 | 0.488 | 2.778 | 0.806 | 0.511 | 2.803 | 0.845 | 0.536 |
| 80 | 2.717 | 0.71 | 0.45 | 2.728 | 0.727 | 0.461 | 2.75 | 0.762 | 0.483 | 2.776 | 0.803 | 0.509 |
| 79 | 2.693 | 0.672 | 0.426 | 2.702 | 0.685 | 0.434 | 2.724 | 0.721 | 0.457 | 2.751 | 0.763 | 0.484 |
| 78 | 2.667 | 0.631 | 0.4 | 2.676 | 0.645 | 0.409 | 2.699 | 0.68 | 0.431 | 2.726 | 0.724 | 0.459 |
| 77 | 2.643 | 0.593 | 0.376 | 2.652 | 0.607 | 0.385 | 2.674 | 0.642 | 0.407 | 2.702 | 0.686 | 0.435 |
| 76 | 2.618 | 0.553 | 0.351 | 2.628 | 0.57 | 0.361 | 2.65 | 0.604 | 0.383 | 2.679 | 0.65 | 0.412 |
| 75 | 2.597 | 0.52 | 0.33 | 2.606 | 0.534 | 0.339 | 2.627 | 0.568 | 0.36 | 2.657 | 0.614 | 0.389 |
| 74 | 2.577 | 0.489 | 0.31 | 2.584 | 0.499 | 0.317 | 2.605 | 0.532 | 0.338 | 2.635 | 0.58 | 0.368 |
| 73 | 2.559 | 0.461 | 0.292 | 2.562 | 0.466 | 0.295 | 2.583 | 0.498 | 0.316 | 2.613 | 0.546 | 0.346 |
| 72 | 2.537 | 0.425 | 0.27 | 2.542 | 0.433 | 0.274 | 2.562 | 0.465 | 0.295 | 2.593 | 0.513 | 0.325 |
| 71 | 2.513 | 0.388 | 0.246 | 2.521 | 0.401 | 0.254 | 2.541 | 0.432 | 0.274 | 2.572 | 0.481 | 0.305 |
| 70 | 2.495 | 0.359 | 0.227 | 2.501 | 0.37 | 0.234 | 2.521 | 0.4 | 0.254 | 2.552 | 0.45 | 0.285 |
| 69 | 2.479 | 0.335 | 0.212 | 2.482 | 0.339 | 0.215 | 2.501 | 0.369 | 0.234 | 2.533 | 0.419 | 0.265 |
| 68 | 2.459 | 0.302 | 0.191 | 2.463 | 0.309 | 0.196 | 2.482 | 0.339 | 0.215 | 2.513 | 0.388 | 0.246 |
| 67 | 2.44 | 0.272 | 0.173 | 2.445 | 0.28 | 0.177 | 2.463 | 0.309 | 0.196 | 2.495 | 0.359 | 0.227 |
| 66 | 2.423 | 0.245 | 0.155 | 2.426 | 0.251 | 0.159 | 2.445 | 0.28 | 0.178 | 2.476 | 0.329 | 0.209 |
| 65 | 2.407 | 0.221 | 0.14 | 2.409 | 0.223 | 0.141 | 2.427 | 0.252 | 0.159 | 2.458 | 0.301 | 0.191 |
| 64 | 2.389 | 0.192 | 0.122 | 2.391 | 0.195 | 0.124 | 2.409 | 0.223 | 0.142 | 2.44 | 0.272 | 0.173 |
| 63 | 2.373 | 0.167 | 0.106 | 2.374 | 0.168 | 0.107 | 2.391 | 0.196 | 0.124 | 2.422 | 0.244 | 0.155 |
| 62 | 2.354 | 0.138 | 0.087 | 2.357 | 0.141 | 0.09 | 2.374 | 0.169 | 0.107 | 2.405 | 0.217 | 0.137 |

TABLE CONTINUED ON THE NEXT PAGE

Quantile Statistics and Value at Risk (Acc Year: Total)

| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 61 | 2.335 | 0.108 | 0.068 | 2.34 | 0.115 | 0.073 | 2.357 | 0.142 | 0.09 | 2.387 | 0.19 | 0.12 |
| 60 | 2.318 | 0.08 | 0.051 | 2.324 | 0.089 | 0.056 | 2.34 | 0.115 | 0.073 | 2.37 | 0.163 | 0.103 |
| 59 | 2.302 | 0.054 | 0.034 | 2.307 | 0.064 | 0.04 | 2.324 | 0.089 | 0.057 | 2.354 | 0.136 | 0.086 |
| 58 | 2.287 | 0.031 | 0.019 | 2.292 | 0.038 | 0.024 | 2.308 | 0.064 | 0.04 | 2.337 | 0.11 | 0.07 |
| 57 | 2.273 | 0.009 | 0.005 | 2.276 | 0.014 | 0.009 | 2.292 | 0.038 | 0.024 | 2.32 | 0.084 | 0.053 |
| 56 | 2.258 | -0.015 | -0.009 | 2.26 | -0.011 | -0.007 | 2.276 | 0.013 | 0.008 | 2.304 | 0.058 | 0.037 |
| 55 | 2.245 | -0.036 | -0.023 | 2.245 | -0.035 | -0.022 | 2.26 | -0.012 | -0.007 | 2.288 | 0.032 | 0.021 |
| 54 | 2.227 | -0.064 | -0.04 | 2.23 | -0.059 | -0.038 | 2.244 | -0.036 | -0.023 | 2.272 | 0.007 | 0.004 |
| 53 | 2.215 | -0.083 | -0.052 | 2.215 | -0.083 | -0.053 | 2.229 | -0.06 | -0.038 | 2.256 | -0.018 | -0.011 |
| 52 | 2.2 | -0.106 | -0.067 | 2.2 | -0.106 | -0.067 | 2.214 | -0.085 | -0.054 | 2.24 | -0.043 | -0.027 |
| 51 | 2.185 | -0.13 | -0.083 | 2.185 | -0.13 | -0.082 | 2.198 | -0.108 | -0.069 | 2.224 | -0.068 | -0.043 |
| 50 | 2.171 | -0.152 | -0.097 | 2.17 | -0.153 | -0.097 | 2.183 | -0.132 | -0.084 | 2.208 | -0.093 | -0.059 |
| 49 | 2.158 | -0.172 | -0.109 | 2.156 | -0.176 | -0.111 | 2.169 | -0.156 | -0.099 | 2.193 | -0.117 | -0.074 |
| 48 | 2.143 | -0.196 | -0.124 | 2.142 | -0.198 | -0.126 | 2.154 | -0.179 | -0.113 | 2.177 | -0.142 | -0.09 |
| 47 | 2.128 | -0.22 | -0.139 | 2.127 | -0.221 | -0.14 | 2.139 | -0.202 | -0.128 | 2.162 | -0.166 | -0.105 |
| 46 | 2.113 | -0.243 | -0.154 | 2.113 | -0.243 | -0.154 | 2.124 | -0.225 | -0.143 | 2.146 | -0.191 | -0.121 |
| 45 | 2.099 | -0.266 | -0.169 | 2.099 | -0.266 | -0.168 | 2.11 | -0.249 | -0.158 | 2.131 | -0.215 | -0.136 |
| 44 | 2.086 | -0.286 | -0.181 | 2.085 | -0.288 | -0.182 | 2.095 | -0.272 | -0.172 | 2.116 | -0.239 | -0.152 |
| 43 | 2.072 | -0.309 | -0.196 | 2.071 | -0.31 | -0.196 | 2.081 | -0.294 | -0.187 | 2.1 | -0.264 | -0.167 |
| 42 | 2.056 | -0.333 | -0.211 | 2.057 | -0.332 | -0.21 | 2.066 | -0.317 | -0.201 | 2.085 | -0.288 | -0.182 |
| 41 | 2.042 | -0.355 | -0.225 | 2.043 | -0.354 | -0.224 | 2.052 | -0.34 | -0.216 | 2.069 | -0.312 | -0.198 |
| 40 | 2.03 | -0.375 | -0.237 | 2.029 | -0.376 | -0.238 | 2.037 | -0.363 | -0.23 | 2.054 | -0.336 | -0.213 |
| 39 | 2.017 | -0.395 | -0.25 | 2.015 | -0.397 | -0.252 | 2.023 | -0.386 | -0.245 | 2.039 | -0.361 | -0.229 |
| 38 | 2.004 | -0.416 | -0.263 | 2.002 | -0.419 | -0.266 | 2.008 | -0.409 | -0.259 | 2.023 | -0.385 | -0.244 |
| 37 | 1.991 | -0.436 | -0.276 | 1.988 | -0.441 | -0.279 | 1.994 | -0.431 | -0.273 | 2.008 | -0.409 | -0.259 |
| 36 | 1.977 | -0.457 | -0.29 | 1.974 | -0.462 | -0.293 | 1.979 | -0.454 | -0.288 | 1.992 | -0.434 | -0.275 |
| 35 | 1.963 | -0.48 | -0.304 | 1.961 | -0.484 | -0.307 | 1.965 | -0.477 | -0.303 | 1.977 | -0.459 | -0.291 |
| 34 | 1.946 | -0.506 | -0.321 | 1.947 | -0.505 | -0.32 | 1.95 | -0.5 | -0.317 | 1.961 | -0.483 | -0.306 |
| 33 | 1.934 | -0.525 | -0.333 | 1.933 | -0.527 | -0.334 | 1.935 | -0.523 | -0.332 | 1.945 | -0.508 | -0.322 |
| 32 | 1.922 | -0.545 | -0.345 | 1.92 | -0.549 | -0.348 | 1.921 | -0.547 | -0.346 | 1.929 | -0.533 | -0.338 |
| 31 | 1.909 | -0.565 | -0.358 | 1.906 | -0.57 | -0.361 | 1.906 | -0.57 | -0.361 | 1.913 | -0.558 | -0.354 |
| 30 | 1.896 | -0.586 | -0.371 | 1.892 | -0.592 | -0.375 | 1.891 | -0.593 | -0.376 | 1.897 | -0.584 | -0.37 |
| 29 | 1.882 | -0.607 | -0.385 | 1.878 | -0.614 | -0.389 | 1.876 | -0.617 | -0.391 | 1.881 | -0.61 | -0.386 |
| 28 | 1.868 | -0.631 | -0.4 | 1.864 | -0.636 | -0.403 | 1.861 | -0.641 | -0.406 | 1.864 | -0.636 | -0.403 |
| 27 | 1.855 | -0.651 | -0.412 | 1.85 | -0.658 | -0.417 | 1.846 | -0.665 | -0.421 | 1.848 | -0.662 | -0.419 |
| 26 | 1.842 | -0.671 | -0.425 | 1.836 | -0.681 | -0.432 | 1.83 | -0.689 | -0.437 | 1.831 | -0.688 | -0.436 |
| 25 | 1.827 | -0.695 | -0.44 | 1.821 | -0.704 | -0.446 | 1.815 | -0.714 | -0.452 | 1.814 | -0.715 | -0.453 |
| 24 | 1.813 | -0.717 | -0.455 | 1.806 | -0.727 | -0.461 | 1.799 | -0.739 | -0.468 | 1.796 | -0.743 | -0.471 |
| 23 | 1.799 | -0.739 | -0.469 | 1.791 | -0.751 | -0.476 | 1.783 | -0.764 | -0.484 | 1.779 | -0.77 | -0.488 |
| 22 | 1.783 | -0.764 | -0.484 | 1.776 | -0.775 | -0.491 | 1.767 | -0.79 | -0.5 | 1.761 | -0.799 | -0.506 |
| 21 | 1.768 | -0.787 | -0.499 | 1.761 | -0.799 | -0.507 | 1.75 | -0.816 | -0.517 | 1.743 | -0.828 | -0.525 |
| 20 | 1.753 | -0.812 | -0.515 | 1.745 | -0.825 | -0.523 | 1.733 | -0.842 | -0.534 | 1.724 | -0.857 | -0.543 |
| 19 | 1.737 | -0.837 | -0.53 | 1.728 | -0.85 | -0.539 | 1.716 | -0.869 | -0.551 | 1.705 | -0.887 | -0.562 |
| 18 | 1.718 | -0.866 | -0.549 | 1.711 | -0.877 | -0.556 | 1.699 | -0.897 | -0.569 | 1.685 | -0.918 | -0.582 |
| 17 | 1.7 | -0.894 | -0.567 | 1.694 | -0.904 | -0.573 | 1.681 | -0.925 | -0.587 | 1.665 | -0.95 | -0.602 |
| 16 | 1.682 | -0.924 | -0.585 | 1.676 | -0.932 | -0.591 | 1.662 | -0.955 | -0.605 | 1.645 | -0.982 | -0.623 |
| 15 | 1.665 | -0.951 | -0.603 | 1.658 | -0.962 | -0.61 | 1.643 | -0.985 | -0.624 | 1.623 | -1.016 | -0.644 |
| 14 | 1.644 | -0.983 | -0.623 | 1.639 | -0.992 | -0.629 | 1.623 | -1.016 | -0.644 | 1.601 | -1.051 | -0.666 |
| 13 | 1.627 | -1.01 | -0.64 | 1.619 | -1.023 | -0.649 | 1.603 | -1.048 | -0.664 | 1.578 | -1.087 | -0.689 |
| 12 | 1.609 | -1.038 | -0.658 | 1.598 | -1.056 | -0.669 | 1.582 | -1.081 | -0.685 | 1.554 | -1.125 | -0.713 |
| 11 | 1.585 | -1.076 | -0.682 | 1.576 | -1.09 | -0.691 | 1.56 | -1.116 | -0.708 | 1.529 | -1.165 | -0.738 |
| 10 | 1.56 | -1.115 | -0.707 | 1.554 | -1.126 | -0.714 | 1.536 | -1.153 | -0.731 | 1.502 | -1.207 | -0.765 |
| 9 | 1.537 | -1.152 | -0.73 | 1.53 | -1.163 | -0.737 | 1.511 | -1.192 | -0.756 | 1.474 | -1.251 | -0.793 |
| 8 | 1.513 | -1.189 | -0.754 | 1.504 | -1.203 | -0.763 | 1.485 | -1.234 | -0.782 | 1.444 | -1.299 | -0.824 |
| 7 | 1.485 | -1.235 | -0.783 | 1.477 | -1.246 | -0.79 | 1.456 | -1.279 | -0.811 | 1.411 | -1.351 | -0.856 |
| 6 | 1.456 | -1.279 | -0.811 | 1.448 | -1.293 | -0.819 | 1.425 | -1.328 | -0.842 | 1.375 | -1.408 | -0.892 |
| 5 | 1.425 | -1.328 | -0.842 | 1.415 | -1.344 | -0.852 | 1.39 | -1.383 | -0.877 | 1.335 | -1.471 | -0.933 |
| 4 | 1.388 | -1.388 | -0.88 | 1.378 | -1.404 | -0.89 | 1.351 | -1.446 | -0.917 | 1.288 | -1.544 | -0.979 |
| 3 | 1.347 | -1.451 | -0.92 | 1.332 | -1.475 | -0.935 | 1.303 | -1.521 | -0.964 | 1.233 | -1.632 | -1.034 |
| 2 | 1.28 | -1.558 | -0.987 | 1.274 | -1.567 | -0.993 | 1.243 | -1.616 | -1.024 | 1.162 | -1.744 | -1.105 |
| 1 | 1.205 | -1.676 | -1.062 | 1.192 | -1.696 | -1.075 | 1.153 | -1.757 | -1.114 | 1.056 | -1.911 | -1.211 |
|  |  |  |  | 2.26 | = 0 |  | 267, | \$ |  |  |  |  |


| Marine Cargo Net Data: Quantiles \& VAR Mean $=2.144$, S.D. $=0.613$, Provision $=2.144,1$ Unit $=\$ 1,000$ |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V -a-R | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | V-a-R |
| 99.995 | 7.015 | 7.952 | 4.871 | 7.257 | 8.347 | 5.113 | 6.131 | 6.508 | 3.987 | 5.384 | 5.29 | 3.2 |
| 99.99 | 6.959 | 7.859 | 4.815 | 7.096 | 8.083 | 4.952 | 5.843 | 6.039 | 3.699 | 5.196 | 4.983 | 3.052 |
| 99.98 | 6.949 | 7.843 | 4.805 | 6.836 | 7.659 | 4.692 | 5.558 | 5.573 | 3.414 | 5.005 | 4.669 | 2.8 |
| 99.97 | 6.477 | 7.073 | 4.333 | 6.545 | 7.185 | 4.401 | 5.391 | 5.301 | 3.247 | 4.89 | 4.483 | 2.746 |
| 99.96 | 5.853 | 6.054 | 3.709 | 6.147 | 6.535 | 4.003 | 5.274 | 5.109 | 3.13 | 4.809 | 4.35 | 2.66 |
| 99.95 | 5.788 | 5.948 | 3.644 | 5.88 | 6.099 | 3.736 | 5.183 | 4.96 | 3.039 | 4.744 | 4.245 | 2.6 |
| 99.94 | 5.699 | 5.803 | 3.555 | 5.712 | 5.825 | 3.568 | 5.108 | 4.839 | 2.964 | 4.692 | 4.159 | 2.548 |
| 99.93 | 5.484 | 5.451 | 3.339 | 5.587 | 5.621 | 3.443 | 5.045 | 4.736 | 2.901 | 4.647 | 4.086 | 2.503 |
| 99.92 | 5.311 | 5.17 | 3.167 | 5.492 | 5.465 | 3.348 | 4.991 | 4.647 | 2.847 | 4.608 | 4.022 | 2.464 |
| 99.91 | 5.224 | 5.028 | 3.08 | 5.417 | 5.342 | 3.273 | 4.943 | 4.569 | 2.799 | 4.573 | 3.965 | 2.429 |
| 99.9 | 5.221 | 5.023 | 3.077 | 5.356 | 5.244 | 3.212 | 4.9 | 4.499 | 2.756 | 4.542 | 3.914 | 2.398 |
| 99.8 | 4.998 | 4.659 | 2.854 | 5.016 | 4.688 | 2.872 | 4.617 | 4.037 | 2.473 | 4.334 | 3.574 | 2.189 |
| 99.7 | 4.795 | 4.327 | 2.651 | 4.8 | 4.335 | 2.656 | 4.452 | 3.767 | 2.308 | 4.208 | 3.37 | 2.06 |
| 99.6 | 4.588 | 3.989 | 2.443 | 4.629 | 4.057 | 2.485 | 4.334 | 3.575 | 2.19 | 4.118 | 3.222 | 1.974 |
| 99.5 | 4.445 | 3.756 | 2.301 | 4.497 | 3.84 | 2.352 | 4.242 | 3.425 | 2.098 | 4.047 | 3.106 | 1.903 |
| 99.4 | 4.354 | 3.608 | 2.21 | 4.39 | 3.666 | 2.246 | 4.167 | 3.303 | 2.023 | 3.988 | 3.01 | . 34 |
| 99.3 | 4.28 | 3.487 | 2.136 | 4.3 | 3.519 | 2.156 | 4.104 | 3.199 | 1.96 | 3.938 | 2.928 | 1.794 |
| 99.2 | 4.205 | 3.364 | 2.061 | 4.222 | 3.392 | 2.078 | 4.049 | 3.109 | 1.905 | 3.894 | 2.857 | 1.75 |
| 99.1 | 4.122 | 3.229 | 1.978 | 4.153 | 3.28 | 2.009 | 4 | 3.029 | 1.856 | 3.855 | 2.793 | 1.711 |
| 99 | 4.034 | 3.085 | 1.89 | 4.093 | 3.181 | 1.949 | 3.956 | 2.958 | 1.812 | 3.82 | 2.735 | 1.676 |
| 98 | 3.71 | 2.556 | 1.566 | 3.731 | 2.591 | 1.587 | 3.665 | 2.483 | 1.521 | 3.581 | 2.346 | 1.437 |
| 97 | 3.501 | 2.215 | 1.357 | 3.529 | 2.261 | 1.385 | 3.492 | 2.2 | 1.348 | 3.434 | 2.107 | 1.29 |
| 96 | 3.36 | 1.986 | 1.216 | 3.387 | 2.029 | 1.243 | 3.367 | 1.996 | 1.223 | 3.327 | 1.931 | 1.183 |
| 95 | 3.255 | 1.813 | 1.111 | 3.277 | 1.85 | 1.133 | 3.269 | 1.836 | 1.124 | 3.241 | 1.791 | 1.097 |
| 94 | 3.169 | 1.674 | 1.025 | 3.187 | 1.703 | 1.043 | 3.187 | 1.703 | 1.043 | 3.169 | 1.673 | 1.025 |
| 93 | 3.095 | 1.552 | 0.951 | 3.111 | 1.578 | 0.967 | 3.117 | 1.589 | 0.973 | 3.107 | 1.571 | 0.963 |
| 92 | 3.026 | 1.44 | 0.882 | 3.044 | 1.469 | 0.9 | 3.056 | 1.489 | 0.912 | 3.052 | 1.482 | 0.908 |
| 91 | 2.971 | 1.349 | 0.827 | 2.985 | 1.373 | 0.841 | 3.001 | 1.4 | 0.857 | 3.002 | 1.401 | 0.858 |
| 90 | 2.915 | 1.259 | 0.771 | 2.933 | 1.287 | 0.789 | 2.952 | 1.319 | 0.808 | 2.957 | 1.327 | 0.813 |
| 89 | 2.87 | 1.185 | 0.726 | 2.885 | 1.21 | 0.741 | 2.907 | 1.245 | 0.763 | 2.915 | 1.259 | 0.771 |
| 88 | 2.827 | 1.115 | 0.683 | 2.842 | 1.139 | 0.698 | 2.865 | 1.177 | 0.721 | 2.877 | 1.196 | 0.733 |
| 87 | 2.785 | 1.046 | 0.641 | 2.802 | 1.073 | 0.658 | 2.826 | 1.114 | 0.682 | 2.841 | 1.137 | 0.697 |
| 86 | 2.75 | 0.99 | 0.606 | 2.765 | 1.013 | 0.621 | 2.79 | 1.055 | 0.646 | 2.807 | 1.082 | 0.663 |
| 85 | 2.719 | 0.939 | 0.575 | 2.73 | 0.957 | 0.586 | 2.756 | 0.999 | 0.612 | 2.775 | 1.03 | 0.631 |
| 84 | 2.685 | 0.883 | 0.541 | 2.698 | 0.904 | 0.554 | 2.724 | 0.946 | 0.58 | 2.744 | 0.98 | 0.6 |
| 83 | 2.652 | 0.829 | 0.508 | 2.667 | 0.854 | 0.523 | 2.693 | 0.896 | 0.549 | 2.715 | 0.933 | 0.571 |
| 82 | 2.625 | 0.784 | 0.481 | 2.639 | 0.808 | 0.495 | 2.664 | 0.849 | 0.52 | 2.688 | 0.887 | 0.544 |
| 81 | 2.6 | 0.744 | 0.456 | 2.612 | 0.763 | 0.468 | 2.636 | 0.803 | 0.492 | 2.661 | 0.844 | 0.517 |
| 80 | 2.575 | 0.703 | 0.431 | 2.586 | 0.721 | 0.442 | 2.61 | 0.76 | 0.466 | 2.635 | 0.802 | 0.491 |
| 79 | 2.547 | 0.658 | 0.403 | 2.561 | 0.681 | 0.417 | 2.584 | 0.718 | 0.44 | 2.611 | 0.762 | 0.467 |
| 78 | 2.527 | 0.625 | 0.383 | 2.537 | 0.642 | 0.393 | 2.559 | 0.678 | 0.415 | 2.587 | 0.723 | 0.443 |
| 77 | 2.504 | 0.588 | 0.36 | 2.515 | 0.605 | 0.371 | 2.535 | 0.639 | 0.391 | 2.564 | 0.685 | 0.42 |
| 76 | 2.487 | 0.56 | 0.343 | 2.493 | 0.569 | 0.349 | 2.512 | 0.601 | 0.368 | 2.541 | 0.648 | 0.397 |
| 75 | 2.466 | 0.525 | 0.322 | 2.471 | 0.534 | 0.327 | 2.49 | 0.565 | 0.346 | 2.52 | 0.613 | 0.375 |
| 74 | 2.445 | 0.491 | 0.301 | 2.45 | 0.5 | 0.306 | 2.468 | 0.53 | 0.324 | 2.498 | 0.578 | 0.354 |
| 73 | 2.424 | 0.458 | 0.28 | 2.43 | 0.467 | 0.286 | 2.447 | 0.495 | 0.303 | 2.478 | 0.544 | 0.334 |
| 72 | 2.406 | 0.428 | 0.262 | 2.41 | 0.435 | 0.266 | 2.427 | 0.462 | 0.283 | 2.457 | 0.512 | 0.313 |
| 71 | 2.389 | 0.4 | 0.245 | 2.391 | 0.403 | 0.247 | 2.407 | 0.429 | 0.263 | 2.438 | 0.479 | 0.294 |
| 70 | 2.37 | 0.369 | 0.226 | 2.372 | 0.372 | 0.228 | 2.388 | 0.397 | 0.243 | 2.418 | 0.448 | 0.274 |
| 69 | 2.351 | 0.338 | 0.207 | 2.354 | 0.342 | 0.209 | 2.368 | 0.366 | 0.224 | 2.399 | 0.417 | 0.255 |
| 68 | 2.334 | 0.31 | 0.19 | 2.335 | 0.312 | 0.191 | 2.35 | 0.336 | 0.206 | 2.381 | 0.386 | 0.237 |
| 67 | 2.313 | 0.276 | 0.169 | 2.317 | 0.283 | 0.173 | 2.332 | 0.306 | 0.188 | 2.363 | 0.357 | 0.219 |
| 66 | 2.297 | 0.25 | 0.153 | 2.3 | 0.254 | 0.156 | 2.314 | 0.277 | 0.17 | 2.345 | 0.327 | 0.201 |
| 65 | 2.278 | 0.219 | 0.134 | 2.282 | 0.226 | 0.138 | 2.296 | 0.248 | 0.152 | 2.327 | 0.299 | 0.183 |
| 64 | 2.26 | 0.189 | 0.116 | 2.265 | 0.198 | 0.121 | 2.279 | 0.22 | 0.135 | 2.31 | 0.27 | 0.166 |
| 63 | 2.245 | 0.165 | 0.101 | 2.249 | 0.171 | 0.105 | 2.262 | 0.193 | 0.118 | 2.293 | 0.242 | 0.148 |
| 62 | 2.229 | 0.138 | 0.085 | 2.232 | 0.144 | 0.088 | 2.245 | 0.166 | 0.101 | 2.276 | 0.215 | 0.132 |

TABLE CONTINUED ON THE NEXT PAGE

| Quantile Statistics and Value at Risk (Acc Year: Total) |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \% | Sample |  |  | Kernel |  |  | LogNormal |  |  | Gamma |  |  |
|  | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | $\mathrm{V}-\mathrm{a}-\mathrm{R}$ | Quantile | \# S.D.'s | V-a-R | Quantile | \# S.D.'s | $\mathrm{V}-\mathrm{a}-\mathrm{R}$ |
| 61 | 2.212 | 0.112 | 0.068 | 2.216 | 0.117 | 0.072 | 2.229 | 0.139 | 0.085 | 2.259 | 0.188 | 0.115 |
| 60 | 2.197 | 0.086 | 0.053 | 2.2 | 0.091 | 0.056 | 2.213 | 0.112 | 0.069 | 2.242 | 0.161 | 0.098 |
| 59 | 2.183 | 0.063 | 0.039 | 2.184 | 0.066 | 0.04 | 2.197 | 0.086 | 0.053 | 2.226 | 0.134 | 0.082 |
| 58 | 2.166 | 0.036 | 0.022 | 2.169 | 0.04 | 0.025 | 2.181 | 0.061 | 0.037 | 2.21 | 0.108 | 0.066 |
| 57 | 2.152 | 0.013 | 0.008 | 2.153 | 0.015 | 0.009 | 2.166 | 0.035 | 0.022 | 2.194 | 0.082 | 0.05 |
| 56 | 2.137 | -0.012 | -0.007 | 2.138 | -0.009 | -0.006 | 2.15 | 0.01 | 0.006 | 2.178 | 0.056 | 0.034 |
| 55 | 2.121 | -0.037 | -0.023 | 2.123 | -0.034 | -0.021 | 2.135 | -0.014 | -0.009 | 2.163 | 0.03 | 0.019 |
| 54 | 2.107 | -0.06 | -0.037 | 2.108 | -0.058 | -0.036 | 2.12 | -0.039 | -0.024 | 2.147 | 0.005 | 0.003 |
| 53 | 2.092 | -0.085 | -0.052 | 2.094 | -0.082 | -0.05 | 2.105 | -0.063 | -0.039 | 2.132 | -0.02 | -0.012 |
| 52 | 2.079 | -0.106 | -0.065 | 2.079 | -0.106 | -0.065 | 2.091 | -0.087 | -0.053 | 2.116 | -0.045 | -0.028 |
| 51 | 2.066 | -0.127 | -0.078 | 2.065 | -0.129 | -0.079 | 2.076 | -0.111 | -0.068 | 2.101 | -0.07 | -0.043 |
| 50 | 2.052 | -0.15 | -0.092 | 2.051 | -0.153 | -0.094 | 2.062 | -0.135 | -0.082 | 2.086 | -0.095 | -0.058 |
| 49 | 2.039 | -0.172 | -0.105 | 2.036 | -0.176 | -0.108 | 2.047 | -0.158 | -0.097 | 2.071 | -0.119 | -0.073 |
| 48 | 2.023 | -0.197 | -0.121 | 2.022 | -0.199 | -0.122 | 2.033 | -0.181 | -0.111 | 2.056 | -0.144 | -0.088 |
| 47 | 2.009 | -0.221 | -0.136 | 2.008 | -0.222 | -0.136 | 2.019 | -0.205 | -0.125 | 2.041 | -0.168 | -0.103 |
| 46 | 1.994 | -0.246 | -0.151 | 1.994 | -0.244 | -0.15 | 2.004 | -0.228 | -0.14 | 2.026 | -0.193 | -0.118 |
| 45 | 1.978 | -0.271 | -0.166 | 1.981 | -0.267 | -0.163 | 1.99 | -0.251 | -0.154 | 2.011 | -0.217 | -0.133 |
| 44 | 1.967 | -0.289 | -0.177 | 1.967 | -0.289 | -0.177 | 1.976 | -0.274 | -0.168 | 1.996 | -0.241 | -0.148 |
| 43 | 1.951 | -0.315 | -0.193 | 1.954 | -0.311 | -0.19 | 1.962 | -0.297 | -0.182 | 1.981 | -0.265 | -0.163 |
| 42 | 1.938 | -0.336 | -0.206 | 1.94 | -0.333 | -0.204 | 1.948 | -0.319 | -0.196 | 1.967 | -0.29 | -0.177 |
| 41 | 1.925 | -0.358 | -0.219 | 1.927 | -0.355 | -0.217 | 1.934 | -0.342 | -0.21 | 1.952 | -0.314 | -0.192 |
| 40 | 1.911 | -0.38 | -0.233 | 1.914 | -0.376 | -0.23 | 1.921 | -0.365 | -0.224 | 1.937 | -0.338 | -0.207 |
| 39 | 1.899 | -0.399 | -0.245 | 1.9 | -0.398 | -0.244 | 1.907 | -0.388 | -0.237 | 1.922 | -0.362 | -0.222 |
| 38 | 1.888 | -0.418 | -0.256 | 1.887 | -0.419 | -0.257 | 1.893 | -0.41 | -0.251 | 1.907 | -0.387 | -0.237 |
| 37 | 1.875 | -0.439 | -0.269 | 1.874 | -0.44 | -0.27 | 1.879 | -0.433 | -0.265 | 1.892 | -0.411 | -0.252 |
| 36 | 1.862 | -0.46 | -0.282 | 1.861 | -0.462 | -0.283 | 1.865 | -0.456 | -0.279 | 1.877 | -0.435 | -0.267 |
| 35 | 1.85 | -0.48 | -0.294 | 1.848 | -0.483 | -0.296 | 1.851 | -0.479 | -0.293 | 1.862 | -0.46 | -0.282 |
| 34 | 1.838 | -0.499 | -0.306 | 1.835 | -0.504 | -0.309 | 1.837 | -0.502 | -0.307 | 1.847 | -0.485 | -0.297 |
| 33 | 1.826 | -0.52 | -0.318 | 1.822 | -0.525 | -0.322 | 1.823 | -0.524 | -0.321 | 1.832 | -0.509 | -0.312 |
| 32 | 1.813 | -0.54 | -0.331 | 1.809 | -0.546 | -0.335 | 1.809 | -0.548 | -0.335 | 1.817 | -0.534 | -0.327 |
| 31 | 1.801 | -0.561 | -0.343 | 1.796 | -0.568 | -0.348 | 1.794 | -0.571 | -0.35 | 1.801 | -0.56 | -0.343 |
| 30 | 1.789 | -0.58 | -0.356 | 1.783 | -0.589 | -0.361 | 1.78 | -0.594 | -0.364 | 1.786 | -0.585 | -0.358 |
| 29 | 1.774 | -0.604 | -0.37 | 1.77 | -0.611 | -0.374 | 1.766 | -0.618 | -0.378 | 1.77 | -0.611 | -0.374 |
| 28 | 1.761 | -0.626 | -0.383 | 1.757 | -0.633 | -0.387 | 1.751 | -0.641 | -0.393 | 1.754 | -0.636 | -0.39 |
| 27 | 1.748 | -0.646 | -0.396 | 1.743 | -0.655 | -0.401 | 1.737 | -0.665 | -0.408 | 1.738 | -0.663 | -0.406 |
| 26 | 1.734 | -0.67 | -0.41 | 1.729 | -0.677 | -0.415 | 1.722 | -0.689 | -0.422 | 1.722 | -0.689 | -0.422 |
| 25 | 1.72 | -0.692 | -0.424 | 1.716 | -0.699 | -0.428 | 1.707 | -0.714 | -0.437 | 1.705 | -0.716 | -0.439 |
| 24 | 1.707 | -0.713 | -0.437 | 1.702 | -0.722 | -0.442 | 1.692 | -0.739 | -0.452 | 1.689 | -0.743 | -0.455 |
| 23 | 1.694 | -0.735 | -0.45 | 1.688 | -0.745 | -0.456 | 1.676 | -0.764 | -0.468 | 1.672 | -0.771 | -0.472 |
| 22 | 1.68 | -0.757 | -0.464 | 1.673 | -0.768 | -0.471 | 1.661 | -0.789 | -0.483 | 1.655 | -0.799 | -0.49 |
| 21 | 1.664 | -0.784 | -0.48 | 1.659 | -0.792 | -0.485 | 1.645 | -0.815 | -0.499 | 1.637 | -0.828 | -0.507 |
| 20 | 1.649 | -0.807 | -0.495 | 1.644 | -0.817 | -0.5 | 1.629 | -0.841 | -0.515 | 1.619 | -0.857 | -0.525 |
| 19 | 1.635 | -0.831 | -0.509 | 1.629 | -0.841 | -0.515 | 1.612 | -0.868 | -0.532 | 1.601 | -0.887 | -0.543 |
| 18 | 1.618 | -0.859 | -0.526 | 1.613 | -0.866 | -0.531 | 1.595 | -0.896 | -0.549 | 1.582 | -0.918 | -0.562 |
| 17 | 1.604 | -0.882 | -0.54 | 1.598 | -0.892 | -0.546 | 1.578 | -0.924 | -0.566 | 1.563 | -0.949 | -0.582 |
| 16 | 1.588 | -0.908 | -0.556 | 1.581 | -0.919 | -0.563 | 1.56 | -0.953 | -0.584 | 1.543 | -0.982 | -0.601 |
| 15 | 1.569 | -0.938 | -0.575 | 1.565 | -0.946 | -0.579 | 1.542 | -0.983 | -0.602 | 1.522 | -1.015 | -0.622 |
| 14 | 1.554 | -0.963 | -0.59 | 1.548 | -0.974 | -0.596 | 1.523 | -1.013 | -0.621 | 1.501 | -1.05 | -0.643 |
| 13 | 1.539 | -0.988 | -0.605 | 1.53 | -1.003 | -0.614 | 1.504 | -1.045 | -0.64 | 1.479 | -1.086 | -0.665 |
| 12 | 1.522 | -1.015 | -0.622 | 1.511 | -1.033 | -0.633 | 1.483 | -1.079 | -0.661 | 1.456 | -1.124 | -0.688 |
| 11 | 1.499 | -1.052 | -0.645 | 1.492 | -1.065 | -0.652 | 1.462 | -1.113 | -0.682 | 1.431 | -1.163 | -0.713 |
| 10 | 1.478 | -1.088 | -0.666 | 1.471 | -1.098 | -0.673 | 1.44 | -1.15 | -0.704 | 1.406 | -1.205 | -0.738 |
| 9 | 1.459 | -1.118 | -0.685 | 1.45 | -1.134 | -0.694 | 1.416 | -1.189 | -0.728 | 1.379 | -1.249 | -0.765 |
| 8 | 1.437 | -1.155 | -0.707 | 1.426 | -1.172 | -0.718 | 1.391 | -1.23 | -0.753 | 1.35 | -1.297 | -0.794 |
| 7 | 1.408 | -1.201 | -0.736 | 1.401 | -1.213 | -0.743 | 1.363 | -1.274 | -0.781 | 1.318 | -1.348 | -0.826 |
| 6 | 1.382 | -1.244 | -0.762 | 1.374 | -1.257 | -0.77 | 1.334 | -1.323 | -0.811 | 1.284 | -1.404 | -0.86 |
| 5 | 1.355 | -1.289 | -0.79 | 1.343 | -1.308 | -0.801 | 1.3 | -1.377 | -0.844 | 1.245 | -1.467 | -0.899 |
| 4 | 1.317 | -1.349 | -0.827 | 1.308 | -1.365 | -0.836 | 1.262 | -1.439 | -0.882 | 1.201 | -1.54 | -0.943 |
| 3 | 1.277 | -1.416 | -0.867 | 1.266 | -1.434 | -0.878 | 1.217 | -1.513 | -0.927 | 1.148 | -1.626 | -0.996 |
| 2 | 1.223 | -1.503 | -0.921 | 1.211 | -1.523 | -0.933 | 1.16 | -1.607 | -0.985 | 1.08 | -1.737 | -1.064 |
| 1 | 1.146 | -1.63 | -0.998 | 1.131 | -1.653 | -1.013 | 1.074 | -1.746 | -1.07 | 0.979 | -1.902 | -1.165 |
|  |  |  |  | ean $=2.144$, | S.D. $=0.613$ | , Provisi | = 2.144, 1 U | nit $=\$ 1,000$ |  |  |  |  |



The model for the Gross reserves is shown above. To be conservative, the $2.48 \%+-1.7 \%$ calendar trend was retained even though it is not statistically significant. Process variance is very high from development year 3 onward.


The model for the Net reserves is shown above. To be conservative, the 2.48\% +- 1.7\%
calendar trend was retained even though it is not statistically significant. Process variance is very high from development year 3 onward and is higher than for the Gross data!

Marine Cargo Gross Net:Composite DS:MPTF[optimal2-1]:Weighted Residual Covariances Between Datasets:Final Correlations

| Final Weighted Residual Correlations Between Datasets |  |  |  |
| :---: | ---: | ---: | :---: |
|  | Marine Cargo GrossAll_Total:PL(I) | Marine Cargo NetAll_Total:PL(I) |  |
| Marine Cargo GrossAll_Total:PL(I) | $\mathbf{1}$ | 0.902865 |  |
| Marine Cargo NetAll_Total:PL(I) | 0.902865 | 1 |  |
| 4 |  |  |  |

The process correlation (90\%) between Gross and Net data is very high (as we would expect).

Marine Cargo Gross Net:Composite DS:MPTF[optimal2-1]:Forecast:Marine Cargo GrossAll_Total:PL(I):Reserve Forecast Table

| Accident Period vs Development Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cal. Per. Total | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | Reserve | Ultimate |
| 1988 | 338 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 1,326 |
|  | 359 | 359 | 750 | 172 | 27 | 25 | 2 | 2 | 0 | 2 | -11 | 0 | 1 | -3 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1989 | 1,121 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1,538 |
|  | 1,075 | 325 | 1,026 | 180 | 45 | -5 | 2 | -6 | 0 | 1 | 1 | 1 | -29 | 1 | 1 | 1 | -6 | 0 | 1 | 1 | 1 |
| 1990 | 1,301 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 1,550 |
|  | 1,619 | 422 | 864 | 276 | 12 | -6 | -13 | -2 | 1 | -5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 1 | 2 | 2 |
| 1991 | 1,384 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 2 | 1,544 |
|  | 1,425 | 354 | 850 | 252 | 47 | 42 | 2 | -11 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 1 | 1 | 3 | 3 |
| 1992 | 1,408 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 3 | 1,154 |
|  | 1,470 | 275 | 652 | 170 | 22 | -2 | 0 | 36 | -3 | 1 | 0 | 0 | 0 | 0 | 0 | 2 | 2 | 1 | 1 | 3 | 3 |
| 1993 | 1,415 | 338 | 783 | 180 | 83 | 24 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 5 | 1,318 |
|  | 1,198 | 285 | 829 | 127 | 36 | 38 | 2 | 5 | -7 | 2 | 0 | 0 | 0 | -4 | 2 | 2 | 2 | 1 | 1 | 4 | 4 |
| 1994 | 1,421 | 338 | 783 | 180 | 83 | 25 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 7 | 1,337 |
|  | 1,342 | 297 | 647 | 190 | 236 | 32 | 7 | -38 | 25 | -54 | 11 | -21 | 0 | 3 | 2 | 2 | 2 | 1 | 1 | 5 | 5 |
| 1995 | 1,425 | 338 | 783 | 180 | 85 | 26 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 9 | 1,266 |
|  | 1,060 | 242 | 681 | 194 | 74 | 58 | -2 | 28 | 0 | -2 | -15 | 0 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 7 | 7 |
| 1996 | 1,429 | 338 | 783 | 184 | 87 | 26 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 12 | 1,166 |
|  | 1,093 | 187 | 710 | 197 | 23 | 20 | 5 | 12 | -4 | 2 | 3 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 1 | 8 | 8 |
| 1997 | 1,432 | 338 | 802 | 189 | 89 | 27 | 8 | 7 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 16 | 1,144 |
|  | 1,310 | 152 | 908 | 90 | 1 | -6 | 23 | 23 | -10 | -53 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 1 | 10 | 10 |
| 1998 | 1,469 | 346 | 822 | 194 | 92 | 28 | 8 | 7 | 5 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 21 | 951 |
|  | 1,429 | 180 | 668 | 88 | 2 | 0 | -16 | 8 | 1 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 1 | 13 | 13 |
| 1999 | 1,507 | 355 | 842 | 199 | 94 | 28 | 9 | 7 | 6 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 27 | 1,689 |
|  | 1,284 | 433 | 1,136 | 96 | -1 | 3 | -1 | -5 | 8 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 15 | 15 |
| 2000 | 1,547 | 364 | 864 | 204 | 96 | 29 | 9 | 7 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 35 | 1,332 |
|  | 1,560 | 393 | 708 | 222 | 35 | -44 | -16 | 10 | 8 | 7 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 19 | 19 |
| 2001 | 1,588 | 373 | 886 | 209 | 99 | 30 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 45 | 1,431 |
|  | 1,348 | 488 | 788 | 63 | 31 | 16 | 13 | 11 | 8 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 24 | 24 |
| 2002 | 1,629 | 383 | 909 | 215 | 102 | 31 | 9 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 77 | 1,666 |
|  | 1,263 | 271 | 1,041 | 210 | 68 | 44 | 14 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 52 | 52 |
| 2003 | 1,673 | 393 | 933 | 220 | 104 | 32 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 183 | 1,855 |
|  | 1,534 | 381 | 1,231 | 60 | 147 | 45 | 14 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 2 | 158 | 158 |
| 2004 | 1,718 | 403 | 958 | 226 | 107 | 32 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 415 | 1,947 |
|  | 2,039 | 653 | 879 | 104 | 152 | 47 | 15 | 12 | 9 | 7 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 2 | 198 | 198 |
| 2005 | 1,764 | 414 | 984 | 233 | 110 | 33 | 10 | 8 | 7 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1,410 | 1,948 |
|  | 1,486 | 538 | 460 | 108 | 156 | 48 | 15 | 12 | 10 | 8 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 2 | 521 | 521 |
|  | Total Fitted/Paid |  | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | Total Reserve | Total Ultimate |
| Cal. Per. Total | 25,568 |  | 1,387 | 413 | 186 | 78 | 46 | 36 | 29 | 23 | 18 | 14 | 11 | 9 | 7 | 5 | 3 | 2 | 1 | 2,267 | 26,162 |
|  | 23,895 |  | 510 | 199 | 167 | 55 | 26 | 21 | 18 | 15 | 13 | 11 | 9 | 8 | 6 | 5 | 4 | 3 | 2 | 634 | 634 |
| 1 Unit = \$1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

The above table shows the forecasted reserve distribution for each cell (mean is in black, standard deviation in red) along with the data (mean in black, observed in blue).

| Accident Yr Summary |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Acc. Yr | Mean |  | Standard Dev. | cV |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1988 | 0 | 1,326 | 0 | *** | 0 |
| 1989 | 1 | 1,538 | 1 | 1.84 | 0 |
| 1990 | 1 | 1,550 | 2 | 1.35 | 0 |
| 1991 | 2 | 1,544 | 3 | 1.13 | 0 |
| 1992 | 3 | 1,154 | 3 | 0.99 | 0 |
| 1993 | 5 | 1,318 | 4 | 0.89 | 0 |
| 1994 | 7 | 1,337 | 5 | 0.81 | 0 |
| 1995 | 9 | 1,266 | 7 | 0.75 | 0.01 |
| 1996 | 12 | 1,166 | 8 | 0.69 | 0.01 |
| 1997 | 16 | 1,144 | 10 | 0.65 | 0.01 |
| 1998 | 21 | 951 | 13 | 0.61 | 0.01 |
| 1999 | 27 | 1,689 | 15 | 0.58 | 0.01 |
| 2000 | 35 | 1,332 | 19 | 0.56 | 0.01 |
| 2001 | 45 | 1,431 | 24 | 0.55 | 0.02 |
| 2002 | 77 | 1,666 | 52 | 0.67 | 0.03 |
| 2003 | 183 | 1,855 | 158 | 0.86 | 0.09 |
| 2004 | 415 | 1,947 | 198 | 0.48 | 0.1 |
| 2005 | 1,410 | 1,948 | 521 | 0.37 | 0.27 |
| Total | 2,267 | 26,162 | 634 | 0.28 | 0.02 |
|  |  |  |  |  |  |

Marine Cargo Gross Net:Composite DS:MPTF[optimal2-2]:Forecast:Marine Cargo NetAll_Total:PL(I):Reserve Forecast Table

| Accident Period vs Development Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cal. Per. Total | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | Reserve | Ultimate |
| 1988 | 318 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 0 | 1,125 |
|  | 297 | 297 | 628 | 155 | 25 | 25 | 2 | 2 | 2 | 2 | -11 | 0 | 1 | -3 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1989 | 1,055 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 1 | 1 | 1,255 |
|  | 910 | 282 | 788 | 176 | 43 | -6 | 2 | -7 | 0 | 1 | 1 | 1 | -30 | 8 | 1 | 1 | -6 | 0 | 2 | 2 | 2 |
| 1990 | 1,224 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 2 | 1,354 |
|  | 1,325 | 382 | 755 | 233 | 7 | -6 | -13 | -2 | 0 | -5 | 0 | 0 | 0 | 0 | 2 | 0 | 0 | 2 | 2 | 3 | 3 |
| 1991 | 1,263 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 3 | 1,283 |
|  | 1,243 | 288 | 723 | 242 | -4 | 32 | 0 | -6 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 2 | 2 | 2 | 4 | 4 |
| 1992 | 1,289 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 1 | 5 | 1,025 |
|  | 1,262 | 237 | 572 | 159 | 19 | 2 | 0 | 36 | -3 | 1 | -3 | 0 | 0 | 0 | 0 | 3 | 2 | 2 | 2 | 5 | 5 |
| 1993 | 1,299 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 7 | 1,149 |
|  | 1,061 | 245 | 705 | 120 | 38 | 37 | 4 | 5 | -11 | 2 | 0 | 0 | 0 | -4 | 3 | 3 | 2 | 2 | 2 | 7 | 7 |
| 1994 | 1,306 | 318 | 737 | 169 | 39 | 26 | 10 | 8 | 7 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 9 | 1,187 |
|  | 1,063 | 205 | 588 | 222 | 71 | 32 | 7 | -8 | 27 | 6 | 45 | -19 | 0 | 4 | 3 | 3 | 3 | 2 | 2 | 8 | 8 |
| 1995 | 1,312 | 318 | 737 | 169 | 40 | 27 | 10 | 8 | 7 | 5 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 12 | 1,132 |
|  | 959 | 218 | 618 | 163 | 72 | 56 | -3 | 27 | -18 | -19 | 13 | -7 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 10 | 10 |
| 1996 | 1,317 | 318 | 737 | 174 | 41 | 28 | 11 | 9 | 7 | 6 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 16 | 1,029 |
|  | 1,047 | 167 | 612 | 175 | 14 | 15 | 25 | 12 | -33 | 23 | 3 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 2 | 13 | 13 |
| 1997 | 1,320 | 318 | 755 | 178 | 42 | 28 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 1 | 21 | 882 |
|  | 1,016 | 148 | 712 | 51 | -5 | -1 | 8 | 11 | -9 | -53 | 8 | 6 | 5 | 5 | 4 | 3 | 3 | 2 | 2 | 16 | 16 |
| 1998 | 1,356 | 326 | 774 | 182 | 43 | 29 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 28 | 932 |
|  | 1,206 | 175 | 667 | 69 | 0 | 0 | -16 | 8 | 1 | 10 | 8 | 7 | 6 | 5 | 4 | 3 | 3 | 3 | 2 | 19 | 19 |
| 1999 | 1,392 | 334 | 793 | 187 | 44 | 30 | 12 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1 | 36 | 1,504 |
|  | 1,234 | 433 | 967 | 73 | -1 | 3 | -1 | -5 | 13 | 10 | 8 | 7 | 6 | 5 | 4 | 3 | 3 | 3 | 2 | 24 | 24 |
| 2000 | 1,429 | 342 | 813 | 192 | 45 | 30 | 12 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 46 | 1,262 |
|  | 1,385 | 393 | 629 | 219 | 35 | -44 | -16 | 16 | 13 | 11 | 9 | 7 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 30 | 30 |
| 2001 | 1,467 | 351 | 834 | 197 | 47 | 31 | 12 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 60 | 1,438 |
|  | 1,273 | 486 | 784 | 61 | 31 | 16 | 21 | 17 | 14 | 11 | 9 | 7 | 6 | 5 | 4 | 4 | 3 | 3 | 2 | 38 | 38 |
| 2002 | 1,506 | 360 | 856 | 202 | 48 | 32 | 13 | 10 | 8 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 2 | 1 | 1 | 93 | 1,679 |
|  | 1,284 | 271 | 1,038 | 210 | 68 | 55 | 22 | 17 | 14 | 11 | 9 | 8 | 6 | 5 | 4 | 4 | 3 | 3 | 3 | 69 | 69 |
| 2003 | 1,547 | 370 | 879 | 207 | 49 | 33 | 13 | 10 | 8 | 7 | 5 | 4 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 145 | 1,815 |
|  | 1,508 | 381 | 1,230 | 60 | 26 | 56 | 23 | 18 | 14 | 12 | 9 | 8 | 6 | 5 | 5 | 4 | , | 3 | 3 | 77 | 77 |
| 2004 | 1,588 | 380 | 902 | 213 | 50 | 34 | 13 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 2 | 2 | 2 | 1 | 1 | 362 | 1,893 |
|  | 2,087 | 652 | 878 | 111 | 27 | 58 | 23 | 19 | 15 | 12 | 10 | 8 | 7 | 6 | 5 | 4 | 4 | 3 | 3 | 142 | 142 |
| 2005 | 1,632 | 390 | 927 | 219 | 52 | 35 | 14 | 11 | 9 | 7 | 6 | 5 | 4 | 3 | 3 | 2 | 2 | 1 | 1 | 1,299 | 1,837 |
|  | 1,478 | 538 | 489 | 115 | 28 | 60 | 24 | 19 | 15 | 12 | 10 | 8 | 7 | 6 | 5 | 4 | 4 | 3 | 3 | 525 | 525 |
|  | Total Fitted/Paid |  | 2006 | 2007 | 2008 | 2009 | 2010 | 2011 | 2012 | 2013 | 2014 | 2015 | 2016 | 2017 | 2018 | 2019 | 2020 | 2021 | 2022 | Total Reserve | Total Ultimate |
| Cal. Per. Total | 23,620 |  | 1,276 | 358 | 143 | 94 | 61 | 49 | 39 | 31 | 24 | 19 | 15 | 12 | 9 | 6 | 4 | 3 | 1 | 2,144 | 23,782 |
|  | 21,638 |  | 518 | 141 | 77 | 73 | 41 | 33 | 27 | 23 | 19 | 16 | 14 | 11 | 9 | 8 | 6 | 4 | 3 | 613 | 613 |
| 1 Unit $=$ \$1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

The above table shows the forecasted reserve distribution for each cell (mean is in black, standard deviation in red) along with the data (mean in black, observed in blue).

| Accident Yr Summary |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Acc. Yr | Mean |  | Standard Dev. | cv |  |
|  | Reserve | Ultimate |  | Reserve | Ultimate |
| 1988 | 0 | 1,125 | 0 | *** | 0 |
| 1989 | 1 | 1,255 | 2 | 2.15 | 0 |
| 1990 | 2 | 1,354 | 3 | 1.56 | 0 |
| 1991 | 3 | 1,283 | 4 | 1.3 | 0 |
| 1992 | 5 | 1,025 | 5 | 1.13 | 0.01 |
| 1993 | 7 | 1,149 | 7 | 1.02 | 0.01 |
| 1994 | 9 | 1,187 | 8 | 0.93 | 0.01 |
| 1995 | 12 | 1,132 | 10 | 0.85 | 0.01 |
| 1996 | 16 | 1,029 | 13 | 0.79 | 0.01 |
| 1997 | 21 | 882 | 16 | 0.75 | 0.02 |
| 1998 | 28 | 932 | 19 | 0.71 | 0.02 |
| 1999 | 36 | 1,504 | 24 | 0.68 | 0.02 |
| 2000 | 46 | 1,262 | 30 | 0.65 | 0.02 |
| 2001 | 60 | 1,438 | 38 | 0.64 | 0.03 |
| 2002 | 93 | 1,679 | 69 | 0.73 | 0.04 |
| 2003 | 145 | 1,815 | 77 | 0.53 | 0.04 |
| 2004 | 362 | 1,893 | 142 | 0.39 | 0.07 |
| 2005 | 1,299 | 1,837 | 525 | 0.4 | 0.29 |
|  |  |  |  |  |  |
| Total | 2,144 | 23,782 | 613 | 0.29 | 0.03 |
|  |  |  |  |  |  |


[^0]:    Note that the CVs increase as we go down the accident years.
    This is because the model has not related the numbers in the triangle in any meaningful way to each other.

    Typically for a parsimonious model that capture the features of the data the CVs will decrease.

    The latest accident years have more cells to be forecast and therefore the CV of the aggregate should decrease rather than increase.

[^1]:    h respect of future calendar year trends we have adopted a consertive view and assumed that the
    future calendar trend reverts to the $8.39 \%+\quad 0.76 \%$ for six years (2005-2011) before reverting to zero.
    This assumption is critical to the forecast distributions.

