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Analysis of Bonus Systems in Motor Insurance

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#### Introduction

A continuous time Markov Chain model is proposed for the analysis of bonus systems in motor insurance which as a particular merit takes erplicitely into account the effect of cancellation of policies. Most analysis of bonus systems treat these as "closed" systems while in reality they are very "open" systems with a considerable change of the portfolio within one year. Relating to Danish experiences there is a yearly number of cancellations (excluding lapses i.e. failure to renew) corresponding to 15-20 % of the portfolios and a level of new business of approximately the same size making the real changes of the portfolios rather small.

The model is being used for tariffing purposes in deriving the overall level of the loss ratio - or rather the denominator part of it - by estimating expected total premium income for the whole portfolic or part of it given the values of certain external parameters. These are the number of new policies issued and trends in the claims frequency and the rate of cancellation. Special attention will be given to the analysis of data from a portfolic comprising about 150.000 policies which is one third of the whole motor portfolic of the Baltica Insurance Company. In this analysis stress will be put on the macro-aspect of the bonus system i.e. on the behaviour of the portfolic as a whole and not on the behaviour of the individual policy. As will be evident from the theoretical set-up this macro-view of the bonus system is also the core of the model.

#### 1. Description of the bonus system to be analyzed.

The bonus system for which a case study has been done is the one presently used for the motor portfolio in the Baltica Insurance Company for third party and own damage coverage. The bonus system consists of 11 classes, 0,1,...,10. New policies are normally placed in class 4 but can - depending on their past claims experience - be placed in any of the ten other classes. The transition rules are:

> in case of a claimfree insurance period the policy is transferred one step forward in the bonus system
>  (i→i+1) with the exception of class 4 (as far as new business is concerned) where the policy in this case remains for two years

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- for each claim in the insurance period the policy is transferred two steps backward in the bonus system  $(i \rightarrow i-2, i \rightarrow i-4 \text{ etc.})$ .

Below are shown in Diagram 1 and 2 the classes of the bonus system and the transition rules for the case of claimfree experience and the case of exactly one claim in the insurance period, respectively. For reasons that will become apparent when we look at the actual data we divide the classes into two subsystems: one for new business and one for the portfolio (excluding new business). Furthermore, we introduce an interm diary class 4.1 corresponding to the above transition rule for class 4. Thereby we operate with 23 classes which will be referred to as

with I=23 for the Baltica case.

1. Claimfree in the insurance period

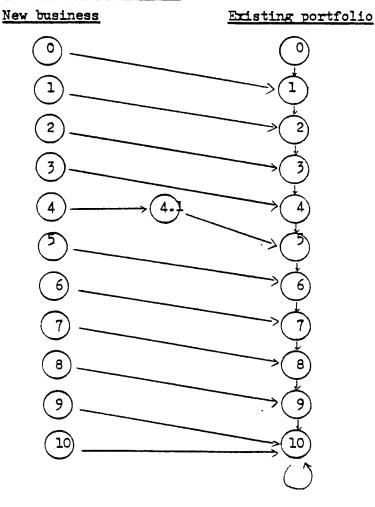
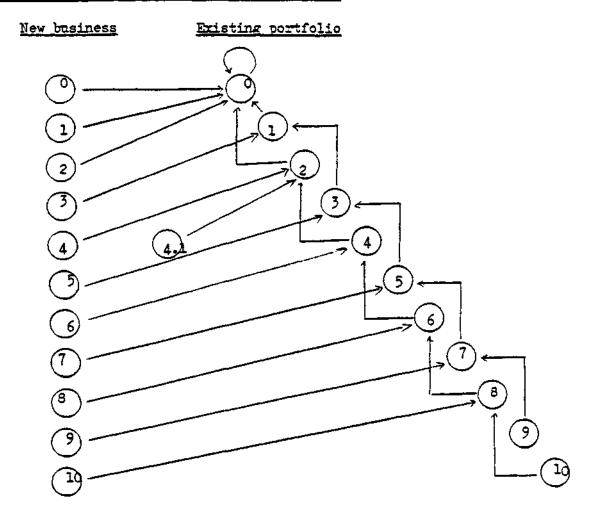


Diagram 1



## 2. Exactly one claim in the insurance period

Diagram 2

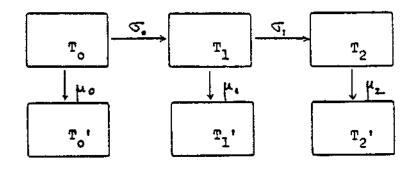
## 2. Definition of states and transitions rules of a bonus system.

The movement of a policy in a bonus system can be described by means of

- a) the events taking place during the insurance period
  - (= period between renewals), and
- b) the transition rules applied at each time of renewal.

a. Events taking place during the insurance period.

At the beginning of an insurance period a given policy is by definition claimfree in the bonus class to which it was transferred at the time of renewal. By the end of the insurance period this policy can be in one of the following six states: The possible direct transitions are given in the following diagram:



The model will be described as a time-homogeneous Markov Chain. The transition intensities are:

- 1. claim-rate for policies with exactly k claims:  $\Im_{k}$  (k=0,1)
- 2. cancellation rate for policies with exactly k claims:  $\mu_{k}$  (k=0,1)

3. cancellation rate for policies with more than one claim:  $\mu_{2}$ . The intensities depend on the bonus class considered and may e.g. vary from calendar year to calendar year.

Let  $\alpha_3 = \alpha_4 + \mu_4$  (k=0,1) and  $\alpha_3 = \mu_3$ . Then the transition probabilities are given by

$$\begin{split} \dot{P}_{oo}(z) &= \exp\left[-x_{o}z\right] \\ \dot{P}_{oo}(z) &= \frac{\mu_{o}}{\kappa_{o}}\left(1 - \exp\left[-x_{o}z\right]\right) \\ \dot{P}_{ok}(z) &= \int_{0}^{z} \dot{P}_{o,k-1}(u) + \exp\left[-x_{o}(z-u)\right] du \quad (k = 1, 2) \\ \dot{P}_{ok}'(z) &= \int_{0}^{z} \dot{P}_{o,k-1}(u) + \det \quad (k' = 1, 2') \end{split}$$

or - if we assume that  $\mu_2 = \mu_1$  -

$$\begin{split} & \oint_{OI}(z) = \frac{\sigma_{o}\sigma_{1}}{\kappa_{1}-\kappa_{o}} \left( \exp\left[-\kappa_{o}z\right] - \exp\left[-\kappa_{1}z\right] \right) \\ & \oint_{OZ}(z) = \frac{\sigma_{o}\sigma_{1}}{\kappa_{1}-\kappa_{o}} \left( \frac{\exp\left[-\kappa_{o}z\right] - \exp\left[-\kappa_{1}z\right]}{\kappa_{1}-\kappa_{o}} - z \exp\left[-\kappa_{1}z\right] \right) \\ & f_{OI}'(z) = \frac{\sigma_{o}\mu_{1}}{\kappa_{o}-\kappa_{1}} \left( \left[ -\frac{1}{\kappa_{1}} \exp\left[-\kappa_{1}z\right] + \frac{1}{\kappa_{o}} \exp\left[-\kappa_{o}z\right] \right) \right) \\ & \oint_{OZ'}(z) = \frac{\sigma_{o}\sigma_{1}\mu_{1}}{\kappa_{1}-\kappa_{o}} \left( \frac{1}{\kappa_{o}(\kappa_{1}-\kappa_{o})} \left( \left[ -\exp\left[-\kappa_{o}z\right] \right] \right) - \left[ -\left[ -\frac{1}{\kappa_{1}} \exp\left[-\kappa_{1}z\right] \right] \right) \right] \\ & - \left( \left[ -\exp\left[-\kappa_{1}z\right] \right] \left( \frac{1}{\kappa_{1}'(\kappa_{1}-\kappa_{o})} + \frac{1}{\kappa_{1}} + \frac{1}{\kappa_{1}'^{2}} \right) \right) \right) \end{split}$$

If one gets slightly different expressions. Let  $A = \{i, 2\}$  and  $B = \{i', 2'\}$ . Then

$$\oint_{OA} (z) = \sigma_0 \frac{\exp[-\mu_1 z] - \exp[-\mu_1 z]}{\alpha_0} - \frac{\mu_1}{\mu_1}$$

$$\oint_{OB} (z) = \frac{\sigma_0 \mu_1}{\alpha_0 - \mu_1} \left( \frac{1 - \exp[-\mu_1 z]}{\mu_1} - \frac{1 - \exp[-\kappa_0 z]}{\kappa_0} \right)$$

#### b. Transition rules

and

We will assume that the bonus system has been so designed that for each bonusclass is given a set of transition rules that only depend on the bonusclass considered and the state  $T_k$  (k=0,1,2), where the policy is situated at the end of the insurance period. Let

$$\chi_0(C_i) = C_{i+1}$$
  
 $\chi_k(C_i) = C_{i-ak}$  for k=1,2 and some a.

c. Definition of underlying stochastic process.

The state space of the process described in the preceding paragraphs is.

 $\{ T_k(C_i), k=0,1,2,0^1,1^2,2^1,i=1,\ldots,I \}$ 

Let X(t) denote the sample path of the process. We will assume that the process is a time-continuous Markov Chain.with transition probabilities as given on p. 5.

#### 3. Estimation of parameters in the model.

a. Derivation of the likelihoodfunction.

I.	Continuous	observation	over	a f	ixed	period	(complete	sampling
	scheme).		7					

Suppose we observe X(t) continuously over say a calendar year period (ref. Diagram 3 below). Consider a given bonus class and let n denote policy No. n. Define

N<sub>2</sub> = transition frequency of k k+l during the period (k=0.1) $M_{n}^{k}$  = transition frequency of k k' during the period (k=0,1,2) $\frac{k}{V_n}$  = the observed time for which policy No. n has occupied state k (k=0.1.2).

Then the likelihoodfunction is

$$L = \frac{2}{11} \exp(-\frac{\sqrt{k}}{2}) \frac{2}{11} \qquad M^{k} \frac{1}{11} \qquad N^{k}$$
  
with  $\sqrt{k} = \sum_{n} \sqrt{k}_{n}$ ,  $M^{k} = \sum_{n} M^{k}_{n}$ , and  $N^{k} = \sum_{n} N^{k}_{n}$ 

Maximum likelihood estimators are the usual occurence/exposure rates  $\delta_{1} = \frac{N^{k}}{2}$  k=0,1

$$\begin{array}{c}
 \lambda \\
 \mu \\
 k = \frac{M^{k}}{\sqrt{k}} \\
 k = 0, 1, 2
\end{array}$$

II. Observartion at fixed points (incomplte sampling schere). Suppose we observe  $\{X(t)\}$  for policy No. n only at each renewal date, i.e. the end of each insurance period. Consider a given bonus class and define

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 $M_{n}^{k} = \begin{cases} 1 & \text{if policy No. n is in state k at the end} \\ & \text{of the insurance period} \\ 0 & \text{otherwise} \end{cases}$   $(k=0,1,2,0^{\circ},1^{\circ},2^{\circ}). \text{ Then the likelihood function is} \\ L = \frac{1}{n} \quad \frac{1}{k} \quad p_{ok}^{k} \quad p_{ok}^{n}$ with n = n (1) assuming the length of the insurance period

with  $p_{ok} = p_{ok}$  (1) assuming the length of the insurance period is 1. The maximum likelihood estimators of the transition probabilities  $p_{ok}$  are given by

 $\hat{p}_{ok} = M^k/M$ with  $M^k = \sum_n M^k_n$  and  $M = \sum_k M^k = N$ , the number of observed policies. From this one gets the m.l.e. of the parameters and by solving the equations

$$exp[-x_{\bullet}] = M^{O}/M$$

$$= (1 - exp[-x_{\bullet}]) = M^{O'}/M$$

etc.etc. Most of the equations will have to be solved by some iteration procedure. By using the approximation  $1-\exp(-x) \approx x$  for small values of x one easily gets suitable initial values for the iteration procedure (e.g. Newton-Raphson) .The case of unequal insurance periods can be handled too. In that case one will have to work with  $p_{ok}(z_n)$  with  $z_n$ being the exposure of policy No. n. The equations involved become a bit more complicated but can easily be solved by a simple algorithm similar to the one indicated above.

The observations in a given bomus class in 1962 using Lexis diagram

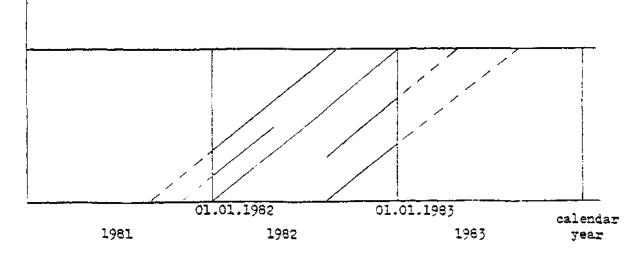


Diagram 3

insurance period

4. Statistical properties of the estimators.

Let  $\hat{\Theta}_N$  denote the vector of maximum likelihood estimators  $(\hat{\sigma}_k, \hat{\mu}_k)$ . Then it is well-known that asymptotically  $\hat{\Theta}_N$  is normally distributed

$$\stackrel{\Theta}{\sim}_{N} \xrightarrow{\sim}_{M} N( \begin{array}{c} \Theta \\ \sim \end{array}, \begin{array}{c} \frac{1}{N} ( \begin{array}{c} \zeta ( \Theta ) \end{array})^{T} )$$

where  $\mathcal{L}(\boldsymbol{\theta}_{n})$  is the Fisher information matrix given by

$$\underbrace{i}_{\mathcal{L}}^{(\mathfrak{g})} = -\left\{ \mathbb{E}_{\mathfrak{g}} \frac{\partial^{2} \mathcal{L}_{\mathcal{L}}}{\partial \theta_{i} \partial \theta_{j}} \right\}_{i,j=0,1,0^{\prime},1^{\prime},2^{\prime}}$$

 $(\theta_{\sigma} = \sigma_{\sigma}, \theta_{\sigma} = \sigma_{\sigma}, \theta_{\sigma}, \sigma_{\mu}, \theta_{\sigma}, \theta_{\sigma}, \sigma_{\mu}, \theta_{\sigma}, \theta_{\sigma$ 

In the complete sampling scheme all covariances vanish while in the incomplte sampling scheme one uses the fact that the  $M^{k}$ 's are Bernoulli s.v. with parameters  $p_{ok}$  in deriving mean values (and in this case the estimators are not asymptotically independent).

If you look at the asymptotic relative efficiency you will in most cases find the the lack of efficiency in estimating the parameters using the incomplete sampling schme is insignificant.

For the complete sampling scheme the following expressions are obtained for the determination of variance terms

$$-E \frac{\partial^2 \underline{f_{u}}}{\partial \theta_{\underline{k}}} = \frac{i}{\theta_{\underline{k}}^2} EN^{\underline{k}} = \frac{i}{\theta_{\underline{k}}^2} \int_{0}^{1} p_{ok}(t) \theta_{\underline{k}} dt \quad \text{for } \underline{k=0,1}$$
  
$$-E \frac{\partial^2 \underline{f_{u}}}{\partial \theta_{\underline{k}}^2} = \frac{i}{\theta_{\underline{k}}^2} EN^{\underline{k}'} = \frac{i}{\theta_{\underline{k}}^2} \int_{0}^{1} p_{ok}(t) \theta_{\underline{k}} dt \quad \text{for } \underline{k=0,1,2}$$

i.e. the information matrix is diagonal with the following non-zero entries

$$\frac{i}{\sigma_{*}^{t}} \int_{0}^{t} p_{00}(t) \sigma_{0} dt$$

$$\frac{i}{\sigma_{*}^{t}} \int_{0}^{t} p_{01}(t) \sigma_{1} dt$$

$$\frac{i}{\mu_{*}^{t}} \int_{0}^{t} p_{00}(t) \mu_{0} dt$$

$$\frac{i}{\mu_{*}^{t}} \int_{0}^{t} p_{01}(t) \mu_{1} dt$$

$$\frac{i}{\mu_{*}^{t}} \int_{0}^{t} p_{02}(t) \mu_{2} dt = \frac{i}{\mu_{*}^{t}} \int_{0}^{t} p_{02}(t) \mu_{1} dt \quad (as \ \mu_{2} = \mu_{1})$$
according to assumptions).

Statistical inference could be based upon the asymptotic distribution of the parameters, e.g. by using linear normal models and treating the estimators as if they were distributed exactly as normal s.v. with known variance-covariance matrix.

It is not the intention to go into detail about these matters but only mention that management - in order to access the efficiency of the bonus system - is interested in testing

if there are significant differences among the k's of the different classes, e.g. hypotheses of the type σ<sub>0</sub>(i) > σ<sub>0</sub>(i) > ... > σ<sub>0</sub>(i) > ... > σ<sub>0</sub>(i) > ... > σ<sub>0</sub>(i)
i referring to bonus class C<sub>1</sub>
if there are significant differences between μ<sub>0</sub>(i) and μ<sub>1</sub>(i) for i=1,...,I.

## 5. The model as an aid in tariffing.

## <u>a. B-values</u>

As the premium in motor insurance in most of the risk categories is linked to the claims frequency experience of the individual policyholder, for tariffing purposes it is necessary to know more than the level and structure of the risk premium in order to access the development of the loss ratio for a future period of say 3-5 years. Focus is put on the loss ratio as this indirect measure of profitability is the one used and understood by management.

In order to access future expected premium income you will have to know the expected claims frequency, the expected level of new business, and the expected cancellation rate during the period of interest. With the bonus system is given a bonus scale  $\mathcal{K}_i$ , i=1,...,I such that the one-period premium is  $\mathcal{K}_i$ \*P for all policies in bonus class  $C_i$ . Here P is a common tariff-premium for the risk group considered. The bonus scale used at present in Baltica is

 $c_1$  0 1 2 3 4 5 6 7 8 9 10  $x_1$  1.113 .9646 .8162 .742 .6678 .5936 .5194 .4452 .371 ,2968 .2226

Let  $N_i$  denote the number of policies in bonus class  $C_i$  and define B as

$$B = \sum_{i}^{\Sigma} P \overline{x}_{i} N_{i} / \sum_{i}^{\Sigma} P N_{i} .$$

The relative measure B is a convenient way of expressing the changes in the level of the premium income during the period of interest. B·P· $\sum N_i$  is the expected premium income.

#### b. Estimation of expected total premium income

Estimation of expected total premium income is done simply by applying the probabilities  $p_{ok}(1)$ 's and  $p_{ok}(1)$ 's to the number of policies in each of the bonus classes at the beginning of the year to obtain the similar number of policies in each of the bonus classes at the end of the year. The process is continued in an obvious manner for each of the years in the period of interest.

You get		
year O	:	portfolio at beginning of year 0
		+ new business in year O
		- cancellations in year 0
		= portfolio at end of year 0
year l	:	= portfolio at beginning of year l
		+ new business in year l
		- cancellations in year l
		= portfolio at end of year l
year 2	:	= portfolio at beginning of year 2
		etc.etc.

Weighing the number of policies (or rather policyyears) in each of the bonus classes with the bonus scale you find the following B-values:

Portfolio at beginning of the year, excluding new business : B<sub>o</sub> do ,including new business : B<sub>o</sub>,n Portfolio at end of the year, excluding cancellations : B<sub>1</sub> do , including cancellations : B<sub>1,c</sub> New business : B<sub>n</sub> Cancellations : B<sub>n</sub> where primarily B<sub>1,c</sub> is the one of importance in estimating the level of premium income.

#### 6. Case study

Below are shown examples based on the data for 1980 (in a complete observational scheme) from a motor portfolio of the Baltica Insurance Company comprising 150.000 policies to illustrate the performance of the model.

We are looking at the biggest single risk category, namely privately owned cars having an insurance coverage including a 300 Dkr deductible. In Appendix 1 the basic data are shown and in Appendix 2 the estimates of the parameters in the model. In order to achieve monotonicity among the parameters of the model the stimates, i.e. claim frequencies and cancellation rates, have been smoothed. The monotonicity requirements - which for most bonus classes are fully justified by the data - consist mainly in the following relations

- the claims frequency in bonus class C<sub>i</sub> decreases with i
- the cancellation rate in bonus class C<sub>i</sub> decreases with i:

The model was implemented on our APL-system. Taking the 31.12.1981 portfolio as starting point we are interested in evaluating the portfolio development during the five-year period 1982-1986 by means of the B-values. We will do that for different sets of assumptions regarding amount of new business and level of claims frequencies and cancellation rates during that period. In one of the cases we will show the portfolio movements for each bonus class. For all other cases we will focus on the B-values alone.

- <u>CASE 1</u> Claims frequencies as in 1980 for all the years 1982-1986 Cancellation rates as in 1980 for all the years 1982-1986 Amount of new business as in 1980 for all the years 1982-1986 - se Appendix 3 for the detailed portfolio movements.
- <u>CASE 2</u> Amount of new business in 1982 is 50% higher than in 1980. Otherwise the same assumptions as in case 1.
- <u>CASE 3</u> Amount of new business in 1982 is 50% lower than in 1980. Otherwise the same assumptions as in case 1.
- CASE 4 Amount of new business in 1982 is 10% higher, in 1983 20% higher, in 1984 30% higher, in 1985 40% higher and in 1986 50% higher than in 1980.

Otherwise the same assumptions as in case 1.

CASE 5 Amount of new business in 1982 is 10% lower, in 1983 20% lower, in 1984 30% lower, in 1985 40% lower and in 1986 50% lower than in 1980.

Otherwise the same assumptions as in case 1.

<u>CASE 6</u> Claims frequencies are for each of the years 1982-1986 10% higher than in 1980.

Otherwise the same assumptions as in case 1.

- CASE 7 As in case 6 but with 25% higher claims frequencies.
- CASE 8 As in case 6 but with 100% higher claims frequencies.
- CASE 9 As in case 6 but with 25% lower claims frequencies.

CASE 10 Cancellation rates are for each of the years 1982-1986 10% higher than in 1980.

Otherwise the same assumptions as in case 1.

- CASE 11 As in case 10 but with 25% higher cancellation rates.
- <u>CASE 12</u> Both claims frequencies and cancellation rates are 25% higher in bonus class 10.

Otherwise the same assumptions as in case 1.

<u>CASE 13</u> As in case 12 but with 25% lower claims frequencies and cancellation rates.

In the tables 1-3 are shown the B-values and the relative differences between the B-values.

It is surprising that the B-values exhibit only rather small deviations even in cases where you alter the assumptions drastically as regards new business, claim frequencies and cancellation rates.

If we e.g. look at case 6 and 7 we note that an overall increase in the claim frequencies and thereby in the risk premiums by 10% and 25% respectively is followed by an overall premium increase of only 1-2% as compared with the "normal" situation (case 1). This shows that the bonus system as a sort of self-adjusting instrument or mechanism is of little or no value at all. The erason is in our view that the level of claim frequency is just simply to small to justify bonus systems as the ones in use by most companies. Statistically the bonus system is justified in the sense that claim frequencies are significantly different from bonus class to bonus class. The qestion is whether this is so in an economic sense of the concept: significance, as seen from a company point of view. Due to competition, however, it is hardly possible to imagine motor insurance without bonus systems.

What is worth noting is in almost all the cases the tendency in the course of time of a decrease in the E-values which is due to the fact that more and more of the policies end up in class 10 and stay there. This general tendency is of course most dangerous if account of it has not been made in the tariff premium calculation.

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#### Conclusion

I have hoped to show that the proposed model is a suitable and practical instrument for gauging different strategies as to new business, cancellations, bonus scales etc.etc. relating to the bonus system of motor insurance. I have stressed the evaluation of the total premium income for a forecasting period of five years through the use of the B-values.

It is possible to work not only on mean values but to consider the model as a "true" stochastic model, e.g. by doing simulation assuming that the parameters as maximum likelihood estimators follow a normal distribution. I am not sure, however, what would be the benefits from such an approach as I am convinced that changes in the level of new business, claim frequencies, and cancellation rates as a consequence of company policies or general changes in society (oil crises etc.) exhibit a far greater effect on the parameters of the model than do the purely stochastic variations in the parameters when dealing with portfolios of the size presented in the present case study.

<u>B-values (1n %)</u>	д°
lable 1	lase F
Ta	Cai

	1986	43.287	43.249	43.718	46.351	39.649	43.685	44.274	47.058	42.273	42.722	42.003	41.896	45.315
	<u>1985</u>	43.436	43.529	43.705	45.898	40.991	43.809	44.359	46.926	42.484	42.818	42.008	41.858	45.654
	1984	43.862	44.187	43.843	45.650	42.426	44.195	44.685	46.937	43.007	43.234	42.384	42.129	46.212
	1983	44.894	45.540	44.445	45.959	44.242	45.147	45.517	47.215	44.244	44.353	43.596	42.983	47-334
° B	1982	46.678	48.018	45.083	47.028	46.435	46.738	46.817	47.075	46.503	46.419	46.049	44.561	49.193
Bn all	years	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427	49.427
	1986	28.103	28.169	28.212	28.986	27.483	28.281	28.556	30.056	27.674	27.807	27.415	28.625	27.575
	<u>1985</u>	28.323	28.470	28.338	28.967	27.978	28.491	28.748	30.124	27.918	28.028	27.630	28.766	27.872
	1984	28.810	29.087	28.668	29.222	28.684	28.958	29.184	30.364	28.449	28.541	28.171	29.165	28.449
	1983	29.754	30.222	29.359	29.948	29.757	29.853	30.004	30.790	29.513	29.568	29.304	29.969	29.535
в о	1982	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311	31.311
Case		Ъ	2	£	4	5	6	7	8	6	10	11	12	13

B-values (in %) Table 2

Case	Bo,n					<sup>B</sup> 1,0	
	1982	1983	1984	1985	<u> 1986</u>	1902	1983
l	32.845	31.506	30.721	30.344	30.203	32.015	30.871
2	33.519	31.882	30.925	30.434	30.223	32.737	31.225
3	32.112	31.205	30.649	30.415	30.361	31.353	30.692
4	32.984	31.985	31.560	31.539	31.741	32.290	31.467
5	32.703	31.195	30.099	29.315	28.701	31.988	30.648
6	32.845	31.600	30.863	30.508	30.381	32.140	31.067
7	32.845	31.742	31.079	30.758	30.654	32.32 <b>7</b>	31.365
8	323845	32.480	32.20 <b>1</b>	32.087	32.124	32.247	32.885
6	32.845	31.277	30.373	29.946	27775	31.702	30.386
10	32.845	31.356	30.517	30.137	30.014	32.004	30.723
11	32.845	31.148	30.245	29.870	29.872	31.992	30.520
12	32.845	31.725	31.094	30.827	30.788	32.104	31.133
13	32.845	31.283	30.341	29.854	29.613	31.916	30.594

	ភា
0	92

<u>1982</u> 32.015	<u>1983</u> 30.871	<u>1984</u> 30.266	<u>1985</u> 30.001	<u>1986</u> 29.852
32.737	31.225	30.448	30.074	29.860
31.353	30.692	30.300	30.166	70 <b>.</b> 05
32.290	31.467	31.198	31.272	31.456
31.988	30.648	29.717	29.050	28.437
32.140	31.067	30.491	30.242	30.105
32.327	31.365	30.834	30.610	30.494
32.247	32.885	32-626	32.573	32.592
31.702	30.386	29.717	29.419	29.242
32.004	30.723	30-074	29-809	29.681
31.992	30.520	29.818	29.563	29.474
32.104	31.133	30.647	30.485	30.423
31.916	30.594	29.869	29.501	29.265

Case	<sup>B</sup> 1,0 :		1982 = 100,0			B1,0	case 1	1 = 100,0	_ 1	
	1982	1983	1984	<u>1985</u>	<u>1986</u>	1982	1983	1984	1985	<u>1986</u>
J	100,0	96.4	94.5	93.7	93.2	100.0	100.0	100.0	100.0	100.0
N	100.0	95.4	93.0	91.9	91.2	102.3	101.1	100.6	100.2	100.0
£	100.0	6-76	9°96	96.2	96.0	6.79	99.4	1001	100.5	100.8
4	100.0	91.5	96.6	96.8	97.4	100.9	<b>9.101</b>	103.1	104.2	105.4
5	100.0	95.8	92.9	90-8	88.9	6.66	6•66	98.2	96.8	95.3
9	100.0	7.96	94.9	94.1	7.56	100.4	100.6	100.7	100.8	100.8
7	100.0	0*16	95.4	94.7	94.3	101.0	101.6	101.9	102.0	102.2
8	100.0	98.9	98.1	98.0	98.0	103.8	106.5	107.8	108.6	109.2
6	100.0	95.8	93.7	992.8	92.2	0*66	98.4	98.2	98.1	98.0
10	100.0	0*96	94.0	93.1	92.7	100.0	<b>66</b> •5	99.4	99.4	99.4
ц	100.0	95.4	93.2	92.4	92.1	6•66	98.9	98.5	98.5	<b>7.</b> 86
12	100.0	0-76	95.5	95.0	94.8	100.3	100.8	101.3	101.6	101.9
13	100.0	95.9	93.6	92.4	7.19	<b>T.</b> 99.7	1.66	<b>7.</b> 86	98.3	98.0

# page 1

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## NUMBER OF POLICYYEARS

1980

Bonus	l Eristin	ng Poztfolio		New	business	
class		after the first claim		before the first claim		total
<u>ہ۔۔۔۔</u>	12	•••••	12	 د		7
1	1443	178	1621	56	8	64
2	2744	216	2960	ខា	13	94
3	2174	1 <b>1</b> 3	2287	f 🕈	í	20
4	2295	115	2410	7210	733	7943
<b>4</b> .i	4709	288	4997	Ø	Θ	e
5	1937	48	1985	123	4	127
6	5875	238	6113	557	35	592
7	8424	299	8723	750	20	770
8	11676	337	12013	798	29	827
9	21429	577	22006	4527	125	4652
10	74051	1248	75299	1144	24	f168
TOTAL	136769	3657	140426	15271	993	16264

NUMBER OF CLAIMS

Bonus	l Ex	isting portfol	lio	l Ne	w business	1
class	<sup> </sup> before the first claim 	after the first claim	total	before the first claim 	after the first clai	total   m
	 5	 0	 5		 ©	**************************************
i	615	126	741	32	10	42
2	822	121	943	4 등	ċ	51
3	418	67	485	4		4
4	370	36	468	2904	508	3412
4.1	1048	127	1175	ø	O	Θ
5	309	14	323	22	2	20
ర	840	71	91i	h 20	25	145
7	1025	88	1113	111	<b>6</b> .	116
8	f117	69	1186	9 T	i ı	108
9	1857	117	1974	461	24	485
10	4471	162	4633	131	2	f33
TOTAL	12897	998	13895	3929	595	4524

## NUMBER OF CANCELLATIONS

Bonus	j E	xisting por	tfolio (	New	business	
class	•	after th m first cl	ne total; .aim ;		after the first claim	
 0	 5	8	 13	i	3	4
4	762	281	1043	32	15	47
2	1198	386	1584	27	22	49
ŧ	705	129	834	<b>f</b> f	Ç	13
4	520	118	638	3276	1874	5150
4.1	1559	289	1848	Θ	Θ	Θ
5	499	88	587	37	8	45
6	1230	213	1443	133	52	185
7	1220	195	1415	126	42	168
8	1211	194	1405	116	33	149
9	1864	253	2117	555	108	663
10	5206	475	5681	101	47	148
TOTAL	15979	2629	18608	4415 1346	2206	6621

## CLAIM FREQUENCIES

# 1980

Bonus	1	Fristin	g portfolio		1	New	business	
class	1	before the first claim	after the first claim	total	1 [	before the first claim	after the first claim	total
0		417	 Θ	417		333	0	286
1		426	708	457		571	1250	656
2		300	560	319		556	462	543
3		192	593	212		211	2000	300
4		161	313	168		403	673	430
4.1		223	<b>4</b> 4 i	235		Ø	Θ	ø
5		160	292	163		179	500	189
6		143	298	149		215	714	245
7		122	294	128		148	250	151
8		96	205	99		122	379	131
9		87	203	90		102	192	f 04
10		60	130	62		115	83	114
TOTAL		94	273	99		257	599	278

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# CANCELLATION RATES

\_\_\_\_\_\*

		ه چک کک کنا خان خان ایند ایندانند چو، دور هن و	**********			~~~~~~~			-
Bonus	I	Eris	ting portfoli	.0	ł	New	business		ł
clas <del>s</del>	ł	before the	after the first claim	total	1	before the	after the first claim	total	1
	1				I				ł
•••••		417	•••••• 0	1083		167	3000	571	-
i		528	1579	643		571	1875	734	
2		437	1787	535		333	1692	521	
3		324	1142	365		579	2000	650	
4		227	1026	265		454	2557	648	
4.1		331	1003	370		ø	Θ	Ũ	
5		258	1833,	296		301	2000	354	
6		209	895	236		239	1486	313	
7		145	652	162		168	2100	218	
8		104	576	117		í 45	1138	180	
9		87	438	96		123	864	143	
f©		70-	381	75		88	1958	127	
TOTAL		117	719	133		289	2222	407	

# CLAIM PROBABILITIES

# 1980

Bonus	Ecc	sting portfolio	1	New	business	
class	first clai	after the m first claim	total   	before the first claim	after the first claim	total
0	435	283	0	607	74	Ø
1	385	91	31	319	79	50
2	479	71	18	411	130	27
3	597	66	19	454	29	27
4	679	71	ίQ	424	65	17
4.1	575	85	18	1000	ø	Θ
5	659	50	ó	619	48	Ŷ
6	703	68	10	635	65	21
7	766	48	10	729	46	4
8	819	59	6	76ó	53	9
9	841	58	6	799	55	5
10	877	44	3	816	43	1
TOTAL	810	53	7	579	59	14

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# CANCELLATION PROBABILITIES

						*****	
Bonus	i	Extis	ting portfolio	I	New	# business	
class	1 !		after the first claim	total   {		after the first claim	total
• 0		283	· ©	283	131	188	3:9
î		340	153	493	341	211	552
2		309	123	432	221	211	432
3		253	65	318	400	90	490
4		188	53	241	305	188	493
4.1		254	68	322	Θ	ø	Ô
5		211	74	285	239	85	324
6		176	43	219	192	88	280
7		127	29	157	144	77	221
8		94	21	115	f 28	45	172
9		80	16	95	110	31	<b>i 4</b> i
10		66	10	76	80	60	140
TOTAL		105	25	130	223	126	349

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# SMOOTHED CLAIM PROBABILITIES

Bonus	l	Ex	isting portfo	olio (	N	ew business	
class	1		after the first claim		before the first claim		after the second claim
		390	84	35	336	52	16
í		391	<del>9</del> 0	33	356	54	18
2		485	80	23	377	58	18
3		598	68	15	399	63	i 8
4		676	68	10	423	68	18
4.1		573	85	t 9	1000	o	0
5		654	69	12	641	47	10
6		701	69	10	677	50	10
7		767	66	8	715	52	5
8		817	58	6	755	55	7
9		841	58	5	798	55	5
10		878	44	3	843	56	3
TOTAL		810	53	6	581	61	13

# AFPEDIX 2

## page

# SMOOTHED CANCELLATION PROBABILITIES

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Bomis	ł	Esci	sting portfol:	io l	ľ	lew business	
class	1 1		after the first claim		before the first claim	after the first claim	tot.al
0		324	177	501	325	271	596
i		328	158	486	321	251	572
2		305	108	412	316	230	546
3		250	70	320	311	208	519
4		193	53	246	306	186	491
4.1		256	67	323	0	ø	0
5		214	51	265	207	94	301
6		177	42	219	184	79	263
7		126	33	159	159	54	224
8		96	23	119	133	50	183
9		80	17	97	106	35	141
10		66	9	75	77	22	99
TOTAL		105	26	131	219	126	345

# PORTFOLIO MOVEMENTS IN 1982

Bonus	Number	or bor	icies a	as at							
class	01.01. 1982	31.12	.1982								
	1962	•	1	2	З	4	5	6	7	8	9 10
o	926	110	352								
1	518	64		203							
2	2468	253			1198						
3	2255	33	152			1349					
4	1916	20		130			1296				
4.1	3411	65		289			1956				
5	6970		83		480			4559			
6	2068			22		143			1450		
7	7528				64		495			5773	
8	12804					79		744			10457
9	11566						57		667		9724
0	100600							280		4453	882 <b>8</b> 4
Θ	86	6	29								
1	13	1		5							
2	218	17			82						
3	17	Θ	i			7					
4	7026	127		476		2972					
5	334		3		16			214			
6	382			4		19			259		
7	503				4		26			360	
8	685					5		37			517
9	579						З		32		461
Θ	4315							11		241	<b>3</b> 636
TAL	167188	695	620	1126	1843	4574	3832	5847	2408	10827	10975102105

# PORTFOLIO MOVEMENTS IN 1983

Bonu				s at							
clas	s 01.01. 1983	<u>31.12</u> 0	2.198 <u>3</u>	2	3	4	5	6	7	8	9 10
0	695	- 83	264	-	-		_				
1	620	76		243							
2	1126	115			547						
3	1843	27	124			1103					
4	1602	16		108			1083				
4.1	2972	57		251			1704				
5	3832		45		264			2507			
ć	5847			61		404			4100		
7	2408				20		158			1847	
8	10827					67		629			8842
9	10975						54		633		9227
o	102105							284		4520	89605
0	86	6	29								
1	13	í		5							
2	218	17			82						
3	17	Θ	1			7					
4	7026	127		476		2972					
5	334		3		16			214			
6	382			4		19			259		
7	503				4		26			360	
8	685					5		37			517
9	579						3		32		462
Θ	4315							11		241	3636
TAL	159011	525	438	1147	933	4577	3029	3684	5024	6967	9360102929

# PORTFOLIO MOVENEUTS IN 1984

\_\_\_

Bonus	Number	of poli	cies at	at								
class	01.01. 1984	31.12	.1984									<u>.</u>
	2/04	O	ſ	2	3	4	5	6	7	8	9	<b>∮</b> €
0	525	- 62	200									
í	468	57		183								
2	1147	118			557							
3	933	14	63			558						
4	1604	16		108			1085					
4.1	2972	57		251			1704					
-	3029		36		209			1981				
6	3684			38		254			2583			
7	5024				42		330			3853		
8	6967					43		405			5690	
9	9360						45		540			<b>78</b> 0
10	102929							287		4556		903:
Θ	86	6	29									
1	13	1		5								
2	218	17			82							
3	17	0	i			7						
4	7026	127		476		2972						
5	334		3		16			214				
6	382			4		19			259			
7	503				4		26			360		
8	685					5		37			517	
9	579						3		32			4
10	4315							11		241		30
OTAL	152800	475	332	1065	910	3859	3195	2936	3414	9010	62071	02

# PORTFOLIO MOVENEITS EL 1985

Bonus		of poli	cies as	; at								
class	01.01. 1985	31.12	.1985									
	2,0)	ø	1	2	3	4	5	6	7	8	9	1 0
0	475	56	180									
i	332	<b>4</b> í		130								
2	1045	109			517							
3	910	13	61			544						
4	885	9		60			599					
4.1	2972	57		251			1704					
5	3195		38		220			2090				
6	2936			31		203			2059			
7	3414				29		224			2618		
8	9010					56		524			7358	
9	6207						30		328			5219
10	102295							285		4528		89771
Θ	85	6	29									
i	13	1		5								
2	218	17			82							
3	17	ø	1			7						
4	7026	127		476		2972						
5	334		3		16			214				
6	382			4		19			259			
7	503				4		26			360		
8	685					5		37			517	
9	579						3		32			462
1 0	4315							1 1		241		3636
OTAL	147855	436	313	956	838	3803	2588	3161	2708	7747	7876	99087

## PORTFOLIO MOVEMENTS DI 1986

Bonus	Number			s at								
class	01.01. 1986	31.12.		••••••		· · · · <u>-</u> · · · <u>- · · ·</u>		<del>.</del>				ter melle atter geter big og s
		0	4	2	3	4	5	6	7	8	9	1 0
0	436	52	166									
1	313	38		123								
2	956	98			464							
3	868	13	59			519						
4	834	9		56			564					
4.1	2972	57		251			1704					
5	2588		31		178			1693				
5	3161			33		218			2217			
7	2708				23		178			2077		
8	7747					48		450			6327	
9	7876						39		454			6622
0	99087							276		4386		86956
o	86	6	29									
٩	13	1		5								
2	218	17			82							
3	17	0	1			7						
4	7026	127		476		2972						
5	334		3		16			214				
6	382			4		19			259			
7	503				4		23			360		
8	685					5		37			517	
9	579						3		32			482
e	4315							11		241		3636
ITAL	143704	417	289	947	767	3789	2514		2962		6844	