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## INTRODUCTION

The Executive Committee of the Continuous Mortality Investigation Bureau of the Institute of Actuaries and the Faculty of Actuaries has pleasure in presenting this, the seventeenth number of its reports.

This number is devoted entirely to the new range of standard tables based on the 1991-94 experience. This range, to be know as the " 92 " Series, is even more extensive than the " 80 " Series introduced by C.M.I.R. 10, the innovations being standard tables for female temporary assurances, amounts tables for immediate annuitants, and tables based for the first time on the experience of retirement annuitants. Only the vested section of this last group warrants standard table status, but tables based on the deferred and on the combined vested and deferred experience are included in this volume for completeness. The official publication date of the new tables is 30 June 1999.

The proposed new standard tables were discussed by the profession at a special seminar held at Staple Inn in December 1998, and were further discussed at the Current Issues in Life Assurance seminar in April 1999. In the light of this, it is not proposed that further meetings be held to discuss this report.

The data upon which the tables were based was included in C.M.I.R. 16, 1-82, and the base table for the life office pensioners' experience was published in C.M.I.R. 16, 113-141.

A Working Party consisting of A D Wilkie, J J McCutcheon and D O Forfar, assisted by P A Leandro of the Bureau Secretariat, prepared the reports. The Committee is grateful to them for all the work undertaken, for which the Committee takes responsibility.

As on the previous occasion when a new suite of standard tables was produced, a computer package is being launched, and will be a Windows version of the Bureau's highly successful Standard Tables Program and will include all the new tables. A companion hardback volume will also be produced, and, as on the last time, will include mortality functions and a sample of monetary functions.

The committee structure of the Bureau was changed earlier this year and this is reflected in the details on the front inside cover. I retire from the Bureau on 1 July 1999, and will be succeeded by Peter Nowell, whom I wish well at this exciting time in the Bureau's history. Once again I would like to thank all those involved in the work of preparing these reports, including the contributing offices, the Secretariat of the Bureau, especially Tony Leandro who has had an exceptionally busy first year as Secretary, Alden Press, and, last but not least,

## CONTENTS

Introduction ..... iii
Standard Tables of Mortality based on the 1991-94 Experiences

1. Permanent Assurances, Males and Females ..... 1
2. Immediate Annuitants, Males and Females, Lives and Amounts ..... 27
3. Self-Employed Retirement Annuitants, Males and Females ..... 44
4. Widows of Life Office Pensioners, Lives and Amounts ..... 59
5. Temporary Assurances, Males and Females ..... 68
6. Projection Factors for Mortality Improvement ..... 89
7. Specimen Monetary Values and Comparisons with Other Tables ..... 109
8. Publication of the New Tables ..... 146
References ..... 148
Appendix A Values of Mortality Rates for the New Standard Tables ..... 149
Appendix B Values of Mortality Rates for Other Tables ..... 218
Appendix C Formulae for the New Standard Tables ..... 225
The Distribution of Policies per Life Assured ..... 229

# STANDARD TABLES OF MORTALITYBASED ON THE 1991-94EXPERIENCES 

## 1. PERMANENTASSURANCES: MALESAND FEMALES

### 1.1 Introduction

1.1.1 The investigations into the mortality of permanent assurances (i.e. whole of life and endowment assurances) are the oldest and largest carried out by the CMI Bureau. They are now carried out for both sexes, and are subdivided by select duration since entry. The data for durations $2,3,4$ and 5 and over are available separately; however, they have been amalgamated to form a section for durations 2 and over (denoted "durations $2+$ "). Data for the United Kingdom and the Republic of Ireland are collected separately. Only the former is considered here. In recent years subdivisions by smoker and nonsmoker status have been included, but these are not considered in this report.

### 1.2 The data

1.2.1 The experience is a large one. The total numbers of exposed to risk and of deaths are shown in Table 1.1, and compared with the corresponding numbers for the 1979-82 experience.
1.2.2 One can see that the male experience, as measured by the number of lives exposed to risk, has declined substantially since 1979-82, though the female experience has increased and at duration 0 is not far short of that of males. Presumably the reduction is the result of larger sales of unit linked or unitised with profits policies in recent years, which are not included in this investigation.
1.2.3 The data covers all adult ages, though there is a considerable reduction in the exposed to risk at around retirement ages, as endowment assurance policies mature. The age range of the data, and the continuous age range within which the central exposed to risk is greater than or equal to 100 , and the continuous age range within which the number of deaths is greater than or equal to 10 are shown in Table 1.2. There are also isolated ages where the number of deaths exceeds nine. These are not noted.
1.2.4 The data at age 10 is clearly faulty, and there is very little exposed to risk between ages 11 and 16, so the graduations have been started at age 17. The select data effectively runs out at about age 80 . The data for durations $2+$ appears to have too few deaths recorded above age 90 , so (as has been the practice for previous graduations) the graduation has been based only on data up to about age 90 .

### 1.3 Variance ratios

1.3.1 The investigations are carried out on the basis of policies, rather than lives. Evidence of the distribution of policies per life is obtained from the cause of death investigation (though this is the last occasion on which this will be possible, because the cause of death investigation is to be discontinued). Investigations showed that it was only necessary to consider the distribution of policies per life for males, durations $2+$. The available data in the cause of death investigation showed 41,478 deaths with 49,305 policies, an average of 1.189 policies per death. This is only a fraction, though a large one, of the total number of deaths, which amount to 68,963 policies, as noted above. For males, durations 0 and 1 , there were only a few lives with duplicate policies. The corresponding information for females is not available.
1.3.2 For each age or group of ages "variance ratios" can be calculated from the distribution of policies per life, as $m_{2} / m_{1}$, where $m_{2}$ and $m_{1}$ are the second and first moments, respectively, of the distribution of policies among lives (or rather deaths). The variance ratio for all ages combined for males durations $2+$ is 1.45 , but the variance ratios are applied for each age separately. They are applied by dividing the exposed to risk and the actual deaths at each age by the ratio applicable to that age. Values of the variance ratios are shown in Table 1.15. In the discussion below certain values are quoted both without and with variance ratios, but the graduation has been carried out with variance ratios. This makes rather little difference to the fitted graduated rates, but a lot of difference to the values of $\chi^{2}$.

### 1.4 Comparison with AM80 and AF80

1.4.1 The experience for 1991-94 has been compared with the AM80 and AF80 tables, which were based on the data for 1979-82. Table 1.3 shows the overall values of $100 \mathrm{~A} / \mathrm{E}$, i.e. actual deaths as a percentage of expected deaths, for durations 0,1 and $2+$ separately, in each case the experience being compared with the corresponding rates from the AM80 and AF80 tables. For this calculation, initial exposures have been used, and the expected deaths have been calculated using values of $q_{x}$.
1.4.2 From this table one can see that the experience for females for 1991-94, at all durations, is much lighter than that in 1979-82. For males the experience at durations $2+$ is also much lighter, though the experience for select durations has hardly improved. There is therefore a good reason to construct new tables for both sexes appropriate to the years 1991-94.

### 1.5 Comparison with pensioners

1.5.1 It is of interest to compare the experiences for durations $2+$ with the
graduated rates for pensioners, produced on the basis of the 1991-94 experience, and reported on in C.M.I.R. 16. The overall comparisons are shown below; the full age range is used; the male experience is compared with the graduated rates for pensioners, males lives, and the female experience is compared with the corresponding experience for females (in this case using central exposures and $\mu_{x}$ ).

|  | $100 \mathrm{~A} / \mathrm{E}$ |
| :--- | :---: |
| Males | 84.6 |
|  | (83.6 with variance ratios) |
| Females | 87.1 |

1.5.2 For both sexes the value of $100 \mathrm{~A} / \mathrm{E}$ is well below 100 . The ratio is far from uniform over the ages, being well above 100 at younger ages, and well below at higher ages. It should be remembered that the great bulk of the pensioners data is for ages above 60 , and the extension to lower ages has been somewhat speculative. In any case it is clear that a simple adjustment to the pensioners tables would not be appropriate for the permanent assurances, a not unexpected conclusion.

### 1.6 Methodology

1.6.1 As for the 1979-82 graduations, the methodology was to use central exposed to risk, fit a formula of the $\mu_{x}=\mathrm{GM}(r, s)$ class, and in the first place choose the parameters by maximum likelihood, taking account also of the usual diagnostic tests (numbers of positive and negative deviations, runs, Kol-mogorov-Smirnov, serial correlations and $\chi^{2}$ ). However, as for the graduation of the life office pensioners experience, which has been considered in C.M.I.R. 16, it was decided to adjust the parameters of the graduation formula so that the resulting mortality rates 'behaved properly', that is, the rates for the select durations rose with increasing duration, the rates for the highest and lowest ages were reasonable, and the rates for the two sexes were not unreasonable in relation to each other.

### 1.7 The experience for males, durations $2+$

1.7.1 We consider first the experience for males, and start with durations $2+$. The data for ages 17 to 91 inclusive is used, and variance ratios are applied. When the corresponding data for $1979-82$, for both sexes, were graduated it was found that the formula $\mu_{x}=\operatorname{GM}(2,2)$ was the most satisfactory. This formula was also the best for the females experience for durations $2+$. It was therefore reasonable to try this formula first on this occasion. However, it was also appropriate to consider alternative orders of formula, i.e. different
values of $r$ and $s$ in the $\operatorname{GM}(r, s)$ formula. The following pairs of values were used: $(0,2),(0,3),(1,2),(0,4),(1,3),(2,2),(0,5),(1,4),(2,3),(3,2)$, i.e. each combination for which $r+s \leq 5$ and $s \geq 2$. (There is no point in trying formulae with $s=0$ or 1 for a mortality graduation, because the underlying shape is always of Gompertz, $\operatorname{GM}(0,2)$, type, i.e. to a first approximation $\log \mu_{x}$ is linear in $x$.)
1.7.2 Note that, for example, the $\operatorname{GM}(2,3)$ formula is parameterised as:

$$
a_{1}+a_{2} t+\exp \left\{b_{1}+b_{2} t+b_{3}\left(2 t^{2}-1\right)\right\}
$$

where $t=(x-70) / 50$. The first two terms of this formula can be described as the " $r$ " part, with two terms, and the exponential can be described as the " $s$ " part, with three terms inside the parentheses.
1.7.3 It was quickly clear that the lower order formulae, with $r+s \leq 3$, provided an unsatisfactory fit to the data; that all the formulae with $r+s=4$ and also $(1,4)$ and $(3,2)$ gave poorish fits, and that $(0,4)$ gave a quite unsatisfactory shape. This left $\mathrm{GM}(0,5)$ and $\mathrm{GM}(2,3)$ as possible contenders, both with high values of the log likelihood (low values of the negative of the log likelihood) and the lowest values of $\chi^{2}$. Although both had reasonable shapes, the values of $q_{x}$ for the $\operatorname{GM}(0,5)$ formula rose above those for English Life Tables No. 15, Males at low and at high ages, whereas those for $\mathrm{GM}(2,3)$ did not. The latter was therefore preferred, without any further adjustments. Critical values for several formulae are shown in Table 1.4.
1.7.4 Values of $q_{x}$ for ages from 10 to 120 are calculated by integrating $\mu_{x}$ using repeated Simpson's rule with successively shorter steps until the result converges. Note that the values of all the parameters are rounded to six decimal places before calculating $\mu_{x}$, as quoted in the tables in this report.
1.7.5 The same graduation statistics as were shown for the 1979-82 graduations are shown in Table 1.6 (for all three durations, 0,1 and $2+$; the first two are described below). Specimen values of $q_{x}$, at decennial ages from 20 to 110, and percentage standard errors, are shown in Table 1.7. One can observe that for durations $2+$ the $T$-ratio for the first serial correlation coefficient is just over 2.0, at 2.07 (note that these values should be compared with a unit normal distribution). Further, the value of $\chi^{2}$ (at 105.8) is still outside what might be expected for 68 degrees of freedom.
1.7.6 Details of the exposed to risk, actual deaths, variance ratios, and adjusted exposure and deaths are shown in Table 1.15. Details of the graduation, with exposed to risk, actual deaths, expected deaths, deviations, and standardised deviations $\left(z_{x}\right)$ are shown in Table 1.16.

### 1.8 The experience for males, durations 0 and $I$

1.8.1 We consider next the experience for males for durations 0 and 1. First we compare the experience for the whole age range with the graduated rates for durations $2+$, with the following results:

|  | $100 \mathrm{~A} / \mathrm{E}$ |
| :--- | :---: |
| Duration 0 | 76.5 |
| Duration 1 | 93.0 |

1.8.2 The overall values of $100 \mathrm{~A} / \mathrm{E}$ are, as expected, below 100 , and that for duration 0 is below that for duration 1. However, the pattern across the ages is not uniform, the values being somewhat lower at higher ages than at lower ages, again as might be expected.
1.8.3 We start again by fitting formulae of the $\mathrm{GM}(r, s)$ series to the data, omitting the data for the lowest ages and using an age range of 17 to 89 for duration 0 and 17 to 100 for duration 1 . Note that variance ratios were not used for these durations. The results for various $\operatorname{GM}(r, s)$ formulae are shown in Table 1.5.
1.8.4 For neither duration do the low order formulae fit the data satisfactorily, and generally the higher order formulae produce a poor 'shape' of graduated rates, with the exception of the $\mathrm{GM}(2,2)$ formula. However, for both durations these have unsatisfactory features in relation to the rates for durations $2+$, especially at the younger ages. An approach similar to that used for the pensioners was therefore adopted, to set some of the parameters to produce approximately the desired shape of curve, and then to optimise the rest, i.e. provide a conditional maximum likelihood fit, conditional on the values of the selected parameters.
1.8.5 For duration 1 many combinations of parameters were tried. Eventually it was decided to use for both durations a $\operatorname{GM}(2,3)$ formula, as for durations $2+$, with in each case the values of three of the parameters fixed ( $a_{1}, a_{2}$ and $b_{3}$ for duration $0 ; a_{2}, b_{2}$ and $b_{3}$ for duration 1), allowing the other two ( $b_{1}$ and $b_{2}$ for duration $0 ; a_{1}$ and $b_{1}$ for duration 1) to be fitted by optimisation. The fixed parameter values are as shown below.

|  | Duration 0 | Duration 1 |
| :--- | :---: | :---: |
| $100 a_{1}$ | 0.02 |  |
| $100 a_{2}$ | -0.02 | -0.03 |
| $b_{2}$ |  | 5.0 |
| $b_{3}$ | -0.6 | -1.2 |

1.8.6 The results appeared to be reasonably satisfactory, and the summary statistics for these constrained $\operatorname{GM}(2,3)$ graduations are shown in Table 1.5.

However, the values of $q_{x}$ for duration 1 approach very close to those for durations $2+$ in the 50 s of age; this is a feature of the data, and has not been eliminated.
1.8.7 The same graduation statistics as for durations $2+$ are shown in Table 1.6 and values of $q_{x}$ at 10 -year intervals are shown in Table 1.7. It should be noted that by fixing three of the parameters, the flexibility lies in the level and shape of the rates as affected by the free parameters, so the maximum percentage standard error in the value of $q_{x}$ is not necessarily at the ends of the range. For duration 0 , for example, the parameters have been chosen so that the values of $q_{x}$ at the youngest ages are almost fixed.
1.8.8 Although the values of $q_{x}$ for higher ages are shown in Table 1.7 (and also in Table 1.11 for females), the select rates for individual ages have been limited to age 90 for duration 0 and to age 91 for duration 1 , as for the AM80 and AF80 tables.
1.8.9 A substantial part of the excess $\chi^{2}$ in the data for duration 1 comes from three adjacent ages, 46,47 and 48 , where the actual deaths ( 135 for the three ages in total) are substantially higher than those expected ( 86.3 for the proposed graduation), whereas the experience for neighbouring ages is more normal. For duration 0 the excess $\chi^{2}$ comes substantially from the youngest ages included, from 17 to 22 inclusive, where the actual deaths ( 90 in total) are also much higher than those expected ( 45.2 for the proposed graduation).

### 1.9 The experience for females, durations $2+$

1.9.1 The experience for females, durations $2+$, is now considered. The values of the $\log$ likelihood and of $\chi^{2}$ for $\mathrm{GM}(r, s)$ formulae $(1,2),(0,4),(1,3)$ and $(2,2)$ are shown in Table 1.8. Although the GM( 1,2 ) graduation (a Makeham formula) has rather poorer values of the statistics, its shape is rather better than that of any of the others, and it provides quite a satisfactory fit to the data, as seen from the statistics shown in Table 1.10. No adjustments needed to be made.
1.9.2 Details of the graduation, with exposed to risk, actual deaths, expected deaths, deviations, and standardised deviations ( $z_{x}$ ) are shown in Table 1.17.

### 1.10 The experience for females, durations 0 and 1

1.10.1 The experience for females at durations 0 and 1 is now considered. One can see from the statistics in Table 1.9 that the differences in the log likelihood as between the GM formulae shown are not very large. However, for duration 0 only the GM(2,2) formula produces a reasonable shape, and for duration 1 only $\mathrm{GM}(1,2)$ and $\mathrm{GM}(2,2)$. Because a $\mathrm{GM}(1,2)$ formula had been found satisfactory for durations $2+$, this form was chosen as the basis for modifications.

Both durations were refitted using a GM(1,2) formula, with a fixed value of $100 a_{1}$ of 0.008 for each in order that the values of $q_{x}$ at the youngest ages were conformable with those for durations $2+$.
1.10.2 The usual graduation statistics and specimen values of $q_{x}$ for all three durations are shown in Tables 1.10 and 1.11.

### 1.11 Comparisons

1.11.1 Several sets of comparative ratios are shown in Tables 1.12, 1.13 and 1.14, and graphs of these ratios are shown in Figures 1.2 to 1.6. Table 1.12 shows ratios of the values of $q_{x}$ for duration 0 to those for duration 1 and the corresponding ratios for durations 1 and $2+$, for both sexes. Figures 1.2 and 1.3 show the same ratios. It can be seen that none of these ratios exceeds 1.0 . However, the rates for males duration 1 are much closer to those for durations $2+$ at all except high ages than are those for females.
1.11.2 Table 1.13 and Figure 1.4 show a comparison between the rates for males and those for females. The female rates are all about half the male rates except at the very highest ages.
1.11.3 Table 1.14 and Figures 1.5 and 1.6 show a comparison of the graduated rates and those from the AM80 and AF80 tables. Many of the rates are much lower than in 1979-82. Others are higher. The shapes of the mortality experiences have changed substantially.

### 1.12 The proposed tables

1.12.1 Values of $q_{x}$ for the proposed tables for permanent assurances are shown in Appendix A in Table A1 for males and in Table A2 for females. All six sets of rates are graphed in Figure 1.1.

Table 1.1. Permanent assurances, males and females: comparison of (central) exposed to risk and deaths for

1991-94 and for 1979-82, durations 0,1 and 2+.

|  | 1979-82 | 1991~94 |
| :---: | :---: | :---: |
| Males |  |  |
| Duration 0 |  |  |
| Central exposed | 1,799,039.7 | 837,360.3 |
| Deaths | 1,795 | 1,345 |
| Duration 1 |  |  |
| Central exposed | 1,776,058.3 | 835,252.4 |
| Deaths | 2,287 | 1,774 |
| Durations 2+ |  |  |
| Central exposed | 22,239,148.0 | 15,139,004.8 |
| Deaths | 90,941 | 68,963 |
| Females |  |  |
| Duration 0 |  |  |
| Central exposed | 719,974.5 | 765,376.5 |
| Deaths | 421 | 613 |
| Duration 1 |  |  |
| Central exposed | 664,893.5 | 733.713.3 |
| Deaths | 601 | 802 |
| Durations $2+$ |  |  |
| Central exposed | 3,375,844.5 | 4,931,581.6 |
| Deaths | 6,368 | 12,214 |

Table 1.2. Permanent assurances, males and females: age ranges.

|  | Range of data | Exposed $\geq 100$ | Deaths $\geq 10$ |
| :--- | :---: | :---: | :---: |
| Mates |  |  |  |
| $\quad$ Duration 0 | $10-89$ | $12-80$ | $38-79$ |
| Duration 1 | $10-100^{*}$ | $13-81$ | 3581 |
| Durations 2+ | $10-108$ | $10-101$ | $19-100$ |
|  |  |  |  |
| Females |  |  |  |
| Duration 0 | $10-91$ | $13-81$ | $60-72$ |
| Duration 1 | 1092 | $13-82$ | $43-70$ |
| $\quad$ Durations 2+ | $10-108$ | $10-94$ | $25-95$ |

[^0]Table 1.3. Values of $100 \mathrm{~A} / \mathrm{E}$ when the experience for 1991-94 is compared with the AM80 and AF80 tables
(using initial exposures and $q_{x}$ ).

|  | Duration 0 | Duration 1 | Durations 2+ |
| :--- | :---: | :---: | :---: |
| Males | 98.7 | 96.2 | 72.2 |
| Females | 87.1 | 68.8 | (72.6 with variance ratios) <br> 79.9 |

Table 1.4. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of males durations $2+$ (age range used: 17-91; with variance ratios).

|  |  | GM(0,4) | GM(1,3) | GM(2,2) | GM(0,5) | GM(1.4) | GM(2,3) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | GM(3,2) |  |  |  |  |  |  |
| -Log likelihood | $259,082.9$ | $259,067.4$ | $259,082.3$ | $259,061.5$ | $259,066.4$ | $259,064.6$ | $259,072.1$ |
| $\chi^{2}$ | 144.7 | 114.1 | 139.0 | 100.4 | 110.9 | 105.8 | 120.8 |

Table 1.5. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of males duration 0 (ages 17 to 89 ) and duration (ages 17 to 100 ).

|  | GM(0,4) | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ | $\mathrm{GM}(2,3)$ <br> (unconstrained) | $\mathrm{GM}(2,3)$ <br> (as fitted) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Duration 0: <br> -Log likelihood | $8,984.4$ | $8,992.8$ | $8,985.1$ | $8,983.6$ | $9,002.3$ |
| $\chi^{2}$ | 48.7 | 67.7 | 52.5 | 49.1 | 100.9 |
| Duration 1: <br> -Log likelihood <br> $\chi^{2}$ | $11,164.5$ | $11,165.5$ | $11,162.7$ | $11,162.7$ | $11,168.0$ |

10 Standard Tables of Mortality based on the 1991-94 Experiences
Table 1.6. Permanent assurances, males, durations 0,1 and $2+$ : statistics for graduations of $\mu_{x}=\operatorname{GM}(r, s)$.

| Duration: <br> Formula | $\begin{gathered} 0 \\ \mathrm{GM}(2,3) \end{gathered}$ | $\begin{gathered} \quad \mathrm{I} \\ \mathrm{GM}(2,3) \end{gathered}$ | $\begin{gathered} 2+ \\ \mathrm{GM}(2,3) \end{gathered}$ |
| :---: | :---: | :---: | :---: |
| Ages used | 17-89 | 17-100 | 17-91 |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | 0.02 | 0.022960 | 0.005887 |
| $T$-ratio |  | 7.0 | 0.3 |
| $100 a_{2}$ | -0.02 | -0.03 | -0.049883 |
| $T$-ratio |  |  | -2.2 |
| $b_{1}$ | -4.755647 | -5.056244 | -4.363378 |
| $T$-ratio | -109.7 | -173.1 | -43.0 |
| $b_{2}$ | 5.236521 | 5.0 | 5.544956 |
| T-ratio | 32.5 |  | 92.9 |
| $b_{3}$ | -0.6 | -1.2 | -0.620345 |
| $T$-ratio |  |  | $-6.3$ |
| Sign test: $p$ (pos) | 0.7338 | 0.3101 | 0.4076 |
| Runs test: $p$ (runs) | 0.1962 | 0.4199 | 0.0800 |
| K-S test: $p(K S)$ | 0.1320 | 0.6071 | 0.6833 |
| Serial corrclation test: |  |  |  |
| $T$-ratio 1 | 2.85 | 1.23 | 2.07 |
| $T$-ratio 2 | 1.80 | 0.39 | -1.21 |
| T-ratio 3 | 1.60 | -0.94 | 0.35 |
| $\chi^{2}$ test: |  |  |  |
| $\chi^{2}$ | 100.87 | 89.24 | 105.79 |
| Degrees of | 59 | 60 | 68 |
| freedom |  |  |  |
| $p\left(\chi^{2}\right)$ | 0.000563 | 0.0085 | 0.0023 |

Table 1.7. Permanent assurances, males, durations 0,1 and $2+:$ specimen values of $q_{x}$ and percentage standard errors.

| Duration: | 0 | 1 | $2+$ <br> Formula |
| :--- | ---: | ---: | ---: |
| $\mathrm{GM}(2,3)$ | $\mathrm{GM}(2,3)$ | $\mathrm{GM}(2,3)$ |  |
| Age 20 | 0.000425 | 0.000541 | 0.000582 |
| Percentage s.e. | 0.27 | 6.06 | 6.71 |
| Age 30 | 0.000476 | 0.000558 | 0.000590 |
| Percentage s.e. | 1.04 | 5.88 | 2.77 |
| Age 40 | 0.000788 | 0.000887 | 0.000937 |
| Percentage s.e. | 2.50 | 3.69 | 1.80 |
| Age 50 | 0.001971 | 0.002434 | 0.002508 |
| Percentage s.e. | 3.59 | 1.34 | 0.85 |
| Age 60 | 0.005774 | 0.007760 | 0.008022 |
| Percentage s.e. | 4.00 | 0.42 | 0.60 |
| Age 70 | 0.016582 | 0.022210 | 0.024783 |
| Percentage s.e. | 4.10 | 0.14 | 0.75 |
| Age 80 | 0.043833 | 0.053078 | 0.069303 |
| Percentage s.e. | 4.08 | 0.06 | 0.76 |
| Age 90 | 0.103990 | 0.104031 | 0.170247 |
| Percentage s.e. | 3.96 | 0.03 | 1.88 |
| Age 100 | 0.217127 | 0.167115 | 0.355505 |
| Percentage s.c. | 3.70 | 0.02 | 4.12 |
| Age 110 | 0.391042 | 0.222233 | 0.607918 |
| Percentage s.e. | 3.23 | 0.01 | 5.93 |

Table 1.8. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of females durations $2+$ (ages 17 to 89).

|  | GM(1,2) | GM (0,4) | GM(1,3) | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: |
| Log likelihood | $74,402.5$ | $74,394.8$ | $74,394.9$ | $74,397.4$ |
| $\chi^{2}$ | 91.0 | 74.9 | 74.9 | 79.8 |

12 Standard Tables of Mortality based on the 1991-94 Experiences
Table 1.9. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of females durations 0 and 1.

|  | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0,4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ |
| :--- | :---: | :---: | :---: | :---: |
| Duration 0: <br> -Log likelihood | $4,516.7$ | $4,509.2$ | $4,512.5$ | $4,509.3$ |
| $\chi^{2}$ <br> Duration 1: <br> -Log likelihood <br> $\chi^{2}$ | 55.6 | 52.8 | 61.4 | 47.7 |

Table 1.10. Permanent assurances, females, durations 0,1 and $2+$ : statistics for graduations of $\mu_{x}=\operatorname{GM}(r, s)$.

| Duration: <br> Formula | $\begin{gathered} 0 \\ \operatorname{GM}(1,2) \end{gathered}$ | $\begin{gathered} 1 \\ \operatorname{GM}(1,2) \end{gathered}$ | $\begin{gathered} 2+ \\ \mathrm{GM}_{( }(1,2) \end{gathered}$ |
| :---: | :---: | :---: | :---: |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | 0.008 | 0.008 | 0.011189 |
| $T$-ratio |  |  | 5.7 |
| $b_{1}$ | -4.978100 | -4.763844 | -4.331121 |
| $T$-ratio | -81.5 | -91.7 | -384.6 |
| $b_{2}$ | 5.078780 | 4.854450 | 5.135803 |
| $T$-ratio | 28.1 | 32.5 | 99.3 |
| Sign test: $p$ (pos) | 0.2950 | 0.2175 | 0.5000 |
| Runs test: $p$ (runs) | 0.5251 | 0.3365 | 0.6861 |
| K-S test: $p(K S)$ | 0.7477 | 0.8805 | 0.4476 |
| Serial correlation test: |  |  |  |
| $T$-ratio 1 | 1.11 | 1.74 | 0.55 |
| $T$-ratio 2 | 0.42 | -0.74 | 1.62 |
| $T$-ratio 3 | 0.67 | -1.15 | 0.55 |
| $\chi_{2}^{2} \text { test: }$ |  |  |  |
| Degrces of | 51 | 55 | 66 |
| freedom |  |  |  |
| $p\left(\chi^{2}\right)$ | 0.0053 | 0.20 | 0.0226 |

Table 1.11. Permanent assurances, females, durations 0,1 and $2+$ : specimen values of $q_{x}$ and percentage standard errors.

| Duration: <br> Formula | 0 | 1 <br> $\mathrm{GM}(1,2)$ | $2+$ <br> $\mathrm{GM}(1,2)$ |
| :--- | ---: | ---: | ---: |
| Age 20 | 0.000125 | 0.000150 | 0.000193 |
| Percentage s.e. | 2.23 | 2.45 | 8.70 |
| Age 30 | 0.000205 | 0.000264 | 0.000339 |
| Percentage s.e. | 3.77 | 3.66 | 4.07 |
| Age 40 | 0.000424 | 0.000567 | 0.000747 |
| Percentage s.e. | 5.02 | 4.51 | 1.70 |
| Age 50 | 0.001030 | 0.001364 | 0.001886 |
| Percentage s.e. | 5.70 | 4.94 | 1.33 |
| Age 60 | 0.002701 | 0.003468 | 0.005058 |
| Percentage s.e. | 6.00 | 5.13 | 1.06 |
| Age 70 | 0.007302 | 0.009000 | 0.013867 |
| Percentage s.e. | 6.10 | 5.18 | 1.10 |
| Age 80 | 0.019897 | 0.023459 | 0.038059 |
| Percentage s.e. | 6.10 | 5.17 | 1.72 |
| Age 90 | 0.053854 | 0.060631 | 0.102533 |
| Percentage s.e. | 6.01 | 5.08 | 2.52 |
| Age 100 | 0.141638 | 0.152114 | 0.260630 |
| Percentage s.e. | 5.72 | 4.83 | 3.13 |
| Age 110 | 0.344012 | 0.353086 | 0.569670 |
| Percentage s.e. | 4.95 | 4.18 | 2.98 |

Table 1.12. Permanent assurances: ratios of values of $q_{x}$ in proposed tables: comparison of durations.

|  | Males |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Age | Duration 0/ <br> Duration 1 | Duration 1/ <br> Durations 2+ | Furation 0/ <br> Duration 1 | Duration 1/ <br> Durations 2+ |
| 20 | 0.7856 | 0.9296 | 0.8333 | 0.7772 |
| 25 | 0.8161 | 0.9417 | 0.8031 | 0.7782 |
| 30 | 0.8530 | 0.9458 | 0.7765 | 0.7788 |
| 35 | 0.8863 | 0.9448 | 0.7553 | 0.7724 |
| 40 | 0.8884 | 0.9466 | 0.7478 | 0.7590 |
| 45 | 0.8566 | 0.9570 | 0.7486 | 0.7419 |
| 50 | 0.8098 | 0.9705 | 0.7551 | 0.7232 |
| 55 | 0.7697 | 0.9673 | 0.7655 | 0.7044 |
| 60 | 0.7441 | 0.9405 | 0.7788 | 0.6856 |
| 65 | 0.7363 | 0.8962 | 0.7944 | 0.6669 |
| 70 | 0.7466 | 0.8366 | 0.8113 | 0.6490 |
| 75 | 0.7759 | 0.7659 | 0.8293 | 0.6321 |
| 80 | 0.8258 | 0.6890 | 0.8482 | 0.6164 |
| 85 | 0.9992 |  | 0.8678 | 0.6025 |
| 90 |  |  | 0.8882 | 0.5913 |

Table 1.13. Permanent assurances: ratios of values of $q_{x}$ in proposed tables: comparison of sexes.

| Age | Duration 0 <br> Females/Males | Duration 1 <br> Females/Males | Durations 2+ <br> Females/Males |
| :---: | :---: | :---: | :---: |
| 20 | 0.2941 | 0.2773 | 0.3316 |
| 25 | 0.3563 | 0.3621 | 0.4382 |
| 30 | 0.4307 | 0.4731 | 0.5746 |
| 35 | 0.4974 | 0.5837 | 0.7141 |
| 40 | 0.5381 | 0.6392 | 0.7972 |
| 45 | 0.5429 | 0.6213 | 0.8014 |
| 50 | 0.5226 | 0.5604 | 0.7520 |
| 55 | 0.4937 | 0.4964 | 0.6881 |
| 60 | 0.4678 | 0.4469 | 0.6305 |
| 65 | 0.4493 | 0.4165 | 0.5873 |
| 70 | 0.4404 | 0.4052 | 0.5595 |
| 75 | 0.4416 | 0.4132 | 0.5469 |
| 80 | 0.4539 | 0.4420 | 0.5492 |
| 85 | 0.4786 | 0.4960 | 0.5671 |
| 90 | 0.5179 | 0.5828 | 0.6023 |
| 95 |  |  | 0.6570 |
| 100 |  |  | 0.7331 |
| 105 |  |  | 0.8296 |
| 110 |  |  | 0.9371 |
| 115 |  |  | 1.0322 |

16 Standard Tables of Mortality based on the 1991-94 Experiences
Table 1.14. Ratios of values of $q_{x}$ in proposed tables to those in corresponding AM80 and AF80 tables.

|  | Males |  |  |  |  | Females |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Duration 0 | Duration 1 | Durations 2+ |  | Duration 0 | Duration 1 | Durations 2+ |  |  |
| 20 | 0.6833 | 0.7483 | 0.7358 |  | 0.5631 | 0.5660 | 0.6725 |  |  |
| 25 | 0.9315 | 0.9621 | 0.9264 |  | 0.7908 | 0.7338 | 0.8702 |  |  |
| 30 | 1.1226 | 1.1318 | 1.0650 |  | 0.9234 | 0.8462 | 0.9391 |  |  |
| 35 | 1.0805 | 1.0416 | 0.9957 |  | 0.9054 | 0.8617 | 0.9028 |  |  |
| 40 | 0.9216 | 0.8512 | 0.8212 |  | 0.8314 | 0.8112 | 0.8616 |  |  |
| 45 | 0.8187 | 0.7595 | 0.7030 |  | 0.7780 | 0.7522 | 0.8297 |  |  |
| 50 | 0.7960 | 0.7678 | 0.6577 |  | 0.7618 | 0.7038 | 0.8060 |  |  |
| 55 | 0.8330 | 0.8394 | 0.6582 |  | 0.7810 | 0.6714 | 0.7915 |  |  |
| 60 | 0.9117 | 0.9495 | 0.6834 |  | 0.8313 | 0.6524 | 0.7836 |  |  |
| 65 | 1.0212 | 1.0784 | 0.7195 |  | 0.9112 | 0.6433 | 0.7808 |  |  |
| 70 | 1.1530 | 1.2085 | 0.7575 |  | 1.0205 | 0.6419 | 0.7816 |  |  |
| 75 | 1.2994 | 1.3217 | 0.7908 |  | 1.1493 | 0.6462 | 0.7852 |  |  |
| 80 | 1.4512 | 1.4011 | 0.8148 |  | 1.1869 | 0.6549 | 0.7909 |  |  |
| 85 | 1.5980 | 1.4335 | 0.8358 |  | 1.2303 | 0.6673 | 0.7987 |  |  |
| 90 | 1.7277 | 1.4123 | 0.8581 |  | 1.2773 | 0.6831 | 0.8088 |  |  |
| 95 |  |  | 0.8802 |  |  | 0.8217 |  |  |  |
| 100 |  |  | 0.9012 |  |  |  | 0.8387 |  |  |
| 105 |  |  | 0.9204 |  |  |  | 0.8611 |  |  |
| 110 |  |  | 0.9375 |  |  | 0.8904 |  |  |  |
| 115 |  |  | 0.9525 |  |  | 0.9264 |  |  |  |

Table 1.15. Permanent assurances, males, durations $2+$ : exposed to risk $\left(\mathrm{R}_{x}\right)$, actual deaths $\left(\mathrm{A}_{x}\right)$, variance ratios, adjusted $\mathrm{R}_{x}$ and $\mathrm{A}_{x}$ and crude $\mu_{x}$.

| Agc $x$ | $\mathrm{R}_{x}$ | $\mathrm{A}_{x}$ | Variance Ratio | $\begin{gathered} \text { Adjusted } \\ \mathbf{R}_{x} \end{gathered}$ | Adjusted $\mathrm{A}_{x}$ | Crude $\mu_{x}=\mathbf{A}_{x} / \mathbf{R}_{x}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 10 | 1,142.9 | 53 | 1.00 | 1,142.9 | 53.00 | 0.046373 |
| 11 | 431.3 | 1 | 1.00 | 431.3 | 1.00 | 0.002319 |
| 12 | 419.8 | 0 | 1.00 | 419.8 | 0.00 | 0.000000 |
| 13 | 455.5 | 0 | 1.00 | 455.5 | 0.00 | 0.000000 |
| 14 | 546.9 | 0 | 1.00 | 546.9 | 0.00 | 0.000000 |
| 15 | 785.8 | 0 | 1.00 | 785.8 | 0.00 | 0.000000 |
| 16 | 1,103.1 | 0 | 1.00 | 1,103.1 | 0.00 | 0.000000 |
| 17 | 1,669.1 | 0 | 1.00 | 1,669.1 | 0.00 | 0.000000 |
| 18 | 4,035.1 | 3 | 1.00 | 4,035.1 | 3.00 | 0.000743 |
| 19 | 11,732.6 | 13 | 1.00 | 11,732.6 | 13.00 | 0.001108 |
| 20 | $23,730.1$ | 21 | 1.00 | 23.730 .1 | 21.00 | 0.000885 |
| 2) | 37,068.9 | 25 | 1.15 | 32,233.8 | 21.74 | 0.000674 |
| 22 | 50,629.3 | 34 | 1.26 | 40,182.0 | 26.98 | 0.000672 |
| 23 | 65,520.4 | 51 | 1.29 | 50,791.0 | 39.53 | 0.000778 |
| 24 | 81,412.3 | 38 | 1.08 | 75,381.8 | 35.19 | 0.000467 |
| 25 | 98,035.3 | 49 | 1.00 | 98,035.3 | 49.00 | 0.000500 |
| 26 | 115,749.5 | 51 | 1.07 | 108,177.1 | 47.66 | 0.000441 |
| 27 | 133,482.3 | 85 | 1.28 | 104,283.1 | 66.41 | 0.000637 |
| 28 | 150,968.9 | 74 | 1.42 | 106,316.1 | 52.11 | 0.000490 |
| 29 | 167,686.6 | 76 | 1.58 | 106,130.8 | 48.10 | 0.000453 |
| 30 | 180,835.6 | 93 | 1.18 | 153,250.5 | 78.81 | 0.000514 |
| 31 | 191,086.2 | 105 | 1.13 | 169,102.8 | 92.92 | 0.000549 |
| 32 | 200,923.9 | 143 | 1.36 | 147,738.2 | 105.15 | 0.000712 |
| 33 | 212,018.3 | 127 | 1.40 | 151,441.6 | 90.71 | 0.000599 |
| 34 | 225,011.1 | 144 | 1.44 | $156,257.7$ | 100.00 | 0.000640 |
| 35 | 240,211.3 | 167 | 1.47 | 163,409.1 | 113.61 | 0.000695 |
| 36 | 256,522.1 | 210 | 1.97 | 130,214.3 | 106.60 | 0.000819 |
| 37 | 275,593.1 | 201 | 1.42 | 194,051.5 | 141.55 | 0.000729 |
| 38 | 298,477.1 | 238 | 1.46 | 204,436.4 | 163.01 | 0.000797 |
| 39 | 323,673.6 | 307 | 1.62 | 199,798.5 | 189.51 | 0.000948 |
| 40 | 351,831.3 | 308 | 1.39 | 253,116.0 | 221.58 | 0.000875 |
| 41 | 384,724.8 | 357 | 1.36 | 282,885.9 | 262.50 | 0.000928 |
| 42 | 424,734.8 | 491 | 2.10 | 202,254.7 | 233.81 | 0.001156 |
| 43 | 475,237.6 | 525 | 1.56 | 304,639.5 | 336.54 | 0.001105 |
| 44 | 537,761.4 | 645 | 1.51 | 356,133.4 | 427.15 | 0.001199 |
| 45 | 580,177.6 | 825 | 1.70 | 341,280.9 | 485.29 | 0.001422 |
| 46 | 592,243.3 | 963 | 1.63 | 363,339.5 | 590.80 | 0.001626 |
| 47 | 593, 39.6 | 1,001 | 1.60 | 370,712.3 | 625.63 | 0.001688 |
| 48 | 574,704.2 | 1,087 | 1.81 | 317,516.1 | 600.55 | 0.001891 |
| 49 | 548,810.1 | 1,134 | 1.75 | 313,605.8 | 648.00 | 0.002066 |
| 50 | 519,668.5 | 1,281 | 1.70 | 305,687.4 | 753.53 | 0.002465 |

Table 1.15. (Continued).

| Age $x$ | $\mathrm{R}_{x}$ | $A_{x}$ | Variance Ratio | $\begin{aligned} & \text { Adjusted } \\ & \mathrm{R}_{x} \end{aligned}$ | Adjusted $\mathrm{A}_{x}$ | Crude $\mu_{x}=\mathrm{A}_{x} / \mathrm{R}_{x}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 51 | 487,844.6 | 1,333 | 1.72 | 283,630.6 | 775.00 | 0.002732 |
| 52 | 466,787.8 | 1,447 | 1.68 | 277,849.9 | 861.31 | 0.003100 |
| 53 | 458,539.4 | 1.634 | 1.73 | 265,051.7 | 944.51 | 0.003563 |
| 54 | 458,637.1 | 1,664 | 1.62 | 283,109.3 | 1,027.16 | 0.003628 |
| 55 | 450,022.4 | 1,765 | 1.59 | 283,033.0 | 1,110.06 | 0.003922 |
| 56 | 436,441.1 | 2,010 | 1.70 | 256,730.1 | 1,182.35 | 0.004605 |
| 57 | 423,193.6 | 2,187 | 1.60 | 264,496.0 | 1,366.88 | 0.005168 |
| 58 | 408,695.6 | 2,446 | 1.67 | 244,727.9 | 1,464.67 | 0.005985 |
| 59 | 390,958.6 | 2,532 | 1.60 | 244,349.1 | 1,582.50 | 0.006476 |
| 60 | 337,641.6 | 2,445 | 1.48 | 228,136.2 | 1,652.03 | 0.007241 |
| 61 | 301,243.3 | 2,662 | 1.57 | 191,874.7 | 1,695.54 | 0.008837 |
| 62 | 283,631.1 | 2,917 | 1.47 | 192,946.3 | 1,984.35 | 0.010284 |
| 63 | 263,928.9 | 2,982 | 1.54 | 171,382.4 | 1,936.36 | 0.011299 |
| 64 | 234,917.5 | 2,789 | 1.41 | 166,608.2 | 1,978.01 | 0.011872 |
| 65 | 135,496.6 | 1,866 | 1.28 | 105,856.7 | 1,457.81 | 0.013772 |
| 66 | 81,428.6 | 1,286 | 1.35 | 60,317.5 | 952.59 | 0.015793 |
| 67 | 69,644.6 | 1,149 | 1.19 | 58,524.9 | 965.55 | 0.016498 |
| 68 | 62,079.6 | 1,183 | 1.27 | 48,881.6 | 931.50 | 0.019056 |
| 69 | 55,106.3 | 1,156 | 1.16 | 47,505.4 | 996.55 | 0.020978 |
| 70 | 49,027.7 | 1.134 | 1.18 | 41,548.9 | 961.02 | 0.023130 |
| 71 | 44,957.6 | 1,110 | 1.15 | 39,093.6 | 965.22 | 0.024690 |
| 72 | 40,870.1 | 1,220 | 1.14 | 35,851.0 | 1,070.18 | 0.029851 |
| 73 | 35,189.3 | 1,194 | 1.33 | 26,458.1 | 897.74 | 0.033931 |
| 74 | 29,862.1 | 1,092 | 1.13 | 26,426.6 | 966.37 | 0.036568 |
| 75 | 24,832.1 | 1,034 | 1.31 | 18,955.8 | 789.31 | 0.041640 |
| 76 | 22,169.6 | 1,007 | 1.25 | 17,735.7 | 805.60 | 0.045423 |
| 77 | 21,344.8 | 1,024 | 1.26 | 16,940.3 | 812.70 | 0.047974 |
| 78 | 20,260.9 | 1,123 | 1.20 | 16,884.1 | 935.83 | 0.055427 |
| 79 | 18,463.3 | 1,115 | 1.23 | 15,010.8 | 906.50 | 0.060390 |
| 80 | 15,487.4 | 1,070 | 1.20 | 12,906.2 | 891.67 | 0.069088 |
| 81 | 13,436.6 | 1,036 | 1.16 | 11.583 .3 | 893.10 | 0.077103 |
| 82 | 11,795.9 | 987 | 1.17 | 10,082.0 | 843.59 | 0.083673 |
| 83 | 10,273.6 | 988 | 1.31 | 7,842.4 | 754.20 | 0.096169 |
| 84 | 8,727.6 | 953 | 1.26 | 6,926.7 | 756.35 | 0.109194 |
| 85 | 6,880.7 | 344 | 1.18 | 5,831.1 | 630.51 | 0.108129 |
| 86 | $5,469.2$ | 618 | 1.18 | 4,634.9 | 523.73 | 0.112996 |
| 87 | 4,422.1 | 627 | 1.18 | 3,747.5 | 531.36 | 0.141788 |
| 88 | 3,439.8 | 508 | 1.34 | 2,567.0 | 379.10 | 0.147683 |
| 89 | 2,730.0 | 406 | 1.34 | 2,037.3 | 302.99 | 0.148718 |
| 90 | 2,140.6 | 374 | 1.29 | 1,659.4 | 289.92 | 0.174717 |
| 91 | 1.625 .3 | 331 | 1.42 | 1,144.6 | 233.10 | 0.203655 |
| 92 | 1,251.0 | 218 | 1.22 | 1,025.4 | 178.69 | 0.174261 |

Table 1.15. (Continued).

| Age $x$ | $\mathbf{R}_{x}$ | $\mathbf{A}_{x}$ | Variance <br> Ratio | Adjusted <br> $\mathbf{R}_{x}$ | Adjusted <br> $\mathbf{A}_{x}$ | Crude <br> $\mu_{x}=\mathbf{A}_{x} / \mathbf{R}_{x}$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| 93 | 946.0 | 178 | 1.19 | 795.0 | 149.58 | 0.188161 |
| 94 | 703.8 | 167 | 1.28 | 549.8 | 130.47 | 0.237283 |
| 95 | 516.5 | 106 | 1.13 | 457.1 | 93.81 | 0.205227 |
| 96 | 400.0 | 71 | 1.08 | 370.4 | 65.74 | 0.177500 |
| 97 | 297.5 | 47 | 1.21 | 245.9 | 38.84 | 0.157983 |
| 98 | 241.8 | 23 | 1.00 | 241.8 | 23.00 | 0.095120 |
| 99 | 194.3 | 27 | 2.70 | 72.0 | 10.00 | 0.138960 |
| 100 | 300.3 | 28 | 1.00 | 300.3 | 28.00 | 0.093240 |
| 101 | 113.5 | 9 | 1.00 | 113.5 | 9.00 | 0.079295 |
| 102 | 80.3 | 5 | 1.00 | 80.3 | 5.00 | 0.062267 |
| 103 | 64.0 | 1 | 1.40 | 45.7 | 0.71 | 0.015625 |
| 104 | 52.5 | 2 | 1.00 | 52.5 | 2.00 | 0.038095 |
| 105 | 46.0 | 2 | 1.00 | 46.0 | 2.00 | 0.043478 |
| 106 | 40.0 | 0 | 1.00 | 40.0 | 0.00 | 0.000000 |
| 107 | 34.0 | 2 | 1.00 | 34.0 | 2.00 | 0.058824 |
| 108 | 127.5 | 3 | 1.00 | 127.5 | 3.00 | 0.023529 |
| Totals | $15,139,004.8$ | 68,963 |  | $9,853,930.8$ | $48,591.64$ |  |

Table 1.16. Details of graduation for males durations $2+$ : exposed to risk and actual deaths adjusted by variance ratios.

| Age $x$ | $\begin{aligned} & \text { Adjusted } \\ & \mathrm{R}_{x} \end{aligned}$ | $\mu_{x}$ | $\begin{aligned} & \text { Adjusted } \\ & \qquad \mathbf{A}_{x} \end{aligned}$ | $\mathbf{E}_{x}$ | $\mathrm{Dev}_{x}$ | $\left(\mathrm{V}_{s}\right)^{1 / 2}$ | $z_{x}$ | 100A/E |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 17 | 1,669.1 | 0.000604 | 0.00 | 1.01 |  |  |  |  |
| 18 | 4,035.1 | 0.000597 | 3.00 | 2.41 |  |  |  |  |
| 19 | 11,732.6 | 0.000590 | 13.00 | 6.93 |  |  |  |  |
| 17-19 | 17,436.8 | 0.000593 | 16.00 | 10.34 | 5.66 | 3.22 | 1.76 | 154.7 |
| 20 | 23,730.1 | 0.000584 | 21.00 | 13.87 | 7.13 | 3.72 | 1.91 | 151.4 |
| 21 | 32,233.8 | 0.000579 | 21.74 | 18.67 | 3.07 | 4.32 | 0.71 | 116.5 |
| 22 | 40,182.0 | 0.000575 | 26.98 | 23.09 | 3.90 | 4.80 | 0.81 | 116.9 |
| 23 | 50,791.0 | 0.000571 | 39.53 | 29.00 | 10.54 | 5.38 | 1.96 | 136.3 |
| 24 | 75,381.8 | 0.000568 | 35.19 | 42.84 | -7.65 | 6.54 | -1.17 | 82.1 |
| 25 | 98.035 .3 | 0.000567 | 49.00 | 55.57 | -6.57 | 7.45 | -0.88 | 88.2 |
| 26 | 108,177.1 | 0.000567 | 47.66 | 61.30 | -13.64 | 7.83 | -1.74 | 77.7 |
| 27 | 104,283.1 | 0.000568 | 66.41 | 59.25 | 7.15 | 7.70 | 0.93 | 112.1 |
| 28 | 106,316.1 | 0.000572 | 52.11 | 60.76 | -8.65 | 7.79 | -1.11 | 85.8 |
| 29 | 106,130.8 | 0.000577 | 48.10 | 61.23 | -13.13 | 7.82 | -1.68 | 78.6 |
| 30 | 153,250.5 | 0.000585 | 78.81 | 89.61 | -10.79 | 9.47 | $-1.14$ | 88.0 |
| 31 | 169,102.8 | 0.000595 | 92.92 | 100.66 | -7.74 | 10.03 | -0.77 | 92.3 |
| 32 | 147,738.2 | 0.000609 | 105.15 | 89.97 | 15.18 | 9.49 | 1.60 | 116.9 |
| 33 | 151,441.6 | 0.000626 | 90.71 | 94.85 | -4.14 | 9.74 | -0.42 | 95.6 |
| 34 | 156,257.7 | 0.000648 | 100.00 | 101.22 | -1.22 | 10.06 | -0.12 | 98.8 |
| 35 | 163,409.1 | 0.000674 | 113.61 | 110.13 | 3.48 | 10.49 | 0.33 | 103.2 |
| 36 | 130,214.3 | 0.000705 | 106.60 | 91.87 | 14.73 | 9.58 | 1.54 | 116.0 |
| 37 | 194,051.5 | 0.000743 | 141.55 | 144.22 | -2.67 | 12.01 | -0.22 | 98.2 |
| 38 | 204,436.4 | 0.000788 | 163.01 | 161.07 | 1.95 | 12.69 | 0.15 | 101.2 |
| 39 | 199,798.5 | 0.000840 | 189.51 | 167.93 | 21.58 | 12.96 | 1.67 | 112.9 |
| 40 | 253,116.0 | 0.000902 | 221.58 | 228.34 | $-6.76$ | 15.11 | -0.45 | 97.0 |
| 41 | 282,885.9 | 0.000974 | 262.50 | 275.53 | -13.03 | 16.60 | -0.78 | 95.3 |
| 42 | 202,254.7 | 0.001057 | 233.81 | 213.88 | 19.93 | 14.62 | 1.36 | 109.3 |
| 43 | 304,639.5 | 0.001154 | 336.54 | 351.58 | -15.04 | 18.75 | -0.80 | 95.7 |
| 44 | 356,133.4 | 0.001266 | 427.15 | 450.71 | -23.56 | 21.23 | -1.11 | 94.8 |
| 45 | 341,280.9 | 0.001394 | 485.29 | 475.68 | 9.61 | 21.81 | 0.44 | 102.0 |
| 46 | 363,339.5 | 0.001541 | 590.80 | 559.91 | 30.89 | 23.66 | 1.31 | 105.5 |
| 47 | 370,712.3 | 0.001710 | 625.63 | 633.75 | -8.13 | 25.17 | -0.32 | 98.7 |
| 48 | 317,516.1 | 0.001902 | 600.55 | 603.95 | $-3.40$ | 24.58 | -0.14 | 99.4 |
| 49 | 313,605.8 | 0.002122 | 648.00 | 665.38 | -17.38 | 25.80 | -0.67 | 97.4 |
| 50 | 305,687.4 | 0.002372 | 753.53 | 724.99 | 28.54 | 26.93 | 1.06 | 103.9 |
| 51 | 283,630.6 | 0.002656 | 775.00 | 753.23 | 21.77 | 27.45 | 0.79 | 102.9 |
| 52 | 277,849.9 | 0.002978 | 861.31 | 827.40 | 33.91 | 28.76 | 1.18 | 104.1 |
| 53 | 265,051.7 | 0.003343 | 944.51 | 886.01 | 58.49 | 29.77 | 1.97 | 106.6 |
| 54 | 283,109.3 | 0.003756 | 1,027.16 | 1,063.23 | -36.07 | 32.61 | -1.11 | 96.6 |
| 55 | 283,033.0 | 0.004222 | 1,110.06 | 1,194.87 | -84.81 | 34.57 | $-2.45$ | 92.9 |

Table 1.16. (Continued).

| Age $x$ | Adjusted $\mathrm{R}_{x}$ | $\mu_{x}$ | Adjusted $\mathrm{A}_{x}$ | $\mathrm{E}_{x}$ | $\mathrm{Dev}_{x}$ | $\left(\mathrm{V}_{x}\right)^{1 / 2}$ | $z_{x}$ | 100A/E |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 56 | 256,730.1 | 0.004747 | 1,182.35 | 1,218.79 | -36.43 | 34.91 | $-1.04$ | 97.0 |
| 57 | 264,496.0 | 0.005339 | 1,366.88 | 1,412.26 | -45.38 | 37.58 | -1.21 | 96.8 |
| 58 | 244,727.9 | 0.006005 | 1,464.67 | 1,469.69 | -5.02 | 38.34 | $-0.13$ | 99.7 |
| 59 | 244,349.J | 0.006754 | 1,582.50 | 1,650.22 | $-67.72$ | 40.62 | $-1.67$ | 95.9 |
| 60 | 228,136.2 | 0.007593 | 1,652.03 | 1,732.22 | -80.20 | 41.62 | $-1.93$ | 95.4 |
| 61 | 191,874.7 | 0.008534 | 1,695.54 | 1,637.37 | 58.17 | 40.46 | 1.44 | 103.6 |
| 62 | 192,946.3 | 0.009586 | 1,984.35 | 1,849.65 | 134.70 | 43.01 | 3.13 | 107.3 |
| 63 | 171,382.4 | 0.010763 | 1,936.36 | 1,844.64 | 91.73 | 42.95 | 2.14 | 105.0 |
| 64 | 166,608.2 | 0.012077 | 1,978.01 | 2,012.20 | -34.19 | 44.86 | -0.76 | 98.3 |
| 65 | 105,856.7 | 0.013543 | 1,457.81 | 1,433.64 | 24.18 | 37.86 | 0.64 | 101.7 |
| 66 | 60,317.5 | 0.015176 | 952.59 | 915.38 | 37.21 | 30.26 | 1.23 | 104.1 |
| 67 | 58,524.9 | 0.016993 | 965.55 | 994.51 | -28.96 | 31.54 | $-0.92$ | 97.1 |
| 68 | 48,881.6 | 0.019012 | 931.50 | 929.36 | 2.14 | 30.49 | 0.07 | 100.2 |
| 69 | 47,505.4 | 0.021255 | 996.55 | 1,009.71 | $-13.16$ | 31.78 | $-0.41$ | 98.7 |
| 70 | 41,548.9 | 0.023741 | 961.02 | 986.41 | -25.40 | 31.41 | -0.81 | 97.4 |
| 71 | 39,093.6 | 0.026495 | 965.22 | 1,035.80 | -70.58 | 32.18 | $-2.19$ | 93.2 |
| 72 | 35,851.0 | 0.029543 | 1,070.18 | 1,059.15 | 11.02 | 32.54 | 0.34 | 101.0 |
| 73 | 26,458.1 | 0.032912 | 897.74 | 870.79 | 26.96 | 29.51 | 0.91 | 103.1 |
| 74 | 26,426.6 | 0.036531 | 966.37 | 968.04 | -1.66 | 31.11 | $-0.05$ | 99.8 |
| 75 | 18,955.8 | 0.040733 | 789.31 | 772.12 | 17.19 | 27.79 | 0.62 | 102.2 |
| 76 | 17,735.7 | 0.045251 | 805.60 | 802.56 | 3.04 | 28.33 | 0.11 | 100.4 |
| 77 | 16,940.3 | 0.050223 | 812.70 | 850.80 | -38.10 | 29.17 | -1.31 | 95.5 |
| 78 | 16,884.1 | 0.055689 | 935.83 | 940.25 | -4.42 | 30.66 | -0.14 | 99.5 |
| 79 | 15,010.8 | 0.061690 | 906.50 | 926.01 | -19.51 | 30.43 | -0.64 | 97.9 |
| 80 | 12,906.2 | 0.068271 | 891.67 | 881.12 | 10.55 | 29.68 | 0.36 | 101.2 |
| 81 | 11,583.3 | 0.075482 | 893.10 | 874.32 | 18.78 | 29.57 | 0.64 | 102.1 |
| 82 | 10,082.0 | 0.083372 | 843.59 | 840.56 | 3.03 | 28.99 | 0.10 | 100.4 |
| 83 | 7,842.4 | 0.091998 | 754.20 | 721.49 | 32.71 | 26.86 | 1.22 | 104.5 |
| 84 | 6,926.7 | 0.101417 | 756.35 | 702.48 | 53.87 | 26.50 | 2.03 | 107.7 |
| 85 | 5,831.1 | 0.111691 | 630.51 | 651.28 | $-20.77$ | 25.52 | $-0.81$ | 96.8 |
| 86 | 4,634.9 | 0.122884 | 523.73 | 569.56 | -45.83 | 23.87 | $-1.92$ | 92.0 |
| 87 | 3,747.5 | 0.135066 | 531.36 | 506.17 | 25.19 | 22.50 | 1.12 | 105.0 |
| 88 | 2,567.0 | 0.148309 | 379.10 | 380.71 | -1.61 | 19.51 | -0.08 | 99.6 |
| 89 | 2,037.3 | 0.162690 | 302.99 | 331.45 | -28.47 | 18.21 | -1.56 | 91.4 |
| 90 | 1,659.4 | 0.178289 | 289.92 | 295.85 | -5.93 | 17.20 | -0.34 | 98.0 |
| 91 | 1,144.6 | 0.195190 | 233.10 | 223.41 | 9.69 | 14.95 | 0.65 | 104.3 |
| Totals | 9,844,448.3 |  | 47,795.80 | 47,795.82 | $-0.02$ |  |  | 100.0 |

Table 1.17. Details of graduation for females durations $2+$.

| Age $x$ | $\mathrm{R}_{x}$ | $\mu_{5}$ | $\mathrm{A}_{x}$ | $\mathrm{E}_{x}$ | $\mathrm{Dev}_{x}$ | $\left(V_{x}\right)^{1 / 2}$ | $z_{x}$ | 100A/E |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 17 | 1,438.6 | 0.000169 | 1 | 0.24 |  |  |  |  |
| 18 | 3,004.9 | 0.000175 | 1 | 0.53 |  |  |  |  |
| 19 | 7,489.1 | 0.000182 | 1 | 1.36 |  |  |  |  |
| 20 | 13,761.4 | 0.000189 | 3 | 2.60 |  |  |  |  |
| 21 | 22,677.2 | 0.000198 | 5 | 4.48 |  |  |  |  |
| 17-21 | 48,371.2 | 0.000191 | 11 | 9.22 | 1.78 | 3.04 | 0.59 | 119.4 |
| 22 | 32,217.1 | 0.000207 | 5 | 6.67 | $-1.67$ | 2.58 | -0.65 | 75.0 |
| 23 | 40,989.9 | 0.000217 | 7 | 8.90 | $-1.90$ | 2.98 | -0.64 | 78.6 |
| 24 | 48,532.7 | 0.000229 | 8 | 11.09 | -3.09 | 3.33 | $-0.93$ | 72.1 |
| 25 | 56,384.8 | 0.000241 | 16 | 13.60 | 2.40 | 3.69 | 0.65 | 117.7 |
| 26 | 65,565.9 | 0.000255 | 12 | 16.73 | $-4.73$ | 4.09 | $-1.16$ | 71.7 |
| 27 | 74,324.1 | 0.000271 | 21 | 20.12 | 0.88 | 4.49 | 0.20 | 104.4 |
| 28 | 82,633.1 | 0.000288 | 18 | 23.79 | -5.79 | 4.88 | $-1.19$ | 75.7 |
| 29 | 89,886.1 | 0.000307 | 24 | 27.59 | -3.59 | 5.25 | -0.68 | 87.0 |
| 30 | 94,863.0 | 0.000328 | 31 | 31.11 | -0.11 | 5.58 | -0.02 | 99.6 |
| 31 | 98,774.2 | 0.000351 | 27 | 34.71 | -7.71 | 5.89 | -1.31 | 77.8 |
| 32 | 102,439.6 | 0.000377 | 32 | 38.65 | -6.65 | 6.22 | $-1.07$ | 82.8 |
| 33 | 106,179.3 | 0.000406 | 44 | 43.11 | 0.89 | 6.57 | 0.14 | 102.1 |
| 34 | 110,276.6 | 0.000438 | 60 | 48.28 | 11.72 | 6.95 | 1.69 | 124.3 |
| 35 | 114,046.4 | 0.000473 | 46 | 53.95 | -7.95 | 7.35 | $-1.08$ | 85.3 |
| 36 | 118,249.1 | 0.000512 | 66 | 60.56 | 5.44 | 7.78 | 0.70 | 109.0 |
| 37 | 122,599.9 | 0.000555 | 56 | 68.09 | $-12.09$ | 8.25 | $-1.47$ | 82.2 |
| 38 | 128,432.3 | 0.000603 | 65 | 77.50 | $-12.50$ | 8.80 | -1.42 | 83.9 |
| 39 | 134,074.3 | 0.000657 | 98 | 88.03 | 9.97 | 9.38 | 1.06 | 111.3 |
| 40 | 139,153.1 | 0.000715 | 124 | 99.56 | 24.44 | 9.98 | 2.45 | 124.5 |
| 41 | 144,206.4 | 0.000781 | 122 | 112.59 | 9.41 | 10.61 | 0.89 | 108.4 |
| 42 | 149,323:3 | 0.000853 | 125 | 127.39 | $-2.39$ | 11.29 | -0.21 | 98.1 |
| 43 | 156,880.6 | 0.000933 | 155 | 146.42 | 8.58 | 12.10 | 0.71 | 105.9 |
| 44 | 165,740.6 | 0.001022 | 175 | 169.42 | 5.58 | 13.02 | 0.43 | 103.3 |
| 45 | 169,793.1 | 0.001121 | 216 | 190.28 | 25.72 | 13.79 | 1.86 | 113.5 |
| 46 | 167,221.0 | 0.001230 | 218 | 205.65 | 12.35 | 14.34 | 0.86 | 106.0 |
| 47 | 162,429.6 | 0.001351 | 268 | 219.40 | 48.60 | 14.81 | 3.28 | 122.2 |
| 48 | 153,393.6 | 0.001485 | 213 | 227.75 | -14.75 | 15.09 | -0.98 | 93.5 |
| 49 | 144,497.6 | 0.001633 | 249 | 236.00 | 13.00 | 15.36 | 0.85 | 105.5 |
| 50 | 137,280.5 | 0.001798 | 253 | 246.80 | 6.20 | 15.71 | 0.39 | 102.5 |
| 51 | 129,184.6 | 0.001980 | 242 | 255.81 | $-13.81$ | 15.99 | -0.86 | 94.6 |
| 52 | 123,498.6 | 0.002182 | 275 | 269.51 | 5.49 | 16.42 | 0.33 | 102.0 |
| 53 | 121,399.4 | 0.002406 | 307 | 292.12 | 14.88 | 17.09 | 0.87 | 105.1 |
| 54 | 120,374.6 | 0.002654 | 329 | 319.53 | 9.47 | 17.88 | 0.53 | 103.0 |
| 55 | 116,525.9 | 0.002930 | 333 | 341.36 | $-8.36$ | 18.48 | -0.45 | 97.6 |
| 56 | 110,757.9 | 0.003234 | 382 | 358.23 | 23.77 | 18.93 | 1.26 | 106.6 |
| 57 | 104,619.1 | 0.003572 | 340 | 373.71 | -33.71 | 19.33 | -1.74 | 91.0 |

Table 1.17. (Continued).

| Age $x$ | $\mathrm{R}_{x}$ | $\mu_{x}$ | $\mathrm{A}_{\boldsymbol{x}}$ | $\mathrm{E}_{x}$ | $\mathrm{Dev}_{x}$ | $\left(\mathrm{V}_{x}\right)^{1 / 2}$ | $z_{x}$ | 100A/E |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 58 | 97,828.4 | 0.003946 | 368 | 386.07 | $-18.07$ | 19.65 | $-0.92$ | 95.3 |
| 59 | 89,950.8 | 0.004361 | 382 | 392.29 | $-10.29$ | 19.81 | $-0.52$ | 97.4 |
| 60 | 74,586.1 | 0.004821 | 330 | 359.57 | -29.57 | 18.96 | $-1.56$ | 91.8 |
| 61 | 62,276.8 | 0.005330 | 302 | 331.95 | $-29.95$ | 18.22 | $-1.64$ | 91.0 |
| 62 | 56,357.6 | 0.005895 | 332 | 332.22 | -0.22 | 18.23 | -0.01 | 99.9 |
| 63 | 50,300.6 | 0.006520 | 319 | 327.98 | -8.98 | 18.11 | $-0.50$ | 97.3 |
| 64 | 43,332.1 | 0.007214 | 299 | 312.58 | $-13.58$ | 17.68 | -0.77 | 95.7 |
| 65 | 34,219.1 | 0.007982 | 277 | 273.13 | 3.87 | 16.53 | 0.23 | 101.4 |
| 66 | 28,227.1 | 0.008833 | 237 | 249.34 | $-12.34$ | 15.79 | -0.78 | 95.1 |
| 67 | 25,149.6 | 0.009777 | 227 | 245.88 | -18.88 | 15.68 | $-1.20$ | 92.3 |
| 68 | 22,799.3 | 0.010822 | 208 | 246.74 | $-38.74$ | 15.71 | $-2.47$ | 84.3 |
| 69 | 20,595.0 | 0.011981 | 247 | 246.74 | 0.26 | 15.71 | 0.02 | 100.1 |
| 70 | 18,508.7 | 0.013265 | 237 | 245.51 | -8.51 | 15.67 | -0.54 | 96.5 |
| 71 | 16,937.3 | 0.014688 | 252 | 248.77 | 3.23 | 15.77 | 0.21 | 101.3 |
| 72 | 15,228.1 | 0.016264 | 245 | 247.67 | -2.67 | 15.74 | -0.17 | 98.9 |
| 73 | 12,948.4 | 0.018012 | 219 | 233.22 | $-14.22$ | 15.27 | $-0.93$ | 93.9 |
| 74 | 10,546.8 | 0.019948 | 220 | 210.39 | 9.61 | 14.50 | 0.66 | 104.6 |
| 75 | 8,372.3 | 0.022094 | 215 | 184.97 | 30.03 | 13.60 | 2.21 | 116.2 |
| 76 | 7,189.9 | 0.024472 | 162 | 175.95 | -13.95 | 13.26 | $-1.05$ | 92.1 |
| 77 | 6,841.7 | 0.027107 | 197 | 185.46 | 11.54 | 13.62 | 0.85 | 106.2 |
| 78 | 6,567.8 | 0.030027 | 197 | 197.21 | -0.21 | 14.04 | -0.02 | 99.9 |
| 79 | 6,099.6 | 0.033263 | 210 | 202.89 | 7.11 | 14.24 | 0.50 | 103.5 |
| 80 | 5,374.3 | 0.036849 | 174 | 198.04 | -24.04 | 14.07 | $-1.71$ | 87.9 |
| 81 | 4,723.1 | 0.040823 | 219 | 192.81 | 26.19 | 13.89 | 1.89 | 113.6 |
| 82 | 4,097.1 | 0.045227 | 192 | 185.30 | 6.70 | 13.61 | 0.49 | 103.6 |
| 83 | 3,405.6 | 0.050108 | 198 | 170.65 | 27.35 | 13.06 | 2.09 | 116.0 |
| 84 | 2,693.7 | 0.055516 | 160 | 149.54 | 10.46 | 12.23 | 0.86 | 107.0 |
| 85 | 1,832.8 | 0.061510 | 107 | 112.74 | -5.74 | 10.62 | $-0.54$ | 94.9 |
| 86 | 1,289.5 | 0.068152 | 95 | 87.88 | 7.12 | 9.37 | 0.76 | 108.1 |
| 87 | 1,007.4 | 0.075512 | 64 | 76.07 | $-12.07$ | 8.72 | -1.38 | 84.1 |
| 88 | 765.6 | 0.083668 | 74 | 64.06 | 9.94 | 8.00 | 1.24 | 115.5 |
| 89 | 573.6 | 0.092707 | 62 | 53.18 | 8.82 | 7.29 | 1.21 | 116.6 |
| Totals | 4,925,744.9 |  | 12,000 | 12,000.01 | -0.01 |  |  | 100.0 |



Figure 1.1. Permanent assurances, values of $q(x)$.


Figure 1.2. Ratios of $q(x)$ to those for durations $2+$, males.


Figure 1.3. Ratios of $q(x)$ to those for durations $2+$, females.


Figure 1.4. Ratios of $q(x)$ for females to those for males.


Figure 1.5. Ratios of $q(x)$ for males for 1991-94 to those of AM80.


Figure 1.6. Ratios of $q(x)$ for females for 1991-94 to those of AF80.

## 2. IMMEDIATEANNUITANTS, MALESANDFEMALES, lives and amounts

### 2.1 Introduction

2.1.1 The investigations into the mortality of immediate annuitants, once one of the most important of the CMI Bureau's investigations, are now relatively small. The exposed to risk for both males and females has reduced since the last graduation exercise in 1979-82, as can be seen from Table 2.1, which also shows the corresponding figures (for lives only) for the five years 192125 , the first experience published under the new permanent investigation. It can be seen that the experience for 1991-94 is smaller than that for 1921-25, even allowing for the extra year's exposure in the earlier investigation.
2.1.2 It will be seen below that the mortality of immediate annuitants (by lives) is no longer lower than that of permanent assurances, as it used to be. For males it is near enough the same, for females rather higher. Why this should be deserves some consideration of the class of person who nowadays buys a purchased life annuity, especially one with no guaranteed period. One might speculate that one of the possible reasons for the purchase of such an annuity is to provide for a lifetime income left as a legacy; the executors of a will with such a provision may prefer to purchase an annuity rather than to keep a trust alive for possibly many years, and, in these circumstances, the state of health of the beneficiary may not be taken into consideration.
2.1.3 The bulk of the data is of course at older ages, as shown in Table 2.2, and for both sexes it is concentrated into the ages 70 to 100, though there is a small amount of rather implausible data at very young ages. The experience for duration 0 , for both sexes, is quite small, but the traditional method of using one year's selection has been maintained on this occasion. It will be seen that the graduation of the duration 0 data is necessarily rather tentative and relies heavily on the graduation for durations 1 and over (durations $1+$ ).
2.1.4 The average amounts per life are much higher than in 1979-82 as might be expected. The increase is much the same for both males and females, as shown in Table 2.3.

### 2.2 Males, lives

2.2.1 We consider first the data for males, lives, durations $1+$, for which there were 2,886 deaths. First, the data were graduated using a variety of GM $(r, s)$ formulae, with the results as shown in Table 2.4. A GM( 0,2 ) formula gives as good a fit as any higher order formula. However, the values of $q_{x}$ at low ages were unreasonable, and it was desirable to try a different approach.
2.2.2 The experience was therefore compared with the graduated rates for permanent assurances, males, durations $2+$, using the rates as fitted to the data without variance ratios. (These rates are very close to, but not identical with, the fitted rates for permanent assurances, males, durations $2+$ with variance ratios, as shown in part 1 of this report.) The value of $100 \mathrm{~A} / \mathrm{E}$ was 98.59 , and the fit was quite satisfactory. It was therefore decided to adopt these graduated rates as the rates for immediate annuitants, males, durations 1+. Details of the graduation statistics are shown in Table 2.5 and specimen values of $q_{x}$ are shown in Table 2.6. Note that standard errors of the estimates of the parameter values and of the values of $q_{x}$ are not available with this method of fitting the rates.
2.2.3 We now consider the data for males, lives, duration 0 . With only 104 deaths it is difficult to fit an independent graduation. Various orders of $\mathrm{GM}(r, s)$ formula gave almost identical values of the $\log$ likelihood and of $\chi^{2}$. When the data were compared with the rates adopted for durations $1+$, the resulting value of $100 \mathrm{~A} / \mathrm{E}$ was 73.801 . If a $\mathrm{GM}(r, s)$ formula is modified by multiplying all values of $\mu_{x}$ by a constant, $k$, the resulting parameters, denoted by ${ }^{*}$, are modified as follows: $a_{i}^{*}=k \cdot a_{i}$ for all $i ; b_{1}^{*}=b_{1}+\ln k ; b_{i}^{*}=b_{i}$ for $i \geq 2$. The graduated values of $\mu_{x}$ for duration 0 were therefore taken as $73.801 \%$ of those for durations $1+$; the fit was reasonably satisfactory, although the value of $\chi^{2}$ is rather on the high side. The usual details are given in Tables 2.5 and 2.6 .

### 2.3 Males, amounts

2.3.1 We now consider the data for males amounts. As with the lives data, for durations $1+$, the various $\mathrm{GM}(r, s)$ formulae tried show very similar values for the $\log$ likelihood and for $\chi^{2}$, so the simplest formula, $\operatorname{GM}(0,2)$, would appear to be the best if we had no other information. However, when we compare the data for amounts, durations $1+$ with the rates adopted for lives, durations $1+$ (the same as those for permanent assurances, males, durations $2+$, without variance ratios), we find a value of $100 \mathrm{~A} / \mathrm{E}$ of 95.344 , and not a bad fit. However, the fit can be improved a little by taking the values of $\mu_{x}$ for amounts, durations $1+$ as $95.344 \%$ of those for lives, durations $1+$. The usual details are shown in Tables 2.5 and 2.6. Note that a high value of $\chi^{2}$ is not unusual for amounts data.
2.3.2 For amounts, duration 0 a similar investigation led to the use of the values of $\mu_{x}$ for amounts, durations $1+$ multiplied by 0.81786 ; this is equivalent to the lives, durations $1+$ values multiplied by 0.77978 . Note that the amounts data for duration 0 shows higher mortality than the lives data for duration 0 , by about $5 \%$, whereas for durations $1+$ the experience by amounts is about $5 \%$
lower than that for lives. There are no conspicuously large amounts at any one age that might account for this, but with only 104 deaths in the lives data, random fluctuations mean that such differences should not be taken as significant. Indeed the amounts data at neither duration is very significantly different from that of the lives, and separate tables would hardly be justified. The usual details of the graduations are shown in Tables 2.5 and 2.6.

### 2.4 Females, lives

2.4.1 We now consider the experience for females, starting with lives, durations $1+$, which, with 5,857 deaths (omitting ages 10 to 16 ) is the largest part of the experience for immediate annuitants. When the data is compared with the graduated rates for permanent assurances, females, durations $2 t$, the value of $100 \mathrm{~A} / \mathrm{E}$ is 130.18 , so there is no justification for using these rates, as was done for the males. But it is notable that the mortality rates for female immediate annuitants are now so much higher than those for permanent assurances.
2.4.2 A variety of $\mathrm{GM}(r, s)$ formulae were fitted. The results are shown in Table 2.7. The most satisfactory formula was GM(1,2), and this was an adequate fit. Higher order formulae showed better values of the log likelihood, but were of an unsuitable shape. However, with this GM(1,2) formula the mortality rates for females rose above those for mates above age 100 , so experiments with a $\mathrm{GM}(1,3)$ formula and various values of $b_{3}$ were tried, allowing the other parameters to be optimised. A negative value of $b_{3}$ reduces the values of $\mu_{x}$ at higher ages. However, it increased the values of $\mu_{x}$ at young ages unreasonably, so it was necessary also to fix the value of $a_{1}$ at a suitable level, optimising on $b_{1}$ and $b_{2}$. The chosen values were:

| $100 a_{1}$ | 0.03 |
| :---: | :---: |
| $b_{3}$ | -0.9 |

The resulting GM $(1,3)$ formula gave a satisfactory fit, rather better than that for GM $(1,2)$ alone. The usual statistics are shown in Tables 2.9 and 2.10.
2.4.3 The experience for females, lives, duration 0 proved difficult. There were only 146 deaths, and each GM $(r, s)$ formula tried gave similar values for the log likelihood and for $\chi^{2}$. However, none was a clearly good fit. When the graduated values of $\mu_{x}$ for durations $1+$ are applied, the resulting value of $100 \mathrm{~A} / \mathrm{E}$ is 71.724 , but this still gives a rather poor fit. Nevertheless it seemed to provide the most reasonable solution, so the values of $\mu_{x}$ for lives, durations $1+$ were multiplied by 0.71724 . This graduation is recommended. The mortality rates lie below those for males, duration 0 throughout. The usual statistics are shown in Tables 2.9 and 2.10.

### 2.5 Females, amounts

2.5.1 The experience for females, amounts, durations $1+$ is considered next. A variety of GM( $r, s)$ formulae were fitted, the statistics for which are shown in Table 2.8. A GM(1,2) formula fitted reasonably, but as for lives, the mortality rates rose above those for males at high ages. The same process as for lives was therefore used, fitting the values of $a_{1}$ and $b_{3}$ as:

| $100 a_{3}$ | 0.04 |
| :---: | :---: |
| $b_{3}$ | -1.5 |

The resulting rates are generally lower than for females, lives, durations $1+$, but are higher at ages up to 39 and from 87 upwards. The usual statistics are shown in Tables 2.9 and 2.10. The overall level of mortality is almost the same as that of lives, durations $1+$, but the fit of these rates would not be satisfactory.
2.5.2 For females, amounts, duration 0 , the best procedure seemed to be, as before, to use a fixed multiple of the values of $\mu_{x}$ for amounts, durations $1+$, in this case $82.226 \%$. The fit is reasonable. The usual statistics are shown in Tables 2.9 and 2.10 .

### 2.6 Comparisons

2.6.1 In Table 2.11 are shown ratios comparing the rates $\left(q_{x}\right)$ for duration 0 with those for durations $1+$. Note that, even when $\mu_{x}$ for duration 0 is taken as a fraction of $\mu_{x}$ for durations $1+$, (a) the ratios of the values of $q_{x}$ are slightly irregular, because the values of $q_{x}$ are rounded to six decimal places, which at young ages is only three significant figures and (b) the ratios turn upwards at higher ages because, as the value of $\mu_{x}$ rises, the effect on $q_{x}$ is not linear. These ratios are graphed in Figure 2.1.
2.6.2 In Table 2.12 are shown ratios comparing the rates for amounts and for lives. Note that, for males, the rates for amounts duration 0 are higher than those for lives throughout, whereas for females at low ages and high ages the rates for amounts are higher than those for lives, for both durations. These ratios are graphed in Figure 2.2.
2.6.3 In Table 2.13 are shown ratios comparing the rates for the sexes. Note that the rates for females durations it rise above those for males, by a very small amount, for a few ages above 100 . This result is a feature of the graduation rather than the data. These ratios are graphed in Figure 2.3. Note that the ratios for amounts, duration 0 and durations $1+$ are very close, as can be seen on the graph, but are not identical.
2.6.4 In Table 2.14 are shown ratios comparing the rates for lives for 1991-94 with the corresponding rates for the IM80 and IF80 tables projected to calendar
year 1992. Note that the new rates are neither consistently lower nor consistently higher than the projected rates. These ratios are graphed in Figure 2.4.

### 2.7 The proposed tables

2.7.1 Values of $q_{x}$ for the proposed tables for immediate annuitants are shown in Appendix A in Table A6 for males and in Table A7 for females. As in the IM80 and IF80 tables, the rates for duration 0 are stopped at age 100 .

Table 2.1. Immediate annuitants, males and females, lives and amounts, durations 0 and $1+:$ comparison of (central) exposed to risk and deaths for 1991-94, for 1979-82 and for the five years 1921-25 (lives only).

|  | 1979-82 | 1991-94 | 1921-25 |
| :---: | :---: | :---: | :---: |
| Males |  |  |  |
| Lives |  |  |  |
| Duration 0 |  |  |  |
| Central exposed | 2,933.5 | 2,933.8 | 4,001.0 |
| Deaths | 122 | 104 | 123 |
| Durations I+ |  |  |  |
| Central exposed | 61,503.5 | 35,689.5 | 49,121.5 |
| Deaths | 4,771 | 2,886 | 3,350 |
| Amounts $£$ |  |  |  |
| Duration 0 |  |  |  |
| Central exposed | 3,463,312.0 | 8,590,467.4 |  |
| Deaths | 174.875 | 451,992 |  |
| Durations 1+ |  |  |  |
| Central exposed | 45,392,287.5 | 52,612,320.8 |  |
| Deaths | 3,857,585 | 4,181,932 |  |
| Females |  |  |  |
| Lives |  |  |  |
| Duration 0 |  |  |  |
| Central exposed | 3,967.0 | 4,033.3 | 9,634.5 |
| Deaths | 154 | 146 | 205 |
| Durations 1+ |  |  |  |
| Central exposed | 142,137.5 | 66,470.3 | 153,625.0 |
| Deaths | 9,789 | 5,863 | 7,738 |
| Amounts $£$ |  |  |  |
| Duration 0 |  |  |  |
| Central exposed | 4,567,483.0 | 13,099,689.2 |  |
| Deaths | 201,549 | 715,195 |  |
| Durations 1- |  |  |  |
| Central exposed | 79,089,613.0 | 97,571,461.7 |  |
| Deaths | 5,342,282 | 8,150,260 |  |

Table 2.2. Age ranges.

|  | Range of data | Exposed $\geq 100$ | Deaths $\geq 10$ |
| :--- | :---: | :---: | :---: |
| Males |  |  |  |
| $\quad$ Duration 0 | $19-100$ | $65-66$ | 84 |
| Durations $1+$ | $22-107$ | $61-96$ | $67-100$ |
|  |  |  |  |
| Females | 32107 | $71-89$ | None |
| $\quad$Duration 0 <br> Durations $1+$ | $10-108$ | $60-101$ | $70-104$ |

Table 2.3. Average amounts per life, 1979-82 and 1991-94.

|  | $1979-82$ | $1991-94$ |
| :--- | ---: | ---: |
| Males |  |  |
| $\quad$ Duration 0 | $£ 1,180.61$ | $£ 2,928.10$ |
| Durations 1+ | $£ 738.04$ | $£ 1,474.17$ |
|  |  |  |
| Females | $£ 1,151.37$ | $£ 3,247.88$ |
| $\quad$ Duration 0 | $£ 556.43$ | $£ 1,467.90$ |

Table 2.4. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of males, lives, durations $1+$ (age range used: $22-107$ ).

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) | GM(2,3) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $9,339.4$ | $9,339.2$ | $9,339.1$ | 9.337 .6 | $9,337.2$ | $9,337.2$ | $9,337.1$ |
| $\chi^{2}$ | 34.4 | 34.0 | 34.5 | 34.2 | 33.6 | 32.8 | 33.2 |

Table 2.5. Immediate annuitants, males, lives and amounts, durations 0 and $1+$ : statistics for graduations of $\mu_{x}=\operatorname{GM}(r, s)$.

|  | Lives |  |  | Amounts |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  |  | 0 | $1-$ |  | 0 |
| Duration: | GM(2,3) | $\operatorname{GM}(2,3)$ |  | $\operatorname{GM}(2,3)$ | $\operatorname{GM}(2,3)$ |

Basis: Permanent assurances, males, durations $2-$, times a factor:

| $100 a_{1}$ | 0.010649 | 0.014429 | 0.011251 | 0.013757 |
| :--- | ---: | ---: | ---: | ---: |
| $100 a_{2}$ | -0.029980 | -0.040629 | -0.031682 | -0.038737 |
| $b_{1}$ | -4.703659 | -4.399861 | -4.648604 | -4.447540 |
| $b_{2}$ | 5.568973 | 5.568973 | 5.568973 | 5.568973 |
| $b_{3}$ | -0.654909 | -0.654909 | -0.654909 | -0.654909 |


| Sign test: $p$ (pos) | 0.4018 | 0.3179 | 0.2517 | 0.2148 |
| :--- | :--- | :--- | :--- | :--- |
| Runs test: $p$ (runs) | 0.9748 | 0.8081 | 0.2966 | 0.4928 |
| K-S test: $p(K S)$ | 0.9996 | 0.8438 | 0.0303 | 0.0073 |

Serial correlation test:

| $T$-ratio 1 | -2.10 | -1.35 | -0.63 | 1.01 |
| :--- | ---: | ---: | ---: | ---: |
| $T$-ratio 2 | 0.38 | 0.68 | 1.17 | -0.71 |
| $T$-ratio 3 | -0.56 | 0.22 | -0.58 | 1.12 |
| $\chi^{2}$ rest: |  |  |  |  |
| $\chi^{2}$ | 25.37 | 37.75 | 80.16 | 157.95 |
| Degrees of freedom | 11 | 35 | 15 | 35 |
| $p\left(\chi^{2}\right)$ | 0.0080 | 0.34 | 0.000000 | 0.00000 |

Note that the method of graduation does not permit the calculation of the standard errors (and hence $T$-ratios) of estimates of the parameter values.

Table 2.6. Immediate annuitants, males, lives and amounts, durations 0 and $1+$ : specimen values of $q_{x}$.

|  | Lives |  | Amounts |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Duration: <br> Formula | 0 <br> GM $(2,3)$ | $1+$ <br> $\mathrm{GM}(2,3)$ | 0 <br> GM $(2,3)$ | GM $(2,3)$ |
| Age 20 | 0.000423 | 0.000573 | 0.000447 | 0.000546 |
| Age 30 | 0.000438 | 0.000593 | 0.000463 | 0.000566 |
| Age 40 | 0.000697 | 0.000945 | 0.000737 | 0.000901 |
| Age 50 | 0.001852 | 0.002508 | 0.001956 | 0.002392 |
| Age 60 | 0.005925 | 0.008019 | 0.006259 | 0.007647 |
| Age 70 | 0.018382 | 0.024826 | 0.019412 | 0.023684 |
| Age 80 | 0.051702 | 0.069406 | 0.054547 | 0.066284 |
| Age 90 | 0.128304 | 0.169777 | 0.135053 | 0.162554 |
| Age 100 | 0.273933 | 0.351927 | 0.286970 | 0.338706 |
| Age 110 | 0.489381 | 0.597771 | 0.508441 | 0.580348 |

Note that the method of graduation docs not permit the calculation of the standard errors of estimates of the values of $q_{x}$.

Table 2.7. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of females, lives, durations $1+$ (age range used: 17-108).

|  | $\mathrm{GM}(0,2)$ | $\mathrm{GM}(0,3)$ | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0.4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $18,359.1$ | $18,354.6$ | $18,359.0$ | $18,344.7$ | $18,345.6$ | $18,354.1$ |
| $\chi^{2}$ | 91.3 | 87.1 | 91.4 | 75.7 | 77.8 | 86.7 |

Table 2.8. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of females, amounts, durations $1+$ (age range used: 17-108).

|  | $\mathrm{GM}(0,2)$ | $\mathrm{GM}(0,3)$ | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0,4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likclihood | $17,567.6$ | $17,560.6$ | $17,567.3$ | $17,549.6$ | 17.550 .3 | 17.556 .4 |
| $\chi^{2}$ | 156.3 | 152.5 | 160.5 | 140.5 | 141.9 | 146.7 |

Table 2.9. Immediate annuitants, females, lives and amounts; durations 0 and $1+:$ statistics for graduations of $\mu_{x}=\mathrm{GM}(r, s)$.

| Duration: <br> Formula | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} 0 \\ \mathrm{GM}(1,3) \end{gathered}$ | $\begin{gathered} 1+ \\ \mathrm{GM}(1,3) \end{gathered}$ | $\begin{gathered} 0 \\ \mathrm{GM}(1,3) \end{gathered}$ | $\begin{gathered} 1+ \\ \mathrm{GM}(1,3) \end{gathered}$ |
| Factor | $\begin{aligned} & 0.71724 \\ & \text { times Lives } \end{aligned}$ |  | $\begin{gathered} 0.82226 \\ \text { times Amounts } \end{gathered}$ |  |
| $100 a_{1}$ | 0.021517 | 0.03 | 0.032890 | 0.04 |
| T-ratio |  |  |  |  |
| $b_{1}$ | -5.597708 | -5.265363 | -6.378326 | -6.182627 |
| $T$-ratio |  | -128.9 |  | -140.1 |
| $b_{2}$ | 6.683129 | 6.683129 | 8.027676 | 8.027676 |
| T-ratio |  | 65.1 |  | 70.9 |
| $b_{3}$ | -0.9 | -0.9 | -1.5 | -1.5 |
| Sign test: $p$ (pos) | 0.3238 | 0.6742 | 0.2517 | 0.7388 |
| Runs test: $p$ (runs) | 0.5467 | 0.0238 | 0.4800 | 0.9114 |
| K-S test: $p(K S)$ | 0.1548 | 0.8126 | 0.2107 | 0.1718 |
| Serial correlation test: |  |  |  |  |
| $T$-ratio 1 | 0.14 | 1.06 | -1.33 | -0.08 |
| $T$-ratio 2 | 1.51 | -0.22 | 0.95 | -1.18 |
| $T$-ratio 3 | -0.12 | -0.14 | -0.79 | -1.27 |
| $\chi^{2}$ test: |  |  |  |  |
| $\chi^{2}$ | 33.94 | 87.37 | 64.71 | 153.81 |
| Degrees of freedom | 15 | 40 | 16 | 35 |
| $p\left(\chi^{2}\right)$ | 0.0035 | 0.000022 | 0.000000 | 0.000000 |

Note that the method of graduation for duration 0 does not permit the calculation of the standard crrors (and hence $T$-ratios) of estimates of the parameter values.

Table 2.10. Immediate annuitants, females, lives and amounts, durations 0 and $1+:$ specimen values of $q_{x}$ and percentage standard errors.

| Duration: <br> Formula | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} 0 \\ \operatorname{GM}(1,3) \end{gathered}$ | $\begin{gathered} 1+ \\ \mathrm{GM}(1,3) \end{gathered}$ | $\begin{gathered} 0 \\ \operatorname{GM}(1,3) \end{gathered}$ | $\begin{gathered} 1+ \\ \mathrm{GM}(1,3) \end{gathered}$ |
| Age 20 | 0.000217 | 0.000303 | 0.000329 | 0.000400 |
| Percentage s.e. |  | 0.04 |  | 0.00 |
| Age 30 | 0.000230 | 0.000321 | 0.000331 | 0.000402 |
| Percentage s.e. |  | 0.27 |  | 0.03 |
| Age 40 | 0.000310 | 0.000432 | 0.000352 | 0.000429 |
| Percentage s.e. |  | 1.25 |  | 0.30 |
| Age 50 | 0.000727 | 0.001014 | 0.000540 | 0.000656 |
| Percentage s.e. |  | 2.89 |  | 1.73 |
| Age 60 | 0.002613 | 0.003642 | 0.001815 | 0.002207 |
| Percentage s.e. |  | 3.76 |  | 3.63 |
| Agc 70 | 0.009917 | 0.013799 | 0.008543 | 0.010381 |
| Percentage s.e. |  | 3.99 |  | 4.24 |
| Age 80 | 0.033910 | 0.046960 | 0.035680 | 0.043224 |
| Percentage s.e. |  | 3.98 |  | 4.30 |
| Age 90 | 0.099302 | 0.135683 | 0.116540 | 0.139889 |
| Percentage s.e. |  | 3.80 |  | 4.10 |
| Age 100 | 0.240692 | 0.318800 | 0.283964 | 0.333841 |
| Percentage s.e. |  | 3.36 |  | 3.59 |
| Age 110 | 0.466470 | 0.583518 | 0.507942 | 0.577873 |
| Percentage s.e. |  | 2.56 |  | 2.78 |

Note that the method of graduation for duration 0 does not permit the calculation of the standard errors of estimates of the values of $q_{x}$.

Table 2.11. Immediate annuitants: ratios of values of $q_{x}$ in proposed tables: comparison of duration 0 and durations $1+$ (duration 0 divided by durations $1+$ ).

|  | Males |  |  | Females |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Age | Lives | Amounts |  | Lives | Amounts |
| 20 | 0.7382 | 0.8187 | 0.7162 | 0.8225 |  |
| 25 | 0.7381 | 0.8178 |  | 0.7175 | 0.8204 |
| 30 | 0.7386 | 0.8180 |  | 0.7165 | 0.8234 |
| 35 | 0.7385 | 0.8178 |  | 0.7175 | 0.8215 |
| 40 | 0.7376 | 0.8180 |  | 0.7176 | 0.8205 |
| 45 | 0.7381 | 0.8181 |  | 0.7173 | 0.8217 |
| 50 | 0.7384 | 0.8177 |  | 0.7170 | 0.8232 |
| 55 | 0.7385 | 0.8182 |  | 0.7170 | 0.8229 |
| 60 | 0.7389 | 0.8185 |  | 0.7175 | 0.8224 |
| 65 | 0.7394 | 0.8188 | 0.7179 | 0.8226 |  |
| 70 | 0.7404 | 0.8196 | 0.7187 | 0.8229 |  |
| 75 | 0.7421 | 0.8209 | 0.7199 | 0.8239 |  |
| 80 | 0.7449 | 0.8229 |  | 0.7221 | 0.8255 |
| 85 | 0.7492 | 0.8261 | 0.7258 | 0.8283 |  |
| 90 | 0.7557 | 0.8308 | 0.7319 | 0.8331 |  |
| 95 | 0.7652 | 0.8377 | 0.7412 | 0.8404 |  |
| 100 | 0.7784 | 0.8473 |  | 0.7550 | 0.8506 |

Table 2.12. Immediate annuitants ratios of values of $q_{x}$ in proposed tables: comparison of rates for lives and amounts (amounts divided by lives).

| Age | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 20 | 1.0567 | 0.9529 | 1.5161 | 1.3201 |
| 25 | 1.0552 | 0.9522 | 1.4887 | 1.3019 |
| 30 | 1.0571 | 0.9545 | 1.4391 | 1.2523 |
| 35 | 1.0564 | 0.9540 | 1.3228 | 1.1554 |
| 40 | 1.0574 | 0.9534 | 1.1355 | 0.9931 |
| 45 | 1.0571 | 0.9537 | 0.9134 | 0.7974 |
| 50 | 1.0562 | 0.9537 | 0.7428 | 0.6469 |
| 55 | 1.0564 | 0.9534 | 0.6746 | 0.5878 |
| 60 | 1.0564 | 0.9536 | 0.6946 | 0.6060 |
| 65 | 1.0563 | 0.9538 | 0.7664 | 0.6690 |
| 70 | 1.0560 | 0.9540 | 0.8615 | 0.7523 |
| 75 | 1.0557 | 0.9544 | 0.9610 | 0.8397 |
| 80 | 1.0550 | 0.9550 | 1.0522 | 0.9204 |
| 85 | 1.0541 | 0.9560 | 1.1255 | 0.9863 |
| 90 | 1.0526 | 0.9575 | 1.1736 | 1.0310 |
| 95 | 1.0505 | 0.9595 | 1.1918 | 1.0512 |
| 100 | 1.0476 | 0.9624 | 1.1798 | 1.0472 |
| 105 |  | 0.9662 |  | 1.0240 |
| 110 |  | 0.9709 |  | 0.9903 |
| 115 |  | 0.9762 |  | 0.9558 |

40 Standard Tables of Mortality based on the 1991-94 Experiences
Table 2.13. Immediate annuitants ratios of values of $q_{x}$ in proposed tables: comparison of rates for males and females (females divided by males).

|  | Lives |  |  | Amounts |  |
| ---: | :---: | :---: | :---: | :---: | :---: |
|  | Age | Duration 0 | Durations 1+ |  |  |
|  |  |  | Duration 0 | Durations 1+ |  |
| 20 | 0.5130 | 0.5288 | 0.7360 | 0.7326 |  |
| 25 | 0.5300 | 0.5451 | 0.7477 | 0.7454 |  |
| 30 | 0.5251 | 0.5413 | 0.7149 | 0.7102 |  |
| 35 | 0.4942 | 0.5086 | 0.6188 | 0.6160 |  |
| 40 | 0.4448 | 0.4571 | 0.4776 | 0.4761 |  |
| 45 | 0.4046 | 0.4163 | 0.3496 | 0.3481 |  |
| 50 | 0.3925 | 0.4043 | 0.2761 | 0.2742 |  |
| 55 | 0.4072 | 0.4194 | 0.2600 | 0.2586 |  |
| 60 | 0.4410 | 0.4542 | 0.2900 | 0.2886 |  |
| 65 | 0.4867 | 0.5012 | 0.3531 | 0.3515 |  |
| 70 | 0.5395 | 0.5558 | 0.4401 | 0.4383 |  |
| 75 | 0.5966 | 0.6150 | 0.5431 | 0.5411 |  |
| 80 | 0.6559 | 0.6766 | 0.6541 | 0.6521 |  |
| 85 | 0.7156 | 0.7387 | 0.7641 | 0.7620 |  |
| 90 | 0.7740 | 0.7992 | 0.8629 | 0.8606 |  |
| 95 | 0.8290 | 0.8558 | 0.9405 | 0.9375 |  |
| 100 | 0.8787 | 0.9059 | 0.9895 | 0.9856 |  |
| 105 |  | 0.9466 |  | 1.0033 |  |
| 110 |  | 0.9762 |  | 0.9957 |  |
| 115 |  |  |  | 0.9940 |  |

Table 2.14. Ratios of values of $q_{x}$ in proposed tables for lives to those in corresponding IM80 and IF80 tables projected to calendar year 1992, i.e. IM80C92 and IF80C92 (1991-94 divided by Ix80C92).

| Age | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations $1+$ | Duration 0 | Durations 1+ |
| 20 | 0.6034 | 0.7860 | 1.2616 | 1.3407 |
| 25 | 0.7694 | 1.0036 | 1.2924 | 1.3689 |
| 30 | 0.8921 | 1.1627 | 1.0599 | 1.1263 |
| 35 | 0.8385 | 1.0909 | 0.7768 | 0.8213 |
| 40 | 0.6894 | 0.8991 | 0.5962 | 0.6316 |
| 45 | 0.5874 | 0.7652 | 0.5165 | 0.5479 |
| 50 | 0.5481 | 0.7137 | 0.5174 | 0.5493 |
| 55 | 0.5484 | 0.7141 | 0.5622 | 0.5997 |
| 60 | 0.5700 | 0.7420 | 0.6367 | 0.6842 |
| 65 | 0.5967 | 0.7687 | 0.7094 | 0.7676 |
| 70 | 0.7300 | 0.8517 | 0.7799 | 0.8500 |
| 75 | 0.8730 | 0.9232 | 0.8368 | 0.9183 |
| 80 | 1.0080 | 0.9770 | 0.8715 | 0.9627 |
| 85 | 1.0398 | 1.0085 | 0.8794 | 0.9773 |
| 90 | 1.0460 | 1.0156 | 0.8688 | 0.9684 |
| 95 | 1.0274 | 0.9995 | 0.9149 | 1.0153 |
| 100 | 0.9879 | 0.9653 | 0.9335 | 1.0293 |
| 105 |  | 0.9214 |  | 0.9945 |
| 110 |  | 0.8795 |  | 0.9349 |
| 115 |  | 0.8691 |  | 0.8984 |



Figure 2.1. Immediate annuitants: ratios of values of $q(x)$ for duration 0 to those for durations $1+$.


Figure 2.2. Immediate annuitants: ratios of values of $q(x)$ for amounts to those for lives.


Figure 2.3. Immediate annuitants: ratios of values of $q(x)$ for females to those for males.


Figure 2.4. Immediate annuitants: ratios of values of $q(x)$ for 1991-94 to projected rates IM80C92 and IF80C92.

### 3.1 Introduction

3.1.1 The investigation into the mortality of those self-employed who have purchased retirement annuities has grown to be a large one. Data is gathered for both males and females, and is subdivided into two sections, deferred and vested, which together form the combined section.
3.1.2 For both sexes there has been a large increase in the exposure since 1979-82, as seen in Table 3.1, especially in the vested section. The experience for males remains much larger than that for females.
3.1.3 The age ranges for the deferred and vested sections overlap substantially in the age range $60-75$, as would be expected, but there are many vested below 60 and a very few deferred above age 75 , as shown in Table 3.2. In both sections there are small amounts of implausible data at the very youngest ages, below 16 ; and it might be questioned whether the data in the vested section at ages a little above 16 is plausible.
3.1.4 It is immediately apparent from inspection of the crude rates of mortality that the mortality in the vested sections, for both sexes, is much higher than that in the deferred at the youngest ages (below age 60 ), and is still rather higher between ages 60 and 75 . Above 75 the combined sections are virtually all vested. The overall mortality of the combined section for both sexes has a 'normal' shape; the mortality for the deferred is almost the same as that for the combined at the youngest ages, and then runs below the combined; the mortality for the vested is well above that for the combined at the youngest ages, running into the combined at high ages. This is consistent with the self-employed choosing to retire when their health is impaired, and choosing to refrain from retiring while they are still healthy.
3.1.5 It is therefore appropriate to start by considering the graduation of the combined experiences, and then to fit the deferred and vested around the graduated rates for the combined. This procedure is based on an assumption that the combined experience contains deferred and vested in the correct proportions. If offices do not contribute correspondingly to both sections, or count lives differently in the two sections, or if many policyholders switch to non-contributing offices, or combine multiple policies in the deferred period to one policy at vesting, then these conditions may not be fulfilled. However, the mortality rates in the combined sections appear reasonably smooth, so it is at least plausible that the two sections are indeed appropriately weighted.

### 3.2 Males, combined

3.2.1 In total there are 37,456 deaths in the males combined experience. The very small amount of data below age 17 has been ignored.
3.2.2 The data is first compared with the graduated rates for permanent assurances, males, durations $2+$. The value of $100 \mathrm{~A} / \mathrm{E}$ is 99.7 , so the overall level is very similar. However, the detailed shape is different, so it is worth graduating the data separately.
3.2.3 A variety of different $\mathrm{GM}(r, s)$ formulae were used, including those with $r+s \leq 4$ and also GM(2,3). Details are given in Table 3.3. GM $(2,3)$ gives the best value of the log likelihood, and provides both a good fit and a reasonable shape, so the $\operatorname{GM}(2,3)$ formula has been adopted. The usual details are given in Tables 3.4 and 3.5, which also show the details for the deferred and vested sections.

### 3.3 Males, deferred

3.3.1 The males deferred section is considered next. When the experience is compared with the newly graduated rates for the combined section the resulting value of $100 \mathrm{~A} / \mathrm{E}$ is 91.6 . At low ages, where the deferred forms the vast bulk of the combined experience, the rates are necessarily very close. However, at rather higher ages, particularly from age 60 upwards, the deferred experience falls below the combined.
3.3.2 The rationale described in paragraph 3.1.4 suggests that the graduated rates for the deferred should be almost the same as those for the combined at the youngest ages, and in fact should be identical at age 17, should always lie below them, and should diverge further below at higher ages. The data effectively ceases at age 75 , so rates above that age are superfluous (although details are shown to indicate where the graduated rates would go).
3.3.3 A further consideration is that it is possible, given the mortality rates for the combined, deferred and vested sections, to derive implied "forces of retirement" from the deferred section to the vested section. These must not be negative.
3.3.4 The procedure used to meet these constraints (while still fitting the experience data adequately) was as follows: the value of $\mu_{x}$ at age 17 in the graduated rates for the combined section is 0.00040096 , and the rates for the deferred were adjusted so that the same value of $\mu_{x}$ at age 17 was obtained. This was done by "nudging" the rates by adjusting the value of $a_{1}$ in stages so that this equality was achieved. The value of $a_{2}$ was chosen to be the same as that for the combined section. Then, by a process of "trial and error", values for $b_{2}$ and $b_{3}$ were found so that the constraints could be met. Then the value of $b_{1}$ was found by optimising, subject to a repeated process of
nudging and optimising. The resulting rates satisfied the criteria, and appear satisfactory. The implied retirement rates are positive at all ages and also increase with increasing age. The usual statistics for the mortality rates are shown in Tables 3.4, 3.5 and 3.6.

### 3.4 Males, vested

3.4.1 When the experience for the vested section for males is compared with the graduated rates for the combined section, the resulting value of $100 \mathrm{~A} / \mathrm{E}$ is 108.5, but the observed rates are exceptionally high at ages below 60 . Since above age 75 the rates must in principle be the same as for the combined, the procedure described below was used.
3.4.2 At age 75 and above the values of $\mu_{x}$ (and hence of $q_{x}$ ) were taken as exactly the same as those for the combined section. The value of $\mu_{x}$ at age 75 on the combined rates was 0.038815 . Only the data up to and including that for age 75 was used for the graduation of the rates below that age. The parameters were adjusted step by step by trial and error so that the value of $\mu_{x}$ at age 75 was reproduced exactly, and a satisfactory fit to the data below that age was achieved. The procedure was similar to that for the deferred section, with the values of $a_{2}$, $b_{2}$ and $b_{3}$ being chosen so that the constraints could be met; the value of $b_{1}$ was used for "nudging" and the value of $a_{1}$ was found by optimising.
3.4.3 The resulting rates at young ages are exceptionally high. The value of $\chi^{2}$ is also high. A large part of this comes from the experience at ages 58 and 59 , where the actual deaths total 156 , as compared with an expected of 75.9. However, no smooth curve will accommodate such outliers. The minimum value of $q_{x}$ is at age 60.
3.4.4 As noted above, the rates above age 75 were taken as exactly the same as those of the combined experience. When these are applied to the data for the vested section at ages 75 and above, they fit quite satisfactorily. Details are not shown.

### 3.5 Females, combined

3.5.1 The experience for females combined is compared with the graduated rates for females permanent assurances durations $2+$, which shows $100 \mathrm{~A} / \mathrm{E}$ of 104.7. As with the males, the rates are broadly at the right level, but are not quite the right shape. The statistics for various $\operatorname{GM}(r, s)$ formulae are shown in Table 3.7. It is found that a GM $(1,2)$ formula (a Makeham formula) fits satisfactorily. The rates from this formula rise above those for the males combined section above age 99 ; which might be considered an objection.
3.5.2 The usual statistics are shown in Tables 3.8 and 3.9, along with those for females deferred and vested.

### 3.6 Females, deferred

3.6.1 When the experience for the females deferred section is compared with the females combined we get a value for $100 \mathrm{~A} / \mathrm{E}$ of 90.2 . As with the males, the experience rates run increasingly below those for the combined, but this starts at ages above about 50 . In order to keep the rates consistent with those for the combined section, a procedure similar to that for males was used. A $\operatorname{GM}(1,3)$ formula was chosen. By trial and error suitable values of $b_{2}$ and $b_{3}$ were found. The value of $a_{1}$ was used to nudge the value of $\mu_{17}$ for the deferred so that it exactly equalled the value of $\mu_{17}$ for the combined ( 0.00018668 ). Then the value of $b_{1}$ was optimised so as to give the best fit to the data subject to the constraints. The resulting rates provide a satisfactory fit to the data, and also produce implied retirement rates that are positive at all ages (and also increase with increasing age). The usual statistics are shown in Tables 3.8 and 3.9.

### 3.7 Females, vested

3.7.1 When the experience for females vested is compared with the females combined rates, the resulting value of $100 \mathrm{~A} / \mathrm{E}$ is 108.5 . The observed rates of mortality are, as for the males, extremely high at low ages. The same process as for males was adopted. From age 75 upwards the combined rates were used. Then the data for ages up to and including age 75 was analysed. A GM $(2,3)$ formula was used. The parameters were adjusted step by step so that the value of $\mu_{x}$ at age $75(0.024050)$ was reproduced exactly, and a satisfactory fit to the data below that age was achieved. Suitable values of $a_{2}, b_{2}$ and $b_{3}$ were chosen, the value of $b_{1}$ was nudged so that the desired value of $\mu_{75}$ was obtained and the value of $a_{1}$ was then optimised. The usual statistics are shown in Tables 3.8 and 3.9.

### 3.8 Comparisons, tables and figures

3.8.1 Various ratios are shown giving comparisons of the graduated rates for retirement annuitants among themselves and with other rates. Table 3.10 shows ratios of deferred to combined and vested to combined for both sexes. Table 3.11 shows ratios of females to males for each of the three sections. Table 3.12 shows ratios of the combined for each sex to the graduated rates for permanent assurances durations $2+$.
3.8.2 Figures 3.1 and 3.2 show the values of $q_{x}$ for each of the three sections, for males and females respectively. Figure 3.3 shows the ratios of the rates for females to those for males for each of the three sections. Figure 3.4 shows the ratios of the rates for the combined sections to those for the graduated tables for permanent assurances durations $2+$.
3.8.3 Values of the graduated rates of mortality, $q_{x}$, at each age for the vested section are shown in Appendix A in Table A8. Rates of mortality for the deferred and combined sections have also been graduated, but these are not designated as standard tables, and are shown in Appendix B in Tables B1 for males and B2 for females.

### 3.9 Age range of the table

3.9.1 For both males and females, the amounts of exposed to risk and deaths below age 40 are small. Although values of $q_{x}$ have been calculated down to age 17 it is doubtful if, in normal use, these tables should be used below age 40 .

Table 3.1. Retirement annuitants, males and females, deferred and vested: comparison of (central) exposed to risk and deaths for 1991-94 and for 1979-82.

|  | 197982 | 1991-94 |
| :---: | :---: | :---: |
| Males |  |  |
| Deferred |  |  |
| Central exposed | 2,997,376.5 | 5,335,263.0 |
| Deaths | 12,328 | 17,256 |
| Vested |  |  |
| Central exposed | 221,898.0 | 630,827.0 |
| Deaths | 8,811 | 20,200 |
| Combined |  |  |
| Central exposed | 3,219,274.5 | 5,966,090.0 |
| Deaths | 21,139 | 37,456 |
| Females |  |  |
| Defcrred |  |  |
| Central exposed | 338,758.0 | 963,447.2 |
| Deaths | 860 | 1,958 |
| Vested |  |  |
| Central exposed | 35,006.0 | 149,663.1 |
| Deaths | 692 | 2,695 |
| Combined |  |  |
| Central exposed | 473,764.0 | 1,113,110.3 |
| Deaths | 1,552 | 4,653 |

Table 3.5. Retirement annuitants, males, combined, deferred and vested: statistics for graduations of $\mu_{x}=\mathrm{GM}(r, s)$.

| Section | Combincd | Deferred | Vested |
| :---: | :---: | :---: | :---: |
| Formula | GM(2,3) | GM $(2,3)$ | GM $(2,3)$ |
| Ages used | 17.108 | 17-106 | 21-75 ${ }^{\text { }}$ |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | 0.016470 | 0.017467 | 0.023761 |
| $T$-ratio | 0.50 |  | 1.13 |
| $100 a_{2}$ | -0.020219 | -0.020219 | -5.0 |
| $T$-ratio | -0.53 |  |  |
| $b_{1}$ | -4.399325 | -5.229113 | --4.713208 |
| $T$-ratio | -35.18 | -634.9 |  |
| $b_{2}$ | 5.219980 | 4.3 | 6.0 |
| T-ratio | 62.12 |  |  |
| $b_{3}$ | -0.637407 | -1.24 | -1.0 |
| $T$-ratio | -5.48 |  |  |
| Sign test: $p$ (pos) | 0.5889 | 0.6061 | 0.8447 |
| Runs test: $p$ (runs) | 0.2485 | 0.0512 | 0.1014 |
| $\mathrm{K}-\mathrm{S}$ test: $p(K S)$ | 0.9967 | 0.0533 | 0.0806 |
| Serial correlation test: |  |  |  |
| $T$-ratio 1 | 0.73 | 1.66 | 2.90 |
| $T$-ratio 2 | -0.67 | 1.27 | 1.29 |
| T-ratio 3 | 1.46 | 1.88 | 0.65 |
| $\chi^{2}$ test: |  |  |  |
| $\chi^{2}$ | 118.60 | 64.96 | 162.61 |
| Degrees of freedom | 74 | 50 | 30 |
| $p\left(\chi^{2}\right)$ | 0.000773 | 0.0759 | 0.0000 |

*Note that the graduated rates for vested above age 75 are assumed to be the same as the graduated rates for combined.

Table 3.2. Retirement annuitants: age ranges.

|  | Range of data | Exposed $\geq 100$ | Deaths $\geq 10$ |
| :--- | :---: | :---: | :---: |
| Males |  |  |  |
| $\quad$ Deferred | $10-106$ | $19-76$ | $25-75$ |
| Vested | $10--108$ | $42-94$ | $50-95$ |
| Combined | $10-108$ | $19-94$ | $25-95$ |
|  |  |  |  |
| Females | $17-103$ | $20-75$ | 3970 |
| $\quad$ Deferred | $10-104$ | $50--91$ | $59-92$ |
| Vested | $10-104$ | $20-91$ | $39-92$ |

Table 3.3. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of retirement annuitants, males, combined.

|  | GM(0,2) | $\mathrm{GM}(0,3)$ | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0,4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ | $\mathrm{GM}(2,3)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $196,533.2$ | $196,521.0$ | $196,531.2$ | $196,471.6$ | $196,465.5$ | $196,478.2$ | $196,465.4$ |
| $\chi^{2}$ | 258.7 | 256.1 | 245.9 | 141.0 | 119.6 | 140.1 | 118.6 |

Table 3.4. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of retirement annuitants, males, deferred.

|  | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) | GM(2,3) <br> (unconstrained) | GM(2,3) <br> (as fitted) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $109,368.8$ | $109,364.0$ | $109,345.5$ | $109,342.3$ | $109,342.4$ | $109,341.1$ | $109,349.1$ |
| $\chi^{2}$ | 112.9 | 95.0 | 61.2 | 53.4 | 50.6 | 48.7 | 65.0 |

Table 3.6. Retirement annuitants, males, combined, deferred and vested: specimen values of $q_{x}$ and percentage standard errors.

| Section <br> Formula | Combincd <br> GM(2,3) | Deferred <br> GM(2,3) | Vested <br> GM |
| :--- | ---: | ---: | ---: |
| Age 20 | 0.000403 | 0.000398 | 0.048530 |
| Percentage s.e. | 17.92 | 0.05 | 0.41 |
| Age 30 | 0.000494 | 0.000466 | 0.039017 |
| Percentage s.e. | 5.31 | 0.23 | 0.52 |
| Age 40 | 0.000969 | 0.000912 | 0.029643 |
| Percentage s.e. | 2.43 | 0.54 | 0.69 |
| Age 50 | 0.002741 | 0.002625 | 0.021244 |
| Percentage s.c. | 1.04 | 0.72 | 0.97 |
| Age 60 | 0.008404 | 0.007665 | 0.016859 |
| Percentage s.c. | 0.69 | 0.77 | 1.23 |
| Age 70 | 0.024353 | 0.019313 | 0.025328 |
| Percentage s.e. | 0.74 | 0.78 | 0.81 |
| Age 80 | 0.063767 | 0.040250 | 0.066228 |
| Percentage s.e. | 0.85 | 0.78 | 0.30 |
| Age 90 | 0.147491 | 0.068761 | 0.167927 |
| Percentage s.e. | 2.15 | 0.77 | 0.10 |
| Age 100 | 0.294827 | 0.096448 | 0.343149 |
| Percentage s.e. | 4.81 | 0.76 | 0.04 |
| Age 110 | 0.498787 | 0.111707 | 0.553955 |
| Percentage s.e. | 7.62 | 0.75 | 0.02 |

Table 3.7. Values of log likelihood and of $\chi^{2}$ for graduation of retirement annuitants, females, combined.

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Log likelihood | $26,416.7$ | $26,408.5$ | $26,413.2$ | $26,408.5$ | $26,407.8$ | $26,408.1$ |
| $\chi^{2}$ | 121.8 | 107.3 | 114.6 | 107.0 | 103.7 | 104.5 |

52 Standard Tables of Mortality based on the 1991-94 Experiences
Table 3.8. Retirement annuitants, females, combined, deferred and vested: statistics for graduations of $\mu_{x}=\mathrm{GM}(r, s)$.

| Section | Combined | Deferred | Vested |
| :---: | :---: | :---: | :---: |
| Formula | GM(1,2) | GM(1,3) | GM(2,3) |
| Ages used | 17-104 | 17-103 | 27-75* |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | 0.014000 | 0.016746 | -0.943368 |
| $T$-ratio | 2.74 |  | -31.4 |
| $100 a_{2}$ |  |  | -5.0 |
| T-ratio |  |  |  |
| $b_{1}$ | -4.271279 | -5.497232 | -4.737523 |
| $T$-ratio | -244.21 | -224.0 |  |
| $\mathrm{H}_{2}$ | 5.378175 | 4.0 | 5.0 |
| $T$-ratio | 62.13 |  |  |
| $b_{3}$ |  | -0.9 | -1.0 |
| T-ratio |  |  |  |
| Sign test: $p$ (pos) | 0.5482 | 0.6170 | 0.7976 |
| Runs test: $p$ (runs) | 0.0890 | 0.1135 | 0.5351 |
| K--S test: $p(K S)$ | 0.3866 | 0.0400 | 0.2256 |

Serial correlation test:

| $T$-ratio 1 | 1.44 | 2.11 | 1.22 |
| :--- | ---: | ---: | ---: |
| $T$-ratio 2 | 1.42 | 1.34 | 0.65 |
| $T$-ratio 3 | 1.73 | 2.44 | 0.53 |
| $\chi^{2}$ test: |  |  |  |
| $\chi^{2}$ | 114.57 | 55.21 | 79.35 |
| Degrees of freedom | 65 | 40 | 18 |
| $p\left(\chi^{2}\right.$ ! | 0.000146 | 0.0553 | 0.0000 |

[^1]Table 3.9. Retirement annuitants, females, combined, deferred and vested: specimen values of $q_{x}$ and percentage standard errors.

| Section <br> Formula | Combined <br> GM(1,2) | $\left.\begin{array}{r}\text { Deferred } \\ \text { GM }\end{array} 1,3\right)$ |
| :--- | ---: | ---: | ---: |

Table 3.10. Retirement annuitants: ratios of values of $q_{x}$ in proposed tables: comparison of sections.

|  | Males |  |  | Females |  |
| :---: | :---: | ---: | :---: | :---: | :---: |
| Age | Deferred/Combined | Vested/Combined |  | Deferred/Combined |  |
|  |  | Vested/Combined |  |  |  |
| 20 | 0.9876 | 120.4218 | 0.9615 | 188.9279 |  |
| 25 | 0.9648 | 102.7582 | 0.9219 | 134.8203 |  |
| 30 | 0.9433 | 78.9818 | 0.9056 | 87.7493 |  |
| 35 | 0.9338 | 52.7477 | 0.9087 | 51.9481 |  |
| 40 | 0.9412 | 30.5913 | 0.9324 | 28.2428 |  |
| 45 | 0.9534 | 15.8672 | 0.9544 | 14.2033 |  |
| 50 | 0.9577 | 7.7505 | 0.9590 | 6.7399 |  |
| 55 | 0.9449 | 3.7829 | 0.9375 | 3.1712 |  |
| 60 | 0.9121 | 2.0061 | 0.8877 | 1.6685 |  |
| 65 | 0.8603 | 1.2835 | 0.8121 | 1.1435 |  |
| 70 | 0.7930 | 1.0400 | 0.7177 | 1.0148 |  |
| 75 | 0.7149 | 1 | 0.6130 | 1 |  |
| 80 | 0.6312 | 1 | 0.5067 | 1 |  |
| 85 | 0.5469 | 1 | 0.4063 | 1 |  |
| 90 | 0.4662 | 1 | 0.3175 | 1 |  |
| 95 | 0.3923 | 1 | 0.2436 | 1 |  |
| 100 | 0.3271 | 1 | 0.1856 | 1 |  |
| 105 | 0.2711 | 1 | 0.1432 | 1 |  |
| 110 | 0.2240 | 1 | 0.1147 | 1 |  |
| 115 | 0.1845 |  |  |  |  |

Table 3.11. Retirement annuitants: ratios of values of $q_{x}$ in proposed tables: comparison of sexes.

| Age | Combined <br> Females/Males | Deferred <br> Females/Males | Vested <br> Females/Males |
| :---: | :---: | :---: | :---: |
| 20 | 0.5161 | 0.5025 | 0.8097 |
| 25 | 0.6009 | 0.5742 | 0.7884 |
| 30 | 0.6862 | 0.6588 | 0.7624 |
| 35 | 0.7415 | 0.7216 | 0.7303 |
| 40 | 0.7482 | 0.7412 | 0.6908 |
| 45 | 0.7181 | 0.7188 | 0.6428 |
| 50 | 0.6760 | 0.6770 | 0.5879 |
| 55 | 0.6388 | 0.6338 | 0.5355 |
| 60 | 0.6134 | 0.5970 | 0.5102 |
| 65 | 0.6024 | 0.5686 | 0.5367 |
| 70 | 0.6066 | 0.5490 | 0.5918 |
| 75 | 0.6271 | 0.5376 | 0.6271 |
| 80 | 0.6655 | 0.5342 | 0.6655 |
| 85 | 0.7246 | 0.5384 | 0.7246 |
| 90 | 0.8078 | 0.5502 | 0.8078 |
| 95 | 0.9177 | 0.5697 | 0.9177 |
| 100 | 1.0529 | 0.5975 | 1.0529 |
| 105 | 1.2011 | 0.6343 | 1.2011 |
| 110 | 1.3304 | 0.6814 | 1.3304 |
| 115 | 1.3915 | 0.7407 | 1.3915 |

Table 3.12. Ratios of values of $q_{x}$ for retirement annuitants, combined to those for permanent assurances, durations $2+$.

| Age | Males | Females |
| :---: | :---: | :---: |
| 20 | 0.6924 | 1.0777 |
| 25 | 0.7527 | 1.0323 |
| 30 | 0.8373 | 1.0000 |
| 35 | 0.9434 | 0.9797 |
| 40 | 1.0342 | 0.9705 |
| 45 | 1.0846 | 0.9719 |
| 50 | 1.0929 | 0.9825 |
| 55 | 1.0761 | 0.9990 |
| 60 | 1.0476 | 1.0192 |
| 65 | 1.0154 | 1.0415 |
| 70 | 0.9826 | 1.0653 |
| 75 | 0.9506 | 1.0900 |
| 80 | 0.9201 | 1.1150 |
| 85 | 0.8917 | 1.1394 |
| 90 | 0.8663 | 1.1620 |
| 95 | 0.8451 | 1.1805 |
| 100 | 0.8293 | 1.1910 |
| 105 | 0.8206 | 1.1881 |
| 110 | 0.8205 | 1.1649 |
| 115 | 0.8297 | 1.1184 |



Figure 3.1. Retirement annuitants: values of $q(x)$ for males.


Figure 3.2. Retirement annuitants: values of $q(x)$ for females.


Figure 3.3. Retirement annuitants: ratios of $q(x)$ for females to those for males.


Figure 3.4. Retirement annuitants: ratios of $q(x)$ for combined to those for permanent assurances durations $2+$.

### 4.1 Introduction

4.1.1 The investigation into the mortality of widows of life office pensioners is one of the newer investigations carried out by the Bureau, and the experience in 1979-82 was quite small. It has increased many-fold since then, as shown in Table 4.1. There were previously only 692 deaths; this number has increased to 5,452 . By contrast, the experience for widowers is still small and has not been considered in this report.
4.1.2 The bulk of the exposure is at older ages, as can be seen from Table 4.2, though there are also some younger widows. Those recorded at ages as young as 10 are presumably errors (possibly orphans), and only the data from age 17 has been used in the graduation process.
4.1.3 The investigation is carried out on an amounts basis as well as on a lives one. The average amounts per life in 1979-82 and 1991-94 are shown in Table 4.3. It can be seen that the average amount has almost doubled. This is a smaller proportionate increase than for pensioners. However, the experience was new in 1979-82, and therefore the average duration since the commencement of pension (though not necessarily the age of the pensioners) was less than in the pensioners investigation, but the impact of this on the experience is not clear.

### 4.2 Comparison with pensioners' rates

4.2.1 The experiences have been compared with the new female pensioners 1991-94 mortality rates. The results are as shown below. Although the overall level of the lives experience is similar to that of the pensioners, the overall shape is very different, with much higher rates for widows at lower ages as compared with the pensioners. The amounts experience is even further adrift.

|  | 100A/E |
| :--- | :---: |
| Lives | 101.9 |
| Amounts | 109.2 |

4.2.2 The experiences have also been compared with the mortality rates projected for 1992 (the centre of the 1991-94 experience quadrennium) from the tables for widows in the " 80 " Series tables, i.e. WL80C92 and WA80C92, using central exposures and $q_{x}$. The results are as shown below. Although the overall levels of both experiences are lower than projected, the rates for younger ages are generally higher, and a simple adjustment to the C92 rates would not be satisfactory.

|  | $100 \mathrm{~A} / \mathrm{E}$ |
| :--- | :---: |
| Lives | 88.0 |
| Amounts | 91.0 |

### 4.3 The lives experience

4.3.1 The values of the $\log$ likelihood and $\chi^{2}$ for various $\mathrm{GM}(r, s)$ graduations are shown in Table 4.4. There is very little difference between all the formulae on this criterion. All are a tolerable shape, and all have some poor features when the fit is considered. All produce relatively high values of $q_{x}$ in the 40 s and 50 s of age, running down to rather low values by age 20 . This reflects the experience, but it would be more compatible with other tables for the values of $q_{x}$ to be rather higher at age 20 . This can be achieved by choosing the $\mathrm{GM}(1,2)$ formula with a value of $100 a_{1}$ of 0.02 , and then optimising for $b_{1}$ and $b_{2}$. The results in the usual form are shown in Tables 4.6 and 4.7.

### 4.4 The amounts experience

4.4.1 Before graduating the amounts experience the numbers of exposed to risk and actual deaths were divided by the average amount per life in the exposed to risk, $£ 1,104.10$, as noted in Table 4.3. The experience was then compared with the newly graduated rates for widows lives. The resulting value of $100 \mathrm{~A} / \mathrm{E}$ of 89.6 justified a separate graduation of the amounts experience.
4.4.2 The statistics for the widows amounts graduations by different $\mathrm{GM}(r, s)$ formulae are shown in Table 4.5. As with the lives, there is very little difference between them. All provide a reasonable fit (bearing in mind the assumption that the value of $\chi^{2}$ is expected to be high), but the $\operatorname{GM}(1,3)$ and $\operatorname{GM}(2,2)$ formulae do not give a comfortable shape of curve. It was therefore decided to use the same method as for lives, i.e. to adopt a $\mathrm{GM}(1,2)$ formula, with a value of $100 a_{1}$ of 0.018 , and then optimising for $b_{1}$ and $b_{2}$. The results are shown in Tables 4.6 and 4.7.

### 4.5 Comparisons, tables and figures

4.5.1 Comparisons of the graduated amounts rates with the rates for lives, comparisons both with the projected rates for calendar year 1992 based on the WL80 and WA80 tables (WL80C92 and WA80C92), and comparisons of both with the graduated rates for female pensioners for 1991-94 are shown in Table 4.8.
4.5.2 The graduated rates for ages from 17 to 120 are shown in Appendix A in Table A9.
4.5.3 Graphs of the graduated rates are shown in Figure 4.1. The ratios of amounts to lives rates, for the widows and for male and female pensioners,
are shown in Figure 4.2. Graphs of the rates as compared with the graduated rates for the projected rates WL80C92 and WA80C92 are shown in Figure 4.3, and ratios of the rates for widows to those for female pensioners are shown in Figure 4.4. The very high ratios of the rates for widows to those for female pensioners in the 30 s and 40 s of age suggest that possibly the graduated rates for the latter are too low.

Table 4.1. Widows, lives and amounts: comparison of (central) exposed to risk and deaths for 1991-94 and for 1979-82.

|  | $1979-82$ | $1991-94$ |
| :--- | :---: | :---: |
| Lives |  |  |
| $\quad$ Central exposed | $28,386.5$ | $162,237.1$ |
| Deaths | 692 | 5,452 |
| Amounts $£$ |  |  |
| Central exposed | $15,892,759.0$ | $179,126,584.6$ |
| Deaths | 238,438 | $4,279,423$ |

Table 4.2. Widows: age ranges.

| Range of data <br> $10-108$ | Exposed $\geq 100$ | Deaths $\geq 10$ |
| :---: | :---: | :---: |
|  | $33-95$ | $56-97$ |

62 Standard Tables of Mortality based on the 1991-94 Experiences
Table 4.3. Widows: average amounts per life, 1979-82 and 1991-94.

| $1979-82$ | $1991-94$ |
| :---: | :---: |
| $£ 559.87$ | $£ 1,104.10$ |

Table 4.4. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of widows lives.

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $22,380.0$ | $22,379.8$ | $22,380.0$ | $22,379.2$ | $22,379.7$ | $22,379.9$ |
| $\chi^{2}$ | 43.8 | 42.9 | 43.9 | 41.2 | 43.4 | 43.4 |

Table 4.5. Values of $\log$ likelihood and of $\chi^{2}$ for graduation of widows amounts.

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $16,999.3$ | $16,998.7$ | $16,999.3$ | $16,997.3$ | $16,996.9$ | $16,998.1$ |
| $\chi^{2}$ | 182.7 | 175.2 | 182.8 | 171.7 | 170.4 | 171.9 |

Table 4.6. Widows, lives and amounts: statistics for graduations

$$
\text { of } \mu_{x}=\mathrm{GM}(1,2) .
$$

| Formula | Lives <br> GM(1,2) | Amounts <br> $G M(1,2)$ |
| :--- | ---: | ---: |
| Values of parameters at optimum point: |  |  |
| $100 a_{1}$ | 0.02 | 0.018 |
| $T$-ratio | -3.795522 | -3.921243 |
| $b_{1}$ | -201.02 | -198.72 |
| $T$-ratio | 4.308854 | 4.444249 |
| $b_{2}$ | 52.93 | 47.75 |
| $T$-ratio | 0.7084 | 0.5000 |
|  | 0.1374 | 0.9873 |
| Sign test: $p$ (pos) | 0.9953 | 0.5214 |
| Runs test: $p$ (runs) |  |  |
| K-S test: $p(K S)$ | 2.13 | -0.16 |
|  | 0.07 | -0.46 |
| Serial correlation test: | -1.03 | -1.29 |
| $T$-ratio 1 |  |  |
| $T$-ratio 2 | 43.9 | 181.44 |
| $T$-ratio 3 | 48 | 48 |
| $\chi^{2}$ test: |  |  |
| $\chi^{2}$ | 0.6408 | 0.0000 |
| Degrees of |  |  |
| freedom |  |  |
| $p\left(\chi^{2}\right)$ |  |  |

Table 4.7. Widows, lives and amounts: specimen values of $q_{x}$ and percentage standard errors.

|  | Lives <br> GM $(1,2)$ | Amounts <br> GM(1,2) |
| :--- | ---: | ---: |
| Formula | 0.000515 | 0.000423 |
| Age 20 | 1.12 | 1.10 |
| Percentage s.e. | 0.000947 | 0.000772 |
| Age 30 | 1.44 | 1.46 |
| Percentage s.e. | 0.001967 | 0.001619 |
| Age 40 | 1.64 | 1.69 |
| Percentage s.e. | 0.004378 | 0.003676 |
| Age 50 | 1.74 | 1.81 |
| Percentage s.e. | 0.010062 | 0.008662 |
| Age 60 | 1.78 | 1.86 |
| Percentage s.e. | 0.023390 | 0.020687 |
| Age 70 | 1.79 | 1.87 |
| Percentage s.e. | 0.054231 | 0.049329 |
| Age 80 | 1.77 | 1.85 |
| Percentage s.e. | 0.123415 | 0.115548 |
| Age 90 | 1.71 | 1.79 |
| Percentage s.e. | 0.267695 | 0.257996 |
| Age 100 | 1.55 | 1.64 |
| Percentage s.e. | 0.521593 | 0.515942 |
| Age 110 | 1.23 | 1.30 |
| Percentage s.e. |  |  |

Table 4.8. Widows: ratios of values of $q_{x}$ in proposed tables: comparison of lives and amounts (1991-94); comparison of both lives and amounts with the projected rates for 1992, WL80C92 and WA80C92; and comparison of both lives and amounts with the graduated rates for female pensioners 1991-94.

| Age | Amounts/Lives <br> $1991-94$ | $1991-94 ;$ <br> WL80C92 <br> Lives | 1991-94/ <br> WA80C92 <br> Amounts | 1991-94 <br> Widows <br> Pensioners <br> Lives | 1991-94 <br> Widows/ <br> Pensioners <br> Amounts |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 20 | 0.8214 | 1.0553 | 1.2405 | 2.0276 | 2.0837 |
| 25 | 0.8161 | 1.4153 | 1.6538 | 2.7075 | 2.7537 |
| 30 | 0.8152 | 1.5423 | 1.7995 | 3.5074 | 3.5576 |
| 35 | 0.8176 | 1.4537 | 1.7048 | 4.0879 | 4.1939 |
| 40 | 0.8231 | 1.3336 | 1.5734 | 4.0809 | 4.2382 |
| 45 | 0.8306 | 1.2119 | 1.4423 | 3.5187 | 3.6601 |
| 50 | 0.8397 | 1.1086 | 1.3328 | 2.8010 | 2.8922 |
| 55 | 0.8499 | 1.0291 | 1.2438 | 2.1815 | 2.2294 |
| 60 | 0.8609 | 1.0179 | 1.1680 | 1.7203 | 1.7446 |
| 65 | 0.8724 | 0.9844 | 1.0728 | 1.3969 | 1.4118 |
| 70 | 0.8844 | 0.9554 | 0.9892 | 1.1763 | 1.1909 |
| 75 | 0.8968 | 0.9298 | 0.9153 | 1.0303 | 1.0508 |
| 80 | 0.9096 | 0.9072 | 0.8501 | 0.9400 | 0.9710 |
| 85 | 0.9228 | 0.8137 | 0.7932 | 0.8936 | 0.9397 |
| 90 | 0.9363 | 0.7725 | 0.7539 | 0.8846 | 0.9512 |
| 95 | 0.9500 | 0.7973 | 0.8428 | 0.9097 | 1.0039 |
| 100 | 0.9638 | 0.9062 | 0.9673 | 0.9672 | 1.0979 |
| 105 | 0.9771 | 0.9993 | 1.0749 | 1.0542 | 1.2317 |
| 110 | 0.9892 | 1.0648 | 1.1496 | 1.1624 | 1.3967 |
| 115 | 0.9986 | 1.1145 | 1.2010 | 1.2739 | 1.5702 |



Figure 4.1. Values of $q(x)$ for widows (lives and amounts) and female pensioners.


Figure 4.2. Ratios of values of $q(x)$ for amounts and lives tables for widows and pensioners.


Figure 4.3. Ratios of values of $q(x)$ for widows 1991-94 to those for WL80C92 and WA80C92.


Figure 4.4. Ratios of values of $q(x)$ for widows 1991-94 to those for female pensioners.

### 5.1 Introduction

5.1.1 From 1971 to 1987 the Bureau carried out separate investigations into the mortality experience of level temporary and decreasing temporary assurances, but after the 1979-82 graduation these were combined into a single investigation for all temporary assurances. The investigation was originally only for male lives, and in 1979-82 the number of female lives included in the investigation was very small. The number of female lives is now significant and fully justifies a separate investigation. In 1979-82 the Committee aggregated the data for durations 1 to 4 (durations $1-4$ ) and kept duration 0 and durations 5 and over (durations $5+$ ) separate. The same has been done on this occasion.
5.1.2 Table 5.1 shows, for 1991-94, the central exposed to risk and the number of deaths for all ages for males and females, for durations $0,1-4$ and $5+$, and, for comparison, the corresponding numbers for males for 1979-82. For males the exposure at durations 0 and $1-4$ has fallen, though it has risen at durations $5+$. The net effect is an overall fall in the exposure, though the number of deaths has risen. It may be that this is a feature of the data supplied by the contributing offices, rather than of the term assurance market as a whole. The size of the experience for females has now grown to a substantial fraction of that for males. 5.1.3 The data cover a wide range of adult ages, though with the peak exposure in the thirties and forties of age. For both sexes the normal data runs out at about age 90 . However, for both sexes there are a few cases recorded, at all durations, from age 100 to 105 . These may well be errors, and in any event they have been ignored in this graduation. The age range of the data (excluding the data at ages over 99 ), the continuous age range within which the central exposed to risk is greater than or equal to 100 , and the continuous age range within which the number of deaths is greater than or equal to 10 are shown in Table 5.2. There are also isolated ages where the number of deaths exceeds nine. These are not noted.
5.1.4 As for other investigations, it seems as if the data at age 10 is faulty, and there is very little exposed to risk between ages 11 and 16 , so the graduations have been started at age 17. Otherwise the full age ranges as shown in Table 5.2 have been used.

### 5.2 Comparison with TM80 and with assured lives rates

5.2.1 The experience for males has been compared with the TM80 table (which was based on the experience for 1979-82, and has rates for durations $0,1-4$ and $5+$ ), and the experience for both sexes has been compared with the graduated rates for permanent assurances for 1991-94 (comparing the experience for
durations 1-4 with the graduated rates for duration 1 , and the experience for durations $5+$ with the graduated rates for durations $2+$ ). The results are as shown in Table 5.3.
5.2.2 It is clear that for males the mortality rates have reduced considerably since the 1979-82 experience, the period on the basis of which the TM80 rates were constructed. Inspection of the comparisons with the graduated rates for permanent assurances for 1991-94 shows that:
(i) whereas the rates for the ultimate durations (D2+ for permanent assurances) might be considered as fitting the experience for temporary assurances (D5+) for both sexes (100A/E $=97$ and 96),
(ii) the experiences for duration 0 for temporary assurances, for both sexes, are distinctly lower than those for the permanent assurances, and
(iii) the experience for durations $1-4$ is conflicting, with $100 \mathrm{~A} / \mathrm{E}=88$ for males and $=120$ for females, which may reflect more stringent underwriting criteria.

### 5.3 The experience for males

5.3.1 The experience for males, durations $5+$, is considered first. The values of the $\log$ likelihood and $\chi^{2}$ for selected $\operatorname{GM}(r, s)$ graduations are shown in Table 5.4. The results for all these formulae except $\operatorname{GM}(0,3)$ are fairly similar. The $\mathrm{GM}(2,3)$ formula has rather the best shape, but the value of $a_{2}$ is positive, so the values of $q_{x}$ do not decrease in the early twenties of age as one might expect. A decision about which formula to use was therefore postponed until the other durations had been considered.
5.3.2 The values of the log likelihood and $\chi^{2}$ for selected $\operatorname{GM}(r, s)$ graduations for males, durations $1-4$, are shown in Table 5.5. There is little to choose in terms of the $\log$ likelihood between the higher-order formulae, but $\operatorname{GM}(2,3)$ has the best overall shape. For example, although GM(0,4) has the 'best' values of the statistics shown, the values of $q_{x}$ are quite unsuitable at higher ages, reaching a peak at too low a level of 0.042185 at age 87 and then falling sharply.
5.3.3 The values of the log likelihood and $\chi^{2}$ for selected $\operatorname{GM}(r, s)$ graduations for males, duration 0, are shown in Table 5.6. Again, there is little to choose between the higher-order formulae, but for consistency with the other durations a $\operatorname{GM}(2,3)$ formula was considered further.
5.3.4 When the graduated rates for the three durations, using the GM $(2,3)$ formulae for each, are compared, one can see uncomfortable features. These are features of the graduation rather than the data. The values of $q_{x}$ cross over at extreme ages; the rates for durations $1-4$ are higher than those for
durations $5+$ below age 35 ; the rates for duration 0 are higher than those for durations $5+$ below age 23 and higher than those for durations 1-4 above age 87 . These are extreme ages, but there is no reason to suppose that these cross-overs represent reality any better than the assumption that the rates lie conformably with one another.
5.3.5 The experience at young ages needs to be considered in detail. Table 5.7 shows a summary of the experience for ages 17-24 and 25-29 at all three duration groups. One can see that, although for all ages combined the exposed to risk for durations $5+$ is larger in total than that for durations 1-4 (see Table 5.1), at young ages the experience for durations $5+$ is relatively small, so the experience for durations 0 and $1-4$ is more 'credible'. For these lower durations the crude death rates for ages 25-29 are smaller than those for ages 17-24. This indicates that falling mortality rates in the early twenties of age, as are found in many other investigations, are more plausible than rising mortality rates.
5.3.6 The graduation for durations $1-4$ was therefore chosen as the 'datum', and those for durations 0 and $5+$ were 'arranged' around it, by choosing suitable values of certain parameters, and then optimising for the others. The parameters found before modification and after modification in this way are shown in Table 5.8. For both durations the values of $100 a_{1}, 100 a_{2}$ and $b_{3}$ were fixed, and the values of $b_{1}$ and $b_{2}$ were optimised. This brought the values of $q_{x}$ at low ages and high ages for all three durations into conformity. The high value of $\chi^{2}$ for durations $1-4$ is unavoidable, since there are high values of $z$ $(=(\mathrm{A}-\mathrm{E}) / \sqrt{ } \mathrm{E})$ at ages 35 and 56 (negative) and 36, 63, 67 and 68 (positive), which no smooth graduation can get rid of.
5.3.7 Details of the usual statistics for the three graduations are shown in Table 5.9 and values of $q_{x}$ at decennial ages are shown in Table 5.10.

### 5.4 The experience for females

5.4.1 We now consider the experience for females, looking first at durations
$5+$. The values of the log likelihood and of $\chi^{2}$ for various $\operatorname{GM}(r, s)$ formulae are shown in Table 5.11. There is not a great deal of difference between the different formulae, but the $\mathrm{GM}(1,2)$ graduation provides a satisfactory shape, and satisfactory statistics. However, it produces rather low rates at low ages, as discussed below.
5.4.2 The values of the $\log$ likelihood and of $\chi^{2}$ for various $\mathrm{GM}(r, s)$ formulae for durations 1-4 are shown in Table 5.12. Excluding GM(0,2), the remainder show rather similar statistics, and again $\mathrm{GM}(1,2)$ provides a satisfactory shape, although the rates at low ages are also rather low. The values of $q_{x}$ for this formula are generally around $90 \%$ of those for durations $2+$, so this formula
gives a good starting point. However, the value of $\chi^{2}$ is rather high. This can be attributed to two ages, 50 and 65 , for which the actual deaths are high relative to those expected, and also relative to the actual deaths at neighbouring ages. It is possible that these can be accounted for by the presence of duplicates, but this has not been investigated.
5.4.3 The values of the log likelihood and of $\chi^{2}$ for various $\mathrm{GM}(r, s)$ formulae for duration 0 are shown in Table 5.13. There were only 86 deaths and all formulae from GM $(0,2)$ upwards show similar results. This time the shape produced by a GM $(1,2)$ formula is not satisfactory; the value of $a_{1}$ is negative, and this makes the values of $q_{x}$ unreasonably low at lower ages. The values of $\chi^{2}$ are remarkably low; although the sparse data fall into rather few cells, there are still usually 14 cells; however, typically none of the values of $z$ exceeds 1 ; the experience conforms almost too closely with the graduation; this is unusual.
5.4.4 Although the GM $(1,2)$ graduations for durations $5+$ and $1-4$ would have seemed satisfactory if they had stood alone, the values of $q_{x}$ at low ages were low relative to those for the males, the female permanent assurances, and ELT 15 females. Whilst this may be a genuine feature of the data, the number of deaths is small and it seemed more satisfactory to adjust the parameters to increase the rates. This was done by choosing values of $100 a_{1}$ as shown in Table 5.14, and then optimising the values of the other two parameters. Duration 0 was treated similarly.
5.4.5 Values of the parameters and selected statistics for the original and the modified graduations for the three durations for females are shown in Table 5.14 .
5.4.6 The usual graduation statistics and specimen values of $q_{x}$ for all three durations are shown in Tables 5.15 and 5.16.

### 5.5 Comparisons

5.5.1 Comparisons of the values of $q_{x}$ in various ways are shown in Tables 5.17 to 5.20 and Figures 5.1 to 5.4 .
5.5.2 Table 5.17 and Figure 5.1 show the ratios of the rates for successive durations. Note that the rates "conform" in the sense that the rates for duration 0 are below those for durations $1-4$, which in turn are below those for durations $5+$, for both sexes. However, the ratios are not constant. The relatively low level of duration 0 for females might suggest that selection is more effective for them than for males, but the data on which this observation is based are very few. It is also possible that this feature is a reflection of fewer suicides and accidental deaths or, perhaps, more mortgage related cases.
5.5.3 Table 5.18 and Figure 5.2 show the ratios of the rates for females to those for males. Note that the rates for females duration 0 are noticeably low
compared with those for males. Note also that the rates for females are well below those for males at young ages. Had the rates for females not been increased, as described in 5.4.4, this effect would have been even more pronounced.
5.5.4 Table 5.19 and Figure 5.3 show the ratios of the graduated rates for temporary assurances for 1991-94 to those for permanent assurances, as discussed in section 1 of this report (comparing durations $1-4$ and $5+$ for the former with durations 1 and $2+$ for the latter). Note that, in general, the rates for temporary assurances are a little below those for permanent assurances, except at young ages (say below age 40) and at high ages (say above 70). However, the rates for temporary assurances females duration 0 are relatively low throughout, and those for females durations 1-4 are relatively high.
5.5.5 Table 5.20 and Figure 5.4 show the ratios of the graduated rates for 1991-94 for temporary assurances, for males only, to the rates in the TM80 tables for temporary assurances based on the 1979-82 experience. Note that the rates for 1991-94 are higher than those based on the data for 1979-82 at younger ages (say below 35 ) and at higher ages (say above 65 ), but are lower in between these ages.

### 5.6 The proposed tables

5.6.1 Values of $q_{x}$ for the proposed tables for temporary assurances are shown in Appendix A in Table A3 for males and in Table A4 for females. All six sets of rates are graphed in Figure 5.5.

Table 5.1. Temporary assurances, males and females: comparison of (central) exposed to risk and deaths for 1991-94 and for 1979-82, durations 0, 1~4 and 5+.

|  | 197982 | 1991-94 |
| :---: | :---: | :---: |
| Males |  |  |
| Duration 0 |  |  |
| Central exposed | 530,080.0 | 343,230.2 |
| Deaths | 507 | 397 |
| Durations 1-4 |  |  |
| Central exposed | 1,679,654.0 | 1,076,950.8 |
| Deaths | 2,180 | 1,816 |
| Durations 5+ |  |  |
| Central exposed | 2,038,825.5 | 2,629,063.5 |
| Deaths | 4,968 | 6,831 |
| Females |  |  |
| Duration 0 |  |  |
| Central exposed |  | 233,720.8 |
| Deaths |  | 86 |
| Durations 1-4 |  |  |
| Central cxposed |  | 702,119.7 |
| Deaths |  | 614 |
| Durations 5+ |  |  |
| Central exposed |  | 2,004,000.2 |
| Deaths |  | 1,299 |

Table 5.2. Temporary assurances, males and females: age ranges.

|  | Range of data | Exposed $\geq 100$ | Deaths $\geq 10$ |
| :--- | :---: | :---: | :---: |
| Males |  |  |  |
| Duration 0 | $10-85$ | $17-72$ | $46-57$ |
| Durations $1-4$ | $10-87$ | 1879 | 2574 |
| Durations 5- | $10-91$ | $21-80$ | $31-77$ |
|  |  |  |  |
| Females | $10-95^{\times}$ | $17-70$ | none |
| Duration 0 | $10-90$ | $18-80$ | $34-55$ |
| Durations 1-4 | $10-91$ | 22.78 | 34.69 |
| Durations 5+ |  |  |  |

[^2]74 Standard Tables of Mortality based on the 1991-94 Experiences
Table 5.3. Temporary assurances: values of $100 \mathrm{~A} / \mathrm{E}$ when experiences are compared with mortality tables shown.

|  | TM80 (for males only) | 1991. 94 graduated rates for permanent assurances (males and females) for durations 0,1 and $2+$ respectively. |
| :---: | :---: | :---: |
| Males |  |  |
| Duration 0 | 86 | 87 |
| Durations 14 | 82 | 88 |
| Durations 5+ | 70 | 97 |
| Females |  |  |
| Duration 0 |  | 78 |
| Durations 1-4 |  | 120 |
| Durations 5+ |  | 96 |

Table 5.4. Temporary assurances, males: values of $\log$ likelihood and of $\chi^{2}$ for graduation of experience for durations $5+$.

|  | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) | GM(2,3) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $44,493.9$ | $44,488.1$ | $44,487.0$ | $44,487.5$ | $44,487.7$ | $44,487.4$ |
| $\chi^{2}$ | 74.3 | 59.3 | 60.9 | 61.1 | 58.7 | 61.3 |

Table 5.5. Temporary assurances, males: values of $\log$ likelihood and of $\chi^{2}$ for graduation of experience for durations 1-4.

|  | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) | GM(2,3) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $12,565.1$ | $12,557.1$ | $12,550.1$ | $12,553.2$ | $12,552.3$ | 12.552 .3 |
| $\chi^{2}$ | 96.3 | 82.1 | 68.9 | 72.8 | 72.3 | 72.1 |

Table 5.6. Temporary assurances, males: values of $\log$ likelihood and of $\chi^{2}$ for graduation of experience for duration 0 .

|  | $\mathrm{GM}(0,3)$ | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0,4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ | $\mathrm{GM}(2,3)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| - Log likelihood | $2,909.8$ | $2,909.1$ | $2,908.9$ | $2,909.0$ | $2,908.7$ | $2,908.4$ |
| $\chi^{2}$ | 40.2 | 39.5 | 38.5 | 39.1 | 36.7 | 36.0 |

Table 5.7. Temporary assurances, males: basic data for ages 17-29, durations $0,1-4$ and $5+$.

|  | Duration 0 | Durations 1-4 | Durations 5+ |
| :--- | ---: | :---: | ---: |
| Ages 17-24 |  |  |  |
| Exposed to risk <br> Deaths <br> Crude $\mu_{x}$ | 20,281 | $29,095.9$ | 20 |
|  |  |  |  |
| Ages 25-29 | 10 | 0.000687 | 0 |
| Exposed to risk <br> Deaths <br> Crude $\mu_{x}$ | 0.000493 |  | 0.0 |

76 Standard Tables of Mortality based on the 1991-94 Experiences
Table 5.8. Temporary assurances, males: values of parameters for graduations using GM(2,3) formulae before and after modification, durations $0,1-4$ and $5+$.

|  | Duration 0 | Durations 1-4 | Durations 5- |
| :--- | ---: | ---: | ---: |
| Before modification |  |  |  |
| $100 a_{1}$ | -0.121308 | -0.113910 | 0.058200 |
| $100 a_{2}$ | 0.057816 | -0.173839 | 0.023011 |
| $b_{1}$ | -2.937319 | -4.170362 | -4.262635 |
| $b_{2}$ | 4.579830 | 4.745634 | 5.970317 |
| $b_{3}$ | 1.429532 | -0.119740 | -0.490863 |
| - Log likelihood | $2,908.36$ | $12,552.3$ | $44,487.44$ |
| $\chi^{2}$ | 36.00 | 72.07 | 61.25 |
|  |  |  |  |
| After modification | -0.12 | unchanged |  |
| $100 a_{1}$ | -0.17 |  | -0.10 |
| $100 a_{2}$ | -4.415430 | -0.17 |  |
| $b_{1}$ | 4.486721 |  | -3.851692 |
| $b_{2}$ | -0.12 |  | 5.433870 |
| $b_{3}$ | $2,911.73$ | -0.12 |  |
| - Log likelihood | 43.40 |  | $44,490.43$ |
| $\chi^{2}$ |  | 67.57 |  |

Table 5.9. Temporary assurances, males, durations $0,1-4$ and $5+$ : statistics for graduations of $\mu_{x}=\operatorname{GM}(2,3)$.

| Duration: | 0 | $1-4$ | $5+$ |
| :--- | ---: | ---: | ---: |
| Ages used | $17-85$ | $17-87$ | $17-91$ |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | -0.12 | -0.113910 | -0.10 |
| $T$-ratio |  | -0.54 |  |
| $100 a_{2}$ | -0.17 | -0.173839 | -0.17 |
| $T$-ratio | -4.415430 | -0.90 |  |
| $b_{1}$ | -48.00 | -4.170362 | -3.851692 |
| $T$-ratio | 4.486721 | -3.45 | -166.20 |
| $b_{2}$ | 20.83 | 4.745634 | 5.433870 |
| $T$-ratio | --0.12 | 11.06 | 81.24 |
| $b_{3}$ |  | -0.119740 | -0.12 |
| $T$-ratio | -0.11 |  |  |
| Sign test: $p$ (pos) | 0.6170 |  |  |
| Runs test: $p($ runs $)$ | 0.6209 | 0.7478 | 0.8251 |
| K-S test: $p(K S)$ | 0.8322 | 0.4620 | 0.9902 |

Serial correlation 1est:

| $T$-ratio 1 | 0.18 | -0.16 | -1.27 |
| :--- | ---: | ---: | ---: |
| $T$-ratio 2 | 1.14 | -0.68 | -0.61 |
| $T$-ratio 3 | -0.86 | -0.54 | 1.18 |
| $\chi^{2}$ test: |  |  |  |
| $\chi^{2}$ | 43.40 | 72.07 | 67.57 |
| Degrees of freedom | 40 | 51 | 51 |
| $p\left(\chi^{2}\right)$ | 0.3284 | 0.0276 | 0.0600 |

Table 5.10. Temporary assurances, males, durations $0,1-4$ and $5+$ : specimen values of $q_{x}$ and percentage standard errors for graduations of $\mu_{x}=\operatorname{GM}(2,3)$.

| Duration: | 0 | $1-4$ | $5+$ |
| :--- | ---: | ---: | ---: |
| Age 20 | 0.000610 | 0.000707 | 0.000770 |
| Percentage s.e. | 1.85 | 473.33 | 0.25 |
| Age 30 | 0.000482 | 0.000587 | 0.000629 |
| Percentage s.e. | 6.27 | 246.14 | 1.01 |
| Age 40 | 0.000691 | 0.000861 | 0.000896 |
| Percentage s.e. | 11.45 | 92.57 | 2.23 |
| Age 50 | 0.001746 | 0.002175 | 0.002436 |
| Percentage s.e. | 11.64 | 22.68 | 2.55 |
| Age 60 | 0.004874 | 0.006181 | 0.007754 |
| Percentage s.e. | 10.49 | 6.08 | 2.43 |
| Age 70 | 0.012959 | 0.016962 | 0.023992 |
| Percentage s.e. | 9.68 | 5.34 | 2.31 |
| Age 80 | 0.032517 | 0.044179 | 0.070274 |
| Percentage s.e. | 9.18 | 16.99 | 2.20 |
| Age 90 | 0.077367 | 0.108844 | 0.190948 |
| Percentage s.e. | 8.76 | 45.01 | 2.03 |
| Age 100 | 0.173276 | 0.248635 | 0.451695 |
| Percentage s.e. | 8.17 | 76.38 | 1.64 |
| Age 110 | 0.355042 | 0.499781 | 0.811264 |
| Percentage s.e. | 7.12 | 66.09 | 0.87 |

Table 5.11. Temporary assurances, females: values of $\log$ likelihood and of $\chi^{2}$ for graduation of experience for durations $5+$.

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM $(1,3)$ | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | $9,464.9$ | $9,464.8$ | $9,464.7$ | $9,462.0$ | $9,462.4$ | $9,463.4$ |
| $\chi^{2}$ | 51.9 | 52.1 | 52.7 | 47.6 | 48.5 | 50.1 |

Table 5.12. Temporary assurances, females: values of log likelihood and of $\chi^{2}$ for graduation of experience for durations 1-4.

|  | GM(0,2) | GM(0,3) | GM(1,2) | GM(0,4) | GM(1,3) | GM(2,2) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelibood | $4,526.7$ | $4,526.6$ | $4,526.5$ | 4.526 .2 | $4,526.4$ | 4.526 .5 |
| $\chi^{2}$ | 87.5 | 75.1 | 74.3 | 73.6 | 74.1 | 74.1 |

Table 5.13. Temporary assurances, females: values of $\log$ likelihood and of $\chi^{2}$ for graduation of experience for duration 0 .

|  | $\mathrm{GM}(0,2)$ | $\mathrm{GM}(0,3)$ | $\mathrm{GM}(1,2)$ | $\mathrm{GM}(0,4)$ | $\mathrm{GM}(1,3)$ | $\mathrm{GM}(2,2)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| -Log likelihood | 709.1 | 709.1 | 709.1 | 709.1 | 709.1 | 709.1 |
| $\chi^{2}$ | 4.1 | 4.1 | 4.2 | 4.2 | 4.2 | 4.1 |

Table 5.14. Temporary assurances, females: values of parameters for graduations using GM(1,2) formulae before and after modification, durations $0,1-4$ and 5+.

|  | Duration 0 | Durations 1-4 | Durations $5+$ |
| :--- | ---: | ---: | ---: |
| Before modification |  |  |  |
| $100 a_{1}$ | -0.001102 | 0.002597 | 0.003781 |
| $b_{1}$ | -5.171077 | -4.489519 | -4.390281 |
| $b_{2}$ | 4.700423 | 4.876527 | 4.922720 |
| -Log likelihood | 709.1 | $4,526.5$ | $9,464.7$ |
| $\chi^{2}$ | 4.2 | 74.3 | 52.7 |
|  |  |  |  |
| After modification | 0.008 | 0.012 | 0.018 |
| $100 a_{1}$ | -5.082158 | -4.462904 | -4.352028 |
| $b_{1}$ | 5.493452 | 5.240908 | 5.347124 |
| $b_{2}$ | 710.4 | $4,528.4$ | $9,467.4$ |
| - Log likelihood | 7.9 | 85.4 | 56.9 |
| $\chi^{2}$ |  |  |  |

Table 5.15. Temporary assurances, females, durations $0,1-4$ and $5+$ : statistics for graduations of $\mu_{x}=\operatorname{GM}(1,2)$.

| Duration: | 0 | 1-4 | $5+$ |
| :---: | :---: | :---: | :---: |
| Ages used | 17-95 | 17-91 | 17-91 |
| Values of parameters at optimum point: |  |  |  |
| $100 a_{1}$ | 0.008 | 0.012 | 0.018 |
| $T$-ratio |  |  |  |
| $b_{1}$ | -5.082158 | -4.462904 | -4.352028 |
| T-ratio | -23.6 | -61.7 | -71.8 |
| $b_{2}$ | 5.493452 | 5.240908 | 5.347124 |
| T-ratio | 10.9 | 30.56 | 36.0 |
| Sign test: $p$ (pos) | 0.1509 | 0.2950 | 0.7121 |
| Runs test: p(runs) | 0.0949 | 0.5251 | 0.7238 |
| K-S test: $p(K S)$ | 0.9373 | 0.9206 | 0.7738 |
| Serial correlation test: |  |  |  |
| T-ratio I | 1.61 | 1.09 | $-0.54$ |
| T-ratio 2 | -0.91 | 0.14 | -0.44 |
| $T$-ratio 3 | -1.46 | 0.88 | 1.52 |
| $\chi_{2}^{2}$ test: |  |  |  |
| $\chi^{2}$ | 7.86 | 85.44 | 56.89 |
| Degrees of freedom | 11 | 51 | 47 |
| $p\left(\chi^{2}\right)$ | 0.7261 | 0.0018 | 0.1529 |

Table 5.16. Temporary assurances, females, durations $0,1-4$ and $5+$ : specimen values of $q_{x}$ and percentage standard errors for graduations of $\mu_{x}=\operatorname{GM}(1,2)$.

| Duration: | 0 | 14 | 5- |
| :---: | :---: | :---: | :---: |
| Age 20 | 0.000107 | 0.000184 | 0.000245 |
| Percentage s.e. | 5.73 | 2.56 | 1.62 |
| Agc 30 | 0.000161 | 0.000304 | 0.000369 |
| Percentage s.e. | 11.43 | 4.44 | 3.14 |
| Age 40 | 0.000323 | 0.000643 | 0.000729 |
| Percentage s.e. | 17.09 | 5.97 | 4.62 |
| Age 50 | 0.000808 | 0.001613 | 0.001780 |
| Percentage s.e. | 20.46 | 6.79 | 5.51 |
| Age 60 | 0.002264 | 0.004372 | 0.004834 |
| Percentage s.e. | 21.89 | 7.12 | 5.89 |
| Age 70 | 0.006618 | 0.012200 | 0.013680 |
| Percentage s.e. | 22.36 | 7.22 | 6.01 |
| Age 80 | 0.019568 | 0.034193 | 0.039009 |
| Percentage s.e. | 22.36 | 7.18 | 5.99 |
| Age 90 | 0.057417 | 0.094273 | 0.109163 |
| Percentage s.e. | 21.89 | 6.96 | 5.77 |
| Age 100 | 0.162429 | 0.245893 | 0.285707 |
| Percentage s.e. | 20.40 | 6.33 | 5.14 |
| Age 110 | 0.412353 | 0.552823 | 0.624691 |
| Percentage s.e. | 16.46 | 4.74 | 3.59 |

82 Standard Tables of Mortality based on the 1991-94 Experiences
Table 5.17. Temporary assurances: ratios of values of $q_{x}$ in proposed tables: comparison of durations.

| Age | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & \text { Duration 0/ } \\ & \text { Durations 1-4 } \end{aligned}$ | Durations 1-4/ <br> Durations 5+ | $\begin{gathered} \text { Duration } 0 / \\ \text { Durations } 1-4 \end{gathered}$ | Durations 1-4/ <br> Durations $5+$ |
| 20 | 0.8628 | 0.9182 | 0.5815 | 0.7510 |
| 25 | 0.8417 | 0.9239 | 0.5546 | 0.7897 |
| 30 | 0.8211 | 0.9392 | 0.5296 | 0.8238 |
| 35 | 0.8071 | 0.9586 | 0.5116 | 0.8566 |
| 40 | 0.8026 | 0.9609 | 0.5023 | 0.8820 |
| 45 | 0.8044 | 0.9355 | 0.4990 | 0.8988 |
| 50 | 0.8028 | 0.8929 | 0.5009 | 0.9062 |
| 55 | 0.7975 | 0.8452 | 0.5080 | 0.9075 |
| 60 | 0.7885 | 0.7971 | 0.5178 | 0.9044 |
| 65 | 0.7771 | 0.7509 | 0.5295 | 0.8988 |
| 70 | 0.7640 | 0.7070 | 0.5425 | 0.8918 |
| 75 | 0.7501 | 0.6660 | 0.5567 | 0.8842 |
| 80 | 0.7360 | 0.6287 | 0.5723 | 0.8765 |
| 85 | 0.7226 | 0.5961 | 0.5895 | 0.8694 |
| 90 | 0.7108 | 0.5700 | 0.6091 | 0.8636 |

Table 5.18. Temporary assurances: ratios of values of $q_{x}$ in proposed tables: comparison of sexes.

| Age | Duration 0 <br> Fcmales/Males | Durations 1 4 <br> Females/Males | Durations 5+ <br> Females/Males |
| :---: | :---: | :---: | :---: |
| 20 | 0.1754 | 0.2603 | 0.3182 |
| 25 | 0.2438 | 0.3700 | 0.4328 |
| 30 | 0.3340 | 0.5179 | 0.5904 |
| 35 | 0.4207 | 0.6636 | 0.7426 |
| 40 | 0.4674 | 0.7468 | 0.8136 |
| 45 | 0.4722 | 0.7612 | 0.7922 |
| 50 | 0.4628 | 0.7416 | 0.7307 |
| 55 | 0.4583 | 0.7195 | 0.6701 |
| 60 | 0.4645 | 0.7073 | 0.6234 |
| 65 | 0.4820 | 0.7074 | 0.5910 |
| 70 | 0.5107 | 0.7193 | 0.5702 |
| 75 | 0.5505 | 0.7417 | 0.5587 |
| 80 | 0.6018 | 0.7740 | 0.5551 |
| 85 | 0.6653 | 0.8156 | 0.5592 |
| 90 | 0.7421 | 0.8661 | 0.5717 |
| 95 |  |  | 0.5949 |
| 100 |  |  | 0.6325 |
| 105 |  |  | 0.6896 |
| 110 |  |  | 0.7700 |
| 115 |  |  | 0.8674 |

84 Standard Tables of Mortality based on the 1991-94 Experiences
Table 5.19. Ratios of values of $q_{x}$ in proposed tables for temporary assurances to those in corresponding tables for permanent assurances for 1991-94.

| Age | Males |  |  | Females |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1-4i $\text { Duration } 1$ | Durations $5+$ <br> Durations $2+$ | Duration 0 | Durations 1-4i Duration 1 | Durations 5+ <br> Durations 2+ |
| 20 | 1.4353 | 1.3068 | 1.3230 | 0.8560 | 1.2267 | 1.2694 |
| 25 | 1.1977 | 1.1614 | 1.1837 | 0.8194 | 1.1865 | 1.1694 |
| 30 | 1.0126 | 1.0520 | 1.0593 | 0.7854 | 1.1515 | 1.0885 |
| 35 | 0.9064 | 0.9954 | 0.9811 | 0.7666 | 1.1316 | 1.0203 |
| 40 | 0.8769 | 0.9707 | 0.9562 | 0.7618 | 1.1340 | 0.9759 |
| 45 | 0.8834 | 0.9408 | 0.9625 | 0.7684 | 1.1527 | 0.9514 |
| 50 | 0.8858 | 0.8936 | 0.9713 | 0.7845 | 1.1826 | 0.9438 |
| 55 | 0.8714 | 0.8409 | 0.9714 | 0.8088 | 1.2188 | 0.9460 |
| 60 | 0.8441 | 0.7965 | 0.9666 | 0.8382 | 1.2607 | 0.9557 |
| 65 | 0.8118 | 0.7693 | 0.9636 | 0.8709 | 1.3067 | 0.9696 |
| 70 | 0.7815 | 0.7637 | 0.9681 | 0.9063 | 1.3556 | 0.9865 |
| 75 | 0.7572 | 0.7833 | 0.9840 | 0.9440 | 1.4062 | 1.0052 |
| 80 | 0.7418 | 0.8323 | 1.0140 | 0.9835 | 1.4576 | 1.0250 |
| 85 | 0.7370 | 0.9170 | 1.0599 | 1.0244 | 1.5080 | 1.0451 |
| 90 | 0.7440 | 1.0463 | 1.1216 | 1.0662 | 1.5549 | 1.0647 |
| 95 |  |  | 1.1954 |  |  | 1.0824 |
| 100 |  |  | 1.2706 |  |  | 1.0962 |
| 105 |  |  | 1.3263 |  |  | 1.1025 |
| 110 |  |  | 1.3345 |  |  | 1.0966 |
| 115 |  |  | 1.2793 |  |  | 1.0749 |

Table 5.20. Ratios of values of $q_{x}$ in proposed tables for temporary assurances 1991-94, for males, to those in the TM80 tables.

| Age |  | Males |  |
| :---: | :---: | :---: | :---: |
|  |  |  |  |
|  | Duration 0 | Durations 1-4 | Durations 5+ |
| 20 | 1.1402 | 1.0584 | 0.9935 |
| 25 | 1.3223 | 1.2581 | 1.1185 |
| 30 | 1.3098 | 1.2789 | 1.1510 |
| 35 | 1.0806 | 1.1002 | 1.0903 |
| 40 | 0.8158 | 0.8849 | 0.8750 |
| 45 | 0.7174 | 0.7251 | 0.7168 |
| 50 | 0.7287 | 0.7066 | 0.6554 |
| 55 | 0.8094 | 0.7357 | 0.6509 |
| 60 | 0.9455 | 0.7949 | 0.6793 |
| 65 | 1.1258 | 0.8772 | 0.7292 |
| 70 | 1.2349 | 0.9798 | 0.7957 |
| 75 | 1.3611 | 1.1017 | 0.8762 |
| 80 | 1.5033 | 1.2426 | 0.9684 |
| 85 | 1.6594 | 1.4012 | 1.0694 |
| 90 | 1.8267 | 1.5744 | 1.1736 |
| 95 |  |  | 1.2710 |
| 100 |  |  | 1.3451 |
| 105 |  |  | 1.3740 |
| 110 |  |  | 1.3385 |
| 115 |  |  | 1.2420 |

86 Standard Tables of Mortality based on the 1991-94 Experiences


Figure 5.1. Ratios comparing values of $q_{x}$ by durations.


Figure 5.2. Ratios comparing values of $q_{x}$ for females and males.


Figure 5.3. Ratios comparing values of $q_{x}$ for temporary assurances 1991-94 with those for permanent assurances 1991-94.


Figure 5.4. Ratios comparing values of $q_{x}$ for temporary assurances, males, 1991-94 with TM80.


Figure 5.5. Temporary assurances, 1991-94, values of $q_{x}$ for males and females, durations $0,1-4$ and $5+$.

## 6. PROJECTION FACTORS FOR MORTALITY IMPROVEMENT

### 6.1 Introduction

6.1.1 In C.M.I.R. 16, 127-129 (1998), the Committee published details of its proposed new tables for life office pensioners. These tables are based on the experiences of normal retirals (i.e. persons who retired at or after the normal retirement age for their scheme) during the 1991-94 quadrennium. For each sex there are separate tables based on the experiences of lives and of amounts. For each of the lives tables the value of $q_{x}$ applies on average to a life attaining age $x$ in the middle of 1992 and gives the probability of death before the attainment of age $x+1$ in the middle of 1993. The 'base' year for these tables is thus 1992. Accordingly, the tables for male lives and female lives are denoted by PML92Base and PFL92Base respectively. The corresponding tables based on the amounts experiences are denoted by PMA92Base and PFA92Base.
6.1.2 When publishing the " 80 " Series of tables, based on the experiences of the 1979-82 quadrennium, the Committee was of the view that it would be imprudent not to incorporate into the new tables for pensioners, annuitants, and widows an allowance for projected improvements in mortality with the passage of time. Analysis of recent trends in the mortality of pensioners confirms that such an allowance is again essential for most practical purposes.

### 6.2 Observed improvements in the mortality of life office pensioners

6.2.1 It is of interest to consider the trends in pensioners' mortality since the setting up by the Bureau of the relevant investigations. For the lives investigations the first quadrennium for which data are available is 1955-58; for amounts it is 1959-62. Tables $6.4,6.5,6.6$, and 6.7 show the values of the $100 \mathrm{~A} / \mathrm{E}$ ratios (i.e. actual deaths as a percentage of those expected), where for each sex a single comparison is used throughout a 40 -year period (19551994). For males (both lives and amounts) the comparison basis is the PMA80C10 table. For females (both lives and amounts) it is the PFA80C10 table. Over the years, of course, different comparison bases have been used by the Committee. When, however, a comparison basis was changed, for at least one quadrennium the values of $100 \mathrm{~A} / \mathrm{E}$ were always given both on the previous basis and on the new basis. These two sets of ratios relating to a single quadrennium permit the calculation of 'bridging factors' which in turn allow all the earlier comparisons to be revised to the most recent basis. (Although an element of approximation is necessarily inherent in the use of
bridging factors, any errors are likely to be very small and of little practical consequence.)
6.2.2 For the first two quadrennia of the pensioners' investigations the data were insufficient to permit meaningful analyses over the age range 91 to 95 , even for males alone. In addition, the initial smallness of the relevant experiences meant that for females analyses above age 85 were feasible only for the 1975-78 and subsequent quadrennia.
6.2.3 The existence of a single comparison basis for each experience enables the crude rates of mortality (for each quinary age group) in the 1967-70 and subsequent quadrennia to be expressed as a percentage of the corresponding rate in the 1963-66 quadrennium. The results of these calculations are shown in Tables 6.8, 6.9, 6.10 and 6.11.
6.2.4 From these tables we may derive the values shown in Table 6.1. Thus, for example, we see that for the 61-65 age group in the male lives experience the crude mortality rate in the 1975-78 quadrennium was $79 \%$ of the corresponding rate in the 1959-62 quadrennium. Similarly, for the same age group the crude rate for the 1991-94 quadrennium was $65 \%$ of that for the period 16 years previously. By combining these figures we see that for the age group 61-65 the crude rate of mortality over the period 1991-94 was $51 \%$ of the corresponding crude rate for the period 1959-62.
6.2.5 Table 6.2 shows the corresponding ratios for the female investigations.
6.2.6 For males a clear pattern emerges from these figures. Within each age group a more rapid rate of improvement in mortality has occurred in recent years than in the more distant past. Improvements in mortality for amounts are significantly greater than for lives. For both lives and amounts the rate of mortality improvement decreases as age increases.
6.2.7 For females, in contrast, no clear pattern can be seen. In general, for both lives and amounts, a slower rate of improvement in mortality has occurred in recent years, but this is not true at all ages. Some age groups show a more rapid mortality improvement for amounts than for lives, but for other age groups the opposite is true.
6.2.8 Figure 6.1 illustrates the reductions in overall mortality for each of the four experiences since the 1963-66 quadrennium. (For male amounts, for example, the total number of deaths in 1963-66 was $162 \%$ of the number expected and in 1991-94 it was $98 \%$ of the number expected (on the same basis - see Table 6.5). Thus, the overall rate of mortality in 1991-94 was in some sense $60.5 \%$ of its value 28 years previously.) This comparison is, of course, influenced by differences in the age distributions of the two experiences. Accordingly, it is desirable to consider the trends in mortality improvement for specific age groups. Figure 6.2 illustrates the position for the male amounts
experience. It shows, for example, that over the 28 year period the mortality rate for the 66-70 age group fell by more than $50 \%$, while for the $76-80$ age group it fell by more than $30 \%$.
6.2.9 Figure 6.3 illustrates the corresponding reductions in mortality for the female amounts experience.
6.2.10 For males the improvements in mortality in recent years have been greater than those implied by the reduction factors recommended by the Committee in conjunction with the " 80 " Series of tables. That this is so is demonstrated by Figure 6.4, which for each year of the period 1983-94 expresses the actual number of deaths as a percentage of the number expected using the " 80 " Series projected mortality rates. For the female experiences, in contrast, significant "swings" occur from year to year and there is no clear pattern. This suggests that the female experiences are in some way unusual.
6.2.11 The experiences for the six most recent years for which data are available are illustrated in Table 6.13. Although not as marked as in previous years, this experience illustrates the fluctuating nature of the female experience.

### 6.3 Allowance for improving mortality

6.3.1 In the light of the above observations, the Committee considers it essential when publishing its new tables for pensioners, immediate annuitants, retirement annuitants, and widows also to publish projection factors to allow for improvements in mortality with the passage of time. As far as possible, the factors should reflect recent trends. In practice the rates at which mortality is improving vary from one experience to another. The Committee has given some thought as to the extent to which this feature should be reflected in its proposed projection factors. On the one hand it is desirable to reflect as far as possible the likely trends in the observed experiences. On the other hand it is desirable to adopt a relatively simple model which avoids obvious anomalies. (For example, it is questionable whether one could justify a mortality improvement model which leads at some future time to lower rates of mortality for males than for females or for lives than for amounts.) Taking all factors into consideration, the Committee is now of the view that for practical purposes the same set of projection factors should be used in conjunction with each of the new tables for pensioners, immediate annuitants, retirement annuitants, and widows.
6.3.2 The Committee's initial thoughts on this matter were presented to the profession at a seminar in London in December 1998. Some speakers agreed with the Committee's proposals while others pointed to the experience of recent years and suggested that more rapid improvements in mortality should be assumed. Having considered both these comments and subsequently
available data relating to the period 1992-97, the Committee has decided to issue revised projection factors, which incorporate more rapid improvement in the early years than was previously envisaged. The revised factors are based on the experiences of the five most recent quadrennia (i.e. 1975-78 to 1991-94). In recommending that a single set of improvement factors be used for all experiences (and, in each case, for males and females, lives and amounts, where relevant) the Committee recognises that it may possibly be overstating the future improvement in female mortality. Given, however, that the female experience has such unusual features, that it is difficult to know how much weight to give to it, and that one should probably err on the side of caution, the Committee feels that the recommended projection factors are reasonable for most practical purposes.
6.3.3 The model adopted to allow for mortality improvement is essentially the same as that described in $\$ 4.3$ of C.M.I.R. 10 (1990). At each age the rate of mortality is assumed to decrease exponentially to a limiting value. With the new factors, however, the speed of convergence to the limit depends on age (in contrast to the projection factors adopted for the " 80 " Series).
6.3.4 Measure time in years from 1992, and for $t=0,1,2, \ldots$ for each experience let $q_{x, t}$ denote the rate of mortality (i.e. " $q_{x}$ ") for a life attaining age $x$ in calendar year $1992+t$. (Similarly for the immediate annuitants let $q_{[x, t}$ be the select rate of mortality (i.e. " $q_{[x]}$ ") for a life newly selected at exact age $x$ in year $1992+t$.) For each experience $q_{x .0}$ is the value of $q_{x}$ in the relevant new 'base' table while $q_{[x], 0}$ is the value of $q_{x]}$ in the new annuitants table.
6.3.5 It is assumed that

$$
q_{x, t}=q_{x, 0} \cdot R F(x, t)
$$

where

$$
R F(x, t)=\frac{q_{x, t}}{q_{x, 0}}
$$

is the "reduction factor" for age $x$ and time $t$.
For immediate annuitants the select rate of mortality in calendar year $1992+t$ is

$$
q_{[x], 2}=q_{x_{i}, 0} \cdot R F(x, t)
$$

6.3.6 The reduction factor $R F(x, t)$ is defined in terms of two subsidiary functions, $\alpha(x)$ and $f(x)$. It is assumed that $q_{x, \infty}$, the long-term rate of mortality at age $x$, will be $\alpha(x)$ times the base rate (i.e. the rate in 1992). In addition, it is assumed that a fraction $f(x)$ of the total fall in the rate of mortality at age $x$ (i.e. $q_{x, 0}-q_{x, \infty}$ ) will occur in the first 20 years.
6.3.7 Both $\alpha(x)$ and $f(x)$ are linear functions of age for $60 \leq x \leq 110$. For $x<60$ and for $x>110$ each function is constant, taking the appropriate values to ensure continuity.
6.3.8 Somewhat arbitrarily the value of $\alpha(110)$ is taken as 1 . (The financial consequences of this choice are of very little significance.) Subject to this single constraint the linear functions $\alpha(x)$ and $f(x)$ were determined to provide a clearly-defined 'best fit' to the observed mortality improvements over the 20 -year period $1975-94$. A further technical adjustment was made so that the finally-adopted projection factors have 1992 as their base year.
6.3.9 The reduction factor recommended for all experiences is

$$
R F(x, t)=\alpha(x)+[1-\alpha(x)] \cdot[1-f(x)]^{\frac{1}{00}}
$$

with

$$
\alpha(x)= \begin{cases}c & x<60 \\ 1+(1-c) \cdot \frac{(x-110)}{50} & 60 \leq x \leq 110 \\ 1 & x>110\end{cases}
$$

and

$$
f(x)= \begin{cases}h & x<60 \\ \frac{(110-x) \cdot h+(x-60) \cdot k}{50} & 60 \leq x \leq 110 \\ k & x>110\end{cases}
$$

where $c=0.13, h=0.55$, and $k=0.29$.
These values of $c, h$, and $k$ are derived from the adjusted 'best fit' procedure described in outline above.
6.3.10 Specimen reduction factors are given in Table 6.12.
6.3.11 One simple measure of the effect of improving mortality is obtained by considering how the expectation of life at a given age increases with the passage of time. Table 6.3 shows the values of $\dot{e}_{x}$ (for $x=60,70,80$ ) on the PMA92Base and PFA92Base tables and at various future times on both a 'true' and a calendar year basis.
6.3.12 It is important to appreciate that the expectations of life reflected in Table 6.3 apply to no one individual, but simply reflect 1992 levels of mortality. For the amounts tables, Figures 6.5 and 6.6 show how, when allowance is made for improving mortality, the 'true' expectation of life at any age increases as the time of attainment of that age moves into the future. Thus, for example, for the male amounts table the 'true' value of $\dot{\varepsilon}_{60}$ for a life attaining age 60 in 1992 is around 1.8 years more than the value of 21.2 derived from the base table
while for a life attaining age 60 in the year 2000 (i.e. time $t=8$ ) the value of $\dot{e}_{60}$ is nearly 2.5 years more than the corresponding base table value. These simple observations are not without financial significance!

Table 6.1. Life office pensioners, males: ratios of crude mortality rates in two sets of quadrennia, each 16 years apart.

|  | $\frac{1975-78}{1959-62}$ |  |  | $\frac{1991-94}{1975 \cdots 78}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Age group | Lives | Amounts |  | Lives |  |
| $61-65$ | .79 | .82 | .65 | Amounts |  |
| $66-70$ | .88 | .83 | .65 | .49 |  |
| $71-75$ | .99 | .94 | .69 | .56 |  |
| $76-89$ | .97 | .75 | .63 |  |  |
| $81-85$ |  |  | .83 | .69 |  |
| All ages | .94 | .92 | .73 | .74 |  |

Note: The ratios for 'all ages' relate to the age range 51-100. At the oldest ages ratios are not available for the 1975-78/1959-62 comparison.

Table 6.2. Life office pensioners, females: ratios of crude mortality rates in two sets of quadrennia, each 16 years apart.

|  | $\frac{1975-78}{1959-62}$ |  | $\frac{1991-94}{1975 \cdot 78}$ |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Age group | Lives | Amounts |  | Lives | Amounts |
| $61-65$ | .76 | .58 | .84 | .91 |  |
| $66-70$ | .88 | .87 | .85 | .71 |  |
| $71-75$ | .78 | .84 | .84 | .89 |  |
| $76-80$ |  |  | .86 | .70 |  |
| $81-85$ | .81 | .78 | .83 | .75 |  |
| All ages |  | .81 | .77 |  |  |

Note: The ratios for 'all ages' relate to the age range $51-100$. At the oldest ages ratios are not available for the 1975-78/1959-62 comparison.

Table 6.3a. Expectation of life $\left(\dot{e}_{x}\right)$, males: calendar year bases.

| Age $x$ | PMA92Base | PMA92C2010 | PMA92C2030 |
| :--- | :---: | :---: | :---: |
| 60 | 21.2 | 23.5 | 25.2 |
| 70 | 13.2 | 14.9 | 16.2 |
| 80 | 7.4 | 8.4 | 9.1 |

Table 6.3b. Expectation of life $\left(\tilde{e}_{x}\right)$, females: calendar year bases.

| Age $x$ | PFA92Base | PFA92C2010 | PFA92C2030 |
| :--- | :---: | :---: | :---: |
| 60 | 24.2 | 26.5 | 28.1 |
| 70 | 16.0 | 17.7 | 19.0 |
| 80 | 9.7 | 10.7 | 11.5 |

Table 6.3c. Expectation of life ( $\stackrel{e}{e}_{x}$ ), males: 'true' (i.e. year of use) bases.

| Age $x$ | PMA92U1992 | PMA92U2000 | PMA92U2010 |
| :--- | :---: | :---: | :---: |
| 60 | 23.0 | 23.9 | 24.8 |
| 70 | 14.0 | 14.7 | 15.5 |
| 80 | 7.7 | 8.1 | 8.5 |

Table 6.3d. Expectation of life $\left(e_{x}\right)$, females: 'true' (i.e. year of use) bases.

| Age $x$ | PFA92U1992 | PFA92U2000 | PFA92U2010 |
| :--- | :---: | :---: | :---: |
| 60 | 26.1 | 26.9 | 27.8 |
| 70 | 16.9 | 17.7 | 18.5 |
| 80 | 10.0 | 10.5 | 11.0 |

Table 6.4. Life office pensioners, male lives, retirements at or after normal pension age: $100 \mathrm{~A} / \mathrm{E}$ ratios where for each quadrennium the comparison basis is PMA80C10.

| Age group | Quadrennium |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1955-58 | 1959-62 | 1963-66 | 1967-70 | 1971-74 | 197578 | 1979 - 82 | 1983-86 | 1987-90 | 1991-94 |
| 51-55 |  |  |  |  |  |  |  | $289 *$ | 333 | 203 |
| 56-60 |  |  |  |  |  |  |  | 226 | 197 | 181 |
| 61-65 | 350 | 271 | 289 | 257 | 235 | 213 | 206 | 175 | 166 | 138 |
| 6670 | 241 | 222 | 217 | 211 | 202 | 196 | 176 | 156 | 140 | 127 |
| 71-75 | 196 | 183 | 181 | 183 | 181 | 181 | 165 | 145 | 138 | 124 |
| 76-80 | 165 | 161 | 161 | 153 | 157 | 159 | 151 | 137 | 129 | 120 |
| 81-85 |  |  | 147 | 135 | 143 | 139 | 137 | 130 | 122 | 116 |
| 86-90 |  |  | 139 | 124 | 128 | 132 | 126 | 123 | 114 | 111 |
| 91-95 |  |  | 114 | 126 | 124 | 115 | 115 | 112 | 109 | 106 |
| 96-100 |  |  |  |  |  |  |  | 109 | 96 | 88 |
| All ages | 184 | 172 | 168 | 162 | 162 | 161 | 151 | 138 | 127 | 118 |
| Total deaths | $11,320{ }^{+}$ | 18,941 ${ }^{\text { }}$ | $30,969^{+}$ | $45,863^{\dagger}$ | $63.481^{\dagger}$ | $76,907^{+}$ | $85,426^{\dagger}$ | 84,267 | 74,842 | 63,558 |

## Notes:

* denotes an age group with fewer than 10 actual deaths.
${ }^{\dagger}$ indicates that the total deaths and the "all agcs" $100 \mathrm{~A} / \mathrm{E}$ ratios include a small number of cases outwith the $51-100$ age group.

Table 6.5. Life office pensioners, male amounts, retirements at or after normal pension age: $100 \mathrm{~A} / \mathrm{E}$ ratios where for each quadrennium the comparison basis is PMA80C10.

| Age group | Quadrennium |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1959-62 | 196366 | 1967-70 | 1971-74 | 1975-78 | 1979-82 | 1983-86 | 1987-90 | 199194 |
| 51-55 |  |  |  |  |  |  | 210 | 207 | 80 |
| 56-60 |  |  |  |  |  |  | 259 | 66 | 101 |
| 61-65 | 233 | 250 | 227 | 218 | 190 | 159 | 153 | 117 | 94 |
| 66-70 | 201 | 198 | 179 | 174 | 167 | 143 | 125 | 107 | 93 |
| 71.75 | 163 | 159 | 160 | 159 | 154 | 140 | 117 | 111 | 97 |
| 76-80 | 145 | 139 | 137 | 141 | 141 | 127 | 117 | 106 | 97 |
| 81-85 |  | 140 | 129 | 133 | 136 | 127 | 119 | 103 | 100 |
| 8690 |  | 139 | 119 | 120 | 125 | 125 | 116 | 110 | 118 |
| 91-95 |  | 114 | 122 | 229 | 111 | 114 | 117 | 105 | 106 |
| 96-100 |  |  |  |  |  |  | 114 | 104 | 89 |
| All ages | 164 | 162 | 155 | 154 | 151 | 135 | 121 | 108 | 98 |
| Total deaths | $1,809^{\dagger}$ | $3,382^{+}$ | $5.758^{\dagger}$ | 9,094 ${ }^{\dagger}$ | 13,595 ${ }^{\text {+ }}$ | $20,021^{+}$ | 28,672 | 37,372 | 46,096 |

Notes:
The total deaths are shown in units of $£ 1,000$.
${ }^{\dagger}$ indicates that the total deaths and the "all ages" $100 \mathrm{~A} / \mathrm{E}$ ratios include a small number of cases outwith the 51-100 age group.

Table 6.6. Life office pensioners, female lives, retirements at or after normal pension age: $100 \mathrm{~A} / \mathrm{E}$ ratios where for each quadrennium the comparison basis is PFA 80 C 10 .

| Age group | Quadrennium |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 195558 | 1959-62 | 1963-66 | 1967-70 | 197174 | 1975-78 | 1979-82 | 1983-86 | 1987-90 | 1991-94 |
| 51-55 |  |  |  |  |  |  |  | $357 *$ | $267^{*}$ | $321{ }^{*}$ |
| 56-60 | 356 | 224 | 194 | 278 | 313 | 278 | 269 | 177 | 201 | 218 |
| 61-65 | 270 | 253 | 231 | 207 | 205 | 193 | 189 | 191 | 179 | 162 |
| 66-70 | 199 | 211 | 203 | 205 | 185 | 185 | 175 | 164 | 168 | 158 |
| 7175 | 185 | 216 | 184 | 181 | 170 | 168 | 156 | 143 | 152 | 141 |
| 76-80 |  |  | 163 | 147 | 157 | 144 | 139 | 130 | 128 | 124 |
| 81-85 |  |  | 139 | 135 | 137 | 141 | 137 | 122 | 117 | 117 |
| 86-90 |  |  |  |  |  | 134 | 132 | 118 | 115 | 108 |
| 91-95 |  |  |  |  |  | 138 | 134 | 134 | 120 | 117 |
| 96-100 |  |  |  |  |  |  |  | 120 | 127 | 121 |
| All ages | 187 | 191 | 169 | 162 | 157 | 154 | 148 | 136 | 132 | 124 |
| Total deaths | 588 | 1,324 ${ }^{\dagger}$ | $2,322^{\dagger}$ | 3,828 ${ }^{+}$ | $5,791^{+}$ | $8,086{ }^{\dagger}$ | 10,536 ${ }^{\dagger}$ | 12,266 | 13,124 | 13,246 |

## Notes:

* denotes an age group with fewer than 10 actual deaths.
${ }^{\dagger}$ indicates that the total deaths and the "all ages" $100 \mathrm{~A} / \mathrm{E}$ ratios include a small number of cases outwith the $51-100$ age group.

Table 6.7. Life office pensioners, female amounts, retirements at or after normal pension age: $100 \mathrm{~A} / E$ ratios where for each quadrenmium the comparison basis is PFA80C10.

| Age group | Quadrennium |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1959-62 | 1963-66 | 1967.70 | 1971-74 | 1975-78 | 1979-82 | 1983-86 | 1987-90 | 1991-94 |
| 51-55 |  |  |  |  |  |  | 221 | 926 | 306 |
| 5660 | 199 | 163 | 288 | 260 | 288 | 224 | 132 | 130 | 123 |
| 61-65 | 273 | 194 | 180 | 199 | 158 | 168 | 171 | 166 | 143 |
| 66-70 | 199 | 188 | 189 | 173 | 173 | 148 | 135 | 132 | 123 |
| 71-75 | 179 | 180 | 153 | 178 | 151 | 140 | 127 | 123 | 134 |
| 76-80 |  | 151 | 153 | 148 | 141 | 126 | 114 | 121 | 99 |
| 81-85 |  | 147 | 142 | 149 | 139 | 126 | 103 | 110 | 104 |
| 86-90 |  |  |  |  | 153 | 133 | 119 | 113 | 101 |
| 91-95 |  |  |  |  | 144 | 124 | 138 | 130 | 117 |
| 96-100 |  |  |  |  |  |  | 142 | 128 | 129 |
| All ages | 193 | 162 | 159 | 162 | 151 | 140 | 127 | 127 | 117 |
| Total deaths | $67^{+}$ | $130^{\dagger}$ | $255^{+}$ | $490{ }^{+}$ | $840^{+}$ | $1,446^{\dagger}$ | 2.470 | 4,177 | 5,598 |

Table 6.9. Life office pensioners, male amounts: mortality rates as a percentage of values in the quadrennium 1963-66.

| Age group | Quadrennium |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1963-66 | 1967-70 | 1971-74 | 1975-78 | 1979-82 | 1983-86 | 198790 | 1991-94 |
| 61-65 | 100 | 91 | 88 | 76 | 64 | 61 | 47 | 38 |
| 66-70 | 100 | 91 | 88 | 84 | 72 | 63 | 54 | 47 |
| 71-75 | 100 | 101 | 100 | 97 | 88 | 74 | 70 | 61 |
| 76-80 | 100 | 98 | 101 | 101 | 91 | 84 | 76 | 70 |
| 81-85 | 100 | 92 | 95 | 97 | 91 | 85 | 74 | 72 |
| 86-90 | 100 | 86 | 87 | 90 | 90 | 84 | 79 | 85 |
| 91-95 | 100 | 107 | 201 | 98 | 100 | 103 | 92 | 93 |

Table 6.11. Life office pensioners, female amounts: mortality rates as a percentage of values in the quadrennium 1963-66.

|  | Quadrennium |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age group | $1963 \cdot 66$ | $1967-70$ | $1971-74$ | $1975-78$ | $1979-82$ | $1983-86$ | $1987-90$ | $1991-94$ |
| $61-65$ | 100 | 93 | 103 | 81 | 87 | 88 | 86 | 74 |
| 6670 | 100 | 101 | 92 | 92 | 79 | 72 | 70 | 66 |
| $71-75$ | 100 | 85 | 99 | 84 | 78 | 71 | 68 | 74 |
| $76-80$ | 100 | 102 | 99 | 94 | 84 | 76 | 80 | 66 |
| $81-85$ | 100 | 97 | 101 | 95 | 86 | 70 | 75 | 71 |

Table 6.12. Reduction factors $\times 1000$ with 1992 as the base year: males and females, lives and amounts.

| Age $x$ | Time $t$ in years, measured from 1992 |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 4 | 8 | 12 | 16 | 20 | 24 | 28 | 32 | 36 | 40 |
| 60 | 872 | 762 | 669 | 589 | 522 | 464 | 414 | 372 | 337 | 306 |
| 65 | 892 | 799 | 719 | 649 | 590 | 538 | 494 | 456 | 423 | 394 |
| 70 | 910 | 832 | 764 | 705 | 653 | 608 | 569 | 535 | 505 | 479 |
| 75 | 927 | 863 | 806 | 756 | 713 | 674 | 640 | 610 | 584 | 561 |
| 80 | 942 | 890 | 844 | 803 | 767 | 735 | 706 | 681 | 658 | 638 |
| 85 | 955 | 915 | 879 | 846 | 817 | 791 | 768 | 747 | 728 | 711 |
| 90 | 967 | 937 | 910 | 885 | 863 | 843 | 825 | 808 | 793 | 780 |
| 95 | 977 | 956 | 937 | 920 | 904 | 889 | 876 | 864 | 853 | 843 |
| 100 | 986 | 973 | 961 | 950 | 940 | 931 | 923 | 915 | 908 | 901 |

Table 6.13. Life office pensioners, all ages: values of $100 \mathrm{~A} / \mathrm{E}$ for calendar years 1992 to 1997 using the base mortality tables published in C.M.I.R. 16, projected to the relevant year using the reduction factors described in 6.3.6, as the comparison basis.

| Ycar | Male Amounts | Males Lives | Female Amounts | Fcmatc Lives |
| :--- | :---: | :---: | :---: | :---: |
| 1992 | 101 | 101 | 109 | 107 |
| 1993 | 99 | 102 | 93 | 95 |
| 1994 | 97 | 98 | 100 | 95 |
| 1995 | 104 | 98 | 94 | 100 |
| 1996 | 98 | 95 | 96 | 96 |
| 1997 | 96 | 93 | 93 | 99 |

106 Standard Tables of Mortality based on the 1991-94 Experiences


Figure 6.1. Life office pensioners, all ages: 100A/E as a percentage of the value for $1963-$ 66: comparison basis PMA80C10 for males, PFA80C10 for females: compare with Tables 6.4, 6.5, 6.6 and 6.7.


Figure 6.2. Life office pensioners, male amounts: $100 \mathrm{~A} / \mathrm{E}$ as a percentage of the value for 1963-66: comparison basis PMA80C10.


Figure 6.3. Life office pensioners, female amounts: $100 \mathrm{~A} / \mathrm{E}$ as a percentage of the value for 1963-66: comparison basis PFA80C10.


Figure 6.4. Life office pensioners, all ages: actual deaths as a percentage of expected deaths using the " 80 " Series projected mortality rates.


Figure 6.5. Life office pensioners, male amounts: increase in 'true' (i.e. year of use) expectation of life over 'base' value (see Table 6.3) with the passing of time.


Figure 6.6. Life office pensioners, female amounts: increase in 'true' (i.e. year of use) expectation of life over 'base' value (see Table 6.3) with the passing of time.

## 7. SPECIMEN MONETARYVALUESAND COMPARISONS WITH <br> other TAbles

### 7.1 Introduction

7.1.1 The tables which follow show specimen monetary functions derived from the new tables and compare these with the corresponding functions from the equivalent " 80 " Series tables (or with the 'least different' " 80 " Series table where a direct equivalent does not exist). The layouts of the tables follow the layouts of the tables in Section 5 of C.M.I.R. 10, 59-85 where the " 80 " Series tables were compared with earlier tables.

### 7.2 Assurances

7.2.1 Tables 7.1 and 7.2 cover permanent (whole life and endowment) assurances and Tables 7.3 and 7.4 cover temporary assurances. For female temporary assurances there is no direct equivalent table in the " 80 " Series, so the AF80 table has been used as the comparison basis. A range of premium rates and policy values are shown.

### 7.3 Pensioners

7.3.1 Tables 7.5 to 7.8 show annuity values on the three projection bases (calendar year, year of birth and year of use) for a variety of years. These are then compared with the annuity values calculated using the corresponding " 80 " Series table. All " 92 " Series tables are projected using the factors described in Section 6 of this report; all " 80 " Series tables are projected using the factors described in Section 4 of C.M.I.R. 10, 44-58.
7.3.2 The section headed Year of Use gives specimen annuity values for cohorts reaching the designated ages in 2000 and 2010. Those values can be picked out from the Year of Birth section where the values appropriate for the year 2010 are highlighted in bold.

### 7.4 Immediate annuitants

7.4.1 Tables 7.9 to 7.16 show specimen annuity values for the new immediate annuitant tables. The format is similar to that used for pensioners. For the new tables based on amounts, the comparison basis is the immediate annuitants " 80 " Series table (based on lives) for males or females and select duration as appropriate.

### 7.5 Retirement annuitants

7.5.1 Tables 7.17 and 7.18 show specimen annuity values for the new vested retirement annuitant tables. The format is similar to that used for pensioners. The comparison basis is the pensioner, lives (males and females respectively) from the " 80 " Series since no exact equivalent was previously produced.

### 7.6 Widows

7.6.1 Tables 7.19 and 7.20 show specimen annuity values for the new widows tables. The format is similar to that used for pensioners.

Table 7.1. Comparison of monetary values on AM92 and AM80: rate of interest $4 \%$.

Sample AM92 monetary functions
Premium rates per $£ 1,000$ sum assured

| $x$ | $P_{\mid x}$ | $P_{\text {[x] } \mid \text { Ts] }}$ | $P_{[x]: \overline{25}}$ | $t$ | ${ }_{1} V_{25}$ | ${ }_{1} V_{45}$ | 1 | ${ }_{\text {t }} V_{25: 251}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 5 | 0.0304 | 0.0732 | 5 | 0.1296 |
| 20 | 4.86 | 48.32 | 23.45 | 10 | 0.0673 | 0.1567 | 10 | 0.2875 |
| 25 | 5.94 | 48.33 | 23.51 | 15 | 0.1116 | 0.2491 | 15 | 0.4798 |
| 30 | 7.33 | 48.39 | 23.64 | 20 | 0.1642 | 0.3478 | 20 | 0.7139 |
| 35 | 9.14 | 48.52 | 23.93 | 25 | 0.2254 | 0.4488 |  |  |
| 40 | 11.52 | 48.79 | 24.48 | 30 | 0.2951 | 0.5471 |  |  |
| 45 | 14.65 | 49.34 | 25.52 | 35 | 0.3724 | 0.6378 | $t$ | ${ }_{\text {, }} V_{45}$ :25 |
| 50 | 18.83 | 50.35 | 27.38 | 40 | 0.4549 | 0.7167 |  |  |
| 55 | 24.47 | 52.19 | 30.60 | 45 | 0.5393 |  | 5 | 0.1349 |
| 60 | 32.12 | 55.41 | 35.96 | 50 | 0.6215 |  | 10 | 0.2948 |
| 65 | 42.60 | 60.90 | 44.56 | 55 | 0.6972 |  | 15 | 0.4844 |
|  |  |  |  | 60 | 0.7632 |  | 20 | 0.7126 |

AM92 monetary functions as a percentage of the equivalent AM80 functions

| Premium rates per $£ 1,000$ sum assured |  |  |  | Policy values |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{[x]}$ | $P_{\text {\{x\| }}$ : 51 | $P_{[x \mid 2} 25$ | $t$ | ${ }_{1}{ }_{25}$ | ${ }_{t} V_{45}$ | $t$ | ${ }_{1} V_{25: 25}$ |
|  |  |  |  | 5 | 88 | 90 | 5 | 100 |
| 20 | 89 | 100 | 100 | 10 | 88 | 91 | 10 | 100 |
| 25 | 89 | 100 | 100 | 15 | 88 | 92 | 15 | 100 |
| 30 | 89 | 100 | 100 | 20 | 89 | 94 | 20 | 100 |
| 35 | 88 | 100 | 99 | 25 | 89 | 95 |  |  |
| 40 | 87 | 99 | 97 | 30 | 91 | 96 |  |  |
| 45 | 87 | 99 | 96 | 35 | 92 | 97 | $t$ | ${ }_{t} V_{45: 25}$ |
| 50 | 87 | 98 | 94 | 40 | 93 | 98 |  |  |
| 55 | 87 | 97 | 92 | 45 | 95 |  | 5 | 98 |
| 60 | 87 | 96 | 91 | 50 | 96 |  | 10 | 99 |
| 65 | 88 | 95 | 90 | 55 | 97 |  | 15 | 100 |
|  |  |  |  | 60 | 98 |  | 20 | 100 |

Table 7.2. Comparison of monetary values on AF92 and AF80: rate of interest $4 \%$.

Sample AF92 monetary functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{[x]}$ |  | $P_{[x]: 20]}^{〕}$ | $P_{i x \mid}$ [2S | $t$ | ${ }_{1} V_{25}$ | ${ }_{1} V_{45}$ | $t$ | , $V_{25: 251}$ | ${ }_{,} V_{25: 25]}^{l}$ |
|  |  |  |  |  | 5 | 0.0256 | 0.0577 | 5 | 0.1302 | 0.0018 |
| 20 | 3.89 | 0.22 | 0.32 | 23.27 | 10 | 0.0561 | 0.1242 | 10 | 0.2884 | 0.0033 |
| 25 | 4.79 | 0.30 | 0.46 | 23.35 | 15 | 0.0924 | 0.1993 | 15 | 0.4806 | 0.0042 |
| 30 | 5.93 | 0.43 | 0.70 | 23.49 | 20 | 0.1352 | 0.2825 | 20 | 0.7145 | 0.0036 |
| 35 | 7.36 | 0.66 | 1.09 | 23.71 | 25 | 0.1851 | 0.3720 |  |  |  |
| 40 | 9.18 | 1.03 | 1.75 | 24.09 | 30 | 0.2426 | 0.4652 |  |  |  |
| 45 | 11.52 | 1.65 | 2.84 | 24.71 | 35 | 0.3076 | 0.5584 | $t$ | ${ }_{1} V_{45}$ :25 | ${ }_{1} V_{45: 251}^{1}$ |
| 50 | 14.56 | 2.67 | 4.63 | 25.75 | 40 | 0.3795 | 0.6472 |  |  |  |
| 55 | 18.53 | 4.38 | 7.56 | 27.47 | 45 | 0.4569 |  | 5 | 0.1324 | 0.0135 |
| 60 | 23.78 | 7.20 | 12.28 | 30.31 | 50 | 0.5375 |  | 10 | 0.2910 | 0.0253 |
| 65 | 30.81 | 11.84 | 19.68 | 34.97 | 55 | 0.6181 |  | 15 | 0.4814 | 0.0322 |
|  |  |  |  |  | 60 | 0.6949 |  | 20 | 0.7125 | 0.0277 |

AF92 monetary functions as a percentage of the equivalent AF80 functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{\|x\|}$ | $P_{\|x\| \text { 南 }}^{1}$ | $P_{[x\|: 20\|}^{1}$ | $P_{[x]} \overline{25}$ | $t$ | ${ }_{1} V_{25}$ | ${ }_{1} V_{45}$ | $t$ | ${ }_{1} V_{25}$ :25i | ${ }_{1} V_{25: 25]}^{1}$ |
|  |  |  |  |  | 5 | 91 | 91 | 5 | 100 | 82 |
| 20 | 91 | 81 | 86 | 100 | 10 | 91 | 91 | 10 | 100 | 81 |
| 25 | 91 | 91 | 88 | 100 | 15 | 91 | 92 | 15 | 100 | 79 |
| 30 | 91 | 89 | 85 | 100 | 20 | 91 | 92 | 20 | 100 | 79 |
| 35 | 90 | 86 | 83 | 99 | 25 | 91 | 93 |  |  |  |
| 40 | 90 | 82 | 80 | 99 | 30 | 92 | 94 |  |  |  |
| 45 | 90 | 80 | 79 | 98 | 35 | 92 | 95 | $t$ | ${ }_{\text {, }} V_{45}$ 25] | ${ }_{1} V_{45}^{1} \overline{25}$ |
| 50 | 89 | 78 | 79 | 97 | 40 | 93 | 96 |  |  |  |
| 55 | 89 | 78 | 79 | 96 | 45 | 94 |  | 5 | 99 | 78 |
| 60 | 88 | 78 | 79 | 94 | 50 | 94 |  | 10 | 100 | 78 |
| 65 | 88 | 78 | 80 | 92 | 55 | 95 |  | 15 | 100 | 78 |
|  |  |  |  |  | 60 | 96 |  | 20 | 100 | 78 |

## 112 Standard Tables of Mortality based on the 1991-94 Experiences

Table 7.3. Comparison of monetary values on TM92 and TM80: rate of interest $4 \%$.

Sample TM92 monetary functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{[x]}$ | $P_{\text {[x] }}^{1}$ :10] | $P_{\text {l }}^{1} \times 2.25$ | $\begin{aligned} & P_{x \mid: n}^{1}: n \\ & x \mid n=6 s \end{aligned}$ | $t$ | ${ }_{1} / 2{ }^{2}$ | ${ }_{4} V_{30: 351}^{1}$ | $t$ | ${ }_{7} V_{25,25}^{1}$ | ${ }_{1} V_{45: 25]}^{1}$ |
|  |  |  |  |  | 5 | 0.0300 | 0.0084 | 5 | 0.0012 | 0.0223 |
| 20 | 4.87 | 0.62 | 0.69 | 1.52 | 10 | 0.0666 | 0.0182 | 10 | 0.0028 | 0.0428 |
| 25 | 5.92 | 0.58 | 0.82 | 1.75 | 15 | 0.1109 | 0.0284 | 15 | 0.0041 | 0.0555 |
| 30 | 7.30 | 0.63 | 1.17 | 2.09 | 20 | 0.1534 | 0.0372 | 20 | 0.0039 | 0.0490 |
| 35 | 9.09 | 0.82 | 1.89 | 2.57 | 25 | 0.2246 | 0.0411 |  |  |  |
| 40 | 11.45 | 1.27 | 3.22 | 3.22 | 30 | 0.2945 | 0.0332 |  |  |  |
| 45 | 14.56 | 2.15 | 5.57 | 4.10 | 35 | 0.3721 |  | $t$ | $1000, V_{25: 70 \mid}^{1}$ | $\left.{ }_{t} V_{45:}^{1} \cdot \underline{\theta}\right]$ |
| 50 | 18.68 | 3.74 | 9.49 | 5.22 | 40 | 0.4557 |  |  |  |  |
| 55 | 24.17 | 6.50 | 15.71 | 6.50 | 45 | 0.5420 |  | 2 | -0.0441 | 0.0018 |
| 60 | 31.55 | 11.18 | 24.97 | 7.03 | 50 | 0.6273 |  | 4 | -0.0427 | 0.0029 |
| 65 | 41.48 | 18.91 | 37.62 |  | 55 | 0.7070 |  | 6 | -0.0156 | 0.0033 |
|  |  |  |  |  | 60 | 0.7772 |  | 8 | 0.0109 | 0.0025 |

TM92 monetary functions as a percentage of the equivalent TM80 functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{\text {ck }}$ |  | $P_{\|x\|: 25 \mid}^{1}$ | $\begin{aligned} & P_{\|x\|: n}^{1} \cdot n \\ & x+n=65 \end{aligned}$ | $t$ | ${ }_{7} F_{25}$ | ${ }_{1} V_{30: 35}^{1}$ | $t$ | ${ }_{2} V_{25: 251}^{1}$ | ${ }_{1} V_{45: 25]}^{1}$ |
|  |  |  |  |  | 5 | 90 | 64 | 5 | 54 | 72 |
| 20 | 93 | 115 | 103 | 78 | 10 | 90 | 64 | 10 | 56 | 74 |
| 25 | 92 | 122 | 90 | 75 | 15 | 90 | 65 | 15 | 55 | 76 |
| 30 | 91 | 109 | 77 | 73 | 20 | 91 | 66 | 20 | 56 | 79 |
| 35 | 91 | 88 | 71 | 71 | 25 | 92 | 68 |  |  |  |
| 40 | 90 | 72 | 69 | 69 | 30 | 94 | 70 |  |  |  |
| 45 | 90 | 67 | 72 | 68 | 35 | 95 |  | 1 | 1000, $V_{25}^{1}$ : 10 | ${ }_{1} V_{45: 10]}^{\prime}$ |
| 50 | 90 | 68 | 76 | 69 | 40 | 97 |  |  |  |  |
| 55 | 92 | 73 | 83 | 73 | 45 | 99 |  | 2 | 89 | 60 |
| 60 | 94 | 79 | 90 | 85 | 50 | 101 |  | 4 | 132 | 61 |
| 65 | 97 | 87 | 96 |  | 55 | 102 |  | 6 | -197 | 61 |
|  |  |  |  |  | 60 | 103 |  | 8 | 44 | 62 |

Table 7.4. Comparison of monetary values on TF92 and AF80: rate of interest $4 \%$.

Sample TF92 monetary functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{\|x\|}$ | $P_{[x]: \text { 团 }}^{1}$ | $P_{[x] 2}^{1} 25$ | $\begin{aligned} & P_{\langle x\| ; \bar{n} \mid}^{1} \bar{i} \\ & x-n ; 65 \end{aligned}$ | $t$ | ${ }_{t} V_{25}$ | , $V_{30: 351}^{1}$ | $t$ | , $V_{25: 25}^{11}$ | ${ }^{\prime} V_{45: 25}^{1}$ |
|  |  |  |  |  | 5 | 0.0254 | 0.0059 | 5 | 0.0016 | 0.0131 |
| 20 | 3.90 | 0.23 | 0.39 | 0.96 | 10 | 0.0558 | 0.0122 | 10 | 0.0030 | 0.0247 |
| 25 | 4.79 | 0.30 | 0.56 | 1.16 | 15 | 0.0921 | 0.0183 | 15 | 0.0038 | 0.0316 |
| 30 | 5.91 | 0.42 | 0.85 | 1.42 | 20 | 0.1350 | 0.0233 | 20 | 0.0033 | 0.0274 |
| 35 | 7.34 | 0.62 | 1.33 | 1.74 | 25 | 0.1852 | 0.0250 |  |  |  |
| 40 | 9.16 | 0.95 | 2.16 | 2.16 | 30 | 0.2432 | 0.0196 |  |  |  |
| 45 | 11.50 | 1.52 | 3.53 | 2.68 | 35 | 0.3092 |  | $t$ | $1000, V_{25: 10}^{\prime}$ | ${ }_{t} V_{45: 10}^{1}$ |
| 50 | 14.55 | 2.49 | 5.82 | 3.34 | 40 | 0.3824 |  |  |  |  |
| 55 | 18.56 | 4.12 | 9.55 | 4.12 | 45 | 0.4615 |  | 2 | 0.1318 | 0.0011 |
| 60 | 23.90 | 6.89 | 15.43 | 4.72 | 50 | 0.5439 |  | 4 | 0.2173 | 0.0018 |
| 65 | 31.08 | 11.53 | 24.28 |  | 55 | 0.6262 |  | 6 | 0.2395 | 0.0020 |
|  |  |  |  |  | 60 | 0.7043 |  | 8 | 0.1759 | 0.0015 |

TF92 monetary functions as a percentage of the equivalent AF80 functions

| Premium rates per $£ 1,000$ sum assured |  |  |  |  | Policy values |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $x$ | $P_{[x]}$ | $P_{1 \times}^{1}: 100$ | $P_{[x!25]}^{1}$ |  | $t$ | ${ }_{4} V_{25}$ | , $V_{30}^{1}$.35 | $t$ | ${ }_{1} V_{25: \overline{25}}^{1}$ | ${ }_{1} V_{45: 235}^{3}$ |
|  |  |  |  |  | 5 | 90 | 73 | 5 | 73 | 75 |
| 20 | 91 | 86 | 87 | 79 | 10 | 90 | 73 | 10 | 72 | 76 |
| 25 | 91 | 91 | 84 | 79 | 15 | 91 | 73 | 15 | 72 | 76 |
| 30 | 90 | 86 | 80 | 78 | 20 | 91 | 74 | 20 | 71 | 77 |
| 35 | 90 | 81 | 77 | 77 | 25 | 91 | 74 |  |  |  |
| 40 | 90 | 76 | 76 | 76 | 30 | 92 | 74 |  |  |  |
| 45 | 89 | 74 | 76 | 75 | 35 | 93 |  | $t$ |  | ${ }_{\text {, }} V_{45: \overline{10}}^{1}$ |
| 50 | 89 | 73 | 77 | 74 | 40 | 94 |  |  |  |  |
| 55 | 89 | 73 | 78 | 73 | 45 | 94 |  | 2 | 92 | 72 |
| 60 | 89 | 74 | 81 | 71 | 50 | 95 |  | 4 | 83 | 72 |
| 65 | 89 | 76 | 83 |  | 55 | 96 |  | 6 | 79 | 72 |
|  |  |  |  |  | 60 | 97 |  | 8 | 77 | 72 |


| 995\％ | $86 \varepsilon^{\circ} \mathrm{E}$ | E06 $\varepsilon^{\circ}$ | $718 . \varepsilon$ | － $0<\underline{L}$ | 995\％ | $899 \%$ | tてs ¢ | $6 \pm \varepsilon \varepsilon$ | ＋81．$¢$ | ¢8 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| LLS＇S | 6で「S | 606 S | $194 \%$ | LLSS | $67 \varepsilon^{\circ} \mathrm{S}$ | 6 L 9 S | $1 / t^{\prime} \mathrm{S}$ | 1Iて＇s | $096{ }^{\text {t }}$ | SL |  |
| sts $L$ | S6でL | L69\％ | Stsel | S6でL |  | 6 LSL | SLEL | $801 / 2$ | 2t8．9 | 59 |  |
| z988 | 9128 | 2988 | 9128 |  |  | 2S88 | 2U8 | 02S．8 | 278．8 | SS | 01 |
| Scet | LOT＇t | L6L＇t | － $290{ }^{\text {¢ }}$ | でS゙ゅ | sce＇t | L9t＇t | 0Lでも | ¢ \％${ }^{\text {t }}$ | 218£ | ¢8 |  |
| $69^{\circ} \mathrm{L}$ | 0669 | 016 L | 2692 | $695 \%$ | 0669 | $615{ }^{\circ} \mathrm{L}$ | $98 \mathrm{I}^{\circ} \mathrm{L}$ | 9LL＇9 | L88\％9 | SL |  |
| SL6\％${ }^{\text {a }}$ | $85^{\circ} 01$ | $80{ }^{\circ} \mathrm{F}$ | SL601 | 88 c 01 |  | 920 II | ャで00 | £ ¢ ${ }^{\text {coiol }}$ | 9196 | ¢9 |  |
| 58151 | ¢E8E1 | 5815 | $\varsigma ¢ 8 \underbrace{\prime}$ |  |  | $080 \downarrow$－ | IZL＇EI | 0¢でをI | 8LLCL | ¢s | $\varsigma$ |
| 29t＇s | LE1＇S | IE［＇9 | ES6 S | ZEL＇S | 290＇s | てt9 ¢ | 8SE＇S | 610 s | $80{ }^{\circ} \mathrm{H}$ | 58 |  |
| sst．01 | E08．6 | し0t 11 | S8601 | sstol | $\varepsilon 086$ | 89901 | 28000 | SLE6 6 | 12L＇8 | SL |  |
| 6EI＇8I | $691^{\circ} \mathrm{LI}$ | L68：81 | 6E1＇8I | $69 \mathrm{I}^{\circ} \mathrm{LI}$ |  | 11181 | $\angle 81^{\circ} \mathrm{LT}$ | 6＋0．91 | LL6け1 | ¢9 |  |
| $06 L \circ L$ | $61 / 29$ | $061 \cdot L Z$ | 614.92 |  |  | $Z L I \cdot L Z$ | 21097 | $95 \varsigma^{\circ}+\tau$ | 6¢1＇Ez | ¢¢ | 0 |
| 0102 | 0002 | S 661 | St61 | ¢£61 | ¢ 265 | 0 zoz | 0102 | 0002 | Z661 |  |  |
|  |  | чإ¢¢ јо лел入 |  |  |  |  |  |  |  | ${ }^{28} \downarrow$ | \％Isajazu！ <br> jo 키름 |


＇Z6TWd＇sən！T＇sə［EW＇shouolsurd SL OqEL

Sample annuity values expressed as a percentage of PML80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 104 | 107 | 110 | 113 |  |  | 112 | 115 | 112 | 115 |
|  | 65 | 104 | 108 | 112 | 115 |  | 112 | 116 | 119 | 112 | 116 |
|  | 75 | 104 | 108 | 112 | 116 | 110 | 114 | 118 | 121 | 110 | 114 |
|  | 85 | 104 | 106 | 110 | 113 | 111 | 114 | 117 | 120 | 107 | 111 |
| 5 | 55 | 103 | 104 | 106 | 108 |  |  | 107 | 109 | 107 | 109 |
|  | 65 | 103 | 106 | 108 | 111 |  | 108 | 111 | 113 | 108 | 111 |
|  | 75 | 104 | 106 | 110 | 113 | 108 | 111 | 114 | 116 | 108 | 111 |
|  | 85 | 103 | 105 | 108 | 111 | 109 | 112 | 115 | 117 | 106 | 109 |
| 10 | 55 | 102 | 103 | 104 | 105 |  |  | 104 | 105 | 104 | 105 |
|  | 65 | 102 | 104 | 106 | 108 |  | 106 | 108 | 109 | 106 | 108 |
|  | 75 | 103 | 105 | 108 | 110 | 106 | 109 | 111 | 113 | 106 | 109 |
|  | 85 | 103 | 105 | 107 | 110 | 108 | 111 | 113 | 115 | 106 | 108 |

Table 7.6. Pensioners, Males, Amounts, PMA92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.233 | 26.452 | 27.717 | 28.722 |  |  | 28.512 | 29.395 | 28.512 | 29.395 |
|  | 65 | 16.443 | 17.438 | 18.485 | 19.332 |  | 18.566 | 19.434 | 20.110 | 18.566 | 19.434 |
|  | 75 | 9.459 | 10.100 | 10.789 | 11.359 | 10.547 | 11.176 | 11.685 | 12.089 | 10.547 | 11.176 |
|  | 85 | 4.941 | 5.254 | 5.592 | 5.876 | 5.700 | 5.969 | 6.189 | 6.367 | 5.376 | 5.700 |
| 5 | 55 | 13.514 | 13.894 | 14.272 | 14.561 |  |  | 14.417 | 14.684 | 14.417 | 14.684 |
|  | 65 | 10.325 | 10.764 | 11.212 | 11.563 |  | 11.175 | 11.548 | 11.832 | 11.175 | 11.548 |
|  | 75 | 6.841 | 7.213 | 7.603 | 7.919 | 7.432 | 7.789 | 8.073 | 8.296 | 7.432 | 7.789 |
|  | 85 | 3.982 | 4.202 | 4.438 | 4.633 | 4.503 | 4.689 | 4.839 | 4.960 | 4.278 | 4.503 |
| 10 | 55 | 8.644 | 8.792 | 8.934 | 9.040 |  |  | 8.956 | 9.060 | 8.956 | 9.060 |
|  | 65 | 7.234 | 7.459 | 7.683 | 7.854 |  | 7.633 | 7.822 | 7.963 | 7.633 | 7.822 |
|  | 75 | 5.261 | 5.497 | 5.740 | 5.933 | 5.616 | 5.839 | 6.014 | 6.150 | 5.616 | 5.839 |
|  | 85 | 3.314 | 3.477 | 3.649 | 3.791 | 3.692 | 3.827 | 3.935 | 4.022 | 3.526 | 3.692 |

Sample annuity values expressed as a percentage of PMA80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 107 | 109 | 111 | 113 |  |  | 113 | 115 | 113 | 115 |
|  | 65 | 107 | 109 | 113 | 115 |  | 113 | 116 | 119 | 113 | 116 |
|  | 75 | 106 | 109 | 112 | 116 | 111 | 115 | 118 | 121 | 111 | 115 |
|  | 85 | 105 | 108 | 111 | 114 | 112 | 115 | 118 | 121 | 109 | 112 |
| 5 | 55 | 104 | 106 | 107 | 108 |  |  | 108 | 109 | 108 | 109 |
|  | 65 | 105 | 107 | 109 | 111 |  | 109 | 111 | 113 | 109 | 111 |
|  | 75 | 105 | 107 | 110 | 113 | 109 | 112 | 114 | 116 | 109 | 112 |
|  | 85 | 104 | 106 | 109 | 112 | 110 | 113 | 116 | 118 | 107 | 110 |
| 10 | 55 | 103 | 104 | 104 | 105 |  |  | 105 | 105 | 105 | 105 |
|  | 65 | 104 | 105 | 107 | 108 |  | 106 | 108 | 109 | 106 | 108 |
|  | 75 | 104 | 106 | 108 | 110 | 107 | 109 | 111 | 113 | 107 | 109 |
|  | 85 | 104 | 106 | 108 | 111 | 109 | 111 | 113 | 115 | 106 | 109 |

Table 7.7. Pensioners, Females, Lives, PFL92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 26.968 | 28.224 | 29.513 | 30.528 |  |  | 30.338 | 31.221 | 30.338 | 31.221 |
|  | 65 | 18.264 | 19.275 | 20.330 | 21.174 |  | 20.467 | 21.320 | 21.979 | 20.467 | 21.320 |
|  | 75 | 11.124 | 11.785 | 12.486 | 13.060 | 12.288 | 12.916 | 13.419 | 13.816 | 12.288 | 12.916 |
|  | 85 | 6.114 | 6.446 | 6.802 | 7.098 | 6.931 | 7.208 | 7.432 | 7.614 | 6.595 | 6.931 |
| 5 | 55 | 13.918 | 14.292 | 14.660 | 14.939 |  |  | 14.797 | 15.055 | 14.797 | 15.055 |
|  | 65 | 11.013 | 11.434 | 11.859 | 12.190 |  | 11.839 | 12.188 | 12.451 | 11.839 | 12.188 |
|  | 75 | 7.735 | 8.095 | 8.470 | 8.771 | 8.327 | 8.664 | 8.930 | 9.137 | 8.327 | 8.664 |
|  | 85 | 4.778 | 5.000 | 5.237 | 5.431 | 5.312 | 5.494 | 5.641 | 5.759 | 5.088 | 5.312 |
| 10 | 55 | 8.756 | 8.899 | 9.036 | 9.136 |  |  | 9.052 | 9.152 | 9.052 | 9.152 |
|  | 65 | 7.539 | 7.747 | 7.952 | 8.108 |  | 7.910 | 8.081 | 8.208 | 7.910 | 8.081 |
|  | 75 | 5.793 | 6.011 | 6.234 | 6.411 | 6.130 | 6.332 | 6.490 | 6.611 | 6.130 | 6.332 |
|  | 85 | 3.885 | 4.044 | 4.210 | 4.346 | 4.257 | 4.385 | 4.488 | 4.569 | 4.098 | 4.257 |

Sample annuity values expressed as a percentage of PFI 80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 100 | 102 | 105 | 107 |  |  | 107 | 109 | 107 | 109 |
|  | 65 | 99 | 102 | 105 | 108 |  | 106 | 108 | 111 | 106 | 108 |
|  | 75 | 102 | 104 | 108 | 111 | 107 | 110 | 113 | 115 | 107 | 110 |
|  | 85 | 108 | 110 | 113 | 115 | 114 | 117 | 119 | 121 | 111 | 114 |
| 5 | 55 | 100 | 101 | 103 | 104 |  |  | 103 | 104 | 103 | 104 |
|  | 65 | 99 | 101 | 103 | 105 |  | 103 | 105 | 106 | 103 | 105 |
|  | 75 | 101 | 103 | 105 | 108 | 104 | 107 | 109 | 111 | 104 | 107 |
|  | 85 | 106 | 108 | 110 | 113 | 111 | 114 | 116 | 118 | 109 | 111 |
| 10 |  | $100$ | $101$ |  | 102 |  |  | 102 | 102 | 102 | 102 |
|  | $65$ | $99$ | $100$ | 102 | 103 |  | 102 | 103 | 104 | 102 | 103 |
|  | $75$ | $100$ | $102$ | $104$ | 106 | $103$ | 105 | 106 | 108 | 103 | 105 |
|  | 85 | 105 | 107 | 109 | 111 | 110 | 112 | 113 | 115 | 107 | 110 |

Table 7.8. Pensioners, Females, Amounts, PFA92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 28.227 | 29.456 | 30.714 | 31.703 |  |  | 31.589 | 32.431 | 31.589 | 32.431 |
|  | 65 | 19.386 | 20.393 | 21.440 | 22.276 |  | 21.628 | 22.461 | 23.103 | 21.628 | 22.461 |
|  | 75 | 12.055 | 12.729 | 13.442 | 14.024 | 13.270 | 13.901 | 14.405 | 14.802 | 13.270 | 13.901 |
|  | 85 | 6.847 | 7.195 | 7.567 | 7.876 | 7.712 | 7.998 | 8.230 | 8.417 | 7.363 | 7.712 |
| 5 | 55 | 14.243 | 14.591 | 14.934 | 15.193 |  |  | 15.077 | 15.313 | 15.077 | 15.313 |
|  | 65 | 11.425 | 11.827 | 12.233 | 12.547 |  | 12.229 | 12.556 | 12.804 | 12.229 | 12.556 |
|  | 75 | 8.195 | 8.550 | 8.917 | 9.211 | 8.788 | 9.116 | 9.373 | 9.574 | 8.788 | 9.116 |
|  | 85 | 5.237 | 5.463 | 5.702 | 5.898 | 5.782 | 5.965 | 6.112 | 6.230 | 5.558 | 5.782 |
| 10 | 55 | 8.868 | 8.997 | 9.120 | 9.210 |  |  | 9.140 | 9.228 | 9.140 | 9.228 |
|  | 65 | 7.725 | 7.918 | 8.108 | 8.251 |  | 8.074 | 8.231 | 8.347 | 8.074 | 8.231 |
|  | 75 | 6.052 | 6.261 | 6.474 | 6.642 | 6.380 | 6.571 | 6.720 | 6.835 | 6.380 | 6.571 |
|  | 85 | 4.194 | 4.351 | 4.516 | 4.649 | 4.565 | 4.690 | 4.790 | 4.869 | 4.409 | 4.565 |

Sample annuity values expressed as a percentage of PFA80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 101 | 103 | 105 | 107 |  |  | 107 | 109 | 107 | 109 |
|  | 65 | 101 | 103 | 106 | 109 |  | 107 | 109 | 112 | 107 | 109 |
|  | 75 | 104 | 107 | 110 | 112 | 109 | 112 | 115 | 117 | 109 | 112 |
|  | 85 | 110 | 112 | 114 | 117 | 116 | 118 | 120 | 122 | 113 | 116 |
| 5 | 55 | 100 | 101 | 103 | 104 |  |  | 103 | 104 | 103 | 104 |
|  | 65 | 100 | 101 | 103 | 105 |  | 103 | 105 | 107 | 103 | 105 |
|  | 75 | 102 | 104 | 107 | 109 | 106 | 108 | 110 | 112 | 106 | 108 |
|  | 85 | 108 | 110 | 112 | 114 | 113 | 115 | 117 | 119 | 111 | 113 |
| 10 | 55 | 100 | 101 | 101 | 102 |  |  | 102 | 102 | 102 | 102 |
|  | 65 | 99 | 101 | 102 | 103 |  | 102 | 103 | 104 | 102 | 103 |
|  | 75 | 101 | 103 | 105 | 106 | 104 | 106 | 107 | 109 | 104 | 106 |
|  | 85 | 107 | 108 | 110 | 112 | 111 | 113 | 114 | 116 | 109 | 111 |

Table 7.9. Immediate Annuitants, Males, Lives, IML92 Ultimate.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.002 | 26.332 | 27.701 | 28.781 |  |  | 28.431 | 29.407 | 28.431 | 29.407 |
|  | 65 | 16.640 | 17.667 | 18.743 | 19.608 |  | 18.797 | 19.689 | 20.381 | 18.797 | 19.689 |
|  | 75 | 9.879 | 10.520 | 11.205 | 11.768 | 10.971 | 11.593 | 12.094 | 12.491 | 10.971 | 11.593 |
|  | 85 | 5.185 | 5.493 | 5.825 | 6.103 | 5.933 | 6.195 | 6.409 | 6.581 | 5.616 | 5.933 |
| 5 | 55 | 13.342 | 13.770 | 14.192 | 14.512 |  |  | 14.309 | 14.616 | 14.309 | 14.616 |
|  | 65 | 10.361 | 10.815 | 11.276 | 11.636 |  | 11.224 | 11.610 | 11.903 | 11.224 | 11.610 |
|  | 75 | 7.078 | 7.446 | 7.831 | 8.141 | 7.666 | 8.016 | 8.294 | 8.512 | 7.666 | 8.016 |
|  | 85 | 4.166 | 4.382 | 4.613 | 4.803 | 4.679 | 4.859 | 5.005 | 5.122 | 4.459 | 4.679 |
| 10 | 55 | 8.534 | 8.708 | 8.875 | 8.996 |  |  | 8.882 | 9.007 | 8.882 | 9.007 |
|  | 65 | 7.225 | 7.459 | 7.692 | 7.868 |  | 7.632 | 7.829 | 7.976 | 7.632 | 7.829 |
|  | 75 | 5.406 | 5.637 | 5.875 | 6.063 | 5.755 | 5.973 | 6.143 | 6.275 | 5.755 | 5.973 |
|  | 85 | 3.456 | 3.616 | 3.784 | 3.922 | 3.827 | 3.958 | 4.063 | 4.147 | 3.666 | 3.827 |

Sample annuity values expressed as a percentage of IM80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Ycar of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 105 | 107 | 110 | 112 |  |  | 112 | 114 | 112 | 114 |
|  | 65 | 104 | 107 | 110 | 113 |  | 110 | 113 | 116 | 110 | 113 |
|  | 75 | 101 | 104 | 107 | 111 | 106 | 110 | 113 | 115 | 106 | 110 |
|  | 85 | 99 | 102 | 105 | 108 | 106 | 109 | 111 | 114 | 103 | 106 |
| 5 | 55 | 103 | 105 | 106 | 108 |  |  | 107 | 108 | 107 | 108 |
|  | 65 | 103 | 105 | 107 | 109 |  | 107 | 109 | 111 | 107 | 109 |
|  | 75 | 101 | 103 | 106 | 109 | 105 | 108 | 110 | 112 | 105 | 108 |
|  | 85 | 99 | 101 | 104 | 107 | 105 | 107 | 110 | 112 | 102 | 105 |
| 10 | 55 | 102 | 103 | 104 | 105 |  |  | 104 | 105 | 104 | 105 |
|  | 65 | 103 | 104 | 106 | 107 |  | 105 | 107 | 108 | 105 | 107 |
|  | 75 | 101 | 103 | 105 | 107 | 104 | 106 | 108 | 110 | 104 | 106 |
|  | 85 | 99 | 101 | 104 | 106 | 104 | 106 | 108 | 110 | 102 | 104 |

Table 7,10. Immediate Annuitants, Males, Lives, IML92 Select.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.031 | 26.355 | 27.719 | 28.795 |  |  | 28.457 | 29.426 | 28.457 | 29.426 |
|  | 65 | 16.703 | 17.720 | 18.786 | 19.644 |  | 18.853 | 19.735 | 20.419 | 18.853 | 19.735 |
|  | 75 | 9.991 | 10.622 | 11.297 | 11.852 | 11.078 | 11.689 | 12.181 | 12.571 | 11.078 | 11.689 |
|  | 85 | 5.347 | 5.648 | 5.973 | 6.245 | 6.084 | 6.339 | 6.548 | 6.716 | 5.775 | 6.084 |
| 5 | 55 | 13.357 | 13.782 | 14.202 | 14.519 |  |  | 14.322 | 14.626 | 14.322 | 14.626 |
|  | 65 | 10.400 | 10.847 | 11.302 | 11.658 |  | 11.258 | 11.637 | 11.925 | 11.258 | 11.637 |
|  | 75 | 7.159 | 7.518 | 7.895 | 8.199 | 7.740 | 8.082 | 8.354 | 8.566 | 7.740 | 8.082 |
|  | 85 | 4.296 | 4.506 | 4.730 | 4.915 | 4.798 | 4.972 | 5.114 | 5.227 | 4.585 | 4.798 |
| 10 | 55 | 8.544 | 8.716 | 8.880 | 9.001 |  |  | 8.890 | 9.013 | 8.890 | 9.013 |
|  | 65 | 7.252 | 7.482 | 7.709 | 7.883 |  | 7.655 | 7.847 | 7.991 | 7.655 | 7.847 |
|  | 75 | 5.467 | 5.692 | 5.923 | 6.107 | 5.811 | 6.022 | 6.187 | 6.315 | 5.811 | 6.022 |
|  | 85 | 3.564 | 3.718 | 3.880 | 4.013 | 3.924 | 4.050 | 4.151 | 4.232 | 3.769 | 3.924 |

Sample annuity values expressed as a percontage of IM80

| Rate of interest \% | Age | Calendar ycar |  |  |  | Year of birth |  |  |  | Yeas of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 105 | 107 | 110 | 112 |  |  | 112 | 114 | 112 | 114 |
|  | 65 | 104 | 107 | 110 | 113 |  | 110 | 113 | 116 | 110 | 113 |
|  | 75 | 101 | 104 | 107 | 110 | 106 | 110 | 113 | 115 | 106 | 110 |
|  | 85 | 99 | 101 | 104 | 107 | 105 | 108 | 111 | 113 | 102 | 105 |
| 5 | 55 | 103 | 105 | 106 | 108 |  |  | 107 | 108 | 107 | 108 |
|  | 65 | 104 | 105 | 108 | 109 |  | 108 | 110 | 111 | 108 | 110 |
|  | 75 | 101 | 103 | 106 | 108 | 105 | 108 | 110 | 112 | 105 | 108 |
|  | 85 | 99 | 101 | 103 | 106 | 104 | 107 | 109 | 111 | 102 | 104 |
| 10 | 55 | 103 | 103 | 104 | 105 |  |  | 104 | 105 | 104 | 105 |
|  | 65 | 103 | 104 | 106 | 107 |  | 106 | 107 | 108 | 106 | 107 |
|  | 75 | 101 | 103 | 105 | 107 | 104 | 106 | 108 | 110 | 104 | 106 |
|  | 85 | 99 | 101 | 103 | 105 | 104 | 106 | 108 | 109 | 101 | 104 |

Table 7.11. Immediate Annuitants, Males, Amounts, IMA92 Ultimate.
Sample unnuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.409 | 26.729 | 28.086 | 29.155 |  |  | 28.831 | 29.791 | 28.831 | 29.791 |
|  | 65 | 17.000 | 18.024 | 19.095 | 19.956 |  | 19.166 | 20.049 | 20.734 | 19.166 | 20.049 |
|  | 75 | 10.163 | 10.807 | 11.494 | 12.058 | 11.269 | 11.891 | 12.390 | 12.786 | 11.269 | 11.891 |
|  | 85 | 5.379 | 5.691 | 6.027 | 6.308 | 6.139 | 6.403 | 6.619 | 6.792 | 5.820 | 6.139 |
| 5 | 55 | 13.459 | 13.878 | 14.291 | 14.604 |  |  | 14.411 | 14.709 | 14.411 | 14.709 |
|  | 65 | 10.508 | 10.955 | 11.409 | 11.763 |  | 11.363 | 11.741 | 12.028 | 11.363 | 11.741 |
|  | 75 | 7.234 | 7.599 | 7.981 | 8.289 | 7.821 | 8.168 | 8.443 | 8.658 | 7.821 | 8.168 |
|  | 85 | 4.299 | 4.516 | 4.747 | 4.938 | 4.815 | 4.996 | 5.141 | 5.258 | 4.595 | 4.815 |
| 10 | 55 | 8.577 | 8.746 | 8.907 | 9.025 |  |  | 8.916 | 9.036 | 8.916 | 9.036 |
|  | 65 | 7.296 | 7.525 | 7.751 | 7.923 |  | 7.696 | 7.887 | 8.030 | 7.696 | 7.887 |
|  | 75 | 5.500 | 5.728 | 5.962 | 6.147 | 5.846 | 6.060 | 6.227 | 6.356 | 5.846 | 6.060 |
|  | 85 | 3.552 | 3.711 | 3.879 | 4.016 | 3.922 | 4.052 | 4.157 | 4.240 | 3.762 | 3.922 |

Sample annuity values expressed as a percentage of IM80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 106 | 109 | 111 | 113 |  |  | 113 | 115 | 113 | 115 |
|  | 65 | 106 | 109 | 112 | 115 |  | 112 | 115 | 118 | 112 | 115 |
|  | 75 | 104 | 107 | 110 | 113 | 109 | 113 | 116 | 118 | 109 | 113 |
|  | 85 | 103 | 105 | 108 | 111 | 109 | 112 | 115 | 117 | 106 | 109 |
| 5 | 55 | 104 | 106 | 107 | 108 |  |  | 108 | 109 | 108 | 109 |
|  | 65 | 105 | 107 | 109 | 110 |  | 109 | 111 | 112 | 109 | 111 |
|  | 75 | 103 | 106 | 108 | 111 | 107 | 110 | 112 | 114 | 107 | 110 |
|  | 85 | 102 | 104 | 107 | 110 | 108 | 110 | 113 | 115 | 105 | 108 |
| 10 | 55 | 103 | 104 | 105 | 105 |  |  | 105 | 105 | 105 | 105 |
|  | 65 | 104 | 105 | 107 | 108 |  | 106 | 108 | 109 | 106 | 108 |
|  | 75 | 103 | 105 | 107 | 109 | 106 | 108 | 110 | 111 | 106 | 108 |
|  | 85 | 102 | 104 | 106 | 108 | 107 | 109 | 111 | 113 | 105 | 107 |

Table 7.12. Immediate Annuitants, Males, Amounts, IMA92 Select.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.429 | 26.745 | 28.098 | 29.164 |  |  | 28.848 | 29.803 | 28.848 | 29.803 |
|  | 65 | 17.042 | 18.059 | 19.124 | 19.980 |  | 19.204 | 20.080 | 20.759 | 19.204 | 20.080 |
|  | 75 | 10.239 | 10.877 | 11.557 | 12.116 | 11.341 | 11.955 | 12.449 | 12.840 | 11.341 | 11.955 |
|  | 85 | 5.490 | 5.797 | 6.128 | 6.405 | 6.242 | 6.502 | 6.713 | 6.884 | 5.928 | 6.242 |
| 5 | 55 | 13.469 | 13.886 | 14.297 | 14.608 |  |  | 14.419 | 14.715 | 14.419 | 14.715 |
|  | 65 | 10.534 | 10.977 | 11.427 | 11.777 |  | 11.386 | 11.759 | 12.043 | 11.386 | 11.759 |
|  | 75 | 7.288 | 7.648 | 8.025 | 8.328 | 7.872 | 8.213 | 8.483 | 8.694 | 7.872 | 8.213 |
|  | 85 | 4.387 | 4.600 | 4.827 | 5.014 | 4.896 | 5.072 | 5.215 | 5.329 | 4.680 | 4.896 |
| 10 | 55 | 8.584 | 8.752 | 8.911 | 9.028 |  |  | 8.921 | 9.040 | 8.921 | 9.040 |
|  | 65 | 7.315 | 7.540 | 7.763 | 7.933 |  | 7.711 | 7.899 | 8.039 | 7.711 | 7.899 |
|  | 75 | $5.541$ | 5.765 | 5.995 | 6.176 | 5.884 | 6.093 | 6.256 | 6.383 | 5.884 | 6.093 |
|  | 85 | 3.625 | 3.780 | 3.944 | 4.078 | 3.988 | 4.115 | 4.216 | 4.297 | 3.832 | 3.988 |

Sample annuity values expressed as a percentage of IM80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 106 | 109 | 111 | 113 |  |  | 113 | 115 | 113 | 115 |
|  | 65 | 106 | 109 | 112 | 115 |  | 113 | 115 | 118 | 113 | 115 |
|  | 75 | 104 | 107 | 110 | 113 | 109 | 112 | 115 | 118 | 109 | 112 |
|  | 85 | 102 | 104 | 107 | 110 | 108 | 111 | 113 | 116 | 105 | 108 |
| 5 | 55 | 104 | 106 | 107 | 108 |  |  | 108 | 109 | 108 | 109 |
|  | 65 | 105 | 107 | 109 | 111 |  | 109 | 111 | 112 | 109 | 111 |
|  | 75 | 103 | 105 | 108 | 110 | 107 | 109 | 112 | 114 | 107 | 109 |
|  | 85 | 101 | 103 | 106 | 108 | 107 | 109 | 111 | 113 | 104 | 107 |
| 10 | 55 | 103 | 104 | 105 | 105 |  |  | 105 | 105 | 105 | 105 |
|  | 65 | 104 | 105 | 107 | 108 |  | 106 | 108 | 109 | 106 | 108 |
|  | 75 | 103 | 104 | 106 | 108 | 105 | 107 | 109 | 111 | 105 | 107 |
|  | 85 | 101 | 102 | 105 | 107 | 105 | 107 | 109 | 111 | 103 | 105 |

Table 7.13. Immediate Annuitants, Females, Lives, IFL92 Ultimate.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 28.981 | 30.042 | 31.124 | 31.972 |  |  | 31.944 | 32.649 | 31.944 | 32.649 |
|  | 65 | 19.859 | 20.747 | 21.667 | 22.399 |  | 21.875 | 22.594 | 23.148 | 21.875 | 22.594 |
|  | 75 | 12.009 | 12.620 | 13.264 | 13.789 | 13.115 | 13.683 | 14.136 | 14.492 | 13.115 | 13.683 |
|  | 85 | 6.251 | 6.566 | 6.902 | 7.181 | 7.026 | 7.287 | 7.498 | 7.668 | 6.710 | 7.026 |
| 5 | 55 | 14.546 | 14.841 | 15.130 | 15.349 |  |  | 15.272 | 15.466 | 15.272 | 15.466 |
|  | 65 | 11.723 | 12.076 | 12.431 | 12.705 |  | 12.446 | 12.728 | 12.941 | 12.446 | 12.728 |
|  | 75 | 8.267 | 8.593 | 8.931 | 9.201 | 8.819 | 9.119 | 9.355 | 9.538 | 8.819 | 9.119 |
|  | 85 | 4.903 | 5.116 | 5.340 | 5.525 | 5.413 | 5.586 | 5.725 | 5.836 | 5.201 | 5.413 |
| 10 | 55 | 9.000 | 9.106 | 9.207 | 9.280 |  |  | 9.229 | 9.300 | 9.229 | 9.300 |
|  | 65 | 7.903 | 8.069 | 8.233 | 8.356 |  | 8.213 | 8.346 | 8.444 | 8.213 | 8.346 |
|  | 75 | 6.136 | 6.330 | 6.527 | 6.682 | 6.444 | 6.620 | 6.757 | 6.863 | 6.444 | 6.620 |
|  | 85 | 3.993 | 4.144 | 4.303 | 4.432 | 4.349 | 4.470 | 4.567 | 4.645 | 4.198 | 4.349 |

Sample annuity values expressed as a percentage of IF80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 103 | 105 | 107 | 108 |  |  | 108 | 110 | 108 | 110 |
|  | 65 | 103 | 105 | 107 | 109 |  | 108 | 110 | 112 | 108 | 110 |
|  | 75 | 102 | 104 | 107 | 109 | 106 | 109 | 111 | 113 | 106 | 109 |
|  | 85 | 101 | 103 | 106 | 108 | 107 | 109 | 111 | 113 | 104 | 107 |
| 5 | 55 | 102 | 103 | 104 | 105 |  |  | 105 | 105 | 105 | 105 |
|  | 65 | 102 | 103 | 105 | 106 |  | 105 | 107 | 108 | 105 | 107 |
|  | 75 | 102 | 103 | 105 | 107 | 105 | 107 | 108 | 110 | 105 | 107 |
|  | 85 | 101 | 103 | 105 | 107 | 106 | 108 | 110 | 111 | 104 | 106 |
| 10 | 55 | 102 | 102 | 103 | 103 |  |  | 103 | 103 | 103 | 103 |
|  | 65 | 102 | 103 | 104 | 105 |  | 104 | 105 | 105 | 104 | 105 |
|  | 75 | 101 | 103 | 104 | 106 | 104 | 105 | 106 | 108 | 104 | 105 |
|  | 85 | 101 | 103 | 104 | 106 | 105 | 107 | 109 | 110 | 103 | 105 |

Table 7.14. Immediate Annuitants, Females, Lives, IFL92 Select.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 28.997 | 30.055 | 31.133 | 31.979 |  |  | 31.957 | 32.659 | 31.957 | 32.659 |
|  | 65 | 19.899 | 20.781 | 21.694 | 22.421 |  | 21.910 | 22.622 | 23.171 | 21.910 | 22.622 |
|  | 75 | 12.098 | 12.700 | 13.336 | 13.854 | 13.199 | 13.757 | 14.202 | 14.554 | 13.199 | 13.757 |
|  | 85 | 6.404 | 6.711 | 7.040 | 7.313 | 7.167 | 7.420 | 7.626 | 7.792 | 6.858 | 7.167 |
| 5 | 55 | 14.554 | 14.847 | 15.135 | 15.353 |  |  | 15.278 | 15.471 | 15.278 | 15.471 |
|  | 65 | 11.747 | 12.095 | 12.446 | 12.718 |  | 12.466 | 12.744 | 12.954 | 12.466 | 12.744 |
|  | 75 | 8.328 | 8.648 | 8.980 | 9.245 | 8.875 | 9.168 | 9.399 | 9.578 | 8.875 | 9.168 |
|  | 85 | 5.023 | 5.229 | 5.447 | 5.626 | 5.522 | 5.689 | 5.823 | 5.931 | 5.316 | 5.522 |
| 10 | 55 | 9.005 | 9.110 | 9.209 | 9.282 |  |  | 9.233 | 9.303 | 9.233 | 9.303 |
|  | 65 | 7.919 | 8.082 | 8.243 | 8.365 |  | 8.226 | 8.356 | 8.453 | 8.226 | 8.356 |
|  | 75 | 6.182 | 6.370 | 6.563 | 6.714 | 6.485 | 6.656 | 6.789 | 6.892 | 6.485 | 6.656 |
|  | 85 | 4.090 | 4.236 | 4.389 | 4.514 | 4.436 | 4.552 | 4.646 | 4.720 | 4.291 | 4.436 |

Sample annuity values expressed as a percentage of IF80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 103 | 105 | 107 | 108 |  |  | 108 | 110 | 108 | 110 |
|  | 65 | 103 | 105 | 107 | 109 |  | 108 | 110 | 112 | 108 | 110 |
|  | 75 | 102 | 104 | 107 | 109 | 106 | 109 | 111 | 113 | 106 | 109 |
|  | 85 | 102 | 104 | 106 | 109 | 107 | 110 | 112 | 114 | 105 | 107 |
| 5 | 55 | 102 | 103 | 104 | 105 |  |  | 105 | 105 | 105 | 105 |
|  | 65 | 102 | 103 | 105 | 106 |  | 105 | 106 | 108 | 105 | 106 |
|  | 75 | 102 | 103 | 105 | 107 | 105 | 107 | 108 | 110 | 105 | 107 |
|  | 85 | 102 | 103 | 105 | 107 | 106 | 108 | 110 | 112 | 104 | 106 |
| 10 | 55 | 102 | 102 | 102 | 103 |  |  | 103 | 103 | 103 | 103 |
|  | 65 | 102 | 103 | 104 | 104 |  | 104 | 104 | 105 | 104 | 104 |
|  | 75 | 101 | 103 | 104 | 106 | 104 | 105 | 106 | 108 | 104 | 105 |
|  | 85 | 102 | 103 | 105 | 107 | 106 | 107 | 109 | 110 | 104 | 106 |

Table 7.15. Immediate Annuitants, Females, Amounts, IFA92 Ultimate.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 30.029 | 30.942 | 31.876 | 32.611 |  |  | 32.700 | 33.284 | 32.700 | 33.284 |
|  | 65 | 20.589 | 21.390 | 22.218 | 22.879 |  | 22.474 | 23.108 | 23.597 | 22.474 | 23.108 |
|  | 75 | 12.256 | 12.838 | 13.452 | 13.951 | 13.327 | 13.865 | 14.293 | 14.630 | 13.327 | 13.865 |
|  | 85 | 6.146 | 6.454 | 6.784 | 7.058 | 6.905 | 7.160 | 7.368 | 7.535 | 6.595 | 6.905 |
| 5 | 55 | 14.901 | 15.140 | 15.375 | 15.555 |  |  | 15.523 | 15.674 | 15.523 | 15.674 |
|  | 65 | 12.076 | 12.384 | 12.694 | 12.935 |  | 12.734 | 12.975 | 13.157 | 12.734 | 12.975 |
|  | 75 | 8.438 | 8.747 | 9.066 | 9.322 | 8.969 | 9.251 | 9.472 | 9.644 | 8.969 | 9.251 |
|  | 85 | 4.844 | 5.053 | 5.274 | 5.456 | 5.346 | 5.517 | 5.654 | 5.764 | 5.137 | 5.346 |
| 10 | 55 | 9.146 | 9.227 | 9.303 | 9.360 |  |  | 9.331 | 9.382 | 9.331 | 9.382 |
|  | 65 | 8.094 | 8.235 | 8.373 | 8.477 |  | 8.367 | 8.477 | 8.558 | 8.367 | 8.477 |
|  | 75 | 6.257 | 6.439 | 6.623 | 6.769 | 6.551 | 6.715 | 6.842 | 6.941 | 6.551 | 6.715 |
|  | 85 | 3.958 | 4.107 | 4.264 | 4.392 | 4.310 | 4.430 | 4.526 | 4.603 | 4.161 | 4.310 |

Sample annuity values expressed as a percentage of IF80

| Rate of interest \% | Agc | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 107 | 108 | 109 | 110 |  |  | 111 | 112 | 111 | 112 |
|  | 65 | 107 | 108 | 110 | 111 |  | 111 | 112 | 114 | 111 | 112 |
|  | 75 | 104 | 106 | 108 | 110 | 108 | 110 | 112 | 114 | 108 | 110 |
|  | 85 | 100 | 101 | 104 | 106 | 105 | 107 | 110 | 111 | 102 | 105 |
| 5 | 55 | 105 | 105 | 106 | 106 |  |  | 106 | 107 | 106 | 107 |
|  | 65 | 105 | 106 | 107 | 108 |  | 108 | 109 | 109 | 108 | 109 |
|  | 75 | 104 | 105 | 107 | 108 | 106 | 108 | 110 | 111 | 106 | 108 |
|  | 85 | 100 | 102 | 104 | 106 | 105 | 107 | 108 | 110 | 102 | 105 |
| 10 | 55 | 103 | 103 | 104 | 104 |  |  | 104 | 104 | 104 | 104 |
|  | 65 | 104 | 105 | 105 | 106 |  | 106 | 106 | 107 | 106 | 106 |
|  | 75 | 103 | 104 | 106 | 107 | 105 | 107 | 108 | 109 | 105 | 107 |
|  | 85 | 100 | 102 | 104 | 105 | 104 | 106 | 108 | 109 | 102 | 104 |

Table 7.16. Immediate Annuitants, Females, Amounts, IFA92 Select.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar ycar |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 30.035 | 30.947 | 31.879 | 32.613 |  |  | 32.705 | 33.288 | 32.705 | 33.288 |
|  | 65 | 20.606 | 21.404 | 22.230 | 22.889 |  | 22.489 | 23.120 | 23.607 | 22.489 | 23.120 |
|  | 75 | 12.304 | 12.882 | 13.490 | 13.986 | 13.372 | 13.904 | 14.328 | 14.663 | 13.372 | 13.904 |
|  | 85 | 6.239 | 6.542 | 6.868 | 7.138 | 6.990 | 7.242 | 7.446 | 7.610 | 6.685 | 6.990 |
| 5 | 55 | 14.904 | 15.142 | 15.377 | 15.556 |  |  | 15.525 | 15.676 | 15.525 | 15.676 |
|  | 65 | 12.086 | 12.392 | 12.701 | 12.940 |  | 12.742 | 12.981 | 13.162 | 12.742 | 12.981 |
|  | 75 | 8.471 | 8.776 | 9.092 | 9.345 | 8.999 | 9.277 | 9.496 | 9.666 | 8.999 | 9.277 |
|  | 85 | 4.917 | 5.122 | 5.339 | 5.518 | 5.412 | 5.579 | 5.714 | 5.821 | 5.207 | 5.412 |
| 10 | 55 | 9.148 | 9.228 | 9.304 | 9.360 |  |  | 9.332 | 9.383 | 9.332 | 9.383 |
|  | 65 | 8.101 | 8.240 | 8.377 | 8.481 |  | 8.373 | 8.481 | 8.562 | 8.373 | 8.481 |
|  | $75$ | 6.282 | 6.460 | 6.642 | 6.786 | 6.573 | 6.734 | 6.860 | 6.956 | 6.573 | 6.734 |
|  | 85 | 4.017 | 4.163 | 4.317 | 4.442 | 4.363 | 4.480 | 4.574 | 4.649 | 4.218 | 4.363 |

Sample annuity values expressed as a percentage of IF80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Ycar of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 107 | 108 | 109 | 110 |  |  | 111 | 112 | 111 | 112 |
|  | 65 | 106 | 108 | 110 | 111 |  | 111 | 112 | 114 | 111 | 112 |
|  | 75 | 104 | 105 | 108 | 110 | 108 | 110 | 112 | 114 | 108 | 110 |
|  | 85 | 99 | 101 | 104 | 106 | 105 | 107 | 109 | 111 | 102 | 105 |
| 5 | 55 | 105 | 105 | 106 | 106 |  |  | 106 | 107 | 106 | 107 |
|  | 65 | 105 | 106 | 107 | 108 |  | 107 | 108 | 109 | 107 | 108 |
|  | $75$ | 103 | 105 | 107 | 108 | 106 | 108 | 109 | 111 | 106 | 108 |
|  | 85 | 100 | 101 | 103 | 105 | 104 | 106 | 108 | 109 | 102 | 104 |
| 10 | 55 | 103 | 103 | 104 | 104 |  |  | 104 | 104 | 104 | 104 |
|  | $65$ | 104 | 105 | 105 | 106 |  | 105 | 106 | 107 | 105 | 106 |
|  | 75 | 103 | 104 | 106 | 107 | 105 | 106 | 108 | 109 | 105 | 106 |
|  | 85 | 100 | 101 | 103 | 105 | 104 | 106 | 107 | 108 | 102 | 104 |

Table 7.17. Retirement Annuitants, Males, Vested, RMV92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 23.328 | 25.082 | 26.877 | 28.277 |  |  | 27.303 | 28.695 | 27.303 | 28.695 |
|  | 65 | 16.962 | 18.059 | 19.206 | 20.126 |  | 19.237 | 20.192 | 20.931 | 19.237 | 20.192 |
|  | 75 | 10.528 | 11.195 | 11.905 | 12.488 | 11.681 | 12.322 | 12.837 | 13.244 | 11.681 | 12.322 |
|  | 85 | 5.818 | 6.144 | 6.495 | 6.788 | 6.618 | 6.892 | 7.115 | 7.295 | 6.286 | 6.618 |
| 5 | 55 | 12.429 | 13.058 | 13.673 | 14.132 |  |  | 13.646 | 14.137 | 13.646 | 14.137 |
|  | 65 | 10.403 | 10.885 | 11.372 | 11.751 |  | 11.300 | 11.711 | 12.021 | 11.300 | 11.711 |
|  | 75 | 7.403 | 7.775 | 8.163 | 8.474 | 8.003 | 8.355 | 8.633 | 8.849 | 8.003 | 8.355 |
|  | 85 | 4.582 | 4.805 | 5.041 | 5.235 | 5.113 | 5.296 | 5.444 | 5.563 | 4.889 | 5.113 |
| 10 | 55 | 7.991 | 8.282 | 8.557 | 8.756 |  |  | 8.484 | 8.711 | 8.484 | 8.711 |
|  | 65 | 7.203 | 7.452 | 7.698 | 7.885 |  | 7.625 | 7.836 | 7.992 | 7.625 | 7.836 |
|  | 75 | 5.588 | 5.817 | 6.052 | 6.238 | 5.936 | 6.151 | 6.318 | 6.447 | 5.936 | 6.151 |
|  | 85 | 3.748 | 3.908 | 4.076 | 4.213 | 4.121 | 4.251 | 4.355 | 4.438 | 3.961 | 4.121 |

Sample annuity values expressed as a percentage of PMLSO

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 105 | 109 | 114 | 118 |  |  | 115 | 119 | 115 | 119 |
|  | 65 | 118 | 121 | 125 | 128 |  | 125 | 129 | 131 | 125 | 129 |
|  | 75 | 126 | 129 | 132 | 135 | 131 | 135 | 138 | 141 | 131 | 135 |
|  | 85 | 128 | 130 | 133 | 136 | 135 | 138 | 140 | 143 | 131 | 135 |
| 5 | 55 | 100 | 103 | 106 | 108 |  |  | 106 | 108 | 106 | 108 |
|  | 65 | 112 | 114 | 116 | 118 |  | 116 | 118 | 120 | 116 | 118 |
|  | 75 | 120 | 122 | 124 | 127 | 124 | 126 | 128 | 130 | 124 | 126 |
|  | 85 | 124 | 126 | 128 | 130 | 129 | 131 | 134 | 136 | 127 | 129 |
| 10 | 55 | 98 | 100 | 102 | 104 |  |  | 102 | 103 | 102 | 103 |
|  | 65 | 108 | 109 | 111 | 112 |  | 111 | 112 | 113 | 111 | 112 |
|  | 75 | 116 | 118 | 119 | 121 | 119 | 120 | 122 | 124 | 119 | 120 |
|  | 85 | 121 | 122 | 124 | 126 | 125 | 127 | 129 | 131 | 123 | 125 |

Table 7.18. Retirement Annuitants, Females, Vested, RFV92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 27.977 | 29.357 | 30.742 | 31.808 |  |  | 31.350 | 32.338 | 31.350 | 32.338 |
|  | 65 | 20.213 | 21.154 | 22.122 | 22.889 |  | 22.295 | 23.057 | 23.642 | 22.295 | 23.057 |
|  | 75 | 12.689 | 13.305 | 13.951 | 14.474 | 13.811 | 14.377 | 14.826 | 15.178 | 13.811 | 14.377 |
|  | 85 | 6.817 | 7.132 | 7.468 | 7.745 | 7.598 | 7.855 | 8.063 | 8.230 | 7.284 | 7.598 |
| 5 | 55 | 13.972 | 14.408 | 14.827 | 15.137 |  |  | 14.877 | 15.191 | 14.877 | 15.191 |
|  | 65 | 11.769 | 12.145 | 12.520 | 12.809 |  | 12.515 | 12.817 | 13.043 | 12.515 | 12.817 |
|  | 75 | 8.596 | 8.918 | 9.251 | 9.516 | 9.144 | 9.438 | 9.668 | 9.846 | 9.144 | 9.438 |
|  | 85 | 5.281 | 5.490 | 5.711 | 5.891 | 5.786 | 5.954 | 6.089 | 6.196 | 5.579 | 5.786 |
| 10 | 55 | 8.662 | 8.845 | 9.016 | 9.138 |  |  | 8.989 | 9.124 | 8.989 | 9.124 |
|  | 65 | 7.881 | 8.061 | 8.237 | 8.369 |  | 8.204 | 8.349 | 8.455 | 8.204 | 8.349 |
|  | 75 | 6.312 | 6.501 | 6.692 | 6.842 | 6.613 | 6.784 | 6.916 | 7.017 | 6.613 | 6.784 |
|  | 85 | 4.259 | 4.406 | 4.560 | 4.684 | 4.606 | 4.722 | 4.815 | 4.889 | 4.461 | 4.606 |

Sample annuity values expressed as a percentage of PFISO

| Rate of intercst \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 103 | 106 | 109 | 111 |  |  | 110 | 113 | 110 | 113 |
|  | 65 | 110 | 112 | 114 | 117 |  | 115 | 117 | 119 | 115 | 117 |
|  | 75 | 116 | 118 | 120 | 123 | 120 | 122 | 124 | 126 | 120 | 122 |
|  | 85 | 120 | 121 | 124 | 126 | 125 | 127 | 129 | 131 | 122 | 125 |
| 5 | 55 | 100 | 102 | 104 | 105 |  |  | 104 | 105 | 104 | 105 |
|  | 65 | 106 | 107 | 109 | 110 |  | 109 | 110 | 111 | 109 | 110 |
|  | 75 | 112 | 113 | 115 | 117 | 114 | 116 | 118 | 119 | 114 | 116 |
|  | 85 | 117 | 119 | 120 | 122 | 121 | 123 | 125 | 126 | 119 | 121 |
| 10 | 55 | 99 | 100 | 101 | 102 |  |  | 101 | 102 | 101 | 102 |
|  | 65 | 103 | 104 | 106 | 106 |  | 105 | 106 | 107 | 105 | 106 |
|  | 75 | 109 | 110 | 111 | 113 | 111 | 112 | 113 | 115 | 111 | 112 |
|  | 85 | 115 | 116 | 118 | 119 | 119 | 120 | 122 | 123 | 117 | 119 |

Table 7.19. Widows, Lives, WL92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 25.968 | 27.477 | 29.019 | 30.223 |  |  | 29.733 | 30.841 | 29.733 | 30.841 |
|  | 65 | 18.159 | 19.264 | 20.412 | 21.328 |  | 20.486 | 21.426 | 22.151 | 20.486 | 21.426 |
|  | 75 | 11.622 | 12.302 | 13.020 | 13.606 | 12.819 | 13.460 | 13.971 | 14.375 | 12.819 | 13.460 |
|  | 85 | 6.664 | 6.999 | 7.357 | 7.653 | 7.492 | 7.768 | 7.991 | 8.170 | 7.156 | 7.492 |
| 5 | 55 | 13.422 | 13.902 | 14.371 | 14.724 |  |  | 14.459 | 14.806 | 14.459 | 14.806 |
|  | 65 | 10.853 | 11.318 | 11.788 | 12.151 |  | 11.733 | 12.123 | 12.417 | 11.733 | 12.123 |
|  | 75 | 7.960 | 8.326 | 8.705 | 9.007 | 8.559 | 8.899 | 9.167 | 9.375 | 8.559 | 8.899 |
|  | 85 | 5.140 | 5.362 | 5.596 | 5.787 | 5.673 | 5.852 | 5.996 | 6.111 | 5.453 | 5.673 |
| 10 | 55 | 8.490 | 8.688 | 8.877 | 9.013 |  |  | 8.866 | 9.011 | 8.866 | 9.011 |
|  | 65 | 7.408 | 7.643 | 7.873 | 8.048 |  | 7.811 | 8.006 | 8.151 | 7.811 | 8.006 |
|  | 75 | 5.905 | 6.124 | 6.348 | 6.524 | 6.242 | 6.445 | 6.602 | 6.723 | 6.242 | 6.445 |
|  | 85 | 4.139 | 4.294 | 4.456 | 4.588 | 4.504 | 4.628 | 4.727 | 4.805 | 4.350 | 4.504 |

Sample annuity values expressed as a percentage of WLSO

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 103 | 106 | 110 | 112 |  |  | 112 | 114 | 112 | 114 |
|  | 65 | 106 | 109 | 112 | 115 |  | 112 | 116 | 118 | 112 | 116 |
|  | 75 | 110 | 113 | 116 | 119 | 115 | 118 | 121 | 124 | 115 | 118 |
|  | 85 | 120 | 122 | 125 | 127 | 126 | 129 | 131 | 133 | 123 | 126 |
| 5 | 55 | 101 | 103 | 105 | 107 |  |  | 105 | 107 | 105 | 107 |
|  | 65 | 103 | 105 | 107 | 109 |  | 107 | 109 | 111 | 107 | 109 |
|  | 75 | 107 | 109 | 111 | 114 | 110 | 113 | 115 | 117 | 110 | 113 |
|  | 85 | 117 | 118 | 120 | 123 | 121 | 124 | 125 | 127 | 119 | 121 |
| 10 | 55 | 100 | 102 | 103 | 104 |  |  | 103 | 104 | 103 | 104 |
|  | 65 | 102 | 103 | 105 | 106 |  | 104 | 106 | 107 | 104 | 106 |
|  | 75 | 105 | 106 | 108 | $110$ | $107$ | 109 | 111 | 112 | 107 | 109 |
|  | 85 | 114 | 115 | 117 | 119 | 118 | 120 | 121 | 123 | 116 | 118 |

Table 7.20. Widows, Amounts, WA92.
Sample annuity values using different types of projected table

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 27.040 | 28.484 | 29.952 | 31.096 |  |  | 30.698 | 31.733 | 30.698 | 31.733 |
|  | 65 | 19.006 | 20.080 | 21.193 | 22.077 |  | 21.305 | 22.204 | 22.895 | 21.305 | 22.204 |
|  | 75 | 12.208 | 12.881 | 13.590 | 14.166 | 13.410 | 14.037 | 14.536 | 14.929 | 13.410 | 14.037 |
|  | 85 | 7.000 | 7.336 | 7.694 | 7.991 | 7.834 | 8.109 | 8.331 | 8.509 | 7.499 | 7.834 |
| 5 | 55 | 13.746 | 14.190 | 14.624 | 14.948 |  |  | 14.721 | 15.036 | 14.721 | 15.036 |
|  | 65 | 11.196 | 11.638 | 12.081 | 12.424 |  | 12.043 | 12.408 | 12.682 | 12.043 | 12.408 |
|  | 75 | 8.268 | 8.623 | 8.990 | 9.283 | 8.858 | 9.185 | 9.442 | 9.641 | 8.858 | 9.185 |
|  | 85 | 5.359 | 5.579 | 5.810 | 5.999 | 5.889 | 6.065 | 6.207 | 6.320 | 5.672 | 5.889 |
| 10 | 55 | 8.619 | 8.798 | 8.968 | 9.091 |  |  | 8.963 | 9.092 | 8.963 | 9.092 |
|  | 65 | 7.576 | 7.794 | 8.008 | 8.169 |  | 7.955 | 8.134 | 8.267 | 7.955 | 8.134 |
|  | 75 | 6.086 | 6.296 | 6.510 | 6.677 | 6.413 | 6.605 | 6.754 | 6.869 | 6.413 | 6.605 |
|  | 85 | 4.290 | 4.443 | 4.602 | 4.732 | 4.650 | 4.771 | 4.868 | 4.944 | 4.500 | 4.650 |

Sample annuity values expressed as a percentage of W A80

| Rate of interest \% | Age | Calendar year |  |  |  | Year of birth |  |  |  | Year of use |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1992 | 2000 | 2010 | 2020 | 1925 | 1935 | 1945 | 1955 | 2000 | 2010 |
| 0 | 55 | 103 | 105 | 109 | 111 |  |  | 110 | 113 | 110 | 113 |
|  | 65 | 106 | 109 | 112 | 115 |  | 112 | 115 | 118 | 112 | 115 |
|  | 75 | 112 | 114 | 117 | 120 | 116 | 119 | 122 | 124 | 116 | 119 |
|  | 85 | 120 | 122 | 124 | 126 | 125 | 128 | 130 | 132 | 123 | 125 |
| 5 | 55 | 100 | 102 | 104 | 105 |  |  | 104 | 106 | 104 | 106 |
|  | 65 | 103 | 105 | 107 | 109 |  | 107 | 109 | 110 | 107 | 109 |
|  | 75 | 108 | 110 | 112 | 114 | 111 | 114 | 116 | 117 | 111 | 114 |
|  | 85 | 117 | 118 | 120 | 122 | 121 | 123 | 125 | 126 | 119 | 121 |
| 10 | 55 | 100 | 101 | 102 | 103 |  |  | 102 | 103 | 102 | 103 |
|  | 65 | 101 | 103 | 104 | 106 |  | 104 | 105 | 106 | 104 | 105 |
|  | 75 | 106 | 107 | 109 | 111 | 108 | 110 | 111 | 113 | 108 | 110 |
|  | 85 | 114 | 115 | 117 | 119 | 118 | 119 | 121 | 122 | 116 | 118 |

### 8.1 Introduction

8.1.1 Prior to the " 80 " Series, new tables had traditionally been published by means of a number of printed volumes containing commutation and monetary functions over a wide range of interest rates, in addition to mortality functions. 8.1.2 When the " 80 " Series of tables was published the Committee, supported by comments received from members of the profession, broke with tradition and decided that due to the widespread use of personal computers the publication of printed volumes of commutation and monetary functions was no longer necessary. Instead, it was decided that a computer package should be made available which would allow the user to generate the basic mortality functions and a wide range of commutation and monetary functions, at any reasonable rate of interest, for any of the new standard tables or, in the case of joint-life functions, a combination of any two of the new standard tables. This computer package is known as the Standard Tables Program.

### 8.2 The "92" Series

8.2.1 Following the favourable reception of the Standard Tables Program and bearing in mind the other information needs of the profession, the Committee is of the view that the new tables should be published by means of:
(a) a single printed volume that contains the basic mortality functions ( $q, \mu$ and $l$ ) for each of the new standard tables for assurances and the base tables for pensioners, immediate annuitants, retirement annuitants and widows, together with corresponding values of $q$ for selected projected tables, as well as examples of monetary functions at only one or two rates of interest for each of the tables;
(b) an updated Windows version of the Standard Tables Program, Version 3.0, containing the new " 92 " Series tables (as well as the " 80 " Series and earlier tables); and
(c) placing the word-processing files that comprise this edition of C.M.I.R. onto the Faculty of Actuaries and Institute of Actuaries web site (www.actuaries.org.uk).
8.2.2 It is the intention of the Committee that future publications will also be put onto the web site so that information can be quickly made available to the profession and other interested parties.
8.2.3 The reader unfamiliar with the Standard Tables Program is referred to C.M.I.R. 10, 86 where a full description of its capabilities is given.

### 8.3 Publication date

8.3.1 The official launch date of the new standard tables was 30 June 1999 when the values of $q$ were put onto the web site. The printed volume, entitled "Standard Tables of Mortality - The " 92 " Series", and Version 3.0 of the Standard Tables Program were available for purchase from July 1999.

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## APPENDIXA

## VALUES OF MORTALITY RATESFOR THE NEW STANDARD TABLES

Basic tables for 1991-94 experience, and projected tables for calendar year 2020 and years of birth 1945 and 1972.
TABLE Page
A1. Permanent assurances, males - AM92 ..... 151
A2. Permanent assurances, females - AF92 ..... 154
A3. Temporary assurances, males - TM92 ..... 157
A4. Temporary assurances, females - TF92 ..... 160
A5. Pensioners - PML92Base, PMA92Base, PFL92Base and PFA92Base ..... 163
A6. Immediate annuitants, males - IML92Base and IMA92Base ..... 166
A7. Immediate annuitants, females - IFL92Base and IFA92Base ..... 169
A8. Retirement annuitants - RMV92Base and RFV92Base ..... 172
A9. Widows - WL92Base and WA92Base ..... 175
A10. Pensioners projected - PML92, PMA92, PFL92 and PFA92 ( $C=2020$ ) ..... 178
A11. Immediate annuitants projected, males - IML92 and IMA92 ( $C=2020$ ) ..... 181
A12. Immediate annuitants projected, females - IFL92 and IFA92 ( $C=2020$ ) ..... 184
A13. Retirement annuitants projected - RMV92 and RFV92 ( $C=2020$ ) ..... 187
A14. Widows projected - WL92 and WA92 $(C=2020)$ ..... 190
A15. Pensioners projected - PML92, PMA92, PFL92 and PFA92 ( $B=1945$ ) ..... 193
A16. Immediate annuitants projected, males - IML92 and IMA92 ( $B=1945$ ) ..... 195
A17. Immediate annuitants projected, females - IFL92 and IFA92 ( $B=1945$ ) ..... 197
A18. Retirement annuitants projected - RMV92 and RFV92 ( $B=1945$ ) ..... 199
A19. Widows projected - WL92 and WA92 $(B=1945)$ ..... 201
A20. Pensioners projected - PML92, PMA92, PFL92 and PFA92 ( $B=1972$ ) ..... 203
A21. Immediate annuitants projected, males - IML92 and IMA92 ( $B=1972$ ) ..... 206
A22. Immediate annuitants projected, females - IFL92 and IFA92 ( $B=1972$ ) ..... 209
A23. Retirement annuitants projected - RMV92 and RFV92 ( $B=1972$ ) ..... 212
A24. Widows projected - WL92 and WA92 $(B=1972)$ ..... 215

Table A1. Permanent assurances, males - AM92 two years select: values of $q_{x-t j+t}$

| Age $x$ | Duration 0 | Duration 1 | Durations 2+ |
| :--- | :---: | :---: | :---: |
| 17 | 0.000427 | 0.000552 | 0.000600 |
| 18 | 0.000426 | 0.000548 | 0.000594 |
| 19 | 0.000425 | 0.000544 | 0.000587 |
| 20 | 0.000425 | 0.000541 | 0.000582 |
| 21 | 0.000425 | 0.000538 | 0.000577 |
| 22 | 0.000427 | 0.000535 | 0.000572 |
| 23 | 0.000429 | 0.000534 | 0.000569 |
| 24 | 0.000431 | 0.000533 | 0.000567 |
| 25 | 0.000435 | 0.000533 | 0.000566 |
| 26 | 0.000440 | 0.000535 | 0.000567 |
| 27 | 0.000447 | 0.000538 | 0.000570 |
| 28 | 0.000455 | 0.000542 | 0.000574 |
| 29 | 0.000465 | 0.000549 | 0.000580 |
| 30 | 0.000476 | 0.000558 | 0.000590 |
| 31 | 0.000490 | 0.000569 | 0.000602 |
| 32 | 0.000507 | 0.000584 | 0.000617 |
| 33 | 0.000527 | 0.000602 | 0.000636 |
| 34 | 0.000550 | 0.000624 | 0.000660 |
| 35 | 0.000577 | 0.000651 | 0.000689 |
| 36 | 0.000608 | 0.000683 | 0.000724 |
| 37 | 0.000644 | 0.000722 | 0.000765 |
| 38 | 0.000685 | 0.000768 | 0.000813 |
| 39 | 0.000733 | 0.000823 | 0.000870 |
| 40 | 0.000788 | 0.000887 | 0.000937 |
| 41 | 0.000851 | 0.000962 | 0.001014 |
| 42 | 0.000922 | 0.001049 | 0.001104 |
| 43 | 0.001003 | 0.001150 | 0.001208 |
| 44 | 0.001096 | 0.001267 | 0.001327 |
| 45 | 0.001201 | 0.001402 | 0.001465 |
| 46 | 0.001320 | 0.001557 | 0.001622 |
| 47 | 0.001455 | 0.0001935 | 0.001802 |
| 48 | 0.001778 | 0.002170 | 0.002008 |
| 49 | 0.001971 | 0.002434 | 0.002241 |
| 50 | 0.002189 | 0.002732 | 0.002508 |
| 51 | 0.002433 | 0.0303070 | 0.002809 |
| 52 | 0.003881 | 0.003152 |  |
| 53 | 54 |  | 0.003976 |
| 54 |  |  |  |

Table A1. (Continued).

| Age $x$ | Duration 0 | Duration 1 | Durations 2+ |
| :---: | :---: | :---: | :---: |
| 55 | 0.003358 | 0.004363 | 0.004469 |
| 56 | 0.003742 | 0.004903 | 0.005025 |
| 57 | 0.004171 | 0.005507 | 0.005650 |
| 58 | 0.004649 | 0.006180 | 0.006352 |
| 59 | 0.005182 | 0.006929 | 0.007140 |
| 60 | 0.005774 | 0.007760 | 0.008022 |
| 61 | 0.006433 | 0.008680 | 0.009009 |
| 62 | 0.007164 | 0.009696 | 0.010112 |
| 63 | 0.007974 | 0.010815 | 0.011344 |
| 64 | 0.008871 | 0.012046 | 0.012716 |
| 65 | 0.009864 | 0.013396 | 0.014243 |
| 66 | 0.010960 | 0.014873 | 0.015940 |
| 67 | 0.012169 | 0.016484 | 0.017824 |
| 68 | 0.013502 | 0.018239 | 0.019913 |
| 69 | 0.014969 | 0.020145 | 0.022226 |
| 70 | 0.016582 | 0.022210 | 0.024783 |
| 71 | 0.018353 | 0.024441 | 0.027606 |
| 72 | 0.020296 | 0.026847 | 0.030718 |
| 73 | 0.022423 | 0.029434 | 0.034144 |
| 74 | 0.024750 | 0.032208 | 0.037911 |
| 75 | 0.027293 | 0.035176 | 0.042046 |
| 76 | 0.030067 | 0.038344 | 0.046578 |
| 77 | 0.033090 | 0.041715 | 0.051538 |
| 78 | 0.036379 | 0.045292 | 0.056956 |
| 79 | 0.039954 | 0.049080 | 0.062867 |
| 80 | 0.043833 | 0.053078 | 0.069303 |
| 81 | 0.048037 | 0.057288 | 0.076300 |
| 82 | 0.052586 | 0.061709 | 0.083893 |
| 83 | 0.057501 | 0.066337 | 0.092117 |
| 84 | 0.062804 | 0.071169 | 0.101007 |
| 85 | 0.068516 | 0.076199 | 0.110600 |
| 86 | 0.074661 | 0.081422 | 0.120929 |
| 87 | 0.081258 | 0.086827 | 0.132028 |
| 88 | 0.088331 | 0.092405 | 0.143929 |
| 89 | 0.095902 | 0.098144 | 0.156660 |
| 90 | 0.103990 | 0.104031 | 0.170247 |
| 91 |  | 0.110052 | 0.184714 |
| 92 |  |  | 0.200079 |
| 93 |  |  | 0.216354 |
| 94 |  |  | 0.233548 |

Table A1. (Continued).

| Age $x$ | Duration 0 | Duration 1 |
| :--- | :---: | :---: |
| 95 |  | Durations 2+ |
| 96 |  | 0.251662 |
| 97 | 0.270688 |  |
| 98 | 0.290613 |  |
| 99 | 0.311414 |  |
| 100 | 0.333058 |  |
| 101 |  | 0.355505 |
| 102 |  | 0.378702 |
| 103 |  | 0.402588 |
| 104 | 0.427090 |  |
| 105 | 0.452127 |  |
| 106 |  | 0.477608 |
| 107 | 0.503432 |  |
| 108 |  | 0.529493 |
| 109 | 0.555674 |  |
| 110 | 0.581857 |  |
| 111 |  | 0.607918 |
| 112 | 0.633731 |  |
| 113 | 0.659171 |  |
| 114 |  | 0.684114 |
| 115 | 0.708442 |  |
| 116 |  | 0.732042 |
| 117 | 0.754809 |  |
| 118 | 0.776648 |  |
| 119 | 0.797477 |  |
| 120 | 0.817225 |  |

Table A2. Permanent assurances, females - AF92 two years select:
values of $q_{[x-t]+t}$

| Age $x$ | Duration 0 | Duration 1 | Durations 2+ |
| :---: | :---: | :---: | :---: |
| 17 | 0.000113 | 0.000132 | 0.000172 |
| 18 | 0.000117 | 0.000138 | 0.000178 |
| 19 | 0.000121 | 0.000143 | 0.000185 |
| 20 | 0.000125 | 0.000150 | 0.000193 |
| 21 | 0.000130 | 0.000157 | 0.000202 |
| 22 | 0.000135 | 0.000165 | 0.000212 |
| 23 | 0.000141 | 0.000173 | 0.000223 |
| 24 | 0.000148 | 0.000183 | 0.000235 |
| 25 | 0.000155 | 0.000193 | 0.000248 |
| 26 | 0.000163 | 0.000205 | 0.000263 |
| 27 | 0.000172 | 0.000218 | 0.000279 |
| 28 | 0.000182 | 0.000232 | 0.000297 |
| 29 | 0.000193 | 0.000247 | 0.000317 |
| 30 | 0.000205 | 0.000264 | 0.000339 |
| 31 | 0.000218 | 0.000283 | 0.000364 |
| 32 | 0.000233 | 0.000304 | 0.000391 |
| 33 | 0.000249 | 0.000327 | 0.000422 |
| 34 | 0.000267 | 0.000352 | 0.000455 |
| 35 | 0.000287 | 0.000380 | 0.000492 |
| 36 | 0.000309 | 0.000410 | 0.000533 |
| 37 | 0.000334 | 0.000444 | 0.000579 |
| 38 | 0.000361 | 0.000481 | 0.000629 |
| 39 | 0.000391 | 0.000522 | 0.000685 |
| 40 | 0.000424 | 0.000567 | 0.000747 |
| 41 | 0.000461 | 0.000616 | 0.000816 |
| 42 | 0.000502 | 0.000671 | 0.000892 |
| 43 | 0.000547 | 0.000731 | 0.000977 |
| 44 | 0.000597 | 0.000798 | 0.001070 |
| 45 | 0.000652 | 0.000871 | 0.001174 |
| 46 | 0.000713 | 0.000951 | 0.001288 |
| 47 | 0.000781 | 0.001040 | 0.001416 |
| 48 | 0.000855 | 0.001138 | 0.001556 |
| 49 | 0.000938 | 0.001246 | 0.001713 |
| 50 | 0.001030 | 0.001364 | 0.001886 |
| 51 | 0.001132 | 0.001495 | 0.002077 |
| 52 | 0.001244 | 0.001640 | 0.002290 |
| 53 | 0.001368 | 0.001798 | 0.002525 |
| 54 | 0.001506 | 0.001973 | 0.002786 |

Table A2. (Continued).

| Age $x$ | Duration 0 | Duration 1 | Durations 2+ |
| :---: | :---: | :---: | :---: |
| 55 | 0.001658 | 0.002166 | 0.003075 |
| 56 | 0.001827 | 0.002379 | 0.003395 |
| 57 | 0.002014 | 0.002613 | 0.003749 |
| 58 | 0.002220 | 0.002871 | 0.004142 |
| 59 | 0.002449 | 0.003155 | 0.004577 |
| 60 | 0.002701 | 0.003468 | 0.005058 |
| 61 | 0.002981 | 0.003813 | 0.005592 |
| 62 | 0.003291 | 0.004192 | 0.006183 |
| 63 | 0.003634 | 0.004611 | 0.006838 |
| 64 | 0.004013 | 0.005071 | 0.007562 |
| 65 | 0.004432 | 0.005579 | 0.008365 |
| 66 | 0.004897 | 0.006138 | 0.009254 |
| 67 | 0.005410 | 0.006754 | 0.010238 |
| 68 | 0.005978 | 0.007431 | 0.011327 |
| 69 | 0.006607 | 0.008178 | 0.012533 |
| 70 | 0.007302 | 0.009000 | 0.013867 |
| 71 | 0.008071 | 0.009905 | 0.015343 |
| 72 | 0.008922 | 0.010901 | 0.016977 |
| 73 | 0.009863 | 0.011998 | 0.018784 |
| 74 | 0.010903 | 0.013205 | 0.020783 |
| 75 | 0.012053 | 0.014534 | 0.022994 |
| 76 | 0.013325 | 0.015995 | 0.025438 |
| 77 | 0.014730 | 0.017604 | 0.028138 |
| 78 | 0.016284 | 0.019373 | 0.031123 |
| 79 | 0.018000 | 0.021319 | 0.034419 |
| 80 | 0.019897 | 0.023459 | 0.038059 |
| 81 | 0.021993 | 0.025812 | 0.042077 |
| 82 | 0.024307 | 0.028398 | 0.046510 |
| 83 | 0.026862 | 0.031240 | 0.051398 |
| 84 | 0.029683 | 0.034362 | 0.056786 |
| 85 | 0.032795 | 0.037791 | 0.062721 |
| 86 | 0.036229 | 0.041555 | 0.069254 |
| 87 | 0.040016 | 0.045686 | 0.076441 |
| 88 | 0.044190 | 0.050217 | 0.084340 |
| 89 | 0.048789 | 0.055186 | 0.093016 |
| 90 | 0.053854 | 0.060631 | 0.102533 |
| 91 |  | 0.066594 | 0.112964 |
| 92 |  |  | 0.124381 |
| 93 |  |  | 0.136862 |
| 94 |  |  | 0.150486 |

Table A2. (Continued).

| Age $x$ | Duration 0 | Duration 1 |
| :--- | :---: | :---: |
| 95 |  | Durations 2+ |
| 96 |  | 0.165332 |
| 97 | 0.181481 |  |
| 98 | 0.199012 |  |
| 99 | 0.218002 |  |
| 100 | 0.238520 |  |
| 101 |  | 0.260630 |
| 102 | 0.284383 |  |
| 103 | 0.309814 |  |
| 104 | 0.336943 |  |
| 105 | 0.365762 |  |
| 106 |  | 0.396238 |
| 107 | 0.428303 |  |
| 108 | 0.461851 |  |
| 109 | 0.496734 |  |
| 110 | 0.532756 |  |
| 111 | 0.569670 |  |
| 112 | 0.607181 |  |
| 113 |  | 0.644942 |
| 114 | 0.682563 |  |
| 115 | 0.719616 |  |
| 116 |  | 0.755647 |
| 117 | 0.790192 |  |
| 118 | 0.822800 |  |
| 119 | 0.853050 |  |
| 120 |  | 0.880579 |

Table A3. Temporary assurances, males - TM92 five years select: values of $q_{[x-t]+t}$

| Age $x$ | Duration 0 | Durations 1-4 | Durations 5+ |
| :---: | :---: | :---: | :---: |
| 17 | 0.000679 | 0.000778 | 0.000846 |
| 18 | 0.000655 | 0.000753 | 0.000820 |
| 19 | 0.000632 | 0.000729 | 0.000794 |
| 20 | 0.000610 | 0.000707 | 0.000770 |
| 21 | 0.000589 | 0.000686 | 0.000747 |
| 22 | 0.000570 | 0.000667 | 0.000725 |
| 23 | 0.000552 | 0.000649 | 0.000705 |
| 24 | 0.000535 | 0.000633 | 0.000687 |
| 25 | 0.000521 | 0.000619 | 0.000670 |
| 26 | 0.000508 | 0.000607 | 0.000656 |
| 27 | 0.000498 | 0.000598 | 0.000644 |
| 28 | 0.000490 | 0.000591 | 0.000634 |
| 29 | 0.000484 | 0.000588 | 0.000628 |
| 30 | 0.000482 | 0.000587 | 0.000625 |
| 31 | 0.000483 | 0.000590 | 0.000626 |
| 32 | 0.000487 | 0.000598 | 0.000630 |
| 33 | 0.000494 | 0.000610 | 0.000640 |
| 34 | 0.000507 | 0.000626 | 0.000655 |
| 35 | 0.000523 | 0.000648 | 0.000676 |
| 36 | 0.000545 | 0.000676 | 0.000703 |
| 37 | 0.000572 | 0.000711 | 0.000738 |
| 38 | 0.000605 | 0.000753 | 0.000781 |
| 39 | 0.000645 | 0.000802 | 0.000833 |
| 40 | 0.000691 | 0.000861 | 0.000896 |
| 41 | 0.000746 | 0.000929 | 0.000970 |
| 42 | 0.000810 | 0.001008 | 0.001057 |
| 43 | 0.000883 | 0.001098 | 0.001158 |
| 44 | 0.000966 | 0.001202 | 0.001275 |
| 45 | 0.001061 | 0.001319 | 0.001410 |
| 46 | 0.001168 | 0.001453 | 0.001565 |
| 47 | 0.001289 | 0.001603 | 0.001743 |
| 48 | 0.001425 | 0.001772 | 0.001945 |
| 49 | 0.001577 | 0.001962 | 0.002175 |
| 50 | 0.001746 | 0.002175 | 0.002436 |
| 51 | 0.001936 | 0.002413 | 0.002731 |
| 52 | 0.002146 | 0.002679 | 0.003064 |
| 53 | 0.002380 | 0.002975 | 0.003441 |
| 54 | 0.002639 | 0.003304 | 0.003864 |

Table A3. (Continued).

| Age $x$ | Duration 0 | Durations 1-4 | Durations 5+ |
| :---: | :---: | :---: | :---: |
| 55 | 0.002926 | 0.003669 | 0.004341 |
| 56 | 0.003244 | 0.004075 | 0.004877 |
| 57 | 0.003594 | 0.004524 | 0.005478 |
| 58 | 0.003980 | 0.005022 | 0.006153 |
| 59 | 0.004406 | 0.005573 | 0.006908 |
| 60 | 0.004874 | 0.006181 | 0.007754 |
| 61 | 0.005390 | 0.006853 | 0.008700 |
| 62 | 0.005956 | 0.007594 | 0.009757 |
| 63 | 0.006578 | 0.008412 | 0.010937 |
| 64 | 0.007260 | 0.009313 | 0.012255 |
| 65 | 0.008008 | 0.010305 | 0.013724 |
| 66 | 0.008828 | 0.011397 | 0.015363 |
| 67 | 0.009726 | 0.012598 | 0.017187 |
| 68 | 0.010709 | 0.013919 | 0.019219 |
| 69 | 0.011784 | 0.015369 | 0.021479 |
| 70 | 0.012959 | 0.016962 | 0.023992 |
| 71 | 0.014243 | 0.018710 | 0.026784 |
| 72 | 0.015645 | 0.020627 | 0.029885 |
| 73 | 0.017176 | 0.022729 | 0.033327 |
| 74 | 0.018846 | 0.025032 | 0.037143 |
| 75 | 0.020667 | 0.027554 | 0.041372 |
| 76 | 0.022652 | 0.030315 | 0.046055 |
| 77 | 0.024814 | 0.033334 | 0.051236 |
| 78 | 0.027169 | 0.036635 | 0.056965 |
| 79 | 0.029730 | 0.040241 | 0.063292 |
| 80 | 0.032517 | 0.044179 | 0.070273 |
| 81 | 0.035546 | 0.048476 | 0.077969 |
| 82 | 0.038837 | 0.053161 | 0.086443 |
| 83 | 0.042410 | 0.058267 | 0.095763 |
| 84 | 0.046289 | 0.063826 | 0.105998 |
| 85 | 0.050495 | 0.069875 | 0.117224 |
| 86 | 0.055054 | 0.076451 | 0.129518 |
| 87 | 0.059992 | 0.083593 | 0.142957 |
| 88 | 0.065337 | 0.091343 | 0.157622 |
| 89 | 0.071118 | 0.099745 | 0.173593 |
| 90 | 0.077367 | 0.108844 | 0.190948 |
| 91 |  | 0.118686 | 0.209761 |
| 92 |  | 0.129320 | 0.230102 |
| 93 |  | 0.140793 | 0.252030 |
| 94 |  | 0.153155 | 0.275595 |

Table A3. (Continued).

| Age $x$ | Duration 0 | Durations 1-4 |
| :--- | :---: | :---: |
| 95 |  | Durations 5+ |
| 96 |  | 0.300831 |
| 97 |  | 0.327753 |
| 98 |  | 0.356353 |
| 99 |  | 0.418410 |
| 100 |  | 0.451695 |
| 101 |  | 0.486303 |
| 102 |  | 0.522043 |
| 103 |  | 0.558677 |
| 104 | 0.595921 |  |
| 105 | 0.633445 |  |
| 106 | 0.670875 |  |
| 107 |  | 0.707804 |
| 108 |  | 0.743804 |
| 109 | 0.778435 |  |
| 110 |  | 0.811264 |
| 111 |  | 0.841891 |
| 112 |  | 0.869961 |
| 113 | 0.895193 |  |
| 114 |  | 0.917393 |
| 115 |  | 0.936470 |
| 116 |  | 0.952446 |
| 117 | 0.965448 |  |
| 118 | 0.975704 |  |
| 119 | 0.983521 |  |
| 120 |  | 1.000000 |

Table A4. Temporary assurances, females - TF92 five years select: values of $q_{[x-t]+t}$

| Age $x$ | Duration 0 | Durations 1-4 | Durations 5+ |
| :---: | :---: | :---: | :---: |
| 17 | 0.000099 | 0.000167 | 0.000227 |
| 18 | 0.000102 | 0.000172 | 0.000232 |
| 19 | 0.000104 | 0.000178 | 0.000238 |
| 20 | 0.000107 | 0.000184 | 0.000245 |
| 21 | 0.000110 | 0.000191 | 0.000252 |
| 22 | 0.000114 | 0.000199 | 0.000260 |
| 23 | 0.000118 | 0.000208 | 0.000269 |
| 24 | 0.000122 | 0.000218 | 0.000279 |
| 25 | 0.000127 | 0.000229 | 0.000290 |
| 26 | 0.000132 | 0.000241 | 0.000303 |
| 27 | 0.000138 | 0.000254 | 0.000317 |
| 28 | 0.000145 | 0.000269 | 0.000332 |
| 29 | 0.000153 | 0.000285 | 0.000349 |
| 30 | 0.000161 | 0.000304 | 0.000369 |
| 31 | 0.000170 | 0.000324 | 0.000390 |
| 32 | 0.000181 | 0.000346 | 0.000414 |
| 33 | 0.000193 | 0.000371 | 0.000440 |
| 34 | 0.000206 | 0.000399 | 0.000469 |
| 35 | 0.000220 | 0.000430 | 0.000502 |
| 36 | 0.000237 | 0.000464 | 0.000538 |
| 37 | 0.000255 | 0.000502 | 0.000579 |
| 38 | 0.000275 | 0.000545 | 0.000624 |
| 39 | 0.000298 | 0.000591 | 0.000674 |
| 40 | 0.000323 | 0.000643 | 0.000729 |
| 41 | 0.000351 | 0.000701 | 0.000791 |
| 42 | 0.000383 | 0.000766 | 0.000860 |
| 43 | 0.000418 | 0.000837 | 0.000937 |
| 44 | 0.000457 | 0.000916 | 0.001022 |
| 45 | 0.000501 | 0.001004 | 0.001117 |
| 46 | 0.000549 | 0.001102 | 0.001223 |
| 47 | 0.000604 | 0.001210 | 0.001341 |
| 48 | 0.000665 | 0.001330 | 0.001472 |
| 49 | 0.000733 | 0.001464 | 0.001618 |
| 50 | 0.000808 | 0.001613 | 0.001780 |
| 51 | 0.000893 | 0.001777 | 0.001960 |
| 52 | 0.000987 | 0.001960 | 0.002161 |
| 53 | 0.001093 | 0.002163 | 0.002384 |
| 54 | 0.001210 | 0.002389 | 0.002633 |

Table A4. (Continued).

| Age $x$ | Duration 0 | Durations 1-4 | Durations 5+ |
| :---: | :---: | :---: | :---: |
| 55 | 0.001341 | 0.002640 | 0.002909 |
| 56 | 0.001488 | 0.002918 | 0.003217 |
| 57 | 0.001651 | 0.003226 | 0.003559 |
| 58 | 0.001834 | 0.003569 | 0.003940 |
| 59 | 0.002037 | 0.003949 | 0.004363 |
| 60 | 0.002264 | 0.004372 | 0.004834 |
| 61 | 0.002517 | 0.004840 | 0.005358 |
| 62 | 0.002800 | 0.005361 | 0.005941 |
| 63 | 0.003115 | 0.005938 | 0.006589 |
| 64 | 0.003467 | 0.006579 | 0.007310 |
| 65 | 0.003860 | 0.007290 | 0.008111 |
| 66 | 0.004298 | 0.008079 | 0.009002 |
| 67 | 0.004787 | 0.008955 | 0.009993 |
| 68 | 0.005332 | 0.009926 | 0.011095 |
| 69 | 0.005940 | 0.011004 | 0.012319 |
| 70 | 0.006618 | 0.012200 | 0.013680 |
| 71 | 0.007375 | 0.013526 | 0.015192 |
| 72 | 0.008219 | 0.014996 | 0.016873 |
| 73 | 0.009159 | 0.016626 | 0.018739 |
| 74 | 0.010208 | 0.018434 | 0.020812 |
| 75 | 0.011378 | 0.020437 | 0.023114 |
| 76 | 0.012682 | 0.022656 | 0.025669 |
| 77 | 0.014135 | 0.025115 | 0.028505 |
| 78 | 0.015754 | 0.027839 | 0.031652 |
| 79 | 0.017559 | 0.030855 | 0.035141 |
| 80 | 0.019568 | 0.034193 | 0.039009 |
| 81 | 0.021807 | 0.037886 | 0.043296 |
| 82 | 0.024299 | 0.041971 | 0.048045 |
| 83 | 0.027074 | 0.046487 | 0.053301 |
| 84 | 0.030161 | 0.051477 | 0.059117 |
| 85 | 0.033595 | 0.056988 | 0.065547 |
| 86 | 0.037414 | 0.063071 | 0.072651 |
| 87 | 0.041658 | 0.069779 | 0.080494 |
| 88 | 0.046373 | 0.077173 | 0.089144 |
| 89 | 0.051608 | 0.085315 | 0.098674 |
| 90 | 0.057417 | 0.094273 | 0.109163 |
| 91 |  | 0.104118 | 0.120693 |
| 92 |  | 0.114925 | 0.133348 |
| 93 |  | 0.126774 | 0.147218 |
| 94 |  | 0.139747 | 0.162392 |

Table A4. (Continued).

| Age $x$ | Duration 0 | Durations I-4 |
| :--- | :---: | :---: |
| 95 | Durations 5+ |  |
| 96 | 0.178963 |  |
| 97 | 0.197018 |  |
| 98 | 0.216645 |  |
| 99 | 0.237923 |  |
| 100 | 0.260924 |  |
| 101 |  | 0.285707 |
| 102 | 0.312311 |  |
| 103 | 0.340754 |  |
| 104 | 0.371027 |  |
| 105 | 0.403085 |  |
| 106 |  | 0.436844 |
| 107 | 0.472174 |  |
| 108 | 0.508892 |  |
| 109 | 0.546759 |  |
| 110 | 0.585478 |  |
| 111 |  | 0.624691 |
| 112 | 0.663985 |  |
| 113 |  | 0.702896 |
| 114 |  | 0.740926 |
| 115 | 0.777553 |  |
| 116 |  | 0.812260 |
| 117 | 0.844557 |  |
| 118 | 0.874011 |  |
| 119 |  | 0.900277 |
| 120 |  | 0.923123 |

Table A5. Pensioners - PML92Base, PMA92Base, PFL92Base and PFA92Base values of $q_{x}$

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92Base } \end{gathered}$ | Amounts PMA92Base | $\begin{gathered} \text { Lives } \\ \text { PFL92Base } \end{gathered}$ | Amounts PFA92Base |
| 20 | 0.000623 | 0.000340 | 0.000254 | 0.000203 |
| 21 | 0.000612 | 0.000338 | 0.000253 | 0.000202 |
| 22 | 0.000601 | 0.000337 | 0.000253 | 0.000202 |
| 23 | 0.000591 | 0.000335 | 0.000252 | 0.000202 |
| 24 | 0.000581 | 0.000334 | 0.000252 | 0.000202 |
| 25 | 0.000573 | 0.000333 | 0.000253 | 0.000203 |
| 26 | 0.000565 | 0.000332 | 0.000255 | 0.000204 |
| 27 | 0.000560 | 0.000331 | 0.000257 | 0.000206 |
| 28 | 0.000555 | 0.000331 | 0.000260 | 0.000209 |
| 29 | 0.000553 | 0.000332 | 0.000265 | 0.000212 |
| 30 | 0.000553 | 0.000333 | 0.000270 | 0.000217 |
| 31 | 0.000556 | 0.000335 | 0.000278 | 0.000222 |
| 32 | 0.000562 | 0.000338 | 0.000287 | 0.000230 |
| 33 | 0.000572 | 0.000343 | 0.000299 | 0.000238 |
| 34 | 0.000586 | 0.000349 | 0.000313 | 0.000249 |
| 35 | 0.000605 | 0.000357 | 0.000330 | 0.000263 |
| 36 | 0.000631 | 0.000367 | 0.000351 | 0.000279 |
| 37 | 0.000663 | 0.000380 | 0.000376 | 0.000298 |
| 38 | 0.000704 | 0.000396 | 0.000405 | 0.000321 |
| 39 | 0.000755 | 0.000417 | 0.000441 | 0.000349 |
| 40 | 0.000817 | 0.000443 | 0.000482 | 0.000382 |
| 41 | 0.000892 | 0.000475 | 0.000532 | 0.000421 |
| 42 | 0.000982 | 0.000514 | 0.000589 | 0.000468 |
| 43 | 0.001089 | 0.000561 | 0.000657 | 0.000522 |
| 44 | 0.001216 | 0.000619 | 0.000736 | 0.000586 |
| 45 | 0.001365 | 0.000689 | 0.000829 | 0.000662 |
| 46 | 0.001541 | 0.000774 | 0.000936 | 0.000750 |
| 47 | 0.001746 | 0.000876 | 0.001061 | 0.000852 |
| 48 | 0.001985 | 0.000997 | 0.001205 | 0.000972 |
| 49 | 0.002262 | 0.001142 | 0.001371 | 0.001110 |
| 50 | 0.002583 | 0.001315 | 0.001563 | 0.001271 |
| 51 | 0.002953 | 0.001519 | 0.001783 | 0.001456 |
| 52 | 0.003378 | 0.001761 | 0.002036 | 0.001670 |
| 53 | 0.003865 | 0.002045 | 0.002326 | 0.001917 |
| 54 | 0.004422 | 0.002379 | 0.002657 | 0.002200 |

Table A5. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Lives PML92Base | Amounts PMA92Base | Lives PFL92Base | Amounts PFA92Base |
| 55 | 0.005058 | 0.002771 | 0.003035 | 0.002524 |
| 56 | 0.005780 | 0.003228 | 0.003466 | 0.002894 |
| 57 | 0.006600 | 0.003759 | 0.003955 | 0.003317 |
| 58 | 0.007527 | 0.004376 | 0.004510 | 0.003799 |
| 59 | 0.008574 | 0.005090 | 0.005139 | 0.004345 |
| 60 | 0.009753 | 0.005914 | 0.005849 | 0.004965 |
| 61 | 0.011076 | 0.006861 | 0.006650 | 0.005667 |
| 62 | 0.012559 | 0.007947 | 0.007553 | 0.006458 |
| 63 | 0.014217 | 0.009189 | 0.008567 | 0.007350 |
| 64 | 0.016065 | 0.010604 | 0.009704 | 0.008352 |
| 65 | 0.018120 | 0.012211 | 0.010977 | 0.009476 |
| 66 | 0.020400 | 0.014032 | 0.012399 | 0.010734 |
| 67 | 0.022923 | 0.016088 | 0.013985 | 0.012138 |
| 68 | 0.025708 | 0.018402 | 0.015751 | 0.013703 |
| 69 | 0.028775 | 0.020998 | 0.017711 | 0.015442 |
| 70 | 0.032142 | 0.023901 | 0.019885 | 0.017371 |
| 71 | 0.035831 | 0.027137 | 0.022289 | 0.019505 |
| 72 | 0.039862 | 0.030732 | 0.024943 | 0.021861 |
| 73 | 0.044253 | 0.034713 | 0.027867 | 0.024455 |
| 74 | 0.049025 | 0.039105 | 0.031081 | 0.027306 |
| 75 | 0.054197 | 0.043935 | 0.034607 | 0.030432 |
| 76 | 0.059787 | 0.049227 | 0.038467 | 0.033849 |
| 77 | 0.065811 | 0.055006 | 0.042682 | 0.037577 |
| 78 | 0.072286 | 0.061292 | 0.047276 | 0.041632 |
| 79 | 0.079225 | 0.068106 | 0.052271 | 0.046035 |
| 80 | 0.086639 | 0.075464 | 0.057690 | 0.050800 |
| 81 | 0.094538 | 0.083379 | 0.063554 | 0.055946 |
| 82 | 0.102929 | 0.091862 | 0.069886 | 0.061488 |
| 83 | 0.111815 | 0.100917 | 0.076705 | 0.067441 |
| 84 | 0.121196 | 0.110544 | 0.084032 | 0.073817 |
| 85 | 0.131071 | 0.120739 | 0.091884 | 0.080629 |
| 86 | 0.141431 | 0.131492 | 0.100278 | 0.087885 |
| 87 | 0.152267 | 0.142786 | 0.109228 | 0.095594 |
| 88 | 0.163563 | 0.154599 | 0.118745 | 0.103761 |
| 89 | 0.175302 | 0.166903 | 0.128838 | 0.112386 |

Table A5. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92Base } \end{gathered}$ | Amounts PMA92Base | $\begin{gathered} \text { Lives } \\ \text { PFL92Base } \end{gathered}$ | Amounts PFA92Base |
| 90 | 0.187460 | 0.179664 | 0.139514 | 0.121470 |
| 91 | 0.200011 | 0.192841 | 0.150773 | 0.131009 |
| 92 | 0.212925 | 0.206389 | 0.162615 | 0.140996 |
| 93 | 0.226167 | 0.220257 | 0.175034 | 0.151420 |
| 94 | 0.239700 | 0.234389 | 0.188021 | 0.162267 |
| 95 | 0.253482 | 0.248727 | 0.201561 | 0.173519 |
| 96 | 0.267471 | 0.263206 | 0.215636 | 0.185155 |
| 97 | 0.281620 | 0.277762 | 0.230224 | 0.197150 |
| 98 | 0.295880 | 0.292327 | 0.245296 | 0.209477 |
| 99 | 0.310203 | 0.306832 | 0.260821 | 0.222103 |
| 100 | 0.324539 | 0.321209 | 0.276762 | 0.234995 |
| 101 | 0.338835 | 0.335389 | 0.293080 | 0.248115 |
| 102 | 0.353042 | 0.349305 | 0.309730 | 0.261424 |
| 103 | 0.367107 | 0.362893 | 0.326664 | 0.274879 |
| 104 | 0.380983 | 0.376091 | 0.343834 | 0.288437 |
| 105 | 0.394621 | 0.388838 | 0.361186 | 0.302054 |
| 106 | 0.407973 | 0.401079 | 0.378665 | 0.315684 |
| 107 | 0.420997 | 0.412763 | 0.396216 | 0.329280 |
| 108 | 0.433649 | 0.423842 | 0.413781 | 0.342795 |
| 109 | 0.445890 | 0.434272 | 0.431304 | 0.356185 |
| 110 | 0.457684 | 0.444014 | 0.448729 | 0.369403 |
| 111 | 0.468997 | 0.453033 | 0.466001 | 0.382406 |
| 112 | 0.479798 | 0.461297 | 0.483064 | 0.395150 |
| 113 | 0.490060 | 0.468780 | 0.499869 | 0.407594 |
| 114 | 0.499756 | 0.475459 | 0.516366 | 0.419700 |
| 115 | 0.508865 | 0.481313 | 0.532508 | 0.431431 |
| 116 | 0.517368 | 0.486326 | 0.548254 | 0.442752 |
| 117 | 0.525248 | 0.490484 | 0.563564 | 0.453630 |
| 118 | 0.532489 | 0.493776 | 0.578402 | 0.464036 |
| 119 | 0.539080 | 0.496194 | 0.592739 | 0.473947 |
| 120 | 1.000000 | 1.000000 | 1.000000 | 1.000000 |

Table A6. Immediate annuitants, males - IML92Base and IMA92Base one year select: values of $q_{[x-t]+t}$

| Age $x$ | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 17 | 0.000433 | 0.000587 | 0.000458 | 0.000560 |
| 18 | 0.000429 | 0.000582 | 0.000454 | 0.000555 |
| 19 | 0.000426 | 0.000577 | 0.000450 | 0.000550 |
| 20 | 0.000423 | 0.000573 | 0.000447 | 0.000546 |
| 21 | 0.000420 | 0.000569 | 0.000444 | 0.000543 |
| 22 | 0.000418 | 0.000567 | 0.000442 | 0.000540 |
| 23 | 0.000417 | 0.000565 | 0.000441 | 0.000539 |
| 24 | 0.000416 | 0.000564 | 0.000440 | 0.000538 |
| 25 | 0.000417 | 0.000565 | 0.000440 | 0.000538 |
| 26 | 0.000418 | 0.000567 | 0.000442 | 0.000540 |
| 27 | 0.000421 | 0.000570 | 0.000445 | 0.000544 |
| 28 | 0.000425 | 0.000576 | 0.000449 | 0.000549 |
| 29 | 0.000430 | 0.000583 | 0.000455 | 0.000556 |
| 30 | 0.000438 | 0.000593 | 0.000463 | 0.000566 |
| 31 | 0.000448 | 0.000606 | 0.000473 | 0.000578 |
| 32 | 0.000460 | 0.000623 | 0.000486 | 0.000594 |
| 33 | 0.000474 | 0.000643 | 0.000501 | 0.000613 |
| 34 | 0.000492 | 0.000667 | 0.000520 | 0.000636 |
| 35 | 0.000514 | 0.000696 | 0.000543 | 0.000664 |
| 36 | 0.000540 | 0.000731 | 0.000570 | 0.000697 |
| 37 | 0.000570 | 0.000772 | 0.000602 | 0.000736 |
| 38 | 0.000606 | 0.000821 | 0.000640 | 0.000783 |
| 39 | 0.000648 | 0.000878 | 0.000685 | 0.000837 |
| 40 | 0.000697 | 0.000945 | 0.000737 | 0.000901 |
| 41 | 0.000754 | 0.001022 | 0.000797 | 0.000974 |
| 42 | 0.000820 | 0.001111 | 0.000867 | 0.001060 |
| 43 | 0.000896 | 0.001215 | 0.000947 | 0.001158 |
| 44 | 0.000984 | 0.001333 | 0.001040 | 0.001271 |
| 45 | 0.001085 | 0.001470 | 0.001147 | 0.001402 |
| 46 | 0.001201 | 0.001627 | 0.001269 | 0.001551 |
| 47 | 0.001333 | 0.001806 | 0.001408 | 0.001722 |
| 48 | 0.001484 | 0.002010 | 0.001568 | 0.001917 |
| 49 | 0.001656 | 0.002243 | 0.001750 | 0.002139 |
| 50 | 0.001852 | 0.002508 | 0.001956 | 0.002392 |
| 51 | 0.002074 | 0.002809 | 0.002191 | 0.002678 |
| 52 | 0.002326 | 0.003150 | 0.002457 | 0.003004 |
| 53 | 0.002611 | 0.003536 | 0.002759 | 0.003372 |
| 54 | 0.002934 | 0.003973 | 0.003099 | 0.003788 |

Table A6. (Continued).

| Age $x$ | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 55 | 0.003298 | 0.004466 | 0.003484 | 0.004258 |
| 56 | 0.003708 | 0.005021 | 0.003917 | 0.004787 |
| 57 | 0.004170 | 0.005646 | 0.004405 | 0.005383 |
| 58 | 0.004689 | 0.006348 | 0.004953 | 0.006053 |
| 59 | 0.005271 | 0.007136 | 0.005569 | 0.006805 |
| 60 | 0.005925 | 0.008019 | 0.006259 | 0.007647 |
| 61 | 0.006656 | 0.009008 | 0.007031 | 0.008590 |
| 62 | 0.007474 | 0.010113 | 0.007895 | 0.009645 |
| 63 | 0.008387 | 0.011347 | 0.008859 | 0.010822 |
| 64 | 0.009405 | 0.012722 | 0.009935 | 0.012134 |
| 65 | 0.010539 | 0.014254 | 0.011132 | 0.013595 |
| 66 | 0.011800 | 0.015956 | 0.012464 | 0.015219 |
| 67 | 0.013201 | 0.017846 | 0.013943 | 0.017022 |
| 68 | 0.014756 | 0.019941 | 0.015584 | 0.019022 |
| 69 | 0.016477 | 0.022261 | 0.017402 | 0.021236 |
| 70 | 0.018382 | 0.024826 | 0.019412 | 0.023684 |
| 71 | 0.020486 | 0.027657 | 0.021633 | 0.026387 |
| 72 | 0.022807 | 0.030778 | 0.024083 | 0.029367 |
| 73 | 0.025365 | 0.034214 | 0.026781 | 0.032647 |
| 74 | 0.028179 | 0.037990 | 0.029749 | 0.036253 |
| 75 | 0.031269 | 0.042133 | 0.033010 | 0.040211 |
| 76 | 0.034660 | 0.046673 | 0.036585 | 0.044549 |
| 77 | 0.038373 | 0.051638 | 0.040500 | 0.049294 |
| 78 | 0.042434 | 0.057061 | 0.044781 | 0.054478 |
| 79 | 0.046868 | 0.062972 | 0.049454 | 0.060130 |
| 80 | 0.051702 | 0.069406 | 0.054547 | 0.066284 |
| 81 | 0.056963 | 0.076395 | 0.060088 | 0.072971 |
| 82 | 0.062680 | 0.083973 | 0.066108 | 0.080224 |
| 83 | 0.068881 | 0.092175 | 0.072634 | 0.088078 |
| 84 | 0.075596 | 0.101034 | 0.079699 | 0.096565 |
| 85 | 0.082853 | 0.110584 | 0.087332 | 0.105718 |
| 86 | 0.090684 | 0.120858 | 0.095563 | 0.115570 |
| 87 | 0.099116 | 0.131886 | 0.104422 | 0.126151 |
| 88 | 0.108179 | 0.143698 | 0.113939 | 0.137491 |
| 89 | 0.117899 | 0.156321 | 0.124140 | 0.149617 |
| 90 | 0.128304 | 0.169777 | 0.135053 | 0.162554 |
| 91 | 0.139418 | 0.184088 | 0.146701 | 0.176322 |
| 92 | 0.151264 | 0.199268 | 0.159106 | 0.190940 |
| 93 | 0.163861 | 0.215328 | 0.172287 | 0.206419 |
| 94 | 0.177225 | 0.232274 | 0.186259 | 0.222767 |

Table A6. (Continued).

|  | Lives |  |  | Amounts |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Age $x$ | Duration 0 | Durations 1+ |  | Duration 0 | Durations 1+ |
| 95 | 0.191370 | 0.250103 |  | 0.201033 | 0.239986 |
| 96 | 0.206304 | 0.268807 |  | 0.216616 | 0.258070 |
| 97 | 0.222031 | 0.288370 |  | 0.233008 | 0.277008 |
| 98 | 0.238551 | 0.308768 |  | 0.250206 | 0.296780 |
| 99 | 0.255856 | 0.329968 |  | 0.268199 | 0.317358 |
| 100 | 0.273933 | 0.351927 |  | 0.286970 | 0.338706 |
| 101 |  | 0.374596 |  | 0.360778 |  |
| 102 |  | 0.397913 |  | 0.383521 |  |
| 103 |  | 0.421809 |  | 0.406871 |  |
| 104 | 0.446206 |  | 0.430756 |  |  |
| 105 |  | 0.471017 |  | 0.455098 |  |
| 106 | 0.496148 |  | 0.479808 |  |  |
| 107 |  | 0.521498 |  | 0.504791 |  |
| 108 | 0.546959 |  | 0.529946 |  |  |
| 109 | 0.572422 |  | 0.555169 |  |  |
| 110 |  | 0.597771 |  | 0.580348 |  |
| 111 |  | 0.622892 |  | 0.605374 |  |
| 112 |  | 0.647670 |  | 0.630135 |  |
| 113 |  | 0.671993 |  | 0.654519 |  |
| 114 | 0.695751 |  | 0.678419 |  |  |
| 115 |  | 0.718842 |  | 0.701731 |  |
| 116 | 0.741170 |  | 0.724359 |  |  |
| 117 | 0.762651 |  | 0.746213 |  |  |
| 118 |  | 0.783207 |  | 0.767212 |  |
| 119 |  | 0.802773 |  | 0.787288 |  |
| 120 |  | 1.000000 |  | 1.000000 |  |

Table A7. Immediate annuitants, females - IFL92Base and IFA92Base one year select: values of $\boldsymbol{q}_{[x-t]+t}$

| Age $x$ | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1 $\perp$ |
| 17 | 0.000216 | 0.000302 | 0.000329 | 0.000400 |
| 18 | 0.000217 | 0.000302 | 0.000329 | 0.000400 |
| 19 | 0.000217 | 0.000302 | 0.000329 | 0.000400 |
| 20 | 0.000217 | 0.000303 | 0.000329 | 0.000400 |
| 21 | 0.000218 | 0.000304 | 0.000329 | 0.000400 |
| 22 | 0.000218 | 0.000304 | 0.000329 | 0.000400 |
| 23 | 0.000219 | 0.000305 | 0.000329 | 0.000400 |
| 24 | 0.000220 | 0.000307 | 0.000329 | 0.000400 |
| 25 | 0.000221 | 0.000308 | 0.000329 | 0.000401 |
| 26 | 0.000222 | 0.000310 | 0.000330 | 0.000401 |
| 27 | 0.000224 | 0.000312 | 0.000330 | 0.000401 |
| 28 | 0.000225 | 0.000314 | 0.000330 | 0.000401 |
| 29 | 0.000228 | 0.000317 | 0.000330 | 0.000402 |
| 30 | 0.000230 | 0.000321 | 0.000331 | 0.000402 |
| 31 | 0.000233 | 0.000325 | 0.000332 | 0.000403 |
| 32 | 0.000237 | 0.000331 | 0.000332 | 0.000404 |
| 33 | 0.000242 | 0.000337 | 0.000333 | 0.000405 |
| 34 | 0.000247 | 0.000345 | 0.000334 | 0.000407 |
| 35 | 0.000254 | 0.000354 | 0.000336 | 0.000409 |
| 36 | 0.000261 | 0.000364 | 0.000338 | 0.000411 |
| 37 | 0.000271 | 0.000377 | 0.000340 | 0.000414 |
| 38 | 0.000281 | 0.000392 | 0.000344 | 0.000418 |
| 39 | 0.000294 | 0.000410 | 0.000347 | 0.000423 |
| 40 | 0.000310 | 0.000432 | 0.000352 | 0.000429 |
| 41 | 0.000328 | 0.000457 | 0.000358 | 0.000436 |
| 42 | 0.000349 | 0.000487 | 0.000366 | 0.000445 |
| 43 | 0.000374 | 0.000522 | 0.000375 | 0.000457 |
| 44 | 0.000404 | 0.000564 | 0.000387 | 0.000471 |
| 45 | 0.000439 | 0.000612 | 0.000401 | 0.000488 |
| 46 | 0.000480 | 0.000669 | 0.000419 | 0.000510 |
| 47 | 0.000528 | 0.000736 | 0.000441 | 0.000536 |
| 48 | 0.000585 | 0.000815 | 0.000467 | 0.000569 |
| 49 | 0.000650 | 0.000907 | 0.000500 | 0.000608 |
| 50 | 0.000727 | 0.001014 | 0.000540 | 0.000656 |
| 51 | 0.000816 | 0.001138 | 0.000588 | 0.000715 |
| 52 | 0.000920 | 0.001283 | 0.000647 | 0.000786 |
| 53 | 0.001041 | 0.001451 | 0.000717 | 0.000872 |
| 54 | 0.001181 | 0.001647 | 0.000803 | 0.000976 |

Table A7. (Continued).

| Age $x$ | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 55 | 0.001343 | 0.001873 | 0.000906 | 0.001101 |
| 56 | 0.001531 | 0.002134 | 0.001029 | 0.001252 |
| 57 | 0.001747 | 0.002435 | 0.001177 | 0.001432 |
| 58 | 0.001997 | 0.002783 | 0.001354 | 0.001647 |
| 59 | 0.002284 | 0.003183 | 0.001565 | 0.001903 |
| 60 | 0.002613 | 0.003642 | 0.001815 | 0.002207 |
| 61 | 0.002991 | 0.004168 | 0.002111 | 0.002567 |
| 62 | 0.003424 | 0.004771 | 0.002462 | 0.002993 |
| 63 | 0.003920 | 0.005461 | 0.002875 | 0.003496 |
| 64 | 0.004485 | 0.006247 | 0.003361 | 0.004086 |
| 65 | 0.005129 | 0.007144 | 0.003931 | 0.004779 |
| 66 | 0.005862 | 0.008164 | 0.004598 | 0.005589 |
| 67 | 0.006695 | 0.009322 | 0.005376 | 0.006534 |
| 68 | 0.007639 | 0.010635 | 0.006280 | 0.007632 |
| 69 | 0.008708 | 0.012120 | 0.007329 | 0.008907 |
| 70 | 0.009917 | 0.013799 | 0.008543 | 0.010381 |
| 71 | 0.011280 | 0.015692 | 0.009944 | 0.012080 |
| 72 | 0.012816 | 0.017824 | 0.011555 | 0.014035 |
| 73 | 0.014544 | 0.020219 | 0.013403 | 0.016276 |
| 74 | 0.016482 | 0.022905 | 0.015516 | 0.018839 |
| 75 | 0.018655 | 0.025913 | 0.017927 | 0.021759 |
| 76 | 0.021084 | 0.029274 | 0.020667 | 0.025078 |
| 77 | 0.023797 | 0.033022 | 0.023774 | 0.028838 |
| 78 | 0.026819 | 0.037193 | 0.027284 | 0.033084 |
| 79 | 0.030180 | 0.041826 | 0.031239 | 0.037862 |
| 80 | 0.033910 | 0.046960 | 0.035680 | 0.043224 |
| 81 | 0.038042 | 0.052639 | 0.040650 | 0.049218 |
| 82 | 0.042610 | 0.058905 | 0.046195 | 0.055897 |
| 83 | 0.047649 | 0.065804 | 0.052360 | 0.063313 |
| 84 | 0.053196 | 0.073381 | 0.059191 | 0.071518 |
| 85 | 0.059289 | 0.081684 | 0.066732 | 0.080561 |
| 86 | 0.065966 | 0.090759 | 0.075028 | 0.090491 |
| 87 | 0.073266 | 0.100653 | 0.084120 | 0.101351 |
| 88 | 0.081230 | 0.111409 | 0.094047 | 0.113183 |
| 89 | 0.089896 | 0.123073 | 0.104843 | 0.126020 |
| 90 | 0.099302 | 0.135683 | 0.116540 | 0.139889 |
| 91 | 0.109485 | 0.149278 | 0.129160 | 0.154808 |
| 92 | 0.120482 | 0.163888 | 0.142720 | 0.170787 |
| 93 | 0.132322 | 0.179541 | 0.157229 | 0.187822 |
| 94 | 0.145037 | 0.196255 | 0.172685 | 0.205901 |

Table A7. (Continued).

| Age $x$ | Lives |  | Amounts |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 95 | 0.158651 | 0.214043 | 0.189079 | 0.224997 |
| 96 | 0.173184 | 0.232906 | 0.206388 | 0.245069 |
| 97 | 0.188651 | 0.252839 | 0.224581 | 0.266063 |
| 98 | 0.205059 | 0.273821 | 0.243613 | 0.287913 |
| 99 | 0.222408 | 0.295822 | 0.263430 | 0.310537 |
| 100 | 0.240692 | 0.318799 | 0.283964 | 0.333841 |
| 101 |  | 0.342696 |  | 0.357722 |
| 102 |  | 0.367443 |  | 0.382063 |
| 103 |  | 0.392956 |  | 0.406741 |
| 104 |  | 0.419139 |  | 0.431627 |
| 105 |  | 0.445883 |  | 0.456587 |
| 106 |  | 0.473069 |  | 0.481486 |
| 107 |  | 0.500566 |  | 0.506192 |
| 108 |  | 0.528237 |  | 0.530573 |
| 109 |  | 0.555937 |  | 0.554505 |
| 110 |  | 0.583518 |  | 0.577873 |
| 111 |  | 0.610832 |  | 0.600570 |
| 112 |  | 0.637731 |  | 0.622502 |
| 113 |  | 0.664072 |  | 0.643586 |
| 114 |  | 0.689719 |  | 0.663752 |
| 115 |  | 0.714545 |  | 0.682946 |
| 116 |  | 0.738437 |  | 0.701125 |
| 117 |  | 0.761294 |  | 0.718260 |
| 118 |  | 0.783030 |  | 0.734334 |
| 119 |  | 0.803579 |  | 0.749341 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A8. Retirement annuitants - RMV92Base and RFV92Base values of $q_{x}$

| Age $x$ | Males <br> RMV92Base | Females <br> RFV92Base |
| :--- | :---: | :---: |
| 17 | 0.051376 | 0.042165 |
| 18 | 0.050428 | 0.041210 |
| 19 | 0.049479 | 0.040254 |
| 20 | 0.048530 | 0.039297 |
| 21 | 0.047580 | 0.038340 |
| 22 | 0.046629 | 0.037384 |
| 23 | 0.045678 | 0.036427 |
| 24 | 0.044726 | 0.035470 |
| 25 | 0.043775 | 0.034514 |
| 26 | 0.042823 | 0.033558 |
| 27 | 0.041871 | 0.032603 |
| 28 | 0.040919 | 0.031649 |
| 29 | 0.039968 | 0.030697 |
| 30 | 0.039017 | 0.029747 |
| 31 | 0.038068 | 0.028798 |
| 32 | 0.037119 | 0.027853 |
| 33 | 0.036173 | 0.026911 |
| 34 | 0.035228 | 0.025973 |
| 35 | 0.034286 | 0.025039 |
| 36 | 0.033348 | 0.024112 |
| 37 | 0.032413 | 0.023190 |
| 38 | 0.031483 | 0.022276 |
| 39 | 0.030560 | 0.021371 |
| 40 | 0.029643 | 0.020476 |
| 41 | 0.028734 | 0.019592 |
| 42 | 0.027835 | 0.018721 |
| 43 | 0.026947 | 0.017865 |
| 44 | 0.026072 | 0.017026 |
| 45 | 0.025213 | 0.016206 |
| 46 | 0.024372 | 0.015407 |
| 47 | 0.023551 | 0.014633 |
| 48 | 0.022753 | 0.013887 |
| 49 | 0.021983 | 0.013171 |
| 50 | 0.021244 | 0.012489 |
| 51 | 0.01989697 | 0.011845 |
| 52 | 0.011244 |  |
| 53 | 0.010689 |  |
| 54 |  | 0.010187 |
|  |  |  |

Table A8. (Continued).

| Age $x$ | Males RMV92Base | Females RFV92Base |
| :---: | :---: | :---: |
| 55 | 0.018192 | 0.009742 |
| 56 | 0.017755 | 0.009361 |
| 57 | 0.017393 | 0.009049 |
| 58 | 0.017116 | 0.008814 |
| 59 | 0.016934 | 0.008662 |
| 60 | 0.016859 | 0.008601 |
| 61 | 0.016903 | 0.008640 |
| 62 | 0.017080 | 0.008786 |
| 63 | 0.017404 | 0.009049 |
| 64 | 0.017892 | 0.009438 |
| 65 | 0.018562 | 0.009962 |
| 66 | 0.019431 | 0.010633 |
| 67 | 0.020520 | 0.011460 |
| 68 | 0.021850 | 0.012455 |
| 69 | 0.023444 | 0.013628 |
| 70 | 0.025328 | 0.014990 |
| 71 | 0.027525 | 0.016554 |
| 72 | 0.030064 | 0.018330 |
| 73 | 0.032972 | 0.020330 |
| 74 | 0.036279 | 0.022565 |
| 75 | 0.039971 | 0.025064 |
| 76 | 0.043988 | 0.027855 |
| 77 | 0.048353 | 0.030953 |
| 78 | 0.053089 | 0.034391 |
| 79 | 0.058219 | 0.038206 |
| 80 | 0.063767 | 0.042436 |
| 81 | 0.069756 | 0.047125 |
| 82 | 0.076213 | 0.052319 |
| 83 | 0.083160 | 0.058070 |
| 84 | 0.090623 | 0.064432 |
| 85 | 0.098625 | 0.071467 |
| 86 | 0.107189 | 0.079238 |
| 87 | 0.116337 | 0.087816 |
| 88 | 0.126091 | 0.097273 |
| 89 | 0.136470 | 0.107689 |
| 90 | 0.147491 | 0.119146 |
| 91 | 0.159170 | 0.131732 |
| 92 | 0.171520 | 0.145535 |
| 93 | 0.184549 | 0.160648 |
| 94 | 0.198265 | 0.177163 |

Table A8. (Continued).

| Age $x$ | Males <br> RMV92Base | Females <br> RFV92Base |
| :--- | :---: | :---: |
| 95 | 0.212671 | 0.195171 |
| 96 | 0.227764 | 0.214760 |
| 97 | 0.243539 | 0.236014 |
| 98 | 0.259986 | 0.259005 |
| 99 | 0.277089 | 0.283793 |
| 100 | 0.294827 | 0.310422 |
| 101 | 0.313175 | 0.338912 |
| 102 | 0.332100 | 0.369254 |
| 103 | 0.351567 | 0.401405 |
| 104 | 0.371532 | 0.435283 |
| 105 | 0.391947 | 0.470757 |
| 106 | 0.412761 | 0.507643 |
| 107 | 0.433916 | 0.545701 |
| 108 | 0.455349 | 0.584630 |
| 109 | 0.476996 | 0.624066 |
| 110 | 0.498787 | 0.663591 |
| 111 | 0.520651 | 0.702735 |
| 112 | 0.542516 | 0.740987 |
| 113 | 0.564306 | 0.777821 |
| 114 | 0.585948 | 0.812707 |
| 115 | 0.607369 | 0.845149 |
| 116 | 0.628496 | 0.874707 |
| 117 | 0.649259 | 0.901033 |
| 118 | 0.669594 | 0.923893 |
| 119 | 0.689438 | 0.943192 |
| 120 | 1.000000 | 1.000000 |
|  |  |  |

Table A9. Widows - WL92Base and WA92Base values of $q_{x}$

| Age $x$ | Lives WL92Base | Amounts WA92Base |
| :---: | :---: | :---: |
| 17 | 0.000444 | 0.000366 |
| 18 | 0.000466 | 0.000384 |
| 19 | 0.000489 | 0.000403 |
| 20 | 0.000515 | 0.000423 |
| 21 | 0.000544 | 0.000446 |
| 22 | 0.000575 | 0.000471 |
| 23 | 0.000609 | 0.000498 |
| 24 | 0.000645 | 0.000527 |
| 25 | 0.000685 | 0.000559 |
| 26 | 0.000729 | 0.000595 |
| 27 | 0.000777 | 0.000633 |
| 28 | 0.000829 | 0.000675 |
| 29 | 0.000885 | 0.000721 |
| 30 | 0.000947 | 0.000772 |
| 31 | 0.001014 | 0.000827 |
| 32 | 0.001087 | 0.000887 |
| 33 | 0.001167 | 0.000953 |
| 34 | 0.001254 | 0.001024 |
| 35 | 0.001349 | 0.001103 |
| 36 | 0.001452 | 0.001189 |
| 37 | 0.001565 | 0.001282 |
| 38 | 0.001687 | 0.001385 |
| 39 | 0.001821 | 0.001497 |
| 40 | 0.001967 | 0.001619 |
| 41 | 0.002126 | 0.001752 |
| 42 | 0.002299 | 0.001899 |
| 43 | 0.002488 | 0.002058 |
| 44 | 0.002693 | 0.002233 |
| 45 | 0.002917 | 0.002423 |
| 46 | 0.003161 | 0.002631 |
| 47 | 0.003428 | 0.002859 |
| 48 | 0.003718 | 0.003108 |
| 49 | 0.004034 | 0.003379 |
| 50 | 0.004378 | 0.003676 |
| 51 | 0.004753 | 0.004000 |
| 52 | 0.005162 | 0.004355 |
| 53 | 0.005607 | 0.004742 |
| 54 | 0.006092 | 0.005165 |

Table A9. (Continued).

| Age $x$ | Lives <br> WL92Base | Amounts <br> WA92Base |
| :--- | :---: | :---: |
| 55 | 0.006621 | 0.005627 |
| 56 | 0.007197 | 0.006132 |
| 57 | 0.007824 | 0.006683 |
| 58 | 0.008507 | 0.007286 |
| 59 | 0.009252 | 0.007944 |
| 60 | 0.010062 | 0.008662 |
| 61 | 0.010945 | 0.009447 |
| 62 | 0.011906 | 0.010304 |
| 63 | 0.012953 | 0.011240 |
| 64 | 0.014093 | 0.012262 |
| 65 | 0.015334 | 0.013378 |
| 66 | 0.016685 | 0.014596 |
| 67 | 0.018155 | 0.015925 |
| 68 | 0.019755 | 0.017376 |
| 69 | 0.021496 | 0.018959 |
| 70 | 0.023390 | 0.020687 |
| 71 | 0.025450 | 0.022572 |
| 72 | 0.027691 | 0.024627 |
| 73 | 0.030128 | 0.026869 |
| 74 | 0.032777 | 0.029314 |
| 75 | 0.035657 | 0.031978 |
| 76 | 0.038786 | 0.034882 |
| 77 | 0.042184 | 0.038046 |
| 78 | 0.045875 | 0.041492 |
| 79 | 0.049882 | 0.045245 |
| 80 | 0.054231 | 0.049329 |
| 81 | 0.058948 | 0.053773 |
| 82 | 0.064063 | 0.058606 |
| 83 | 0.069606 | 0.063861 |
| 84 | 0.075611 | 0.069570 |
| 85 | 0.082112 | 0.075770 |
| 86 | 0.089147 | 0.082500 |
| 87 | 0.096753 | 0.089798 |
| 88 | 0.104971 | 0.097709 |
| 89 | 0.113844 | 0.106277 |
| 90 | 0.123415 | 0.115548 |
| 91 | 0.133730 | 0.125571 |
| 92 | 0.144835 | 0.136395 |
| 94 |  | 0.148073 |
|  |  | 0.160656 |
| 4 |  |  |

Table A9. (Continued).

| Age $x$ | Lives <br> WL92Base | Amounts <br> WA992Base |
| :--- | :---: | :---: |
| 95 | 0.183364 | 0.174195 |
| 96 | 0.198102 | 0.188744 |
| 97 | 0.213864 | 0.204352 |
| 98 | 0.230692 | 0.221067 |
| 99 | 0.248625 | 0.238935 |
| 100 | 0.267695 | 0.257996 |
| 101 | 0.287931 | 0.278283 |
| 102 | 0.309351 | 0.299821 |
| 103 | 0.331967 | 0.322627 |
| 104 | 0.355774 | 0.346705 |
| 105 | 0.380759 | 0.372042 |
| 106 | 0.406891 | 0.398613 |
| 107 | 0.434119 | 0.426369 |
| 108 | 0.462377 | 0.455242 |
| 109 | 0.491573 | 0.485139 |
| 110 | 0.521593 | 0.515942 |
| 111 | 0.552300 | 0.547505 |
| 112 | 0.583529 | 0.579652 |
| 113 | 0.615094 | 0.612181 |
| 114 | 0.646780 | 0.644862 |
| 115 | 0.678355 | 0.677439 |
| 116 | 0.709565 | 0.709637 |
| 117 | 0.740144 | 0.741163 |
| 118 | 0.769821 | 0.771718 |
| 119 | 0.798321 | 0.801003 |
| 120 | 1.000000 | 1.000000 |

Table A10. Pensioners - PML92, PMA92, PFL92 and PFA92 ( $C=2020$ ) values of $q_{x}$ for calendar year 2020

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92C20 } \end{gathered}$ | $\begin{aligned} & \text { Amounts } \\ & \text { PMA92C20 } \end{aligned}$ | $\begin{gathered} \text { Lives } \\ \text { PFL92C20 } \end{gathered}$ | Amounts PFA92C20 |
| 20 | 0.000258 | 0.000141 | 0.000105 | 0.000084 |
| 21 | 0.000254 | 0.000140 | 0.000105 | 0.000084 |
| 22 | 0.000249 | 0.000140 | 0.000105 | 0.000084 |
| 23 | 0.000245 | 0.000139 | 0.000104 | 0.000084 |
| 24 | 0.000241 | 0.000138 | 0.000104 | 0.000084 |
| 25 | 0.000237 | 0.000138 | 0.000105 | 0.000084 |
| 26 | 0.000234 | 0.000138 | 0.000106 | 0.000085 |
| 27 | 0.000232 | 0.000137 | 0.000107 | 0.000085 |
| 28 | 0.000230 | 0.000137 | 0.000108 | 0.000087 |
| 29 | 0.000229 | 0.000138 | 0.000110 | 0.000088 |
| 30 | 0.000229 | 0.000138 | 0.000112 | 0.000090 |
| 31 | 0.000230 | 0.000139 | 0.000115 | 0.000092 |
| 32 | 0.000233 | 0.000140 | 0.000119 | 0.000095 |
| 33 | 0.000237 | 0.000142 | 0.000124 | 0.000099 |
| 34 | 0.000243 | 0.000145 | 0.000130 | 0.000103 |
| 35 | 0.000251 | 0.000148 | 0.000137 | 0.000109 |
| 36 | 0.000262 | 0.000152 | 0.000145 | 0.000116 |
| 37 | 0.000275 | 0.000157 | 0.000156 | 0.000124 |
| 38 | 0.000292 | 0.000164 | 0.000168 | 0.000133 |
| 39 | 0.000313 | 0.000173 | 0.000183 | 0.000145 |
| 40 | 0.000339 | 0.000184 | 0.000200 | 0.000158 |
| 41 | 0.000370 | 0.000197 | 0.000220 | 0.000174 |
| 42 | 0.000407 | 0.000213 | 0.000244 | 0.000194 |
| 43 | 0.000451 | 0.000233 | 0.000272 | 0.000216 |
| 44 | 0.000504 | 0.000257 | 0.000305 | 0.000243 |
| 45 | 0.000566 | 0.000286 | 0.000344 | 0.000274 |
| 46 | 0.000639 | 0.000321 | 0.000388 | 0.000311 |
| 47 | 0.000724 | 0.000363 | 0.000440 | 0.000353 |
| 48 | 0.000823 | 0.000413 | 0.000499 | 0.000403 |
| 49 | 0.000938 | 0.000473 | 0.000568 | 0.000460 |
| 50 | 0.001071 | 0.000545 | 0.000648 | 0.000527 |
| 51 | 0.001224 | 0.000630 | 0.000739 | 0.000603 |
| 52 | 0.001400 | 0.000730 | 0.000844 | 0.000692 |
| 53 | 0.001602 | 0.000848 | 0.000964 | 0.000795 |
| 54 | 0.001833 | 0.000986 | 0.001101 | 0.000912 |

Table A10. (Continued).

| Age $x$ | Mates |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92C20 } \end{gathered}$ | Amounts <br> PMA92C20 | $\begin{gathered} \text { Lives } \\ \text { PFL92C20 } \end{gathered}$ | Amounts PFA92C20 |
| 55 | 0.002096 | 0.001148 | 0.001258 | 0.001046 |
| 56 | 0.002396 | 0.001338 | 0.001437 | 0.001199 |
| 57 | 0.002735 | 0.001558 | 0.001639 | 0.001375 |
| 58 | 0.003120 | 0.001814 | 0.001869 | 0.001575 |
| 59 | 0.003554 | 0.002110 | 0.002130 | 0.001801 |
| 60 | 0.004042 | 0.002451 | 0.002424 | 0.002058 |
| 61 | 0.004770 | 0.002955 | 0.002864 | 0.002441 |
| 62 | 0.005611 | 0.003550 | 0.003374 | 0.002885 |
| 63 | 0.006578 | 0.004251 | 0.003964 | 0.003401 |
| 64 | 0.007685 | 0.005073 | 0.004642 | 0.003996 |
| 65 | 0.008950 | 0.006032 | 0.005422 | 0.004681 |
| 66 | 0.010391 | 0.007147 | 0.006315 | 0.005467 |
| 67 | 0.012025 | 0.008439 | 0.007336 | 0.006367 |
| 68 | 0.013873 | 0.009930 | 0.008500 | 0.007395 |
| 69 | 0.015956 | 0.011644 | 0.009821 | 0.008563 |
| 70 | 0.018296 | 0.013605 | 0.011319 | 0.009888 |
| 71 | 0.020916 | 0.015841 | 0.013011 | 0.011386 |
| 72 | 0.023841 | 0.018380 | 0.014918 | 0.013075 |
| 73 | 0.027094 | 0.021253 | 0.017062 | 0.014973 |
| 74 | 0.030702 | 0.024490 | 0.019465 | 0.017100 |
| 75 | 0.034689 | 0.028121 | 0.022151 | 0.019478 |
| 76 | 0.039082 | 0.032179 | 0.025145 | 0.022127 |
| 77 | 0.043904 | 0.036696 | 0.028474 | 0.025069 |
| 78 | 0.049182 | 0.041702 | 0.032166 | 0.028326 |
| 79 | 0.054939 | 0.047229 | 0.036248 | 0.031923 |
| 80 | 0.061197 | 0.053303 | 0.040749 | 0.035882 |
| 81 | 0.067976 | 0.059952 | 0.045698 | 0.040227 |
| 82 | 0.075297 | 0.067201 | 0.05124 | 0.044981 |
| 83 | 0.083174 | 0.075068 | 0.057057 | 0.050166 |
| 84 | 0.091621 | 0.083569 | 0.063526 | 0.055804 |
| 85 | 0.100650 | 0.092716 | 0.070558 | 0.061915 |
| 86 | 0.110265 | 0.102516 | 0.078180 | 0.068518 |
| 87 | 0.120470 | 0.112969 | 0.086418 | 0.075631 |
| 88 | 0.131262 | 0.124068 | 0.095295 | 0.083270 |
| 89 | 0.142636 | 0.135802 | 0.104830 | 0.091444 |

Table A10. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92C20 } \end{gathered}$ | $\begin{gathered} \text { Amounts } \\ \text { PMA92C20 } \end{gathered}$ | $\begin{gathered} \text { Lives } \\ \text { PFL92C20 } \end{gathered}$ | Amounts PFA92C20 |
| 90 | 0.154580 | 0.148151 | 0.115043 | 0.100164 |
| 91 | 0.167078 | 0.161088 | 0.125947 | 0.109437 |
| 92 | 0.180110 | 0.174581 | 0.137554 | 0.119266 |
| 93 | 0.193650 | 0.188589 | 0.149868 | 0.129650 |
| 94 | 0.207667 | 0.203065 | 0.162894 | 0.140582 |
| 95 | 0.222124 | 0.217957 | 0.176626 | 0.152053 |
| 96 | 0.236984 | 0.233205 | 0.191057 | 0.164051 |
| 97 | 0.252201 | 0.248746 | 0.206174 | 0.176555 |
| 98 | 0.267726 | 0.264511 | 0.221955 | 0.189545 |
| 99 | 0.283510 | 0.280429 | 0.238377 | 0.202991 |
| 100 | 0.299498 | 0.296425 | 0.255408 | 0.216863 |
| 101 | 0.315633 | 0.312423 | 0.273011 | 0.231125 |
| 102 | 0.331857 | 0.328344 | 0.291144 | 0.245737 |
| 103 | 0.348109 | 0.344113 | 0.309759 | 0.260654 |
| 104 | 0.364331 | 0.359653 | 0.328806 | 0.275830 |
| 105 | 0.380462 | 0.374887 | 0.348227 | 0.291217 |
| 106 | 0.396441 | 0.389742 | 0.367961 | 0.306760 |
| 107 | 0.412210 | 0.404148 | 0.387946 | 0.322407 |
| 108 | 0.427710 | 0.418037 | 0.408114 | 0.338100 |
| 109 | 0.442886 | 0.431346 | 0.428398 | 0.353785 |
| 110 | 0.457684 | 0.444014 | 0.448729 | 0.369403 |
| 111 | 0.468997 | 0.453033 | 0.466001 | 0.382406 |
| 112 | 0.479798 | 0.461297 | 0.483064 | 0.395150 |
| 113 | 0.490060 | 0.468780 | 0.499869 | 0.407594 |
| 114 | 0.499756 | 0.475459 | 0.516366 | 0.419700 |
| 115 | 0.508865 | 0.481313 | 0.532508 | 0.431431 |
| 116 | 0.517368 | 0.486326 | 0.548254 | 0.442752 |
| 117 | 0.525248 | 0.490484 | 0.563564 | 0.453630 |
| 118 | 0.532489 | 0.493776 | 0.578402 | 0.464038 |
| 119 | 0.539080 | 0.496194 | 0.592739 | 0.473947 |
| 120 | 1.000000 | 1.000000 | 1.000000 | 1.000000 |

Table Al1. Immediate annuitants, males - IML92 and IMA92
( $C=2020$ ) one year select: values of $q_{[x-t]+i}$ for calendar year 2020

| Age $x$ | Lives - IML92C20 |  | Amounts - IMA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 17 | 0.000179 | 0.000243 | 0.000190 | 0.000232 |
| 18 | 0.000178 | 0.000241 | 0.000188 | 0.000230 |
| 19 | 0.000177 | 0.000239 | 0.000187 | 0.000228 |
| 20 | 0.000175 | 0.000237 | 0.000185 | 0.000226 |
| 21 | 0.000174 | 0.000236 | 0.000184 | 0.000225 |
| 22 | 0.000173 | 0.000235 | 0.000183 | 0.000224 |
| 23 | 0.000173 | 0.000234 | 0.000183 | 0.000223 |
| 24 | 0.000172 | 0.000234 | 0.000182 | 0.000223 |
| 25 | 0.000173 | 0.000234 | 0.000182 | 0.000223 |
| 26 | 0.000173 | 0.000235 | 0.000183 | 0.000224 |
| 27 | 0.000174 | 0.000236 | 0.000184 | 0.000225 |
| 28 | 0.000176 | 0.000239 | 0.000186 | 0.000228 |
| 29 | 0.000178 | 0.000242 | 0.000189 | 0.000230 |
| 30 | 0.000182 | 0.000246 | 0.000192 | 0.000235 |
| 31 | 0.000186 | 0.000251 | 0.000196 | 0.000240 |
| 32 | 0.000191 | 0.000258 | 0.000201 | 0.000246 |
| 33 | 0.000196 | 0.000266 | 0.000208 | 0.000254 |
| 34 | 0.000204 | 0.000276 | 0.000216 | 0.000264 |
| 35 | 0.000213 | 0.000288 | 0.000225 | 0.000275 |
| 36 | 0.000224 | 0.000303 | 0.000236 | 0.000289 |
| 37 | 0.000236 | 0.000320 | 0.000250 | 0.000305 |
| 38 | 0.000251 | 0.000340 | 0.000265 | 0.000325 |
| 39 | 0.000269 | 0.000364 | 0.000284 | 0.000347 |
| 40 | 0.000289 | 0.000392 | 0.000305 | 0.000373 |
| 41 | 0.000313 | 0.000424 | 0.000330 | 0.000404 |
| 42 | 0.000340 | 0.000460 | 0.000359 | 0.000439 |
| 43 | 0.000371 | 0.000504 | 0.000392 | 0.000480 |
| 44 | 0.000408 | 0.000552 | 0.000431 | 0.000527 |
| 45 | 0.000450 | 0.000609 | 0.000475 | 0.000581 |
| 46 | 0.000498 | 0.000674 | 0.000526 | 0.000643 |
| 47 | 0.000552 | 0.000749 | 0.000584 | 0.000714 |
| 48 | 0.000615 | 0.000833 | 0.000650 | 0.000795 |
| 49 | 0.000686 | 0.000930 | 0.000725 | 0.000887 |
| 50 | 0.000768 | 0.001039 | 0.000811 | 0.000991 |
| 51 | 0.000860 | 0.001164 | 0.000908 | 0.001110 |
| 52 | 0.000964 | 0.001306 | 0.001018 | 0.001245 |
| 53 | 0.001082 | 0.001466 | 0.001143 | 0.001398 |
| 54 | 0.001216 | 0.001647 | 0.001284 | 0.001570 |

Table A11. (Continued).

| Age $x$ | Lives - IML92C20 |  | Amounts -- IMA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 55 | 0.001367 | 0.001851 | 0.001444 | 0.001765 |
| 56 | 0.001537 | 0.002081 | 0.001623 | 0.001984 |
| 57 | 0.001728 | 0.002340 | 0.001826 | 0.002231 |
| 58 | 0.001943 | 0.002631 | 0.002053 | 0.002509 |
| 59 | 0.002185 | 0.002958 | 0.002308 | 0.002820 |
| 60 | 0.002456 | 0.003324 | 0.002594 | 0.003169 |
| 61 | 0.002867 | 0.003880 | 0.003028 | 0.003700 |
| 62 | 0.003339 | 0.004518 | 0.003527 | 0.004309 |
| 63 | 0.003880 | 0.005250 | 0.004099 | 0.005007 |
| 64 | 0.004499 | 0.006086 | 0.004753 | 0.005805 |
| 65 | 0.005206 | 0.007041 | 0.005499 | 0.006715 |
| 66 | 0.006010 | 0.008127 | 0.006349 | 0.007752 |
| 67 | 0.006925 | 0.009362 | 0.007314 | 0.008929 |
| 68 | 0.007963 | 0.010761 | 0.008410 | 0.010265 |
| 69 | 0.009137 | 0.012344 | 0.009650 | 0.011776 |
| 70 | 0.010463 | 0.014131 | 0.011050 | 0.013481 |
| 71 | 0.011959 | 0.016145 | 0.012628 | 0.015403 |
| 72 | 0.013641 | 0.018408 | 0.014404 | 0.017564 |
| 73 | 0.015530 | 0.020948 | 0.016397 | 0.019988 |
| 74 | 0.017647 | 0.023791 | 0.018630 | 0.022703 |
| 75 | 0.020014 | 0.026968 | 0.021128 | 0.025737 |
| 76 | 0.022657 | 0.030510 | 0.023915 | 0.029121 |
| 77 | 0.025600 | 0.034449 | 0.027019 | 0.032885 |
| 78 | 0.028872 | 0.038824 | 0.030468 | 0.037066 |
| 79 | 0.032501 | 0.043668 | 0.034294 | 0.041698 |
| 80 | 0.036519 | 0.049024 | 0.038529 | 0.046819 |
| 81 | 0.040958 | 0.054931 | 0.043205 | 0.052469 |
| 82 | 0.045853 | 0.061430 | 0.048361 | 0.058687 |
| 83 | 0.051237 | 0.068565 | 0.054029 | 0.065517 |
| 84 | 0.057149 | 0.076379 | 0.060250 | 0.073001 |
| 85 | 0.063623 | 0.084918 | 0.067063 | 0.081181 |
| 86 | 0.070701 | 0.094225 | 0.074504 | 0.090103 |
| 87 | 0.078418 | 0.104345 | 0.082616 | 0.099807 |
| 88 | 0.086815 | 0.115320 | 0.091438 | 0.110338 |
| 89 | 0.095929 | 0.127192 | 0.101007 | 0.121737 |
| 90 | 0.105799 | 0.139998 | 0.111365 | 0.134042 |
| 91 | 0.116462 | 0.153777 | 0.122546 | 0.147289 |
| 92 | 0.127952 | 0.168558 | 0.134585 | 0.161513 |
| 93 | 0.140302 | 0.184369 | 0.147516 | 0.176741 |
| 94 | 0.153541 | 0.201233 | 0.161367 | 0.192997 |

Table A11. (Continued).

| Age $x$ | Lives - IML92C20 |  | Amounts - IMA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 95 | 0.167696 | 0.219163 | 0.176164 | 0.210298 |
| 96 | 0.182789 | 0.238168 | 0.191926 | 0.228655 |
| 97 | 0.198837 | 0.258246 | 0.208667 | 0.248071 |
| 98 | 0.215852 | 0.279388 | 0.226398 | 0.268541 |
| 99 | 0.233840 | 0.301574 | 0.245120 | 0.290049 |
| 100 | 0.252797 | 0.324773 | 0.264828 | 0.312572 |
| 101 |  | 0.348945 |  | 0.336073 |
| 102 |  | 0.374036 |  | 0.360507 |
| 103 |  | 0.399981 |  | 0.385816 |
| 104 |  | 0.426704 |  | 0.411929 |
| 105 |  | 0.454117 |  | 0.438769 |
| 106 |  | 0.482123 |  | 0.466245 |
| 107 |  | 0.510613 |  | 0.494255 |
| 108 |  | 0.539468 |  | 0.522688 |
| 109 |  | 0.568565 |  | 0.551428 |
| 110 |  | 0.597771 |  | 0.580348 |
| 111 |  | 0.622892 |  | 0.605374 |
| 112 |  | 0.647670 |  | 0.630135 |
| 113 |  | 0.671993 |  | 0.654519 |
| 114 |  | 0.695751 |  | 0.678419 |
| 115 |  | 0.718842 |  | 0.701731 |
| 116 |  | 0.741170 |  | 0.724359 |
| 117 |  | 0.762651 |  | 0.746213 |
| 118 |  | 0.783207 |  | 0.767212 |
| 119 |  | 0.802773 |  | 0.787288 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A12. Immediate annuitants, females - IFL92 and IFA92 ( $C=2020$ ) one year select: values of $q_{[x-t]+t}$ for calendar year 2020

| Age $x$ | Lives - IFL92C20 |  | Amounts - IFA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 17 | 0.000090 | 0.000125 | 0.000136 | 0.000166 |
| 18 | 0.000090 | 0.000125 | 0.000136 | 0.000166 |
| 19 | 0.000090 | 0.000125 | 0.000136 | 0.000166 |
| 20 | 0.000090 | 0.000126 | 0.000136 | 0.000166 |
| 21 | 0.000090 | 0.000126 | 0.000136 | 0.000166 |
| 22 | 0.000090 | 0.000126 | 0.000136 | 0.000166 |
| 23 | 0.000091 | 0.000126 | 0.000136 | 0.000166 |
| 24 | 0.000091 | 0.000127 | 0.000136 | 0.000166 |
| 25 | 0.000092 | 0.000128 | 0.000136 | 0.000166 |
| 26 | 0.000092 | 0.000128 | 0.000137 | 0.000166 |
| 27 | 0.000093 | 0.000129 | 0.000137 | 0.000166 |
| 28 | 0.000093 | 0.000130 | 0.000137 | 0.000166 |
| 29 | 0.000094 | 0.000131 | 0.000137 | 0.000167 |
| 30 | 0.000095 | 0.000133 | 0.000137 | 0.000167 |
| 31 | 0.000097 | 0.000135 | 0.000138 | 0.000167 |
| 32 | 0.000098 | 0.000137 | 0.000138 | 0.000167 |
| 33 | 0.000100 | 0.000140 | 0.000138 | 0.000168 |
| 34 | 0.000102 | 0.000143 | 0.000138 | 0.000169 |
| 35 | 0.000105 | 0.000147 | 0.000139 | 0.000170 |
| 36 | 0.000108 | 0.000151 | 0.000140 | 0.000170 |
| 37 | 0.000112 | 0.000156 | 0.000141 | 0.000172 |
| 38 | 0.000116 | 0.000162 | 0.000143 | 0.000173 |
| 39 | 0.000122 | 0.000170 | 0.000144 | 0.000175 |
| 40 | 0.000128 | 0.000179 | 0.000146 | 0.000178 |
| 41 | 0.000136 | 0.000189 | 0.000148 | 0.000181 |
| 42 | 0.000145 | 0.000202 | 0.000152 | 0.000184 |
| 43 | 0.000155 | 0.000216 | 0.000155 | 0.000189 |
| 44 | 0.000167 | 0.000234 | 0.000160 | 0.000195 |
| 45 | 0.000182 | 0.000254 | 0.000166 | 0.000202 |
| 46 | 0.000199 | 0.000277 | 0.000174 | 0.000211 |
| 47 | 0.000219 | 0.000305 | 0.000183 | 0.000222 |
| 48 | 0.000242 | 0.000338 | 0.000194 | 0.000236 |
| 49 | 0.000269 | 0.000376 | 0.000207 | 0.000252 |
| 50 | 0.000301 | 0.000420 | 0.000224 | 0.000272 |
| 51 | 0.000338 | 0.000472 | 0.000244 | 0.000296 |
| 52 | 0.000381 | 0.000532 | 0.000268 | 0.000326 |
| 53 | 0.000431 | 0.000601 | 0.000297 | 0.000361 |
| 54 | 0.000489 | 0.000683 | 0.000333 | 0.000405 |

Table A12. (Continued).

| Age $x$ | Lives - IFL92C20 |  | Amounts - IFA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 55 | 0.000557 | 0.000776 | 0.000375 | 0.000456 |
| 56 | 0.000635 | 0.000884 | 0.000426 | 0.000519 |
| 57 | 0.000724 | 0.001009 | 0.000488 | 0.000594 |
| 58 | 0.000828 | 0.001153 | 0.000561 | 0.000683 |
| 59 | 0.000947 | 0.001319 | 0.000649 | 0.000789 |
| 60 | 0.001083 | 0.001509 | 0.000752 | 0.000915 |
| 61 | 0.001288 | 0.001795 | 0.000909 | 0.001106 |
| 62 | 0.001530 | 0.002131 | 0.001100 | 0.001337 |
| 63 | 0.001814 | 0.002527 | 0.001330 | 0.001617 |
| 64 | 0.002146 | 0.002989 | 0.001608 | 0.001955 |
| 65 | 0.002534 | 0.003529 | 0.001942 | 0.002361 |
| 66 | 0.002986 | 0.004158 | 0.002342 | 0.002847 |
| 67 | 0.003512 | 0.004890 | 0.002820 | 0.003428 |
| 68 | 0.004122 | 0.005739 | 0.003389 | 0.004118 |
| 69 | 0.004829 | 0.006721 | 0.004064 | 0.004939 |
| 70 | 0.005645 | 0.007855 | 0.004863 | 0.005909 |
| 71 | 0.006585 | 0.009160 | 0.005805 | 0.007052 |
| 72 | 0.007665 | 0.010660 | 0.006911 | 0.008394 |
| 73 | 0.008905 | 0.012379 | 0.008206 | 0.009965 |
| 74 | 0.010322 | 0.014344 | 0.009717 | 0.011798 |
| 75 | 0.011940 | 0.016586 | 0.011474 | 0.013927 |
| 76 | 0.013782 | 0.019136 | 0.013510 | 0.016393 |
| 77 | 0.015876 | 0.022030 | 0.015860 | 0.019239 |
| 78 | 0.018247 | 0.025306 | 0.018564 | 0.022510 |
| 79 | 0.020929 | 0.029005 | 0.021663 | 0.026256 |
| 80 | 0.023952 | 0.033170 | 0.025202 | 0.030531 |
| 81 | 0.027354 | 0.037849 | 0.029229 | 0.035390 |
| 82 | 0.031171 | 0.043091 | 0.033794 | 0.040891 |
| 83 | 0.035444 | 0.048949 | 0.038948 | 0.047096 |
| 84 | 0.040215 | 0.055474 | 0.044747 | 0.054066 |
| 85 | 0.045528 | 0.062725 | 0.051244 | 0.061863 |
| 86 | 0.051429 | 0.070759 | 0.058495 | 0.070550 |
| 87 | 0.057966 | 0.079634 | 0.066554 | 0.080186 |
| 88 | 0.065188 | 0.089407 | 0.075474 | 0.090831 |
| 89 | 0.073144 | 0.100139 | 0.085306 | 0.102537 |
| 90 | 0.081884 | 0.111884 | 0.096099 | 0.115352 |
| 91 | 0.091458 | 0.124698 | 0.107893 | 0.129318 |
| 92 | 0.101914 | 0.138630 | 0.120725 | 0.144466 |
| 93 | 0.113297 | 0.153727 | 0.134623 | 0.160818 |
| 94 | 0.125654 | 0.170028 | 0.149607 | 0.178384 |

Table A12. (Continued).

| Age $x$ | Lives - IFL92C20 |  | Amounts - IFA92C20 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations I+ | Duration 0 | Durations 1+ |
| 95 | 0.139025 | 0.187564 | 0.165688 | 0.197163 |
| 96 | 0.153444 | 0.206359 | 0.182863 | 0.217135 |
| 97 | 0.168944 | 0.226427 | 0.201120 | 0.238269 |
| 98 | 0.185547 | 0.247766 | 0.220433 | 0.260517 |
| 99 | 0.203270 | 0.270366 | 0.240762 | 0.283815 |
| 100 | 0.222121 | 0.294201 | 0.262054 | 0.308083 |
| 101 |  | 0.319230 |  | 0.333227 |
| 102 |  | 0.345394 |  | 0.359137 |
| 103 |  | 0.372621 |  | 0.385692 |
| 104 |  | 0.400820 |  | 0.412762 |
| 105 |  | 0.429885 |  | 0.440205 |
| 106 |  | 0.459697 |  | 0.467876 |
| 107 |  | 0.490118 |  | 0.495626 |
| 108 |  | 0.521002 |  | 0.523306 |
| 109 |  | 0.552191 |  | 0.550769 |
| 110 |  | 0.583518 |  | 0.577873 |
| 111 |  | 0.610832 |  | 0.600570 |
| 112 |  | 0.637731 |  | 0.622502 |
| 113 |  | 0.664072 |  | 0.643586 |
| 114 |  | 0.689719 |  | 0.663752 |
| 115 |  | 0.714545 |  | 0.682946 |
| 116 |  | 0.738437 |  | 0.701125 |
| 117 |  | 0.761294 |  | 0.718260 |
| 118 |  | 0.783030 |  | 0.734334 |
| 119 |  | 0.803579 |  | 0.749341 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A13. Retirement annuitants - RMV92 and RFV92
( $C=2020$ ) values of $q_{x}$ for calendar year 2020

| Age $x$ | Males <br> RMV92C20 | Females <br> RFV92C20 |
| :--- | :---: | :---: |
| 17 | 0.021293 | 0.017476 |
| 18 | 0.020900 | 0.017080 |
| 19 | 0.020507 | 0.016684 |
| 20 | 0.020114 | 0.016287 |
| 21 | 0.019720 | 0.015890 |
| 22 | 0.019326 | 0.015494 |
| 23 | 0.018932 | 0.015097 |
| 24 | 0.018537 | 0.014701 |
| 25 | 0.018143 | 0.014305 |
| 26 | 0.017748 | 0.013908 |
| 27 | 0.017354 | 0.013513 |
| 28 | 0.016959 | 0.013117 |
| 29 | 0.016565 | 0.012723 |
| 30 | 0.016171 | 0.012329 |
| 31 | 0.015778 | 0.011936 |
| 32 | 0.015384 | 0.011544 |
| 33 | 0.014992 | 0.011153 |
| 34 | 0.014600 | 0.010765 |
| 35 | 0.014210 | 0.010378 |
| 36 | 0.013821 | 0.009993 |
| 37 | 0.013434 | 0.009611 |
| 38 | 0.013048 | 0.009232 |
| 39 | 0.012666 | 0.008857 |
| 40 | 0.012286 | 0.008486 |
| 41 | 0.011909 | 0.008120 |
| 42 | 0.0111686 | 0.007759 |
| 43 | 0.010806 | 0.007404 |
| 44 | 0.010450 | 0.007057 |
| 45 | 0.010101 | 0.006717 |
| 46 | 0.009761 | 0.006386 |
| 47 | 0.009430 | 0.006065 |
| 48 | 0.009111 | 0.005756 |
| 49 | 0.00088057749 | 0.005459 |
| 50 |  | 0.005176 |
| 51 | 0.004909 |  |
| 52 | 0.004660 |  |
| 54 |  | 0.004222 |
|  |  |  |

Table A13. (Continued).

| Age $x$ | Males <br> RMV92C20 | $\begin{aligned} & \text { Females } \\ & \text { RFV92C20 } \end{aligned}$ |
| :---: | :---: | :---: |
| 55 | 0.007540 | 0.004038 |
| 56 | 0.007359 | 0.003880 |
| 57 | 0.007209 | 0.003750 |
| 58 | 0.007094 | 0.003653 |
| 59 | 0.007018 | 0.003590 |
| 60 | 0.006987 | 0.003565 |
| 61 | 0.007280 | 0.003721 |
| 62 | 0.007631 | 0.003925 |
| 63 | 0.008052 | 0.004187 |
| 64 | 0.008559 | 0.004515 |
| 65 | 0.009169 | 0.004921 |
| 66 | 0.009897 | 0.005416 |
| 67 | 0.010764 | 0.006012 |
| 68 | 0.011791 | 0.006721 |
| 69 | 0.013000 | 0.007557 |
| 70 | 0.014417 | 0.008532 |
| 71 | 0.016067 | 0.009663 |
| 72 | 0.017981 | 0.010963 |
| 73 | 0.020187 | 0.012447 |
| 74 | 0.022720 | 0.014131 |
| 75 | 0.025584 | 0.016042 |
| 76 | 0.028754 | 0.018208 |
| 77 | 0.032258 | 0.020650 |
| 78 | 0.036121 | 0.023399 |
| 79 | 0.040372 | 0.026494 |
| 80 | 0.045041 | 0.029974 |
| 81 | 0.050157 | 0.033885 |
| 82 | 0.055753 | 0.038273 |
| 83 | 0.061859 | 0.043196 |
| 84 | 0.068509 | 0.048709 |
| 85 | 0.075734 | 0.054880 |
| 86 | 0.083568 | 0.061777 |
| 87 | 0.092043 | 0.069478 |
| 88 | 0.101190 | 0.078063 |
| 89 | 0.111040 | 0.087622 |
| 90 | 0.121621 | 0.098248 |
| 91 | 0.132962 | 0.110041 |
| 92 | 0.145086 | 0.123106 |
| 93 | 0.158015 | 0.137551 |
| 94 | 0.171769 | 0.153487 |

Table A13. (Continued).

| Age $x$ | Males <br> RMV92C20 | Females <br> RFV92C20 |
| :--- | :---: | :---: |
| 95 | 0.186362 | 0.171027 |
| 96 | 0.201803 | 0.190281 |
| 97 | 0.218098 | 0.211359 |
| 98 | 0.235248 | 0.234360 |
| 99 | 0.253245 | 0.259373 |
| 100 | 0.272079 | 0.286471 |
| 101 | 0.291730 | 0.315705 |
| 102 | 0.312172 | 0.347096 |
| 103 | 0.333374 | 0.380632 |
| 104 | 0.355293 | 0.416258 |
| 105 | 0.377884 | 0.453867 |
| 106 | 0.401093 | 0.493293 |
| 107 | 0.424859 | 0.534311 |
| 108 | 0.449113 | 0.576623 |
| 109 | 0.473782 | 0.619861 |
| 110 | 0.498787 | 0.663591 |
| 111 | 0.520651 | 0.702735 |
| 112 | 0.542516 | 0.740987 |
| 113 | 0.564306 | 0.777821 |
| 114 | 0.585948 | 0.812707 |
| 115 | 0.607369 | 0.845149 |
| 116 | 0.628496 | 0.874707 |
| 117 | 0.649259 | 0.901033 |
| 118 | 0.669594 | 0.923893 |
| 119 | 0.689438 | 0.943192 |
| 120 | 1.000000 | 1.000000 |

Table A14. Widows - WL92 and WA92 $(C=2020)$ values of $q_{x}$ for calendar year 2020

| Age $x$ | $\begin{gathered} \text { Lives } \\ \text { WL92C } 20 \end{gathered}$ | Amounts <br> WA92C20 |
| :---: | :---: | :---: |
| 17 | 0.000184 | 0.000152 |
| 18 | 0.000193 | 0.000159 |
| 19 | 0.000203 | 0.000167 |
| 20 | 0.000213 | 0.000175 |
| 21 | 0.000225 | 0.000185 |
| 22 | 0.000238 | 0.000195 |
| 23 | 0.000252 | 0.000206 |
| 24 | 0.000267 | 0.000218 |
| 25 | 0.000284 | 0.000232 |
| 26 | 0.000302 | 0.000247 |
| 27 | 0.000322 | 0.000262 |
| 28 | 0.000344 | 0.000280 |
| 29 | 0.000367 | 0.000299 |
| 30 | 0.000392 | 0.000320 |
| 31 | 0.000420 | 0.000343 |
| 32 | 0.000451 | 0.000368 |
| 33 | 0.000484 | 0.000395 |
| 34 | 0.000520 | 0.000424 |
| 35 | 0.000559 | 0.000457 |
| 36 | 0.000602 | 0.000493 |
| 37 | 0.000649 | 0.000531 |
| 38 | 0.000699 | 0.000574 |
| 39 | 0.000755 | 0.000620 |
| 40 | 0.000815 | 0.000671 |
| 41 | 0.000881 | 0.000726 |
| 42 | 0.000953 | 0.000787 |
| 43 | 0.001031 | 0.000853 |
| 44 | 0.001116 | 0.000925 |
| 45 | 0.001209 | 0.001004 |
| 46 | 0.001310 | 0.001090 |
| 47 | 0.001421 | 0.001185 |
| 48 | 0.001541 | 0.001288 |
| 49 | 0.001672 | 0.001400 |
| 50 | 0.001814 | 0.001524 |
| 51 | 0.001970 | 0.001658 |
| 52 | 0.002139 | 0.001805 |
| 53 | 0.002324 | 0.001965 |
| 54 | 0.002525 | 0.002141 |

Table A14. (Continued).

| Age $x$ | Lives <br> WL92C20 | Amounts <br> WA92C20 |
| :--- | :--- | :---: |
| 55 | 0.002744 | 0.002332 |
| 56 | 0.002983 | 0.002541 |
| 57 | 0.003243 | 0.002770 |
| 58 | 0.003526 | 0.003020 |
| 59 | 0.003835 | 0.003292 |
| 60 | 0.004170 | 0.003590 |
| 61 | 0.004714 | 0.004069 |
| 62 | 0.005319 | 0.004603 |
| 63 | 0.005993 | 0.005200 |
| 64 | 0.006742 | 0.005866 |
| 65 | 0.007574 | 0.006608 |
| 66 | 0.008499 | 0.007435 |
| 67 | 0.009524 | 0.008354 |
| 68 | 0.010660 | 0.009377 |
| 69 | 0.011920 | 0.010513 |
| 70 | 0.013314 | 0.011775 |
| 71 | 0.014856 | 0.013176 |
| 72 | 0.016562 | 0.014729 |
| 73 | 0.018446 | 0.016451 |
| 74 | 0.020527 | 0.018358 |
| 75 | 0.022823 | 0.020468 |
| 76 | 0.025354 | 0.022802 |
| 77 | 0.028142 | 0.025382 |
| 78 | 0.031213 | 0.028231 |
| 79 | 0.034591 | 0.031375 |
| 80 | 0.038306 | 0.034843 |
| 81 | 0.042386 | 0.038665 |
| 82 | 0.046865 | 0.042873 |
| 83 | 0.051777 | 0.047503 |
| 84 | 0.057160 | 0.052593 |
| 85 | 0.063054 | 0.058184 |
| 86 | 0.069502 | 0.064320 |
| 87 | 0.076548 | 0.071046 |
| 88 | 0.084241 | 0.078413 |
| 89 | 0.092630 | 0.086473 |
| 90 | 0.101768 | 0.095281 |
| 91 | 0.111710 | 0.104895 |
| 92 | 0.122514 | 0.126784 |
| 93 | 0.13236 |  |
| 94 |  |  |
|  |  |  |

Table A14. (Continued).

| Age $x$ | Lives <br> WL92C20 | Amounts <br> WA92C20 |
| :--- | :---: | :---: |
| 95 | 0.160680 | 0.152646 |
| 96 | 0.175522 | 0.167230 |
| 97 | 0.191523 | 0.183005 |
| 98 | 0.208741 | 0.200032 |
| 99 | 0.227231 | 0.218375 |
| 100 | 0.247040 | 0.238090 |
| 101 | 0.268215 | 0.259227 |
| 102 | 0.290788 | 0.281830 |
| 103 | 0.314788 | 0.305931 |
| 104 | 0.340224 | 0.331552 |
| 105 | 0.367098 | 0.358693 |
| 106 | 0.395389 | 0.387345 |
| 107 | 0.425058 | 0.417470 |
| 108 | 0.456044 | 0.449007 |
| 109 | 0.488261 | 0.481870 |
| 110 | 0.521593 | 0.515942 |
| 111 | 0.552300 | 0.547505 |
| 112 | 0.583529 | 0.579652 |
| 113 | 0.615094 | 0.612181 |
| 114 | 0.646780 | 0.644862 |
| 115 | 0.678355 | 0.677439 |
| 116 | 0.709565 | 0.709637 |
| 117 | 0.740144 | 0.741163 |
| 118 | 0.769821 | 0.771718 |
| 119 | 0.798321 | 0.801003 |
| 120 | 1.000000 | 1.000000 |

Table A15. Pensioners - PML92, PMA92, PFL92 and PFA92
( $B=1945$ ) values of $q_{x}$ for year of birth 1945

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Lives PML92B45 | Amounts PMA92B45 | Lives PFL92B45 | Amounts PFA92B45 |
| 47 | 0.001746 | 0.000876 | 0.001061 | 0.000852 |
| 48 | 0.001917 | 0.000963 | 0.001164 | 0.000939 |
| 49 | 0.002111 | 0.001066 | 0.001279 | 0.001036 |
| 50 | 0.002329 | 0.001186 | 0.001410 | 0.001146 |
| 51 | 0.002574 | 0.001324 | 0.001554 | 0.001269 |
| 52 | 0.002846 | 0.001484 | 0.001715 | 0.001407 |
| 53 | 0.003149 | 0.001666 | 0.001895 | 0.001562 |
| 54 | 0.003484 | 0.001874 | 0.002093 | 0.001733 |
| 55 | 0.003855 | 0.002112 | 0.002313 | 0.001924 |
| 56 | 0.004262 | 0.002380 | 0.002556 | 0.002134 |
| 57 | 0.004710 | 0.002682 | 0.002822 | 0.002367 |
| 58 | 0.005199 | 0.003023 | 0.003115 | 0.002624 |
| 59 | 0.005734 | 0.003404 | 0.003437 | 0.002906 |
| 60 | 0.006317 | 0.003831 | 0.003789 | 0.003216 |
| 61 | 0.007076 | 0.004383 | 0.004248 | 0.003620 |
| 62 | 0.007932 | 0.005019 | 0.004771 | 0.004079 |
| 63 | 0.008898 | 0.005751 | 0.005362 | 0.004600 |
| 64 | 0.009985 | 0.006590 | 0.006031 | 0.005191 |
| 65 | 0.011206 | 0.007552 | 0.006788 | 0.005860 |
| 66 | 0.012577 | 0.008651 | 0.007644 | 0.006618 |
| 67 | 0.014114 | 0.009906 | 0.008611 | 0.007474 |
| 68 | 0.015834 | 0.011334 | 0.009702 | 0.008440 |
| 69 | 0.017756 | 0.012957 | 0.010929 | 0.009529 |
| 70 | 0.019898 | 0.014797 | 0.012310 | 0.010754 |
| 71 | 0.022283 | 0.016876 | 0.053861 | 0.012130 |
| 72 | 0.024931 | 0.019221 | 0.015600 | 0.013673 |
| 73 | 0.027866 | 0.021858 | 0.017548 | 0.015399 |
| 74 | 0.031110 | 0.024815 | 0.019723 | 0.017328 |
| 75 | 0.034689 | 0.028121 | 0.022151 | 0.019478 |
| 76 | 0.038628 | 0.031805 | 0.024853 | 0.021870 |
| 77 | 0.042951 | 0.035899 | 0.027856 | 0.024524 |
| 78 | 0.047684 | 0.040432 | 0.031186 | 0.027463 |
| 79 | 0.052853 | 0.045435 | 0.034871 | 0.030711 |
| 80 | 0.058481 | 0.050938 | 0.038941 | 0.034290 |
| 81 | 0.064593 | 0.056969 | 0.043423 | 0.038225 |
| 82 | 0.071212 | 0.063555 | 0.048351 | 0.042541 |
| 83 | 0.078358 | 0.070721 | 0.053754 | 0.047262 |
| 84 | 0.086051 | 0.078488 | 0.059664 | 0.052411 |

Table A15. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92B45 } \end{gathered}$ | Amounts PMA92B45 | $\begin{gathered} \text { Lives } \\ \text { PFL92B45 } \end{gathered}$ | Amounts PFA92B45 |
| 85 | 0.094309 | 0.086875 | 0.066113 | 0.058015 |
| 86 | 0.103145 | 0.095896 | 0.073132 | 0.064094 |
| 87 | 0.112571 | 0.105561 | 0.080752 | 0.070673 |
| 88 | 0.122594 | 0.115875 | 0.089002 | 0.077771 |
| 89 | 0.133220 | 0.126837 | 0.097910 | 0.085407 |
| 90 | 0.144447 | 0.138440 | 0.107502 | 0.093599 |
| 91 | 0.156273 | 0.150670 | 0.117802 | 0.102360 |
| 92 | 0.168688 | 0.163510 | 0.128830 | 0.111703 |
| 93 | 0.181680 | 0.176932 | 0.140605 | 0.121636 |
| 94 | 0.195231 | 0.190905 | 0.153139 | 0.132163 |
| 95 | 0.209317 | 0.205391 | 0.166443 | 0.143287 |
| 96 | 0.223914 | 0.220343 | 0.180520 | 0.155003 |
| 97 | 0.238988 | 0.235714 | 0.195372 | 0.167305 |
| 98 | 0.254503 | 0.251446 | 0.210992 | 0.180183 |
| 99 | 0.270419 | 0.267481 | 0.227371 | 0.193618 |
| 100 | 0.286695 | 0.283753 | 0.244489 | 0.207593 |
| 101 | 0.303281 | 0.300196 | 0.262327 | 0.222080 |
| 102 | 0.320128 | 0.316740 | 0.280854 | 0.237052 |
| 103 | 0.337183 | 0.333313 | 0.300037 | 0.252473 |
| 104 | 0.354393 | 0.349843 | 0.319837 | 0.268306 |
| 105 | 0.371701 | 0.366254 | 0.340208 | 0.284510 |
| 106 | 0.389048 | 0.382474 | 0.361100 | 0.301040 |
| 107 | 0.406379 | 0.398431 | 0.382459 | 0.317847 |
| 108 | 0.423634 | 0.414053 | 0.404225 | 0.334878 |
| 109 | 0.440754 | 0.429270 | 0.426336 | 0.352083 |
| 110 | 0.457684 | 0.444014 | 0.448729 | 0.369403 |
| 111 | 0.468997 | 0.453033 | 0.466001 | 0.382406 |
| 112 | 0.479798 | 0.461297 | 0.483064 | 0.395150 |
| 113 | 0.490060 | 0.468780 | 0.499869 | 0.407594 |
| 114 | 0.499756 | 0.475459 | 0.516366 | 0.419700 |
| 115 | 0.508865 | 0.481313 | 0.532508 | 0.431431 |
| 116 | 0.517368 | 0.486326 | 0.548254 | 0.442752 |
| 117 | 0.525248 | 0.490484 | 0.563564 | 0.453630 |
| 118 | 0.532489 | 0.493776 | 0.578402 | 0.464038 |
| 119 | 0.539080 | 0.496194 | 0.592739 | 0.473947 |
| 120 | 1.000000 | 1.000000 | 1.000000 | 1.000000 |

Table A16. Immediate annuitants, males - IML92 and JMA92 ( $B=1945$ ) one year select: values of $q_{[x-z]+t}$ for year of birth 1945

| Age $x$ | Lives - IML92B45 |  | Amounts -- IMA92B45 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 47 | 0.001333 | 0.001806 | 0.001408 | 0.001722 |
| 48 | 0.001433 | 0.001942 | 0.001515 | 0.001852 |
| 49 | 0.001545 | 0.002093 | 0.001633 | 0.001996 |
| 50 | 0.001670 | 0.002262 | 0.001764 | 0.002157 |
| 51 | 0.001808 | 0.002448 | 0.001910 | 0.002334 |
| 52 | 0.001960 | 0.002654 | 0.002070 | 0.002531 |
| 53 | 0.002127 | 0.002881 | 0.002248 | 0.002747 |
| 54 | 0.002312 | 0.003130 | 0.002442 | 0.002984 |
| 55 | 0.002513 | 0.003404 | 0.002655 | 0.003245 |
| 56 | 0.002734 | 0.003702 | 0.002888 | 0.003531 |
| 57 | 0.002976 | 0.004029 | 0.003143 | 0.003841 |
| 58 | 0.003239 | 0.004385 | 0.003421 | 0.004181 |
| 59 | 0.003525 | 0.004773 | 0.003725 | 0.004551 |
| 60 | 0.003838 | 0.005194 | 0.004054 | 0.004953 |
| 61 | 0.004252 | 0.005755 | 0.004492 | 0.005488 |
| 62 | 0.004721 | 0.006387 | 0.004987 | 0.006092 |
| 63 | 0.005249 | 0.007102 | 0.005545 | 0.006773 |
| 64 | 0.005845 | 0.007907 | 0.006175 | 0.007541 |
| 65 | 0.006518 | 0.008815 | 0.006884 | 0.008408 |
| 66 | 0.007275 | 0.009837 | 0.007684 | 0.009383 |
| 67 | 0.008128 | 0.010988 | 0.008585 | 0.010481 |
| 68 | 0.009089 | 0.012282 | 0.009599 | 0.011716 |
| 69 | 0.010167 | 0.013737 | 0.010738 | 0.013104 |
| 70 | 0.011380 | 0.015369 | 0.012018 | 0.014662 |
| 71 | 0.012740 | 0.017200 | 0.013453 | 0.016410 |
| 72 | 0.014265 | 0.019250 | 0.015063 | 0.018367 |
| 73 | 0.015972 | 0.021544 | 0.016864 | 0.020558 |
| 74 | 0.017882 | 0.024108 | 0.018878 | 0.023005 |
| 75 | 0.020014 | 0.026968 | 0.021128 | 0.025737 |
| 76 | 0.022394 | 0.030155 | 0.023637 | 0.028783 |
| 77 | 0.025044 | 0.033701 | 0.026432 | 0.032171 |
| 78 | 0.027992 | 0.037641 | 0.029540 | 0.035937 |
| 79 | 0.031267 | 0.042010 | 0.032992 | 0.040114 |
| 80 | 0.034899 | 0.046849 | 0.036819 | 0.044742 |
| 81 | 0.038920 | 0.052197 | 0.041055 | 0.049857 |
| 82 | 0.043365 | 0.058097 | 0.045737 | 0.055503 |
| 83 | 0.048271 | 0.064595 | 0.050901 | 0.061724 |
| 84 | 0.053674 | 0.071736 | 0.056588 | 0.068563 |

Table A16. (Continued).

| Age $x$ | Lives - IML92B4S |  | Amounts - IMA92B45 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations $1+$ | Duration 0 | Durations $1+$ |
| 85 | 0.059615 | 0.079568 | 0.062838 | 0.076067 |
| 86 | 0.066135 | 0.088141 | 0.069694 | 0.084285 |
| 87 | 0.073276 | 0.097503 | 0.077199 | 0.093263 |
| 88 | 0.081082 | 0.107705 | 0.085400 | 0.103052 |
| 89 | 0.089597 | 0.118795 | 0.094339 | 0.113700 |
| 90 | 0.098864 | 0.130821 | 0.104065 | 0.125256 |
| 91 | 0.108930 | 0.143832 | 0.114620 | 0.137764 |
| 92 | 0.119838 | 0.157868 | 0.126050 | 0.151271 |
| 93 | 0.131630 | 0.172973 | 0.138398 | 0.165816 |
| 94 | 0.144346 | 0.189183 | 0.151704 | 0.181439 |
| 95 | 0.158027 | 0.206527 | 0.166007 | 0.198173 |
| 96 | 0.172708 | 0.225032 | 0.181341 | 0.216044 |
| 97 | 0.188420 | 0.244716 | 0.197735 | 0.235074 |
| 98 | 0.205191 | 0.265588 | 0.215216 | 0.255277 |
| 99 | 0.223042 | 0.287649 | 0.233802 | 0.276657 |
| 100 | 0.241990 | 0.310889 | 0.253507 | 0.299210 |
| 101 |  | 0.335289 |  | 0.322921 |
| 102 |  | 0.360816 |  | 0.347766 |
| 103 |  | 0.387426 |  | 0.373706 |
| 104 |  | 0.415064 |  | 0.400692 |
| 105 |  | 0.443660 |  | 0.428665 |
| 106 |  | 0.473133 |  | 0.457551 |
| 107 |  | 0.503391 |  | 0.487264 |
| 108 |  | 0.534327 |  | 0.517707 |
| 109 |  | 0.565829 |  | 0.548775 |
| 110 |  | 0.597771 |  | 0.580348 |
| 111 |  | 0.622892 |  | 0.605374 |
| 112 |  | 0.647670 |  | 0.630135 |
| 113 |  | 0.671993 |  | 0.654519 |
| 114 |  | 0.695751 |  | 0.678419 |
| 115 |  | 0.718842 |  | 0.701731 |
| 116 |  | 0.741170 |  | 0.724359 |
| 117 |  | 0.762651 |  | 0.746213 |
| 118 |  | 0.783207 |  | 0.767212 |
| 119 |  | 0.802773 |  | 0.787288 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A17. Immediate annuitants, females - IFL92 and IFA92 ( $B=1945$ ) one year select: values of $q_{[x-t]+t}$ for year of birth 1945

| Age $x$ | Lives - IFL92B45 |  | Amounts - IFA92B45 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 47 | 0.000528 | 0.000736 | 0.000441 | 0.000536 |
| 48 | 0.000565 | 0.000787 | 0.000451 | 0.000550 |
| 49 | 0.000607 | 0.000846 | 0.000467 | 0.000567 |
| 50 | 0.000656 | 0.000914 | 0.000487 | 0.000592 |
| 51 | 0.000711 | 0.000992 | 0.000512 | 0.000623 |
| 52 | 0.000775 | 0.001081 | 0.000545 | 0.000662 |
| 53 | 0.000848 | 0.001182 | 0.000584 | 0.000710 |
| 54 | 0.000930 | 0.001298 | 0.000633 | 0.000769 |
| 55 | 0.001024 | 0.001427 | 0.000690 | 0.000839 |
| 56 | 0.001129 | 0.001574 | 0.000759 | 0.000923 |
| 57 | 0.001247 | 0.001738 | 0.000840 | 0.001022 |
| 58 | 0.001379 | 0.001922 | 0.000935 | 0.001138 |
| 59 | 0.001528 | 0.002129 | 0.001047 | 0.001273 |
| 60 | 0.001693 | 0.002359 | 0.001176 | 0.001430 |
| 61 | 0.001911 | 0.002663 | 0.001349 | 0.001640 |
| 62 | 0.002163 | 0.003013 | 0.001555 | 0.001890 |
| 63 | 0.002453 | 0.003418 | 0.001799 | 0.002188 |
| 64 | 0.002787 | 0.003883 | 0.002089 | 0.002539 |
| 65 | 0.003172 | 0.004418 | 0.002431 | 0.002955 |
| 66 | 0.003614 | 0.005033 | 0.002835 | 0.003446 |
| 67 | 0.004122 | 0.005740 | 0.003310 | 0.004023 |
| 68 | 0.004705 | 0.006550 | 0.003868 | 0.004701 |
| 69 | 0.005373 | 0.007479 | 0.004523 | 0.005496 |
| 70 | 0.006139 | 0.008543 | 0.005289 | 0.006427 |
| 71 | 0.007015 | 0.009759 | 0.006184 | 0.007512 |
| 72 | 0.008016 | 0.011148 | 0.007227 | 0.008778 |
| 73 | 0.009158 | 0.012732 | 0.008440 | 0.010249 |
| 74 | 0.010459 | 0.014535 | 0.009846 | 0.011955 |
| 75 | 0.011940 | 0.016586 | 0.011474 | 0.013927 |
| 76 | 0.013622 | 0.018914 | 0.013353 | 0.016203 |
| 77 | 0.015531 | 0.021551 | 0.015516 | 0.018821 |
| 78 | 0.017691 | 0.024535 | 0.017998 | 0.021824 |
| 79 | 0.020134 | 0.027903 | 0.020840 | 0.025259 |
| 80 | 0.022889 | 0.031698 | 0.024084 | 0.029176 |
| 81 | 0.025992 | 0.035966 | 0.027774 | 0.033628 |
| 82 | 0.029480 | 0.040754 | 0.031960 | 0.038673 |
| 83 | 0.033392 | 0.046114 | 0.036693 | 0.044369 |
| 84 | 0.037770 | 0.052102 | 0.042027 | 0.050779 |

Table A17. (Continued).

| Age $x$ | Lives - IFL92B45 |  | Amounts - IFA92B45 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations $1+$ | Duration 0 | Durations 1+ |
| 85 | 0.042660 | 0.058774 | 0.048015 | 0.057966 |
| 86 | 0.048109 | 0.066190 | 0.054717 | 0.065995 |
| 87 | 0.054165 | 0.074413 | 0.062190 | 0.074929 |
| 88 | 0.060884 | 0.083503 | 0.070490 | 0.084833 |
| 89 | 0.068316 | 0.093529 | 0.079675 | 0.095768 |
| 90 | 0.076517 | 0.104550 | 0.089800 | 0.107791 |
| 91 | 0.085543 | 0.116634 | 0.100915 | 0.120955 |
| 92 | 0.095451 | 0.129839 | 0.113069 | 0.135305 |
| 93 | 0.106294 | 0.144225 | 0.126302 | 0.150877 |
| 94 | 0.118130 | 0.159846 | 0.140648 | 0.167702 |
| 95 | 0.131009 | 0.176750 | 0.156135 | 0.185795 |
| 96 | 0.144981 | 0.194978 | 0.172778 | 0.205160 |
| 97 | 0.160093 | 0.214564 | 0.190583 | 0.225786 |
| 98 | 0.176382 | 0.235528 | 0.209545 | 0.247650 |
| 99 | 0.193884 | 0.257883 | 0.229645 | 0.270710 |
| 100 | 0.212625 | 0.281624 | 0.250851 | 0.294912 |
| 101 |  | 0.306737 |  | 0.320186 |
| 102 |  | 0.333187 |  | 0.346444 |
| 103 |  | 0.360925 |  | 0.373587 |
| 104 |  | 0.389886 |  | 0.401502 |
| 105 |  | 0.419986 |  | 0.430068 |
| 106 |  | 0.451125 |  | 0.459151 |
| 107 |  | 0.483186 |  | 0.488616 |
| 108 |  | 0.516037 |  | 0.518319 |
| 109 |  | 0.549534 |  | 0.548118 |
| 110 |  | 0.583518 |  | 0.577873 |
| 111 |  | 0.610832 |  | 0.600570 |
| 112 |  | 0.637731 |  | 0.622502 |
| 113 |  | 0.664072 |  | 0.643586 |
| 114 |  | 0.689719 |  | 0.663752 |
| 115 |  | 0.714545 |  | 0.682946 |
| 116 |  | 0.738437 |  | 0.701125 |
| 117 |  | 0.761294 |  | 0.718260 |
| 118 |  | 0.783030 |  | 0.734334 |
| 119 |  | 0.803579 |  | 0.749341 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A18. Retirement annuitants - RMV92 and RFV92 ( $B=1945$ ) values of $q_{x}$ for year of birth 1945

| Age $x$ | Males <br> RMV92B45 | $\begin{aligned} & \text { Females } \\ & \text { RFV92B45 } \end{aligned}$ |
| :---: | :---: | :---: |
| 47 | 0.023551 | 0.014633 |
| 48 | 0.021978 | 0.013414 |
| 49 | 0.020515 | 0.012292 |
| 50 | 0.019158 | 0.011262 |
| 51 | 0.017903 | 0.010324 |
| 52 | 0.016748 | 0.009474 |
| 53 | 0.015691 | 0.008708 |
| 54 | 0.014731 | 0.008026 |
| 55 | 0.013865 | 0.007425 |
| 56 | 0.013092 | 0.006903 |
| 57 | 0.012412 | 0.006457 |
| 58 | 0.011823 | 0.006088 |
| 59 | 0.011326 | 0.005793 |
| 60 | 0.010920 | 0.005571 |
| 61 | 0.010799 | 0.005520 |
| 62 | 0.010788 | 0.005549 |
| 63 | 0.010893 | 0.005663 |
| 64 | 0.011120 | 0.005866 |
| 65 | 0.011479 | 0.006161 |
| 66 | 0.011980 | 0.006556 |
| 67 | 0.012635 | 0.007056 |
| 68 | 0.013458 | 0.007671 |
| 69 | 0.014467 | 0.008409 |
| 70 | 0.015680 | 0.009280 |
| 71 | 0.017118 | 0.010295 |
| 72 | 0.018803 | 0.011464 |
| 73 | 0.020762 | 0.012802 |
| 74 | 0.023022 | 0.014319 |
| 75 | 0.025584 | 0.016042 |
| 76 | 0.028420 | 0.017997 |
| 77 | 0.031557 | 0.020201 |
| 78 | 0.035021 | 0.022686 |
| 79 | 0.038839 | 0.025488 |
| 80 | 0.043043 | 0.028644 |
| 81 | 0.047661 | 0.032198 |
| 82 | 0.052728 | 0.036197 |
| 83 | 0.058277 | 0.040695 |
| 84 | 0.064344 | 0.045748 |

Table A18. (Continued).

| Age $x$ | Males <br> RMV92B45 | Females <br> RFV92B45 |
| :--- | :---: | :---: |
| 85 | 0.070963 | 0.051422 |
| 86 | 0.078172 | 0.057788 |
| 87 | 0.086008 | 0.064922 |
| 88 | 0.094508 | 0.072908 |
| 89 | 0.103709 | 0.081838 |
| 90 | 0.113649 | 0.091808 |
| 91 | 0.124363 | 0.102925 |
| 92 | 0.135885 | 0.115299 |
| 93 | 0.148248 | 0.129049 |
| 94 | 0.161483 | 0.144296 |
| 95 | 0.175617 | 0.161166 |
| 96 | 0.190673 | 0.179787 |
| 97 | 0.206672 | 0.200286 |
| 98 | 0.223628 | 0.222784 |
| 99 | 0.241552 | 0.247396 |
| 100 | 0.260448 | 0.274224 |
| 101 | 0.280313 | 0.303350 |
| 102 | 0.301139 | 0.334829 |
| 103 | 0.322910 | 0.368686 |
| 104 | 0.345602 | 0.404903 |
| 105 | 0.369182 | 0.443415 |
| 106 | 0.393614 | 0.484095 |
| 107 | 0.418850 | 0.526753 |
| 108 | 0.444833 | 0.571128 |
| 109 | 0.471502 | 0.616878 |
| 110 | 0.498787 | 0.663591 |
| 111 | 0.542516 | 0.702735 |
| 112 | 0.564306 | 0.740987 |
| 113 | 0.585948 | 0.777821 |
| 114 | 0.607369 | 0.812707 |
| 115 | 0.628496 | 0.845149 |
| 116 | 0.689438 | 0.874707 |
| 117 |  | 0.901033 |
| 118 |  | 0.923893 |
| 119 |  | 0.00000000 |
| 120 |  |  |
|  |  |  |

Table A19. Widows - WL92 and WA92 ( $B=1945$ ) values of $q_{x}$ for year of birth 1945

| Age $x$ | $\begin{gathered} \text { Lives } \\ \text { WL92B45 } \end{gathered}$ | Amounts WA92B45 |
| :---: | :---: | :---: |
| 47 | 0.003428 | 0.002859 |
| 48 | 0.003591 | 0.003002 |
| 49 | 0.003765 | 0.003153 |
| 50 | 0.003948 | 0.003315 |
| 51 | 0.004143 | 0.003486 |
| 52 | 0.004349 | 0.003669 |
| 53 | 0.004568 | 0.003863 |
| 54 | 0.004800 | 0.004069 |
| 55 | 0.005046 | 0.004288 |
| 56 | 0.005307 | 0.004522 |
| 57 | 0.005583 | 0.004769 |
| 58 | 0.005876 | 0.005033 |
| 59 | 0.006188 | 0.005313 |
| 60 | 0.006518 | 0.005611 |
| 61 | 0.006992 | 0.006035 |
| 62 | 0.007520 | 0.006508 |
| 63 | 0.008107 | 0.007035 |
| 64 | 0.008759 | 0.007621 |
| 65 | 0.009483 | 0.008273 |
| 66 | 0.010287 | 0.008999 |
| 67 | 0.011178 | 0.009805 |
| 68 | 0.012168 | 0.010702 |
| 69 | 0.013265 | 0.011699 |
| 70 | 0.014480 | 0.012807 |
| 71 | 0.015827 | 0.014037 |
| 72 | 0.017319 | 0.015403 |
| 73 | 0.018971 | 0.016919 |
| 74 | 0.020800 | 0.018602 |
| 75 | 0.022823 | 0.020468 |
| 76 | 0.025059 | 0.022537 |
| 77 | 0.027531 | 0.024830 |
| 78 | 0.030262 | 0.027371 |
| 79 | 0.033278 | 0.030184 |
| 80 | 0.036606 | 0.033297 |
| 81 | 0.040276 | 0.036740 |
| 82 | 0.044322 | 0.040547 |
| 83 | 0.048779 | 0.044753 |
| 84 | 0.053685 | 0.049396 |

Table A19. (Continued).

| Age $x$ | Lives <br> WL92B45 | Amounts <br> WA92B45 |
| :--- | :---: | :---: |
| 85 | 0.059082 | 0.054519 |
| 86 | 0.065014 | 0.060167 |
| 87 | 0.071529 | 0.066388 |
| 88 | 0.078678 | 0.073235 |
| 89 | 0.086515 | 0.080765 |
| 90 | 0.095097 | 0.089035 |
| 91 | 0.104486 | 0.098111 |
| 92 | 0.114744 | 0.108058 |
| 93 | 0.125939 | 0.118947 |
| 94 | 0.138140 | 0.130851 |
| 95 | 0.151416 | 0.143845 |
| 96 | 0.165841 | 0.158007 |
| 97 | 0.181489 | 0.173417 |
| 98 | 0.198431 | 0.190152 |
| 99 | 0.216739 | 0.208291 |
| 100 | 0.236479 | 0.227911 |
| 101 | 0.257718 | 0.249083 |
| 102 | 0.280511 | 0.271869 |
| 103 | 0.304908 | 0.296329 |
| 104 | 0.330943 | 0.322507 |
| 105 | 0.358644 | 0.350433 |
| 106 | 0.388016 | 0.380122 |
| 107 | 0.419046 | 0.411565 |
| 108 | 0.451698 | 0.444728 |
| 109 | 0.485911 | 0.479551 |
| 110 | 0.521593 | 0.515942 |
| 111 | 0.552300 | 0.547505 |
| 112 | 0.583529 | 0.579652 |
| 113 | 0.615094 | 0.612181 |
| 114 | 0.646780 | 0.644862 |
| 115 | 0.678355 | 0.677439 |
| 116 | 0.709565 | 0.709637 |
| 117 | 0.740144 | 0.741163 |
| 118 | 0.798321 | 0.771718 |
| 119 | 000000 | 0.801003 |
| 120 |  | 1.000000 |
|  |  |  |
|  |  |  |

Table A20. Pensioners - PML92, PMA92, PFL92 and PFA92 ( $B=1972$ ) values of $q_{x}$ for year of birth 1972

| Age $x$ | Mates |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92B72 } \end{gathered}$ | Amounts PMA92B72 | $\begin{gathered} \text { Lives } \\ \text { PFL92B72 } \end{gathered}$ | Amounts PFA92B72 |
| 20 | 0.000623 | 0.000340 | 0.000254 | 0.000203 |
| 21 | 0.000591 | 0.000326 | 0.000244 | 0.000195 |
| 22 | 0.000561 | 0.000314 | 0.000236 | 0.000189 |
| 23 | 0.000533 | 0.000302 | 0.000227 | 0.000182 |
| 24 | 0.000506 | 0.000291 | 0.000220 | 0.000176 |
| 25 | 0.000483 | 0.000281 | 0.000213 | 0.000171 |
| 26 | 0.000460 | 0.000270 | 0.000208 | 0.000166 |
| 27 | 0.000441 | 0.000261 | 0.000202 | 0.000162 |
| 28 | 0.000423 | 0.000252 | 0.000198 | 0.000159 |
| 29 | 0.000408 | 0.000245 | 0.000195 | 0.000156 |
| 30 | 0.000395 | 0.000238 | 0.000193 | 0.000155 |
| 31 | 0.000384 | 0.000231 | 0.000192 | 0.000153 |
| 32 | 0.000376 | 0.000226 | 0.000192 | 0.000154 |
| 33 | 0.000371 | 0.000222 | 0.000194 | 0.000154 |
| 34 | 0.000368 | 0.000219 | 0.000196 | 0.000156 |
| 35 | 0.000368 | 0.000217 | 0.000201 | 0.000160 |
| 36 | 0.000372 | 0.000216 | 0.000207 | 0.000164 |
| 37 | 0.000379 | 0.000217 | 0.000215 | 0.000170 |
| 38 | 0.000390 | 0.000219 | 0.000224 | 0.000178 |
| 39 | 0.000406 | 0.000224 | 0.000237 | 0.000188 |
| 40 | 0.000426 | 0.000231 | 0.000251 | 0.000199 |
| 41 | 0.000452 | 0.000240 | 0.000269 | 0.000213 |
| 42 | 0.000483 | 0.000253 | 0.000289 | 0.000230 |
| 43 | 0.000520 | 0.000268 | 0.000314 | 0.000249 |
| 44 | 0.000564 | 0.000287 | 0.000341 | 0.000272 |
| 45 | 0.000615 | 0.000310 | 0.000374 | 0.000298 |
| 46 | 0.000675 | 0.000339 | 0.000410 | 0.000329 |
| 47 | 0.000744 | 0.000373 | 0.000452 | 0.000363 |
| 48 | 0.000823 | 0.000413 | 0.000499 | 0.000403 |
| 49 | 0.000912 | 0.000461 | 0.000553 | 0.000448 |
| 50 | 0.001014 | 0.000516 | 0.000614 | 0.000499 |
| 51 | 0.001129 | 0.000581 | 0.000682 | 0.000557 |
| 52 | 0.001258 | 0.000656 | 0.000758 | 0.000622 |
| 53 | 0.001403 | 0.000742 | 0.000844 | 0.000696 |
| 54 | 0.001565 | 0.000842 | 0.000940 | 0.000778 |

Table A20. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92B72 } \end{gathered}$ | $\begin{aligned} & \text { Amounts } \\ & \text { PMA92B72 } \end{aligned}$ | $\begin{aligned} & \text { Lives } \\ & \text { PFL92B72 } \end{aligned}$ | Amounts PFA92B72 |
| 55 | 0.001746 | 0.000956 | 0.001047 | 0.000871 |
| 56 | 0.001946 | 0.001087 | 0.001167 | 0.000974 |
| 57 | 0.002169 | 0.001235 | 0.001300 | 0.001090 |
| 58 | 0.002415 | 0.001404 | 0.001447 | 0.001219 |
| 59 | 0.002687 | 0.001595 | 0.001610 | 0.001362 |
| 60 | 0.002986 | 0.001811 | 0.001791 | 0.001520 |
| 61 | 0.003514 | 0.002177 | 0.002110 | 0.001798 |
| 62 | 0.004127 | 0.002612 | 0.002482 | 0.002122 |
| 63 | 0.004838 | 0.003127 | 0.002915 | 0.002501 |
| 64 | 0.005658 | 0.003735 | 0.003418 | 0.002942 |
| 65 | 0.006602 | 0.004449 | 0.004000 | 0.003453 |
| 66 | 0.007686 | 0.005287 | 0.004671 | 0.004044 |
| 67 | 0.008925 | 0.006264 | 0.005445 | 0.004726 |
| 68 | 0.010338 | 0.007400 | 0.006334 | 0.005511 |
| 69 | 0.011945 | 0.008717 | 0.007352 | 0.006410 |
| 70 | 0.013765 | 0.010236 | 0.008516 | 0.007439 |
| 71 | 0.015822 | 0.011983 | 0.009842 | 0.008613 |
| 72 | 0.018138 | 0.013984 | 0.011350 | 0.009947 |
| 73 | 0.020738 | 0.016267 | 0.013059 | 0.011460 |
| 74 | 0.023646 | 0.018862 | 0.014991 | 0.013171 |
| 75 | 0.026891 | 0.021799 | 0.017171 | 0.015099 |
| 76 | 0.030497 | 0.025111 | 0.019622 | 0.017266 |
| 77 | 0.034494 | 0.028831 | 0.022371 | 0.019695 |
| 78 | 0.038909 | 0.032991 | 0.025447 | 0.022409 |
| 79 | 0.043770 | 0.037627 | 0.028878 | 0.025433 |
| 80 | 0.049103 | 0.042770 | 0.032696 | 0.028791 |
| 81 | 0.054937 | 0.048452 | 0.036932 | 0.032511 |
| 82 | 0.061297 | 0.054706 | 0.041619 | 0.036618 |
| 83 | 0.068206 | 0.061559 | 0.046789 | 0.041138 |
| 84 | 0.075688 | 0.069036 | 0.052479 | 0.046099 |
| 85 | 0.083763 | 0.077160 | 0.058720 | 0.051527 |
| 86 | 0.092448 | 0.085951 | 0.065548 | 0.057447 |
| 87 | 0.101758 | 0.095422 | 0.072996 | 0.063884 |
| 88 | 0.111703 | 0.105582 | 0.081095 | 0.070862 |
| 89 | 0.122292 | 0.116433 | 0.089878 | 0.078401 |

Table A20. (Continued).

| Age $x$ | Males |  | Females |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \text { Lives } \\ \text { PML92B72 } \end{gathered}$ | $\begin{gathered} \text { Amounts } \\ \text { PMA92B72 } \end{gathered}$ | $\begin{gathered} \text { Lives } \\ \text { PFL92B72 } \end{gathered}$ | Amounts PFA92B72 |
| 90 | 0.133526 | 0.127973 | 0.099374 | 0.086522 |
| 91 | 0.145404 | 0.140191 | 0.109609 | 0.095241 |
| 92 | 0.157920 | 0.153072 | 0.120607 | 0.104572 |
| 93 | 0.171063 | 0.166593 | 0.132388 | 0.114528 |
| 94 | 0.184817 | 0.180722 | 0.144971 | 0.125114 |
| 95 | 0.199161 | 0.195425 | 0.158367 | 0.136334 |
| 96 | 0.214070 | 0.210656 | 0.172584 | 0.148188 |
| 97 | 0.229511 | 0.226367 | 0.187625 | 0.160671 |
| 98 | 0.245449 | 0.242501 | 0.203487 | 0.173773 |
| 99 | 0.261845 | 0.258999 | 0.220161 | 0.187479 |
| 100 | 0.278655 | 0.275796 | 0.237633 | 0.201771 |
| 101 | 0.295830 | 0.292821 | 0.255882 | 0.216624 |
| 102 | 0.313320 | 0.310003 | 0.274881 | 0.232010 |
| 103 | 0.331069 | 0.327269 | 0.294596 | 0.247895 |
| 104 | 0.349023 | 0.344542 | 0.314991 | 0.264241 |
| 105 | 0.367123 | 0.361743 | 0.336018 | 0.281006 |
| 106 | 0.385308 | 0.378797 | 0.357628 | 0.298146 |
| 107 | 0.403518 | 0.395626 | 0.379766 | 0.315609 |
| 108 | 0.421692 | 0.412155 | 0.402372 | 0.333343 |
| 109 | 0.439767 | 0.428309 | 0.425381 | 0.351294 |
| 110 | 0.457684 | 0.444014 | 0.448729 | 0.369403 |
| 111 | 0.468997 | 0.453033 | 0.466001 | 0.382406 |
| 112 | 0.479798 | 0.461297 | 0.483064 | 0.395150 |
| 113 | 0.490060 | 0.468780 | 0.499869 | 0.407594 |
| 114 | 0.499756 | 0.475459 | 0.516366 | 0.419700 |
| 115 | 0.508865 | 0.481313 | 0.532508 | 0.431431 |
| 116 | 0.517368 | 0.486326 | 0.548254 | 0.442752 |
| 117 | 0.525248 | 0.490484 | 0.563564 | 0.453630 |
| 118 | 0.532489 | 0.493776 | 0.578402 | 0.464038 |
| 119 | 0.539080 | 0.496194 | 0.592739 | 0.473947 |
| 120 | 1.000000 | 1.000000 | 1.000000 | 1.000000 |

Table A21. Immediate annuitants, males - IML92 and IMA92 ( $B=1972$ ) one year select: values of $q_{x-t]+t}$ for year of birth 1972

| Age $x$ | Lives - IML92B72 |  | Amounts - IMA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 20 | 0.000423 | 0.000573 | 0.000447 | 0.000546 |
| 21 | 0.000406 | 0.000550 | 0.000429 | 0.000525 |
| 22 | 0.000390 | 0.000529 | 0.000412 | 0.000504 |
| 23 | 0.000376 | 0.000510 | 0.000398 | 0.000486 |
| 24 | 0.000363 | 0.000492 | 0.000383 | 0.000469 |
| 25 | 0.000351 | 0.000476 | 0.000371 | 0.000453 |
| 26 | 0.000341 | 0.000462 | 0.000360 | 0.000440 |
| 27 | 0.000332 | 0.000449 | 0.000351 | 0.000429 |
| 28 | 0.000324 | 0.000439 | 0.000342 | 0.000418 |
| 29 | 0.000317 | 0.000430 | 0.000336 | 0.000410 |
| 30 | 0.000313 | 0.000423 | 0.000330 | 0.000404 |
| 31 | 0.000309 | 0.000419 | 0.000327 | 0.000399 |
| 32 | 0.000308 | 0.000417 | 0.000325 | 0.000397 |
| 33 | 0.000307 | 0.000416 | 0.000325 | 0.000397 |
| 34 | 0.000309 | 0.000419 | 0.000326 | 0.000399 |
| 35 | 0.000313 | 0.000423 | 0.000330 | 0.000404 |
| 36 | 0.000318 | 0.000431 | 0.000336 | 0.000411 |
| 37 | 0.000326 | 0.000441 | 0.000344 | 0.000421 |
| 38 | 0.000336 | 0.000455 | 0.000355 | 0.000434 |
| 39 | 0.000348 | 0.000472 | 0.000368 | 0.000450 |
| 40 | 0.000363 | 0.000493 | 0.000384 | 0.000470 |
| 41 | 0.000382 | 0.000517 | 0.000403 | 0.000493 |
| 42 | 0.000403 | 0.000546 | 0.000426 | 0.000521 |
| 43 | 0.000428 | 0.000580 | 0.000452 | 0.000553 |
| 44 | 0.000456 | 0.000618 | 0.000482 | 0.000589 |
| 45 | 0.000489 | 0.000662 | 0.000517 | 0.000632 |
| 46 | 0.000526 | 0.000713 | 0.000556 | 0.000679 |
| 47 | 0.000568 | 0.000769 | 0.000600 | 0.000734 |
| 48 | 0.000615 | 0.000833 | 0.000650 | 0.000795 |
| 49 | 0.000668 | 0.000905 | 0.000706 | 0.000863 |
| 50 | 0.000727 | 0.000985 | 0.000768 | 0.000939 |
| 51 | 0.000793 | 0.001074 | 0.000838 | 0.001024 |
| 52 | 0.000866 | 0.001173 | 0.000915 | 0.001119 |
| 53 | 0.000948 | 0.001283 | 0.001001 | 0.001224 |
| 54 | 0.001038 | 0.001406 | 0.001097 | 0.001340 |

Table A21. (Continued).

| Age $x$ | Lives - IML92B72 |  | Amounts - IMA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 55 | 0.001138 | 0.001541 | 0.001202 | 0.001469 |
| 56 | 0.001248 | 0.001690 | 0.001319 | 0.001612 |
| 57 | 0.001370 | 0.001855 | 0.001447 | 0.001769 |
| 58 | 0.001504 | 0.002037 | 0.001589 | 0.001942 |
| 59 | 0.001652 | 0.002236 | 0.001745 | 0.002132 |
| 60 | 0.001814 | 0.002455 | 0.001916 | 0.002341 |
| 61 | 0.002112 | 0.002858 | 0.002231 | 0.002725 |
| 62 | 0.002456 | 0.003323 | 0.002594 | 0.003170 |
| 63 | 0.002854 | 0.003861 | 0.003015 | 0.003683 |
| 64 | 0.003312 | 0.004481 | 0.003499 | 0.004274 |
| 65 | 0.003840 | 0.005194 | 0.004056 | 0.004953 |
| 66 | 0.004446 | 0.006011 | 0.004696 | 0.005734 |
| 67 | 0.005140 | 0.006948 | 0.005429 | 0.006627 |
| 68 | 0.005934 | 0.008019 | 0.006267 | 0.007650 |
| 69 | 0.006840 | 0.009241 | 0.007224 | 0.008816 |
| 70 | 0.007872 | 0.010632 | 0.008314 | 0.010143 |
| 71 | 0.009046 | 0.012213 | 0.009553 | 0.011652 |
| 72 | 0.010378 | 0.014005 | 0.010958 | 0.013363 |
| 73 | 0.011886 | 0.016033 | 0.012550 | 0.015299 |
| 74 | 0.013592 | 0.018324 | 0.014349 | 0.017486 |
| 75 | 0.015515 | 0.020905 | 0.016378 | 0.019951 |
| 76 | 0.017680 | 0.023808 | 0.018662 | 0.022724 |
| 77 | 0.020113 | 0.027065 | 0.021228 | 0.025837 |
| 78 | 0.022841 | 0.030714 | 0.024104 | 0.029323 |
| 79 | 0.025893 | 0.034790 | 0.027322 | 0.033220 |
| 80 | 0.029302 | 0.039336 | 0.030915 | 0.037567 |
| 81 | 0.033102 | 0.044394 | 0.034918 | 0.042404 |
| 82 | 0.037327 | 0.050008 | 0.039369 | 0.047775 |
| 83 | 0.042017 | 0.056226 | 0.044306 | 0.053727 |
| 84 | 0.047210 | 0.063097 | 0.049773 | 0.060306 |
| 85 | 0.052949 | 0.070671 | 0.055811 | 0.067561 |
| 86 | 0.059277 | 0.079000 | 0.062466 | 0.075544 |
| 87 | 0.066238 | 0.088138 | 0.069784 | 0.084305 |
| 88 | 0.073880 | 0.098137 | 0.077813 | 0.093898 |
| 89 | 0.082247 | 0.109051 | 0.086601 | 0.104374 |

Table A21. (Continued).

| Age $x$ | Lives - IML92B72 |  | Amounts - IMA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations $1+$ | Duration 0 | Durations 1+ |
| 90 | 0.091389 | 0.120930 | 0.096197 | 0.115785 |
| 91 | 0.101354 | 0.133828 | 0.106648 | 0.128182 |
| 92 | 0.112188 | 0.147791 | 0.118004 | 0.141614 |
| 93 | 0.123937 | 0.162865 | 0.130310 | 0.156126 |
| 94 | 0.136647 | 0.179092 | 0.143612 | 0.171761 |
| 95 | 0.150360 | 0.196506 | 0.157952 | 0.188557 |
| 96 | 0.165115 | 0.215139 | 0.173368 | 0.206545 |
| 97 | 0.180948 | 0.235012 | 0.189894 | 0.225752 |
| 98 | 0.197891 | 0.256140 | 0.207560 | 0.246195 |
| 99 | 0.215970 | 0.278529 | 0.226389 | 0.267884 |
| 100 | 0.235204 | 0.302171 | 0.246398 | 0.290819 |
| 101 |  | 0.327052 |  | 0.314988 |
| 102 |  | 0.353142 |  | 0.340370 |
| 103 |  | 0.380401 |  | 0.366930 |
| 104 |  | 0.408775 |  | 0.394621 |
| 105 |  | 0.438196 |  | 0.423386 |
| 106 |  | 0.468584 |  | 0.453152 |
| 107 |  | 0.499847 |  | 0.483833 |
| 108 |  | 0.531877 |  | 0.515333 |
| 109 |  | 0.564562 |  | 0.547545 |
| 110 |  | 0.597771 |  | 0.580348 |
| 111 |  | 0.622892 |  | 0.605374 |
| 112 |  | 0.647670 |  | 0.630135 |
| 113 |  | 0.671993 |  | 0.654519 |
| 114 |  | 0.695751 |  | 0.678419 |
| 115 |  | 0.718842 |  | 0.701731 |
| 116 |  | 0.741170 |  | 0.724359 |
| 117 |  | 0.762651 |  | 0.746213 |
| 118 |  | 0.783207 |  | 0.767212 |
| 119 |  | 0.802773 |  | 0.787288 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A22. Immediate annuitants, females - IFL92 and IFA92 ( $B=1972$ ) one year select: values of $q_{[x-t+t}$ for year of birth 1972

| Age $x$ | Lives - IFL92B72 |  | Amounts - IFA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1- |
| 20 | 0.000217 | 0.000303 | 0.000329 | 0.000400 |
| 21 | 0.000211 | 0.000294 | 0.000318 | 0.000386 |
| 22 | 0.000203 | 0.000284 | 0.000307 | 0.000373 |
| 23 | 0.000197 | 0.000275 | 0.000297 | 0.000361 |
| 24 | 0.000192 | 0.000268 | 0.000287 | 0.000349 |
| 25 | 0.000186 | 0.000260 | 0.000277 | 0.000338 |
| 26 | 0.000181 | 0.000253 | 0.000269 | 0.000327 |
| 27 | 0.000176 | 0.000246 | 0.000260 | 0.000316 |
| 28 | 0.000171 | 0.000239 | 0.000252 | 0.000306 |
| 29 | 0.000168 | 0.000234 | 0.000243 | 0.000296 |
| 30 | 0.000164 | 0.000229 | 0.000236 | 0.000287 |
| 31 | 0.000161 | 0.000225 | 0.060229 | 0.000278 |
| 32 | 0.000159 | 0.000221 | 0.000222 | 0.000270 |
| 33 | 0.000157 | 0.000218 | 0.000216 | 0.000262 |
| 34 | 0.000155 | 0.000216 | 0.000210 | 0.000255 |
| 35 | 0.000154 | 0.000215 | 0.000204 | 0.000249 |
| 36 | 0.000154 | 0.000215 | 0.000199 | 0.000242 |
| 37 | 0.000155 | 0.000215 | 0.000194 | 0.000237 |
| 38 | 0.000156 | 0.000217 | 0.000191 | 0.000232 |
| 39 | 0.000158 | 0.000220 | 0.000186 | 0.000227 |
| 40 | 0.000162 | 0.000225 | 0.000184 | 0.000224 |
| 41 | 0.000166 | 0.000231 | 0.000181 | 0.000221 |
| 42 | 0.000172 | 0.000239 | 0.000180 | 0.000219 |
| 43 | 0.000179 | 0.000249 | 0.000179 | 0.000218 |
| 44 | 0.000187 | 0.000262 | 0.000179 | 0.000218 |
| 45 | 0.000198 | 0.000276 | 0.000181 | 0.000220 |
| 46 | 0.000210 | 0.000293 | 0.000184 | 0.000223 |
| 47 | 0.000225 | 0.000314 | 0.000188 | 0.000228 |
| 48 | 0.000242 | 0.000338 | 0.000194 | 0.000236 |
| 49 | 0.000262 | 0.000366 | 0.000202 | 0.000245 |
| 50 | 0.000285 | 0.000398 | 0.000212 | 0.000258 |
| 51 | 0.000312 | 0.000435 | 0.000225 | 0.000273 |
| 52 | 0.000343 | 0.000478 | 0.000241 | 0.000293 |
| 53 | 0.000378 | 0.000527 | 0.000260 | 0.000317 |
| 54 | 0.000418 | 0.000583 | 0.000284 | 0.000345 |

Table A22. (Continued).

| Age $x$ | Lives - IFL92B72 |  | Amounts - IFA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1 |
| 55 | 0.000463 | 0.000646 | 0.000313 | 0.000380 |
| 56 | 0.000515 | 0.000718 | 0.000346 | 0.000422 |
| 57 | 0.000574 | 0.000800 | 0.000387 | 0.000471 |
| 58 | 0.000641 | 0.000893 | 0.000434 | 0.000528 |
| 59 | 0.000716 | 0.000997 | 0.000490 | 0.000596 |
| 60 | 0.000800 | 0.001115 | 0.000556 | 0.000676 |
| 61 | 0.000949 | 0.001322 | 0.000670 | 0.000814 |
| 62 | 0.001125 | 0.001568 | 0.000809 | 0.000984 |
| 63 | 0.001334 | 0.001858 | 0.000978 | 0.001190 |
| 64 | 0.001580 | 0.002200 | 0.001184 | 0.001439 |
| 65 | 0.001869 | 0.002603 | 0.001432 | 0.001741 |
| 66 | 0.002208 | 0.003076 | 0.001732 | 0.002106 |
| 67 | 0.002607 | 0.003629 | 0.002093 | 0.002544 |
| 68 | 0.003072 | 0.004277 | 0.002525 | 0.003069 |
| 69 | 0.003615 | 0.005031 | 0.003042 | 0.003697 |
| 70 | 0.004247 | 0.005910 | 0.003659 | 0.004446 |
| 71 | 0.004981 | 0.006929 | 0.004391 | 0.005334 |
| 72 | 0.005832 | 0.008110 | 0.005258 | 0.006386 |
| 73 | 0.006816 | 0.009475 | 0.006281 | 0.007627 |
| 74 | 0.007950 | 0.011048 | 0.007484 | 0.009087 |
| 75 | 0.009256 | 0.012857 | 0.008895 | 0.010796 |
| 76 | 0.010755 | 0.014933 | 0.010542 | 0.012792 |
| 77 | 0.012473 | 0.017308 | 0.012461 | 0.015115 |
| 78 | 0.014436 | 0.020020 | 0.014686 | 0.017808 |
| 79 | 0.016674 | 0.023108 | 0.017259 | 0.020918 |
| 80 | 0.019219 | 0.026615 | 0.020222 | 0.024497 |
| 81 | 0.022107 | 0.030589 | 0.023622 | 0.028601 |
| 82 | 0.025375 | 0.035079 | 0.027510 | 0.033288 |
| 83 | 0.029066 | 0.040140 | 0.031939 | 0.038620 |
| 84 | 0.033221 | 0.045827 | 0.036965 | 0.044664 |
| 85 | 0.037890 | 0.052202 | 0.042646 | 0.051484 |
| 86 | 0.043120 | 0.059326 | 0.049043 | 0.059151 |
| 87 | 0.048963 | 0.067265 | 0.056216 | 0.067732 |
| 88 | 0.055475 | 0.076085 | 0.064228 | 0.077297 |
| 89 | 0.062712 | 0.085857 | 0.073139 | 0.087912 |

Table A22. (Continued).

| Age $x$ | Lives - IFL92B72 |  | Amounts IFA92B72 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Duration 0 | Durations 1+ | Duration 0 | Durations 1+ |
| 90 | 0.070732 | 0.096645 | 0.083010 | 0.099641 |
| 91 | 0.079593 | 0.108522 | 0.093896 | 0.112542 |
| 92 | 0.089358 | 0.121551 | 0.105851 | 0.126667 |
| 93 | 0.100083 | 0.135797 | 0.118921 | 0.142061 |
| 94 | 0.111829 | 0.151320 | 0.133146 | 0.158757 |
| 95 | 0.124652 | 0.168174 | 0.148560 | 0.176780 |
| 96 | 0.138607 | 0.186406 | 0.165182 | 0.196140 |
| 97 | 0.153744 | 0.206055 | 0.183026 | 0.216832 |
| 98 | 0.170108 | 0.227150 | 0.202090 | 0.238840 |
| 99 | 0.187736 | 0.249706 | 0.222363 | 0.262127 |
| 100 | 0.206662 | 0.273726 | 0.243817 | 0.286642 |
| 101 |  | 0.299201 |  | 0.312320 |
| 102 |  | 0.326101 |  | 0.339076 |
| 103 |  | 0.354381 |  | 0.366812 |
| 104 |  | 0.383978 |  | 0.395419 |
| 105 |  | 0.414813 |  | 0.424771 |
| 106 |  | 0.446787 |  | 0.454737 |
| 107 |  | 0.479784 |  | 0.485176 |
| 108 |  | 0.513672 |  | 0.515943 |
| 109 |  | 0.548303 |  | 0.546891 |
| 110 |  | 0.583518 |  | 0.577873 |
| 111 |  | 0.610832 |  | 0.600570 |
| 112 |  | 0.637731 |  | 0.622502 |
| 113 |  | 0.664072 |  | 0.643586 |
| 114 |  | 0.689719 |  | 0.663752 |
| 115 |  | 0.714545 |  | 0.682946 |
| 116 |  | 0.738437 |  | 0.701125 |
| 117 |  | 0.761294 |  | 0.718260 |
| 118 |  | 0.783030 |  | 0.734334 |
| 119 |  | 0.803579 |  | 0.749341 |
| 120 |  | 1.000000 |  | 1.000000 |

Table A23. Retirement annuitants - RMV92 and RFV92 ( $B=1972$ ) values of $q_{x}$ for year of birth 1972
\(\left.$$
\begin{array}{lcc}\hline \text { Age } x & \begin{array}{c}\text { Males } \\
\text { RMV92B72 }\end{array} & \begin{array}{c}\text { Females } \\
\text { RFV }\end{array}
$$ <br>

\hline 20 \& 0.0485372\end{array}\right]\)| 0.039297 |
| :--- |
| 21 |

Table A23. (Continued).

| Age $x$ | Males RMV92B72 | Females RFV92B72 |
| :---: | :---: | :---: |
| 55 | 0.006278 | 0.003362 |
| 56 | 0.005978 | 0.003152 |
| 57 | 0.005715 | 0.002973 |
| 58 | 0.005491 | 0.002828 |
| 59 | 0.005306 | 0.002714 |
| 60 | 0.005162 | 0.002633 |
| 61 | 0.005362 | 0.002741 |
| 62 | 0.005613 | 0.002887 |
| 63 | 0.005922 | 0.003079 |
| 64 | 0.006302 | 0.003324 |
| 65 | 0.006763 | 0.003630 |
| 66 | 0.007321 | 0.004006 |
| 67 | 0.007989 | 0.004462 |
| 68 | 0.008787 | 0.005009 |
| 69 | 0.009732 | 0.005657 |
| 70 | 0.010847 | 0.006420 |
| 71 | 0.012154 | 0.007310 |
| 72 | 0.013680 | 0.008341 |
| 73 | 0.015451 | 0.009527 |
| 74 | 0.017499 | 0.010884 |
| 75 | 0.019832 | 0.012436 |
| 76 | 0.022438 | 0.014209 |
| 77 | 0.025344 | 0.016224 |
| 78 | 0.028576 | 0.018511 |
| 79 | 0.032164 | 0.021108 |
| 80 | 0.036140 | 0.024051 |
| 81 | 0.040536 | 0.027385 |
| 82 | 0.045387 | 0.031157 |
| 83 | 0.050727 | 0.035422 |
| 84 | 0.056595 | 0.040238 |
| 85 | 0.063028 | 0.045672 |
| 86 | 0.070065 | 0.051795 |
| 87 | 0.077747 | 0.058686 |
| 88 | 0.086112 | 0.066431 |
| 89 | 0.095202 | 0.075125 |
| 90 | 0.105056 | 0.084866 |
| 91 | 0.115713 | 0.095766 |
| 92 | 0.127211 | 0.107939 |
| 93 | 0.139585 | 0.121507 |
| 94 | 0.152870 | 0.136599 |

Table A23. (Continued).

| Age $x$ | Males <br> RMV92B72 | Females <br> RFV92B72 |
| :--- | :---: | :---: |
| 95 | 0.167096 | 0.153346 |
| 96 | 0.182290 | 0.171882 |
| 97 | 0.198476 | 0.192344 |
| 98 | 0.215673 | 0.214859 |
| 99 | 0.233893 | 0.239552 |
| 100 | 0.253144 | 0.266534 |
| 101 | 0.273427 | 0.295897 |
| 102 | 0.294734 | 0.327708 |
| 103 | 0.317055 | 0.362000 |
| 104 | 0.340365 | 0.398768 |
| 105 | 0.364636 | 0.437954 |
| 106 | 0.389830 | 0.479440 |
| 107 | 0.415901 | 0.523045 |
| 108 | 0.442793 | 0.568510 |
| 109 | 0.470446 | 0.615496 |
| 110 | 0.498787 | 0.663591 |
| 111 | 0.520651 | 0.702735 |
| 112 | 0.542516 | 0.740987 |
| 113 | 0.564306 | 0.777821 |
| 114 | 0.585948 | 0.812707 |
| 115 | 0.607369 | 0.845149 |
| 116 | 0.628496 | 0.874707 |
| 117 | 0.649259 | 0.901033 |
| 118 | 0.669594 | 0.923893 |
| 119 | 0.689438 | 0.943192 |
| 120 | 1.000000 | 1.000000 |

Table A24. Widows - WL92 and WA92 ( $B=1972$ ) values of $q_{x}$ for year of birth 1972

| Age $x$ | Lives <br> WL92B72 | Amounts <br> WA92B72 |
| :--- | :---: | :---: |
| 20 | 0.000515 | 0.000423 |
| 21 | 0.000525 | 0.000431 |
| 22 | 0.000537 | 0.000440 |
| 23 | 0.000549 | 0.000449 |
| 24 | 0.000562 | 0.000459 |
| 25 | 0.000577 | 0.000471 |
| 26 | 0.000594 | 0.000485 |
| 27 | 0.000612 | 0.000499 |
| 28 | 0.000632 | 0.000514 |
| 29 | 0.000653 | 0.000532 |
| 30 | 0.000676 | 0.000551 |
| 31 | 0.000700 | 0.000571 |
| 32 | 0.000727 | 0.000593 |
| 33 | 0.000756 | 0.000617 |
| 34 | 0.000787 | 0.000643 |
| 35 | 0.000820 | 0.000671 |
| 36 | 0.000856 | 0.000701 |
| 37 | 0.000894 | 0.000732 |
| 38 | 0.000935 | 0.000767 |
| 39 | 0.000979 | 0.000805 |
| 40 | 0.001026 | 0.000844 |
| 41 | 0.001076 | 0.000887 |
| 42 | 0.001130 | 0.000933 |
| 43 | 0.001188 | 0.000982 |
| 44 | 0.001249 | 0.001035 |
| 45 | 0.001315 | 0.001092 |
| 46 | 0.001385 | 0.001153 |
| 47 | 0.001460 | 0.001218 |
| 48 | 0.0017627 | 0.001288 |
| 49 | 0.001817 | 0.001363 |
| 50 |  | 0.001443 |
| 51 |  | 0.001529 |
| 52 |  | 0.001721 |
| 53 |  |  |
| 54 |  |  |
|  |  |  |

Table A24. (Continued).

| Age $x$ | Lives WL92B72 | Amounts WA92B72 |
| :---: | :---: | :---: |
| 55 | 0.002285 | 0.001942 |
| 56 | 0.002423 | 0.002065 |
| 57 | 0.002571 | 0.002196 |
| 58 | 0.002729 | 0.002337 |
| 59 | 0.002899 | 0.002489 |
| 60 | 0.003081 | 0.002652 |
| 61 | 0.003472 | 0.002997 |
| 62 | 0.003913 | 0.003386 |
| 63 | 0.004408 | 0.003825 |
| 64 | 0.004964 | 0.004319 |
| 65 | 0.005587 | 0.004874 |
| 66 | 0.006286 | 0.005499 |
| 67 | 0.007069 | 0.006200 |
| 68 | 0.007944 | 0.006988 |
| 69 | 0.008923 | 0.007870 |
| 70 | 0.010017 | 0.008860 |
| 71 | 0.011238 | 0.009967 |
| 72 | 0.012600 | 0.011206 |
| 73 | 0.014119 | 0.012591 |
| 74 | 0.015809 | 0.014139 |
| 75 | 0.017692 | 0.015866 |
| 76 | 0.019785 | 0.017793 |
| 77 | 0.022110 | 0.019941 |
| 78 | 0.024693 | 0.022334 |
| 79 | 0.027558 | 0.024997 |
| 80 | 0.030736 | 0.027958 |
| 81 | 0.034255 | 0.031248 |
| 82 | 0.038151 | 0.034901 |
| 83 | 0.042459 | 0.038955 |
| 84 | 0.047220 | 0.043447 |
| 85 | 0.052475 | 0.048422 |
| 86 | 0.058272 | 0.053927 |
| 87 | 0.064659 | 0.060011 |
| 88 | 0.071689 | 0.066729 |
| 89 | 0.079418 | 0.074140 |
| 90 | 0.087907 | 0.082303 |
| 91 | 0.097219 | 0.091287 |
| 92 | 0.107420 | 0.101160 |
| 93 | 0.118579 | 0.111996 |
| 94 | 0.130772 | 0.123872 |

Table A24. (Continued).

| Age $x$ | Lives <br> WL92B72 | Amounts <br> WA92B72 |
| :--- | :--- | :--- |
| 95 | 0.144069 | 0.136865 |
| 96 | 0.158550 | 0.151061 |
| 97 | 0.174292 | 0.166540 |
| 98 | 0.191372 | 0.183387 |
| 99 | 0.209866 | 0.201687 |
| 100 | 0.229848 | 0.221520 |
| 101 | 0.251387 | 0.242963 |
| 102 | 0.274545 | 0.266087 |
| 103 | 0.299379 | 0.290956 |
| 104 | 0.325929 | 0.317621 |
| 105 | 0.354227 | 0.346118 |
| 106 | 0.384286 | 0.376468 |
| 107 | 0.416095 | 0.408667 |
| 108 | 0.449628 | 0.442689 |
| 109 | 0.484823 | 0.478477 |
| 110 | 0.521593 | 0.515942 |
| 111 | 0.552300 | 0.547505 |
| 112 | 0.583529 | 0.579652 |
| 113 | 0.615094 | 0.612181 |
| 114 | 0.646780 | 0.644862 |
| 115 | 0.678355 | 0.677439 |
| 116 | 0.709565 | 0.709637 |
| 117 | 0.740144 | 0.741163 |
| 118 | 0.769821 | 0.771718 |
| 119 | 0.798321 | 0.801003 |
| 120 | 1.000000 | 1.000000 |

Table B1. Retirement annuitants, males - RMD92 and RMC92
values of $q_{x}$

| Age $x$ | Deferred <br> RMD92 | $\begin{aligned} & \text { Combined } \\ & \text { RMC92 } \end{aligned}$ |
| :---: | :---: | :---: |
| 17 | 0.000393 | 0.000401 |
| 18 | 0.000392 | 0.000401 |
| 19 | 0.000392 | 0.000401 |
| 20 | 0.000392 | 0.009403 |
| 21 | 0.000394 | 0.000405 |
| 22 | 0.000396 | 0.000408 |
| 23 | 0.000400 | 0.000413 |
| 24 | 0.000404 | 0.000419 |
| 25 | 0.000411 | 0.000426 |
| 26 | 0.000419 | 0.000435 |
| 27 | 0.000429 | 0.000446 |
| 28 | 0.000442 | 0.000459 |
| 29 | 0.000457 | 0.000475 |
| 30 | 0.000475 | 0.000494 |
| 31 | 0.000497 | 0.000516 |
| 32 | 0.000523 | 0.000543 |
| 33 | 0.000553 | 0.000573 |
| 34 | 0.000589 | 0.000609 |
| 35 | 0.000630 | 0.000650 |
| 36 | 0.000679 | 0.000697 |
| 37 | 0.000734 | 0.000752 |
| 38 | 0.000799 | 0.000815 |
| 39 | 0.000872 | 0.000887 |
| 40 | 0.000957 | 0.000969 |
| 41 | 0.001053 | 0.001064 |
| 42 | 0.001162 | 0.001171 |
| 43 | 0.001285 | 0.001293 |
| 44 | 0.001425 | 0.001432 |
| 45 | 0.001583 | 0.001589 |
| 46 | 0.001760 | 0.001767 |
| 47 | 0.001959 | 0.001969 |
| 48 | 0.002181 | 0.002196 |
| 49 | 0.002429 | 0.002452 |
| 50 | 0.002705 | 0.002741 |
| 51 | 0.003011 | 0.003065 |
| 52 | 0.003351 | 0.003430 |
| 53 | 0.003725 | 0.003839 |
| 54 | 0.004138 | 0.004296 |

Table B1. (Continued).

| Age $x$ | Deferred <br> RMD92 | Combined <br> RMC92 |
| :--- | :--- | :---: |
| 55 | 0.004592 | 0.004809 |
| 56 | 0.005090 | 0.005381 |
| 57 | 0.005635 | 0.006019 |
| 58 | 0.006230 | 0.006731 |
| 59 | 0.006877 | 0.007523 |
| 60 | 0.007580 | 0.008404 |
| 61 | 0.008341 | 0.009381 |
| 62 | 0.009164 | 0.010466 |
| 63 | 0.010050 | 0.011666 |
| 64 | 0.011004 | 0.012995 |
| 65 | 0.012026 | 0.014462 |
| 66 | 0.013119 | 0.016081 |
| 67 | 0.014286 | 0.017865 |
| 68 | 0.015528 | 0.019828 |
| 69 | 0.016846 | 0.021985 |
| 70 | 0.018241 | 0.024353 |
| 71 | 0.019714 | 0.026947 |
| 72 | 0.021264 | 0.029787 |
| 73 | 0.022892 | 0.032891 |
| 74 | 0.024597 | 0.036279 |
| 75 | 0.026378 | 0.039971 |
| 76 |  | 0.043988 |
| 77 |  | 0.048353 |
| 78 |  | 0.053089 |
| 79 |  | 0.058219 |
| 80 |  | 0.063767 |
| 81 |  | 0.069756 |
| 82 |  | 0.076213 |
| 83 |  | 0.083160 |
| 84 |  | 0.090623 |
| 85 |  | 0.098625 |
| 86 |  | 0.107189 |
| 87 |  | 0.116337 |
| 88 |  | 0.126091 |
| 89 |  | 0.136470 |
| 90 |  | 0.147491 |
| 91 |  | 0.159170 |
| 4 |  | 0.184549 |
| 4 |  |  |

000000 I ..... OZI
8Et689．0 ..... 61 I
$\pm 65699^{\circ} 0$ ..... 8II
$6526+9^{\circ} 0$ ..... LII
96t8Z90 ..... 911
$69 \varepsilon L 09^{\circ} 0$ ..... SII
$806585^{\circ} 0$ ..... ゅII
90Et950 ..... £II
9ISてtč0 ..... ZII
IS90Zs． 0 ..... ［II
L8L86t 0 ..... 0II
9669んが0 ..... 601
$6 \pm E S S t 0$ ..... 801
916E\＆ち0 ..... 201
19LてIt゚ 0 ..... 901
Lt616e： 0 ..... s0I
てESTLE＇0 ..... t0I
L9SISE： 0 ..... E0I
00IZEE： 0 ..... Z0I
SLIEIE0 ..... I0I
LZ8t6で0 ..... 001
680LLで0 ..... 66
9866さで0 ..... 86
6\＆ऽEャで0 ..... L6
t9LLでく ..... 96
1L9ZIZ゚0 ..... S6

| z6OWY | z6OW\％ |  |
| :---: | :---: | :---: |
| peutquro | рәıəjə | $x \stackrel{3}{ }{ }^{\text {V }}$ |



Table B2. Retirement annuitants, females - RFD92 and RFC92
values of $q_{x}$

| Age $x$ | Deferred <br> RFD92 | Combined <br> RFC92 |
| :--- | :--- | :---: |
| 17 | 0.000176 | 0.000189 |
| 18 | 0.000180 | 0.000195 |
| 19 | 0.000185 | 0.000201 |
| 20 | 0.000190 | 0.000208 |
| 21 | 0.000196 | 0.000216 |
| 22 | 0.000203 | 0.000224 |
| 23 | 0.000211 | 0.000234 |
| 24 | 0.000220 | 0.000245 |
| 25 | 0.000231 | 0.000256 |
| 26 | 0.000242 | 0.000270 |
| 27 | 0.000256 | 0.000284 |
| 28 | 0.000271 | 0.000301 |
| 29 | 0.000288 | 0.000319 |
| 30 | 0.000307 | 0.000339 |
| 31 | 0.000329 | 0.000362 |
| 32 | 0.000354 | 0.000387 |
| 33 | 0.000382 | 0.000415 |
| 34 | 0.000413 | 0.000447 |
| 35 | 0.000449 | 0.000482 |
| 36 | 0.000488 | 0.000520 |
| 37 | 0.000532 | 0.000563 |
| 38 | 0.000582 | 0.000612 |
| 39 | 0.000637 | 0.000665 |
| 40 | 0.000698 | 0.000725 |
| 41 | 0.000766 | 0.000791 |
| 42 | 0.000842 | 0.000865 |
| 43 | 0.000926 | 0.000947 |
| 44 | 0.001019 | 0.001039 |
| 45 | 0.001122 | 0.001141 |
| 46 | 0.001235 | 0.001255 |
| 47 | 0.001360 | 0.001381 |
| 48 | 0.001497 | 0.001522 |
| 49 | 0.001648 | 0.001679 |
| 50 | 0.001813 | 0.001853 |
| 51 | 0.001994 | 0.002048 |
| 52 | 0.002191 | 0.002264 |
| 53 | 0.002505 |  |
| 54 |  | 0.002773 |
|  |  |  |
| 1 |  |  |

Table B2. (Continued).

| Age $x$ | Deferred RFD92 | Combined RFC92 |
| :---: | :---: | :---: |
| 55 | 0.002895 | 0.003072 |
| 56 | 0.003172 | 0.003404 |
| 57 | 0.003471 | 0.003774 |
| 58 | 0.003795 | 0.004186 |
| 59 | 0.004145 | 0.004645 |
| 60 | 0.004522 | 0.005155 |
| 61 | 0.004927 | 0.005723 |
| 62 | 0.005363 | 0.006355 |
| 63 | 0.005830 | 0.007058 |
| 64 | 0.006330 | 0.007841 |
| 65 | 0.006864 | 0.008712 |
| 66 | 0.007434 | 0.009681 |
| 67 | 0.008040 | 0.010758 |
| 68 | 0.008685 | 0.011957 |
| 69 | 0.009368 | 0.013290 |
| 70 | 0.010092 | 0.014772 |
| 71 | 0.010858 | 0.016420 |
| 72 | 0.011665 | 0.018252 |
| 73 | 0.012515 | 0.020288 |
| 74 | 0.013409 | 0.022551 |
| 75 | 0.014346 | 0.025064 |
| 76 |  | 0.027855 |
| 77 |  | 0.030953 |
| 78 |  | 0.034391 |
| 79 |  | 0.038206 |
| 80 |  | 0.042436 |
| 81 |  | 0.047125 |
| 82 |  | 0.052319 |
| 83 |  | 0.058070 |
| 84 |  | 0.064432 |
| 85 |  | 0.071467 |
| 86 |  | 0.079238 |
| 87 |  | 0.087816 |
| 88 |  | 0.097273 |
| 89 |  | 0.107689 |
| 90 |  | 0.119146 |
| 91 |  | 0.131732 |
| 92 |  | 0.145535 |
| 93 |  | 0.160648 |
| 94 |  | 0.177163 |

Table B2. (Continued).

| Age $x$ | Deferred <br> RFD92 | Combined <br> RFC92 |
| :--- | :--- | :--- |
| 95 |  | 0.195171 |
| 96 |  | 0.214760 |
| 97 |  | 0.236014 |
| 98 | 0.259005 |  |
| 99 |  | 0.283793 |
| 100 |  | 0.310422 |
| 101 |  | 0.338912 |
| 102 |  | 0.369254 |
| 103 |  | 0.401405 |
| 104 | 0.435283 |  |
| 105 |  | 0.470757 |
| 106 |  | 0.507643 |
| 107 |  | 0.545701 |
| 108 |  | 0.584630 |
| 109 |  | 0.624066 |
| 110 |  | 0.663591 |
| 111 | 0.702735 |  |
| 112 |  | 0.740987 |
| 113 |  | 0.777821 |
| 114 |  | 0.812707 |
| 115 |  | 0.845149 |
| 116 |  | 0.901033 |
| 117 | 0.923893 |  |
| 118 |  | 0.943192 |
| 119 |  | 1.000000 |
| 120 |  |  |
|  |  |  |

## APPENDIXC

## FORMULAE FORTHENEW STANDARDTABLES

In this appendix the parameters used for the calculation of the adjusted values of $\mu_{x}$ are described. The formulae apply over all ages of the standard tables apart from the self employed retirement annuitants, males and females, where separate formulae are applied up to age 74 and 75 and over, as shown.

It should be noted that the formulae described in this appendix are used to calculate values of $\mu_{x}$ for the various experiences into which the data are classified. Using a numerical integration technique, these values of $\mu_{x}$ can then be used to calculate the values of $q_{x}, q_{[x]-1}, q_{x}$ etc, shown in Appendices A and B. The values of $q_{[x]}, q_{[x]+1}, q_{x}$ etc, rounded to six decimal places, are then used to calculate values of $\mu_{[x]}, \mu_{[x]-1}, \mu_{x}$ etc as described in section 2.6 of C.M.I.R. 10 .

Table C1. Graduation formulae and their parameter values.

| Parameters |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $\mathrm{GM}(r, s)$ |  |  |  |  |  |
|  | $100 a_{0}$ | $100 a_{1}$ | $b_{0}$ | $b_{1}$ | $b_{2}$ |  |

Permanent assurances, males

| Duration 0 | $\mathrm{GM}(2,3)$ | 0.02 | -0.02 | -4.755647 | 5.236521 | -0.6 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Duration 1 | $\mathrm{GM}(2,3)$ | 0.02296 | -0.03 | -5.056244 | 5.0 | -1.2 |
| Durations 2+ | $\mathrm{GM}(2,3)$ | 0.005887 | -0.049883 | -4.363378 | 5.544956 | -0.620345 |

Permanent assurances, females

| Duration 0 | GM(1,2) | 0.008 | - | -4.978100 | 5.07878 | - |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Duration 1 | GM(1,2) | 0.008 | - | -4.763844 | 4.85445 | - |
| Durations 2+ | GM(1,2) | 0.011189 | - | -4.331121 | 5.135803 | - |

Temporary assurances, males
$\begin{array}{lllllll}\text { Duration } 0 & G M(2,3) & -0.12 & -0.17 & -4.415430 & 4.486721 & -0.12\end{array}$
Durations 1-4 GM(2,3) $-0.113910 \quad-0.173839 \quad-4.170362$ 4.745634 -0.119740
Durations 5+ $\begin{array}{lllllll}\mathrm{GM}(2,3) & -0.10 & -0.17 & -3.851692 & 5.43387 & -0.12\end{array}$
Temporary assurances, females
Duration $0 \quad$ GM $(1,2) \quad 0.008 \quad \cdots \quad-5.082158 \quad 5.493452$
$\begin{array}{lllllll}\text { Durations 1-4 } & \text { GM }(1,2) & 0.012 & - & -4.462904 & 5.240908 & - \\ \text { Durations 5+ } & \text { GM }(1,2) & 0.018 & & -4.352028 & 5.347124 & -\end{array}$
Durations 5+ GM(1,2) $0.018 \quad-\quad-4.352028 \quad 5.347124 \quad-$
Pensioners, males
$\begin{array}{lcccccc}\text { Lives } & \text { GM(2,3) } & -0.0081 & -0.07 & -4.67509 & 5.629188 & -1.2 \\ \text { Amounts } & \text { GM(2,3) } & 0.023 & -0.011 & -5.39778 & 6.622746 & -1.6\end{array}$
Pensioners, females
$\begin{array}{lllllll}\text { Lives } & \mathrm{GM}(2,3) & 0.0108 & -0.014 & -4.97225 & 5.884075 & -1.0 \\ \text { Amounts } & \mathrm{GM}(2,3) & 0.0119 & -0.008 & -5.26110 & 5.982521 & -1.15\end{array}$
Immediate annuitants, males

| Lives | Duration 0 | $\mathrm{GM}(2,3)$ | 0.010649 | -0.029980 | -4.703659 | 5.568973 | -0.654909 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Durations 1+ | $\mathrm{GM}(2,3)$ | 0.014429 | -0.040629 | -4.399861 | 5.568973 | -0.654909 |
| Amounts | Duration 0 | $\mathrm{GM}(2,3)$ | 0.011251 | -0.031682 | -4.648604 | 5.568973 | -0.654909 |
|  | Durations I+ | $\mathrm{GM}(2,3)$ | 0.013757 | -0.038737 | -4.447540 | 5.568973 | -0.654909 |

Immediate annuitants, females

| Lives | Duration 0 | GM(1,3) | 0.021517 | - | -5.597708 | 6.683129 | -0.9 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Durations 1+ | GM(1,3) | 0.03 | - | -5.265363 | 6.683129 | -0.9 |
| Amounts | Duration 0 | GM(1,3) | 0.03289 | - | -6.378326 | 8.027676 | -1.5 |
|  | Durations 1+ | GM(1,3) | 0.04 | - | -6.182627 | 8.027676 | -1.5 |

Table C1. (Continued).

|  |  | Parameters |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\mathrm{GM}(r, s)$ | $100 a_{0}$ | $100 a_{1}$ | $b_{0}$ | $b_{1}$ | $b_{2}$ |
| Widows |  |  |  |  |  |  |  |
|  | Lives | GM(1,2) | 0.02 | - | -3.795522 | 4.308854 | - |
|  | Amounts | GM(1,2) | 0.018 | - | -3.921243 | 4.444249 | - |
| Retirement annuitants, males |  |  |  |  |  |  |  |
| Vested | Pre age 75 | GM $(2,3)$ | 0.023761 | -5.0 | -4.713208 | 6.0 | -1.0 |
| Vested | Post age 74 | GM $(2,3)$ | 0.01647 | $-0.02022$ | -4.39933 | 5.21998 | $-0.63741$ |
| Retirement annuitants, females |  |  |  |  |  |  |  |
| Vested | Pre age 75 | GM( 2,3 ) | -0.943368 | $-5.0$ | -4.737523 | 5.0 | -1.0 |
| Vested | Post age 74 | GM(1,2) | 0.014 | - | -4.27128 | 5.378175 | - |

# AN INVESTIGATION INTO THE DISTRIBUTION OF POLICIES PERLIFE ASSUREDIN THE CAUSE OF DEATHINVESTIGATION 1991-94 

The continuous investigation into the mortality of assured lives is based on policies and not lives. Consequently the death of a policyholder carrying $n$ policies appears as $n$ deaths in the data. In order to estimate the standard deviations needed to test differences between actual and expected deaths, information is required about the distribution of policies per life assured. This information is not contained in the data submitted to the main assured lives investigations but is included in the data for deaths in the cause of death investigation. This data has been used as a proxy for direct information about the distribution of duplicates in the assured lives investigations, for the purpose of the graduations described elsewhere in this edition of C.M.I.R.

The last report by the C.M.I. Committee into the distribution of duplicates amongst deaths in the cause of death investigation was contained in C.M.I.R. 8, 49-58 (1986). The work described in that report to analyse the cause of death experience was repeated for the 1991-94 experience. Table 1 shows for U.K. males, at durations 2 and over, the 'variance ratios' in quinary age groups and at all ages combined. This is the ratio of the variance of a distribution where there are duplicates to that of a straightforward binomial variance. As previously reported, this can be approximated by $m_{1} / m_{2}$ where $m_{1}$ and $m_{2}$ are the first and second moments about zero of the duplicates' distribution.

Table 2 shows the number of deaths with $n$ policies contained in the data. The values of $m_{1}$ and $m_{2}$ shown in Table 1 can be calculated directly from this table.

The total number of policies held by lives who died in the period, as reported in Table 1, is 49,353 . This compares with 68,963 deaths in the assured lives investigation in the same period. The difference is represented by offices who did not contribute data to the cause of death investigation and by offices who were not able to provide death certificates for all deaths.

Table 1 shows values that are, in all cases except the two indicated, lower than those reported in C.M.I.R. 8 for the investigation period 1981-82. The reason for the decline in the number of duplicate policies per life is not clear. It does not seem to accord with the perception that the insured population takes out more policies as inflation erodes the value of their existing contracts. The number of lives in both the cause of death and assured lives investigation has been declining and it is not known how this has affected the results shown in Table 2.

Table 1. First and second moments about zero together with the variance ratios of the distribution of duplicates in the cause of death investigation: assured lives (whole-life and endowment), males, U.K., durations 2 and over, 1991-94.

| Age group | $m_{1}$ | $m_{2}$ | $m_{1} / m_{2}$ | Number of lives | Number of policies |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Under 25 | 1.080 | $1.277^{*}$ | $1.18^{*}$ | 112 | 121 |
| $25-29$ | 1.113 | 1.466 | 1.32 | 204 | 227 |
| $30-34$ | 1.155 | 1.532 | 1.33 | 361 | 417 |
| $35-39$ | 1.230 | 1.979 | 1.61 | 662 | 814 |
| $40-44$ | 1.263 | 2.100 | 1.66 | 1,304 | 1,647 |
| $45-49$ | 1.312 | 2.292 | 1.75 | 2,754 | 3,612 |
| $50-54$ | 1.301 | 2.195 | 1.69 | 4,012 | 5,219 |
| $55-59$ | 1.281 | 2.107 | 1.65 | 6,079 | 7,785 |
| $60-64$ | 1.216 | 1.814 | 1.49 | 8,220 | 9,995 |
| $65-69$ | 1.113 | 1.399 | 1.26 | 4,274 | 4,755 |
| $70-74$ | 1.079 | 1.284 | 1.19 | 3,695 | 3,988 |
| $75-79$ | 1.106 | 1.402 | 1.27 | 3,407 | 3,768 |
| $80-84$ | 1.088 | 1.326 | 1.22 | 3,472 | 3,776 |
| $85-89$ | 1.103 | 1.356 | 1.23 | 1,875 | 2,068 |
| 90 and over | 1.104 | 1.429 | 1.29 | 1,052 | 1,161 |
|  |  |  |  |  |  |
| All ages | 1.190 | 1.736 | 1.46 | 41,483 | 49,353 |

Values marked with * are bigger than the equivalent calculated from the 1981-82 data.

Table 2. Numbers of deaths with $n$ policies in the 1991-94 cause of death investigation: assured lives (whole-life and endowment), males, U.K., durations 2 and over.

|  | Number of policies per life ( $n$ ) |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | $10+$ |
| Under 25 | 105 | 5 | 2 | - | - | - | - | - | - | - |
| 25-29 | 189 | 11 | 1 | 2 | 1 | - | - | - | - | - |
| 30-34 | 315 | 38 | 6 | 2 | - | - | - | - | - | - |
| 35-39 | 556 | 79 | 20 | 4 | -- | - | 1 | 1 | 1 | - |
| 40-44 | 1,066 | 176 | 40 | 15 | 3 | - | 1 | 1 | 1 | 1 |
| 45-4¢ | 2,173 | 404 | 125 | 35 | 4 | 6 | 1 | 3 | 1 | 2 |
| 50-54 | 3,176 | 594 | 165 | 45 | 20 | 8 | 2 | 1 | - | 1 |
| 55-59 | 4,871 | 879 | 230 | 63 | 21 | 6 | 4 | 2 | 1 | 2 |
| $60 \cdots 64$ | 6,897 | 1,010 | 228 | 54 | 17 | 8 | 5 | - | -- | 1 |
| 65-69 | 3,873 | 346 | 42 | 8 | 2 | 1 | 1 | .. | 1 | - |
| 70-74 | 3,455 | 203 | 28 | 6 | 2 | - | - | - | 1 | - |
| 75-79 | 3,133 | 216 | 43 | 8 | 3 | 3 | - | .. | 1 | - |
| 80-84 | 3,235 | 197 | 24 | 9 | 4 | 2 | 1 | - | - | - |
| 85-89 | 1,717 | 131 | 20 | 6 | 1 | - | - | - | - | - |
| 90 and over | 975 | 61 | 8 | 4 | 2 | - | 2 | - | - | - |
| All ages | 35,736 | 4,350 | 982 | 261 | 80 | 34 | 18 | 8 | 7 | 7 |

## CONTINUOUS MORTALITY INVESTIGATION REPORTS

## NUMBER 17

Introduction ..... iii
Standard Tables of Mortality based on the 1991-94 Experiences

1. Permanent Assurances, Males and Females ..... 1
2. Immediate Annuitants, Males and Females, Lives and Amounts ..... 27
3. Self-Employed Retirement Annuitants, Males and Females ..... 44
4. Widows of Life Office Pensioners, Lives and Amounts ..... 59
5. Temporary Assurances, Males and Females ..... 68
6. Projection Factors for Mortality Improvement ..... 89
7. Specimen Monetary Values and Comparisons with Other Tables ..... 109
8. Publication of the New Tables ..... 146
References ..... 148
Appendix A Values of Mortality Rates for the New Standard Tables ..... 149
Appendix B Values of Mortality Rates for Other Tables ..... 218
Appendix C Formulae for the New Standard Tables ..... 225
The Distribution of Policies per Life Assured ..... 229

[^0]:    *An isolated case at 100 ; the rest of the data ceases at age 89 .

[^1]:    *Note that the graduated rates for vested above age 75 are assumed to be the same as the graduated rates for combined.

[^2]:    *An isolated case at 95 ; the rest of the data ceases at age 87 .

