

New Directions
in the Modelling of Longevity Risk

Andrew Cairns
Heriot-Watt University,
and The Maxwell Institute, Edinburgh

Joint work (in progress!) with:

George Mavros, Torsten Kleinow, George Streftaris

Life Convention, Brussels, 5 November 2012

Plan

- Motivation
- Genealogy
- New directions in modelling
- Numerical illustrations – single population models
- Remarks on multiple populations

Motivation

- Application focus:
 - risk measurement and management of longevity risk
 - multiple populations
 - life insurance diversification benefits
 - basis risk in standardised longevity contracts
- industry requires robust models

Development of New Models

- Many new stochastic mortality models since Lee-Carter
- Are they fit for purpose?
- Are they robust?

GENEALOGY: 1st GENERATION MODELS

Eilers/Marx
P-splines

Currie/Richards (M4)
2-D P-splines
2002, ...



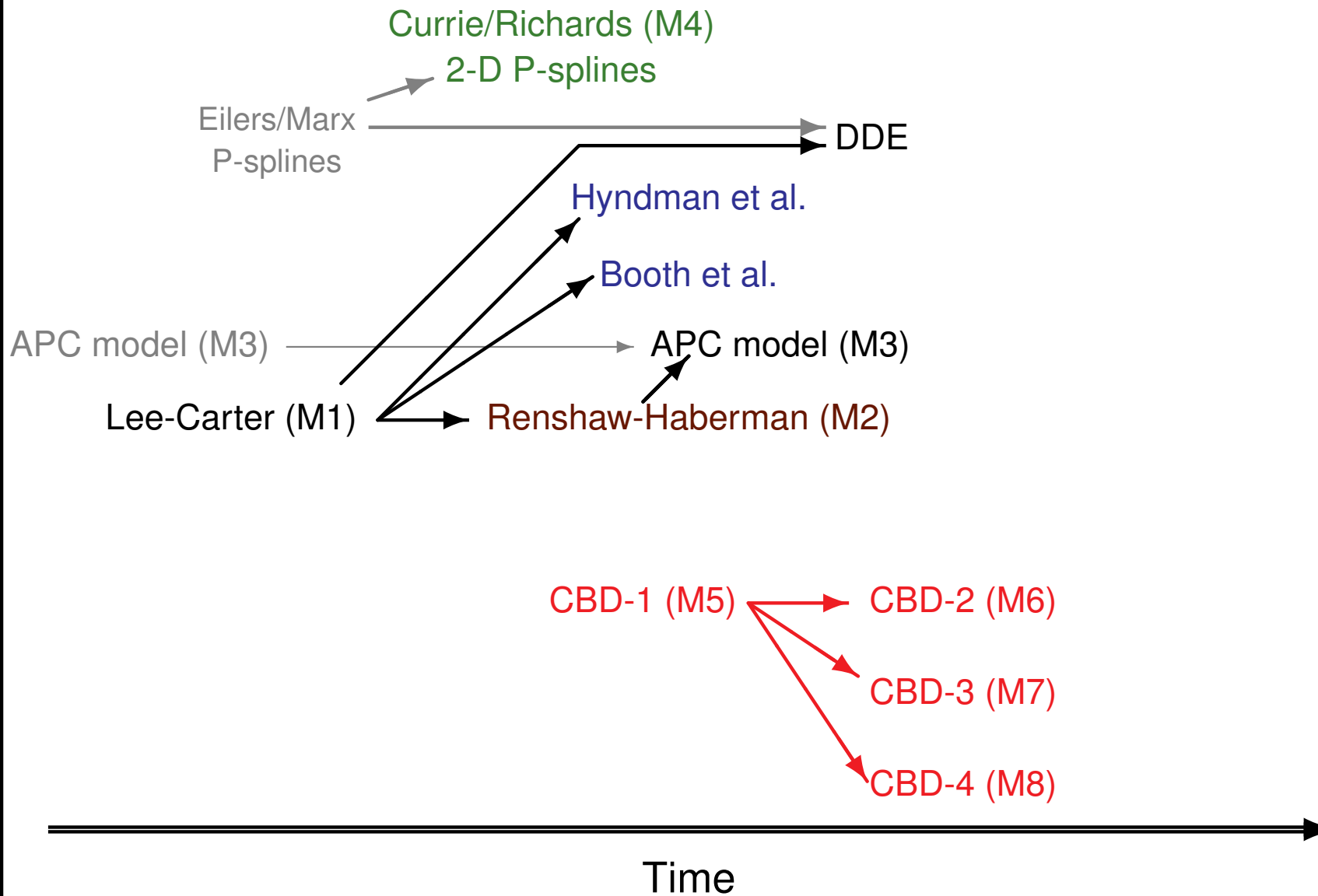
Lee-Carter (M1)
1992

CBD-1 (M5)
2006

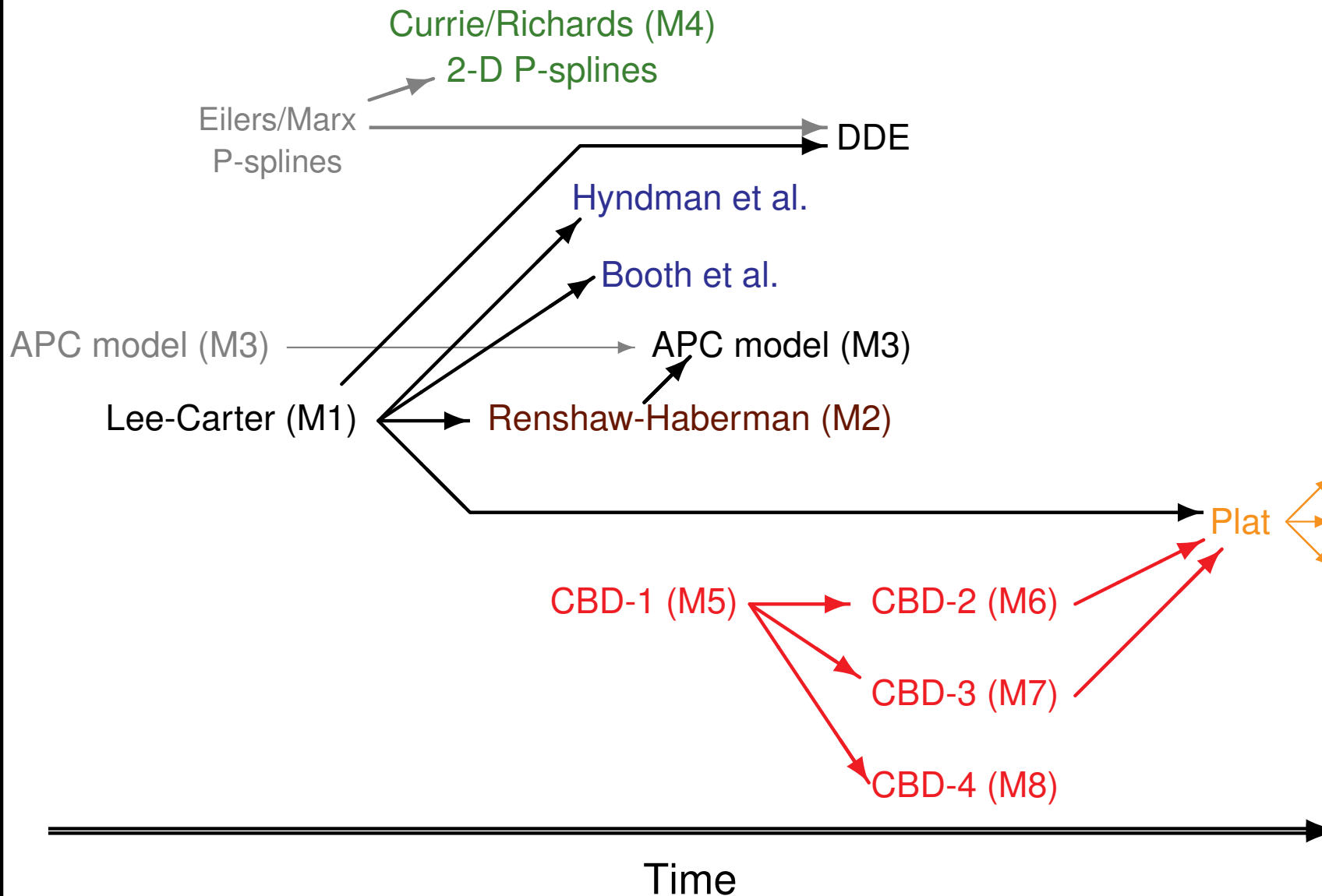


Time

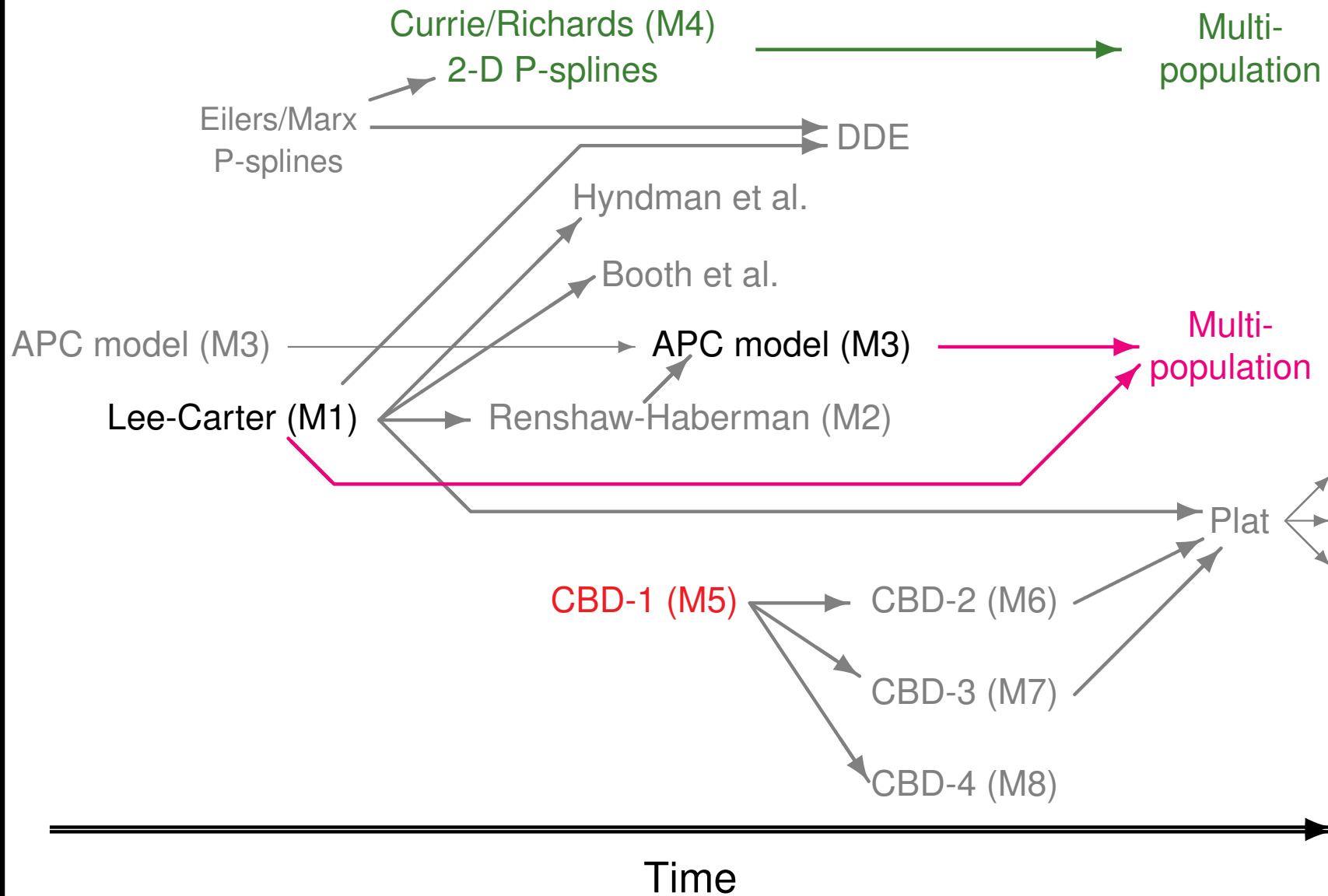
Improvements + more complexity



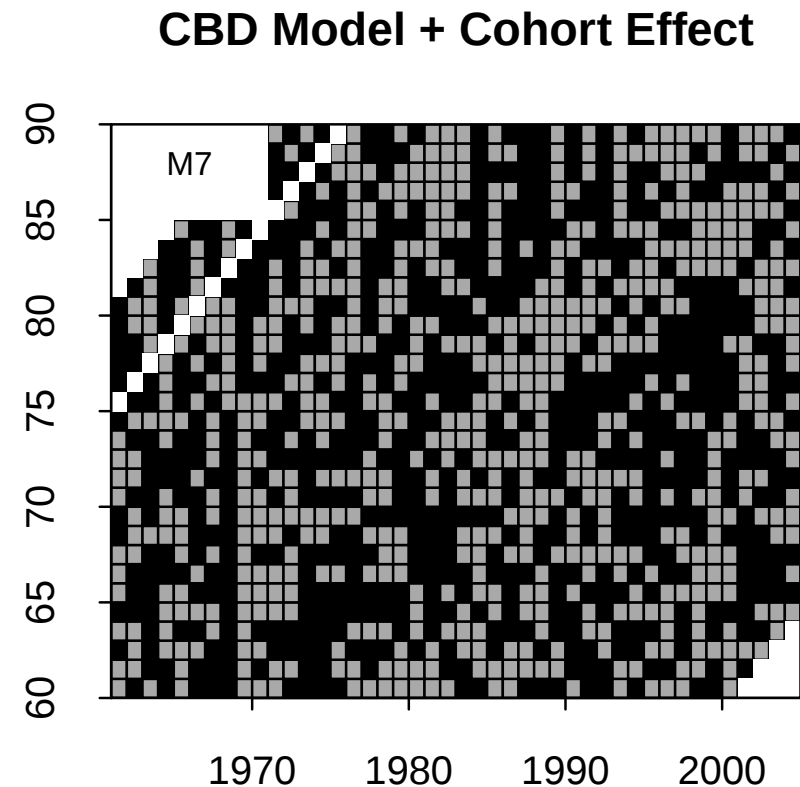
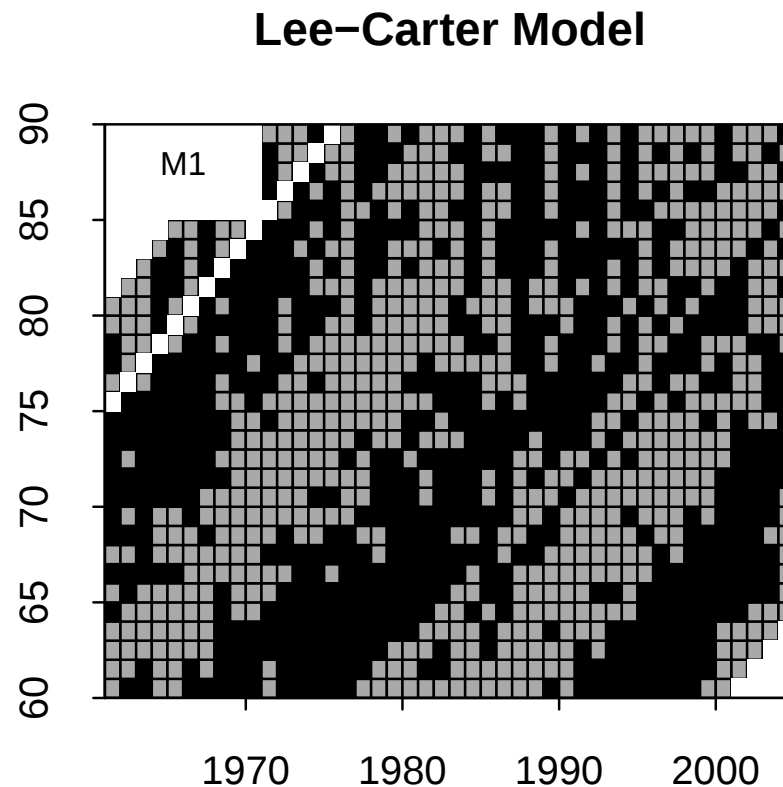
More improvements + even more complexity



Multiple population modelling



Why do we need complexity?



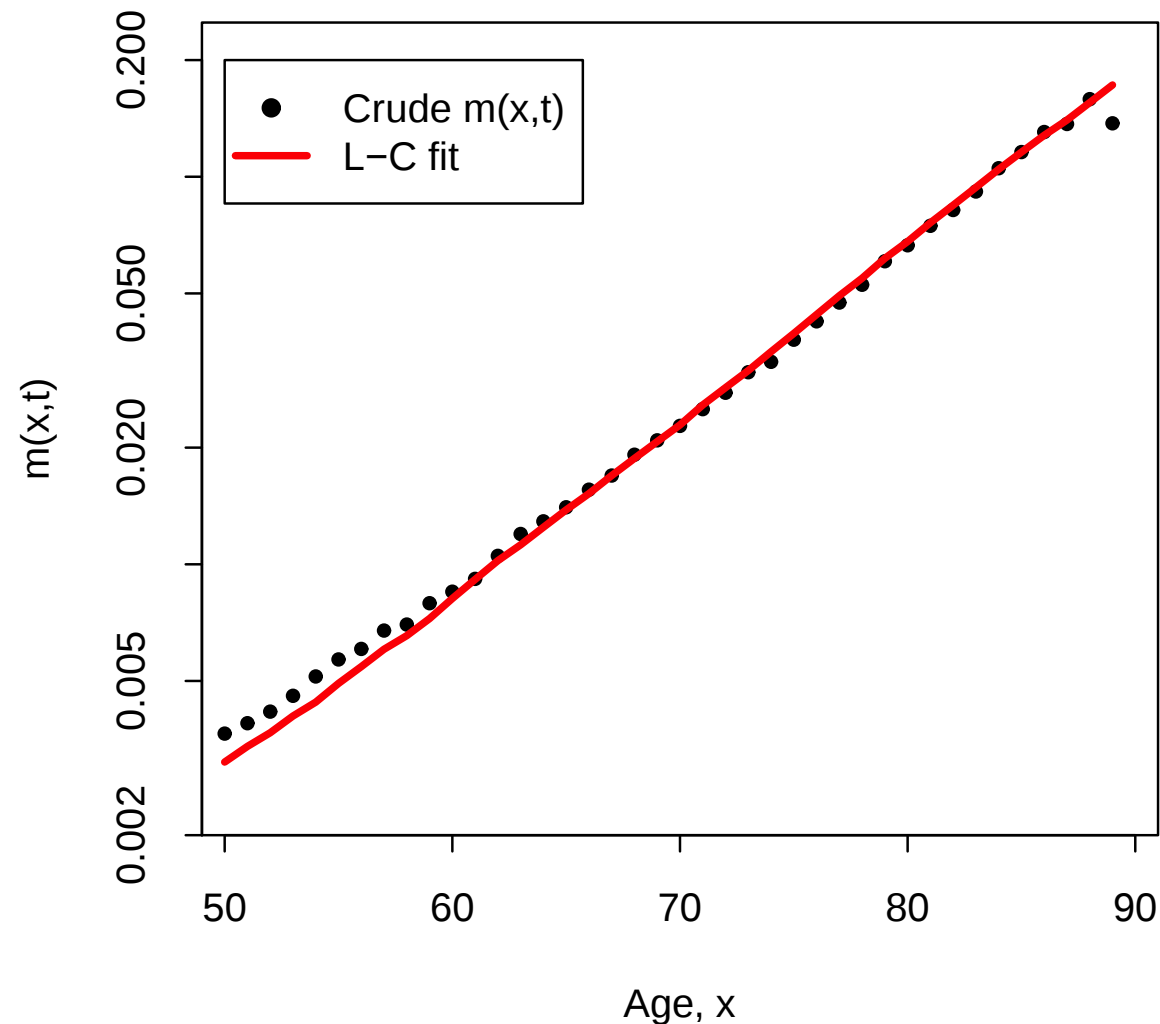
Black $\Rightarrow$ model *over*-estimates $m(x, t)$ death rate

Gray $\Rightarrow$ model *under*-estimates $m(x, t)$ death rate

LC: non-random clusters + errors are too big

Need for complexity: accurate base table for forecasts

EW males 1971–2008: Lee–Carter Fit
 $m(x, 2008)$ (log scale)



Issues on complexity

- More complex $\Rightarrow$ More random processes
- More random processes $\Rightarrow$
MUCH more difficult to model multiple populations
- Excessive complexity $\Rightarrow$
potentially less robust forecasts

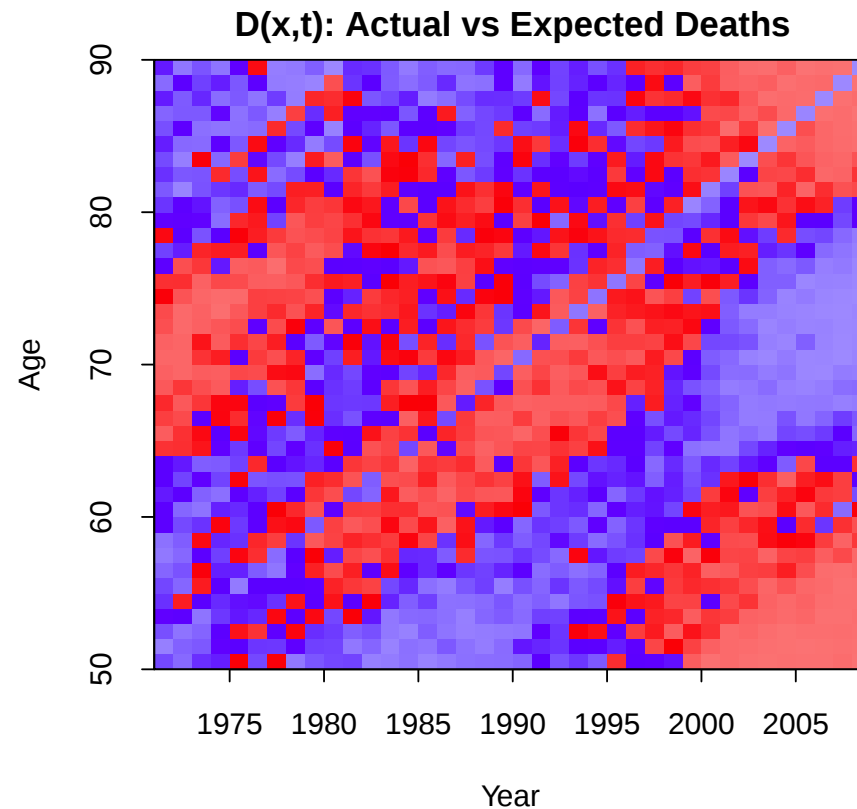
A Possible Way Forward

Single-population models

- Focus on **small number of key drivers**
⇒ much easier to extend to multi-populations
- Focus on greater robustness of forecasts

Case Study: CBD/Plat Revisited

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x})$$



Red $\Rightarrow$ actual deaths $>$ expected deaths

CBD/Plat Revisited: **Key Idea: Possible responses**

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x})$$

Add:

- Cohort effect, $\gamma(t - x)$
- Extra age-period effects
- Do something new

Key Idea: CBD/Plat Revisited

Underlying $\log m(x, t) =$

- $\beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x})$: **two key drivers**

PLUS

$R(x, t)$ *Residuals*

- Assume: vector $R(t) \rightarrow R(t + 1)$ mean reverting process

$\Rightarrow$ long term risk depends on **two key drivers**

Specific Model

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x}) + R(x, t)$$

- $(\kappa_1(t), \kappa_2(t))$: bivariate random walk
- $R(t) = (n_x \times 1 \text{ vector})$ VAR(2), reverting to 0

$$R(t) = AR(t-1) + BR(t-2) + Z(t)$$

- $Z(x, t) \sim \text{i.i.d. } N(0, \sigma_Z^2)$
- $A = A_1 + A_2$ and $B = -A_2A_1$

VAR matrices A_1 and A_2

$$A_i = \begin{pmatrix} a_i & 0 & 0 & \dots & & & \\ c_i & d_i & 0 & 0 & \dots & & \\ d_i/2 & c_i & d_i/2 & 0 & 0 & \dots & \\ 0 & d_i/2 & c_i & d_i/2 & 0 & 0 & \dots \\ 0 & 0 & d_i/2 & c_i & d_i/2 & 0 & 0 & \dots \\ \vdots & & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \end{pmatrix}$$

a_i = AR terms for new members;

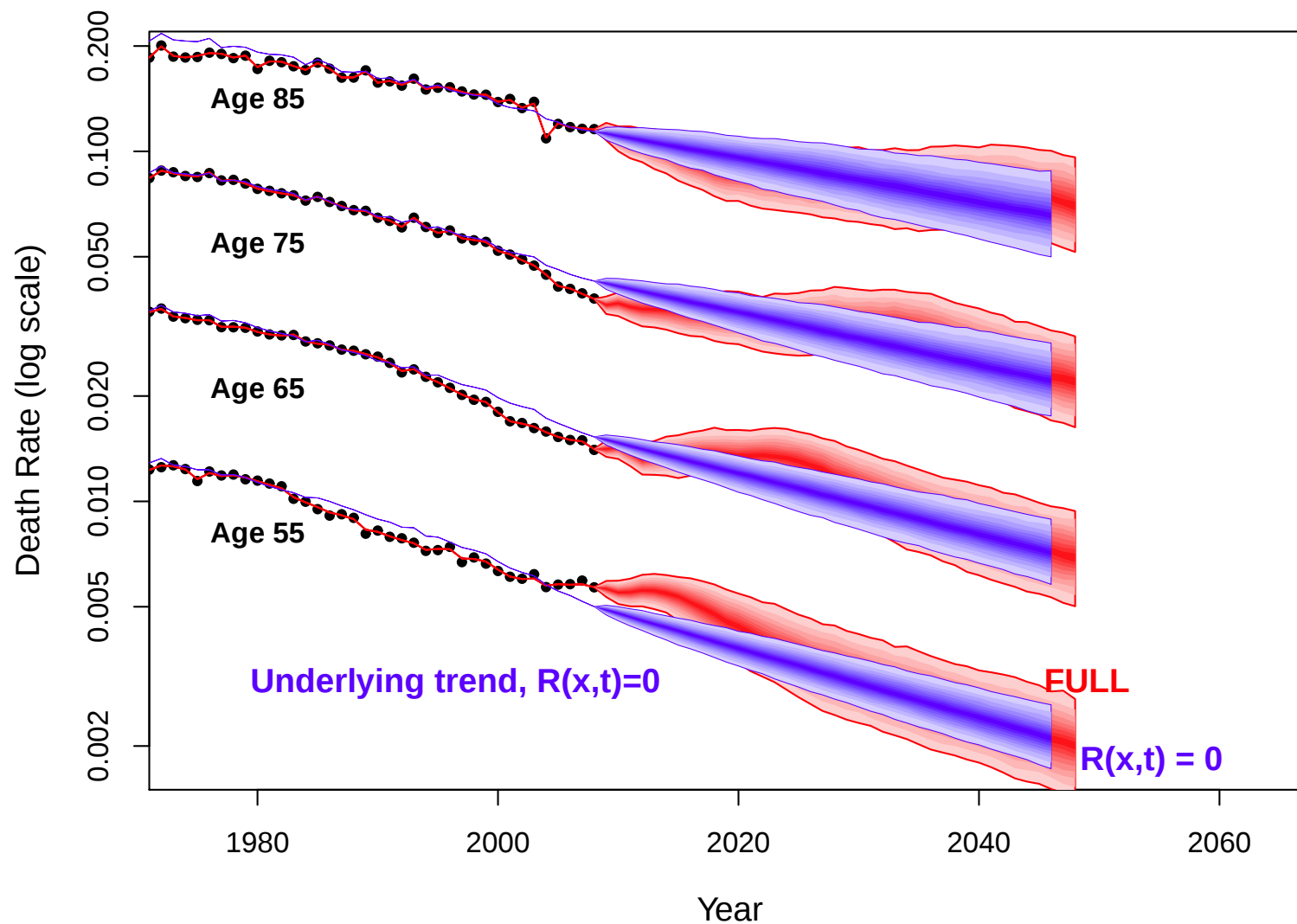
c_i = cohort persistence;

d_i = diffusion coeff.

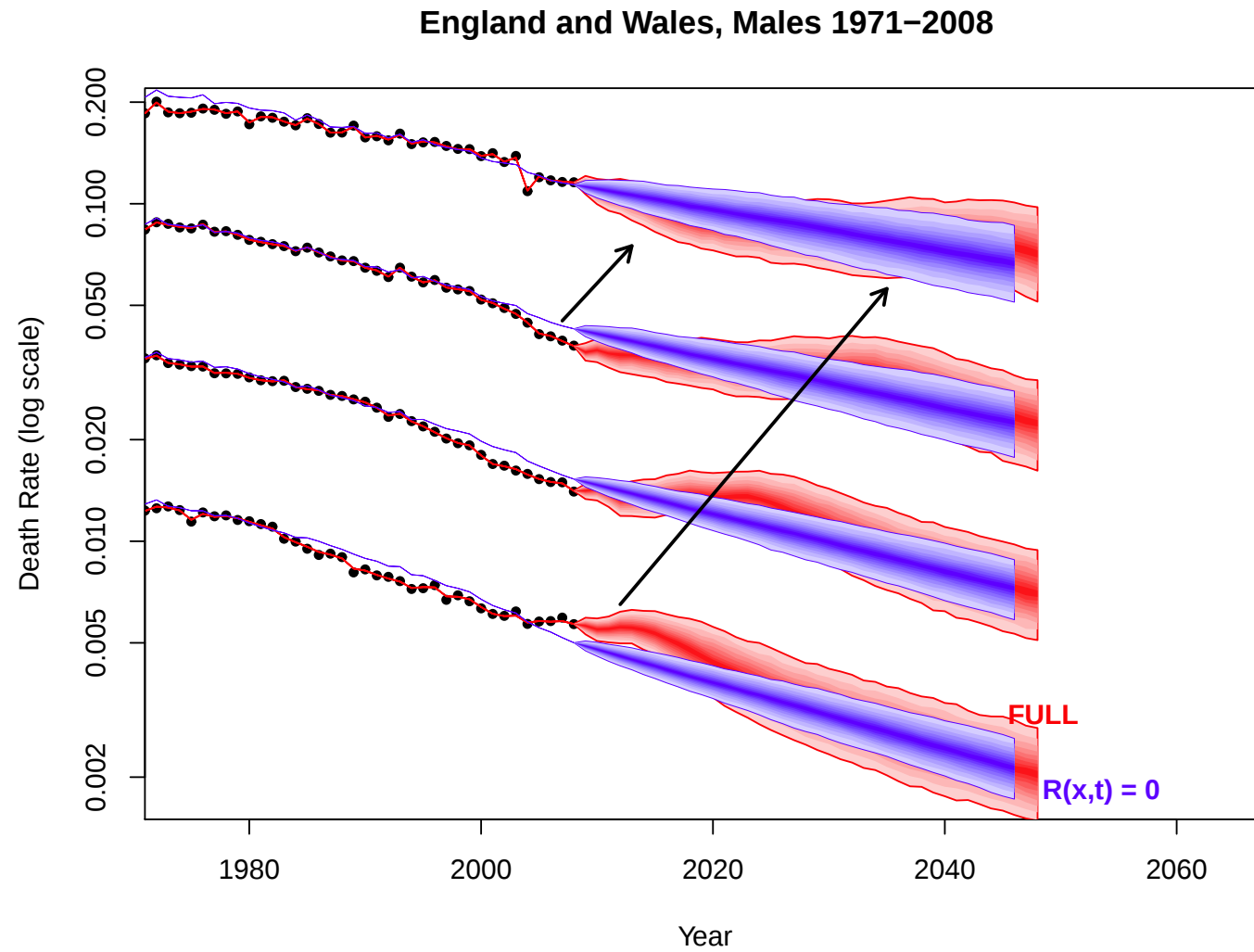
Further details

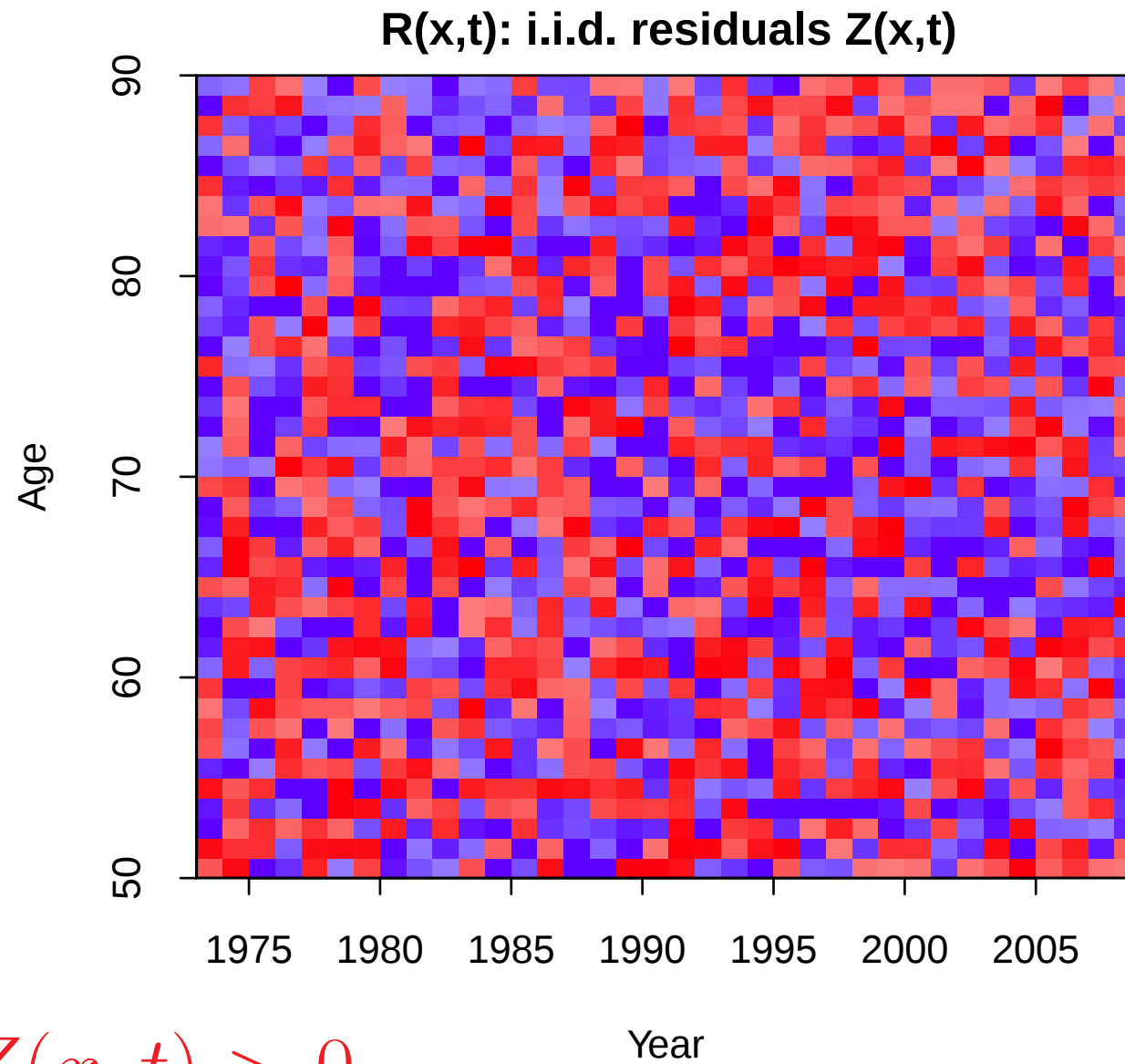
- Deaths: $D(x, t) \sim \text{Poisson}(m(x, t)E(x, t))$
- Bayesian approach:
posterior density = likelihood $\times$ prior
- Upcoming results: mode of posterior density
- Further work: Bayesian parameter uncertainty

England and Wales, Males 1971–2008

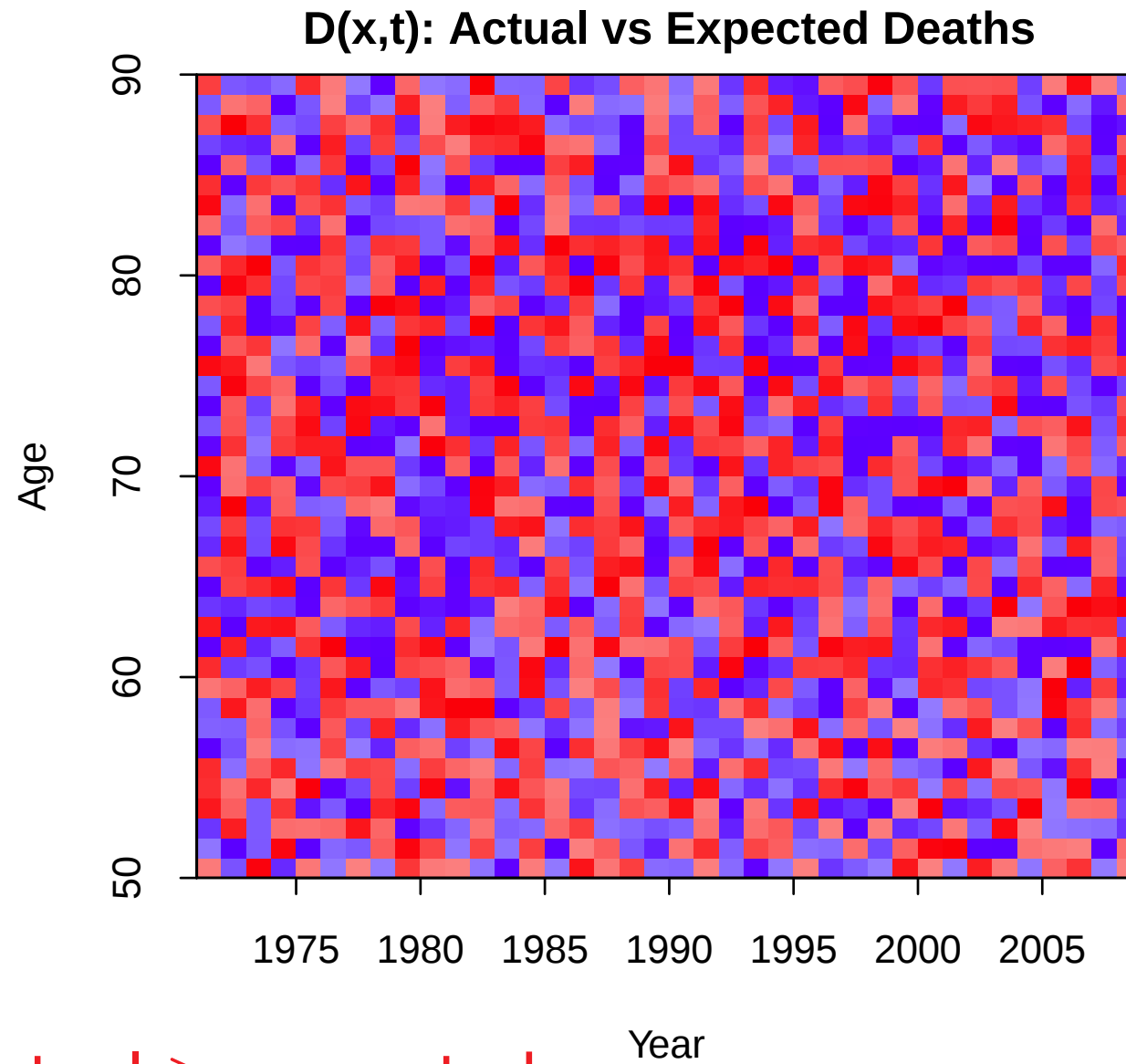


Cohort-type effects





Red $\Rightarrow Z(x, t) > 0$



Red $\Rightarrow$ actual $>$ expected

Comparison with related models

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x})$$

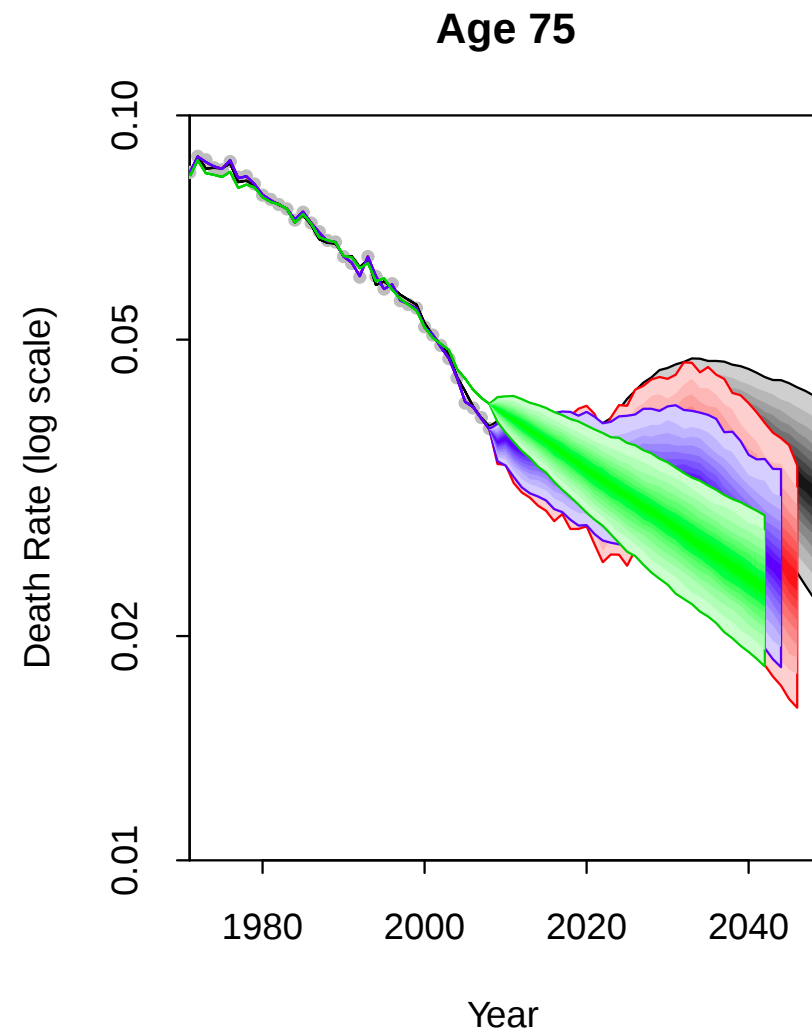
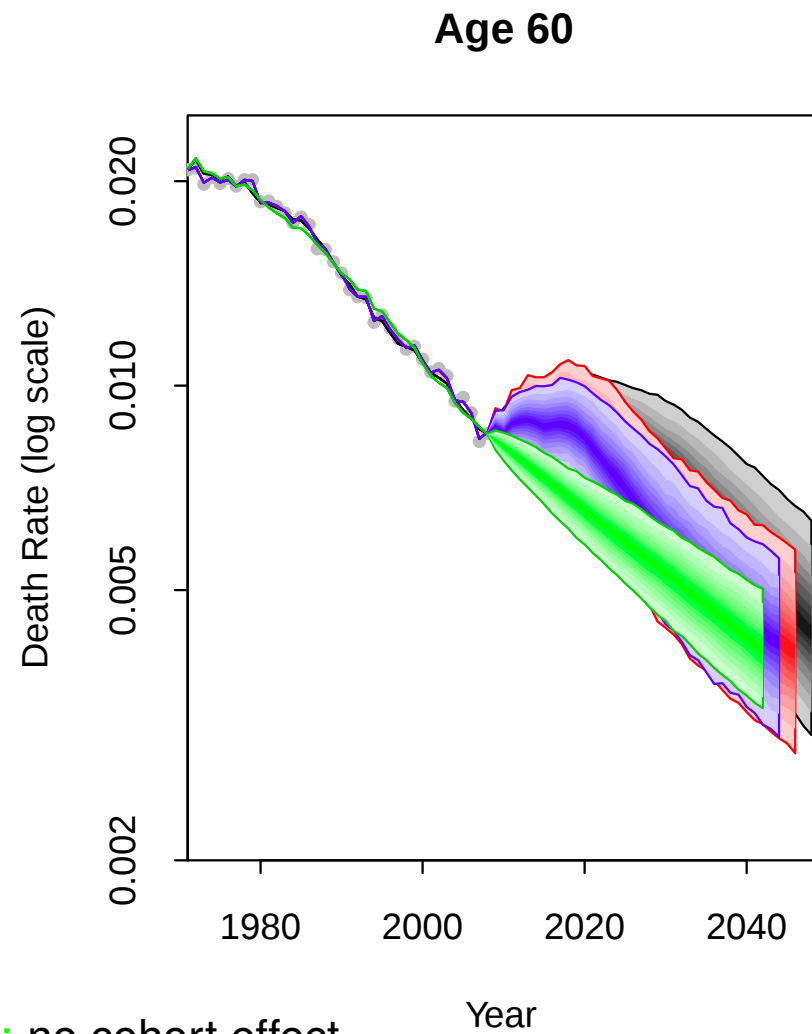
$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x}) + \gamma(t - x)$$

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x}) + R(x, t)$$

$$R(t) = AR(t - 1) + BR(t - 2) + Z(t) \text{ (} A, B \text{ as specified earlier)}$$

$$\log m(x, t) = \beta(x) + \kappa_1(t) + \kappa_2(t)(x - \bar{x}) + R(x, t)$$

$$R(t) = AR(t - 1) + BR(t - 2) + Z(t) \text{ (simplified } A, B)$$

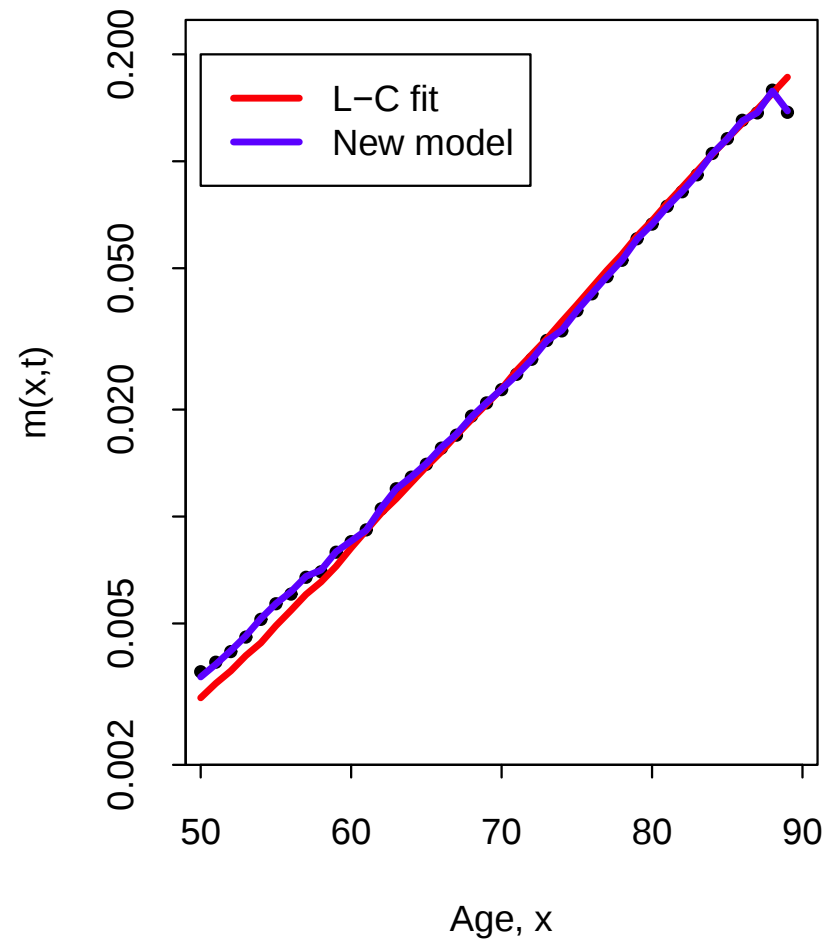


Green: no cohort effect

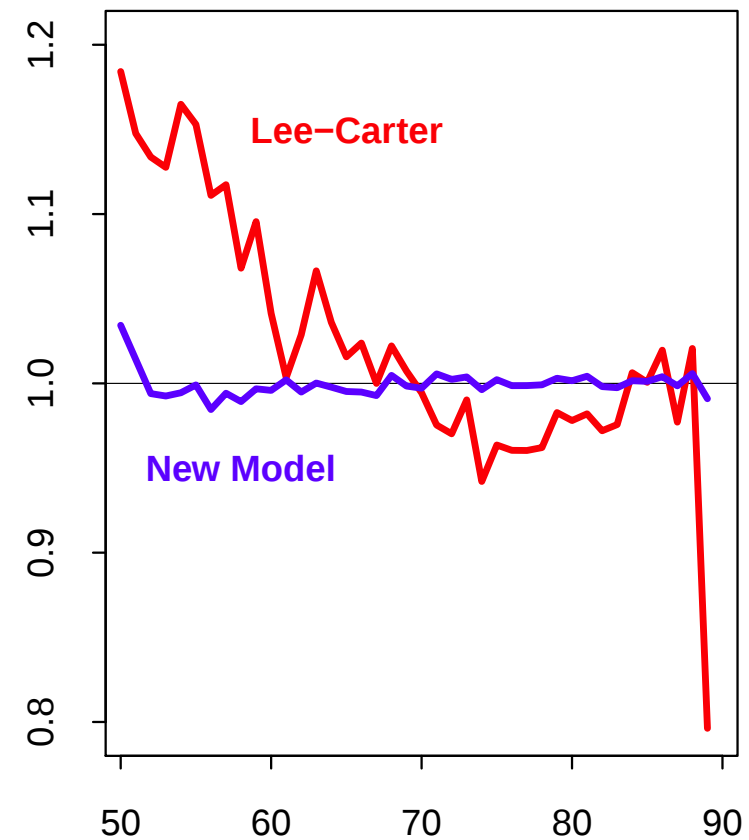
Red, Blue, Gray: similar trends but also differences

Base Table Accuracy

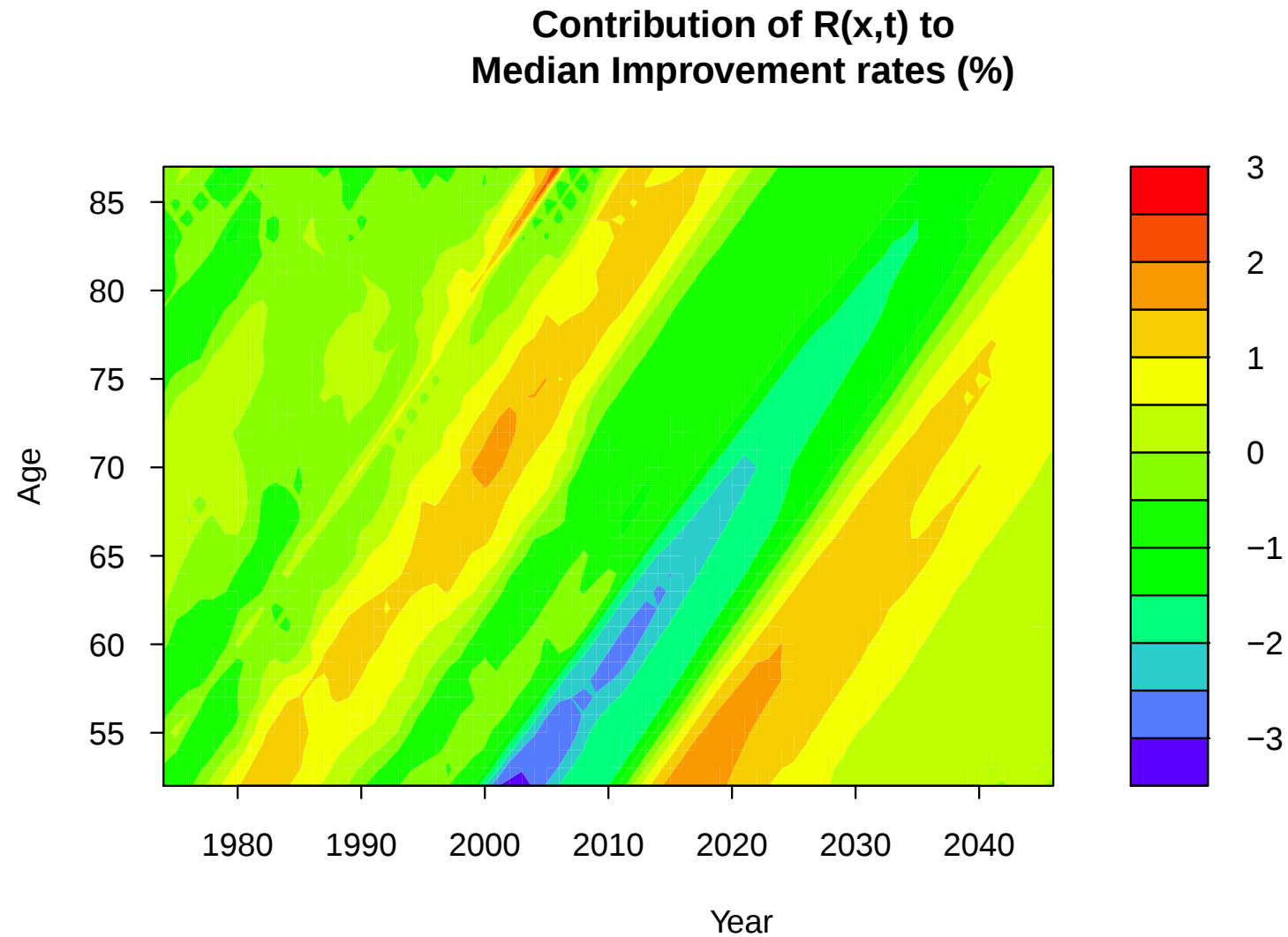
EW males 1971–2008: two models
 $m(x, 2008)$ (log scale)



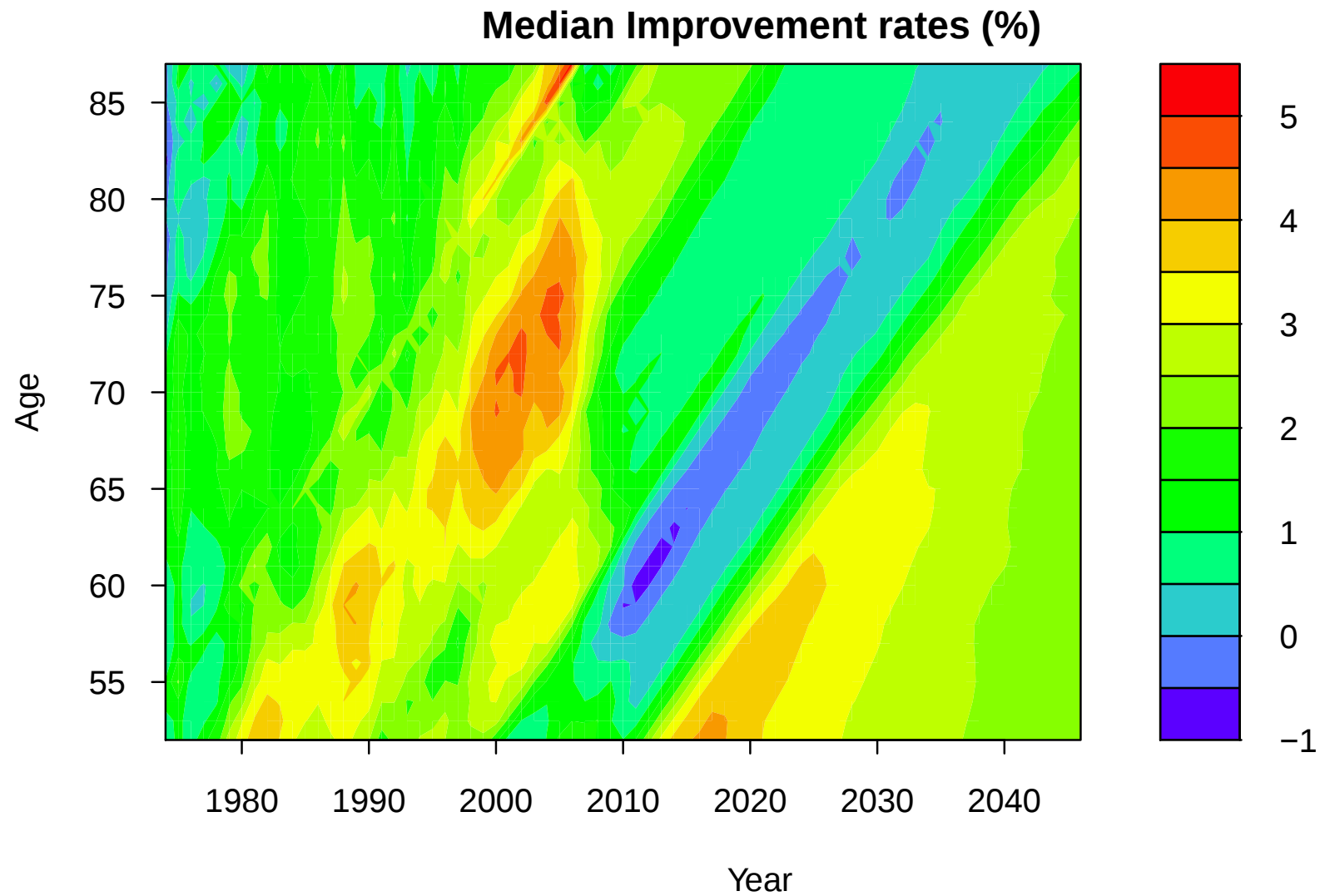
$m_c(x, t) / \hat{m}(x, t)$



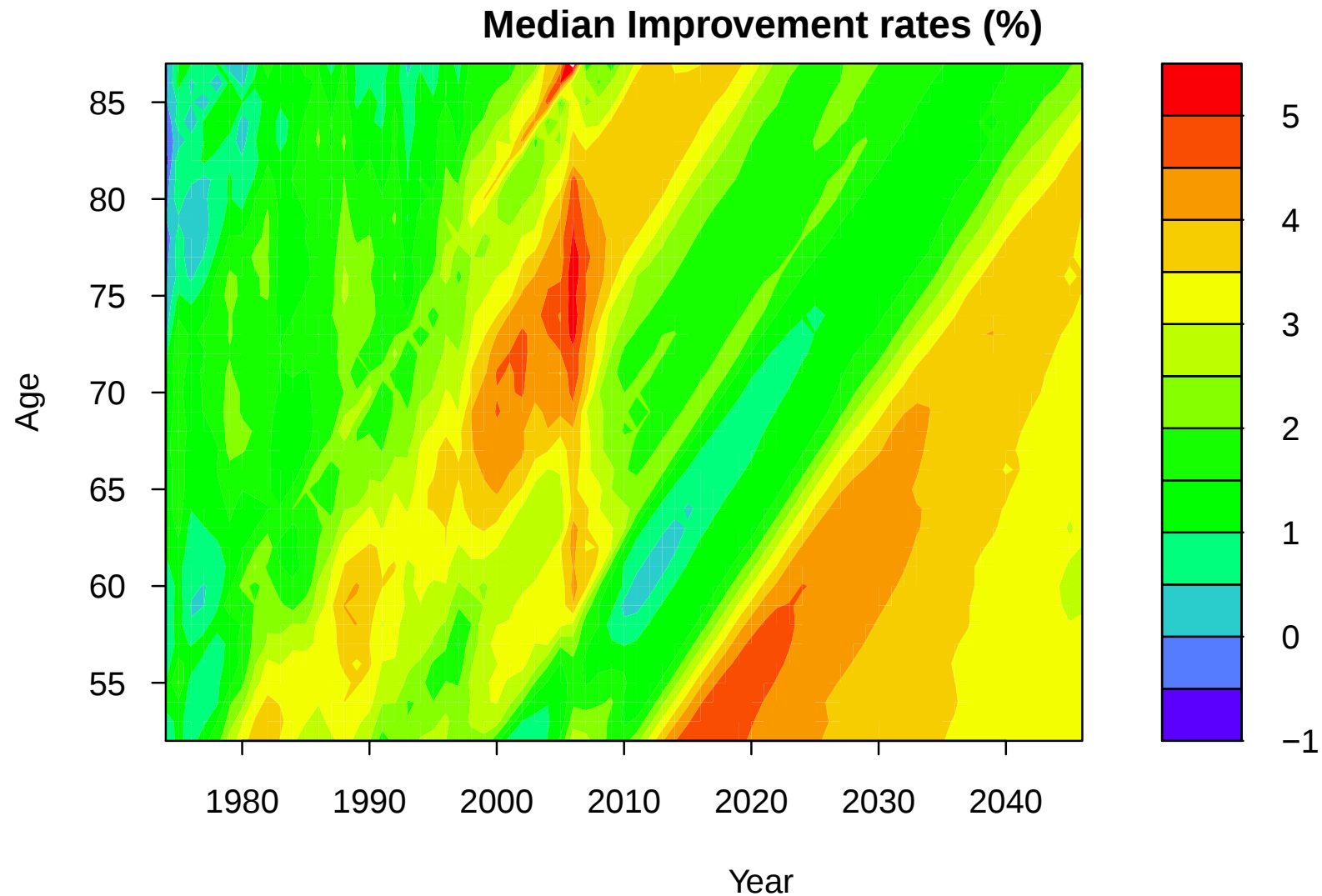
Contribution of $R(x, t)$ to improvement rates



Median total improvement rates



Median total improvement rates: Adjusted $\kappa(t)$



Multiple populations: some thoughts

- Aim for a parsimonious structure
- Items to deal with:
 - Population, P , specific $\kappa_i^{(P)}(t)$
 - Population, P , specific $R^{(P)}(x, t)$

Multiple populations: possible structures

Mortality – version 1:

- Population, P , specific $\kappa_i^{(P)}(t)$ correlated
- $R^{(P)}(x, t)$: assume independent

Mortality – version 2:

- All populations have the same $\kappa_i(t)$
- $R^{(P)}(x, t)$: assume independent
- Greater role for $R^{(P)}(x, t)$ as country specific effect

Conclusions

- New model
 - focus on a small number of core period effects
 - adds alternative $R(x, t)$ to popular cohort effects, $\gamma(t - x)$
- Model risk more evident in the mean reverting $R(x, t)$
- But general framework should prove to be more robust:
long term underlying trends ($\kappa(t)$) are reasonably consistent

Further work

- Bayesian parameter uncertainty
- Multiple populations: focus on underlying $\kappa(t)$
 $\Rightarrow$ less complexity

Questions

W: www.ma.hw.ac.uk/~andrewc

E: A.J.G.Cairns@hw.ac.uk