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Bayesian Inference for Small Population Longevity Risk Modelling

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Introduction

Model: For death counts $D(t, x)$ and Exposure $E(t, x)$:

$$\text{logit } q(\theta, t, x) = \kappa_t^{(1)} + \kappa_t^{(2)}(x - \bar{x}) + \kappa_t^{(3)}((x - \bar{x})^2 - \hat{\sigma}_x^2) + \gamma_c^{(4)}$$

$$D(t, x) | \theta \sim \text{Poi}(m(\theta_1, t, x)E(t, x))$$

$$m(\theta, t, x) = -\log[1 - q(\theta, t, x)]$$

Background: For small population, modelling with Two Stage approach (fit time series for the maximum likelihood (MLE) estimates $\hat{\theta}$) leads to biased estimates of volatility.

- Large sampling variation -> significant noise for latent parameter estimation (Cairns, Blake, Dowd et al. 2011).
- Non-negligible bias to the parameter estimation of the projecting model (Chen, Cairns and Kleinow 2015).
- Over fits the short cohorts with only one observation (Cairns et al. 2009)

Motivation: Bayesian approach offers a way to reduce such bias by

- Combining Poisson and time series likelihood
- Using knowledge of larger population to choose more informative prior
- Balancing the short cohorts with time series likelihood

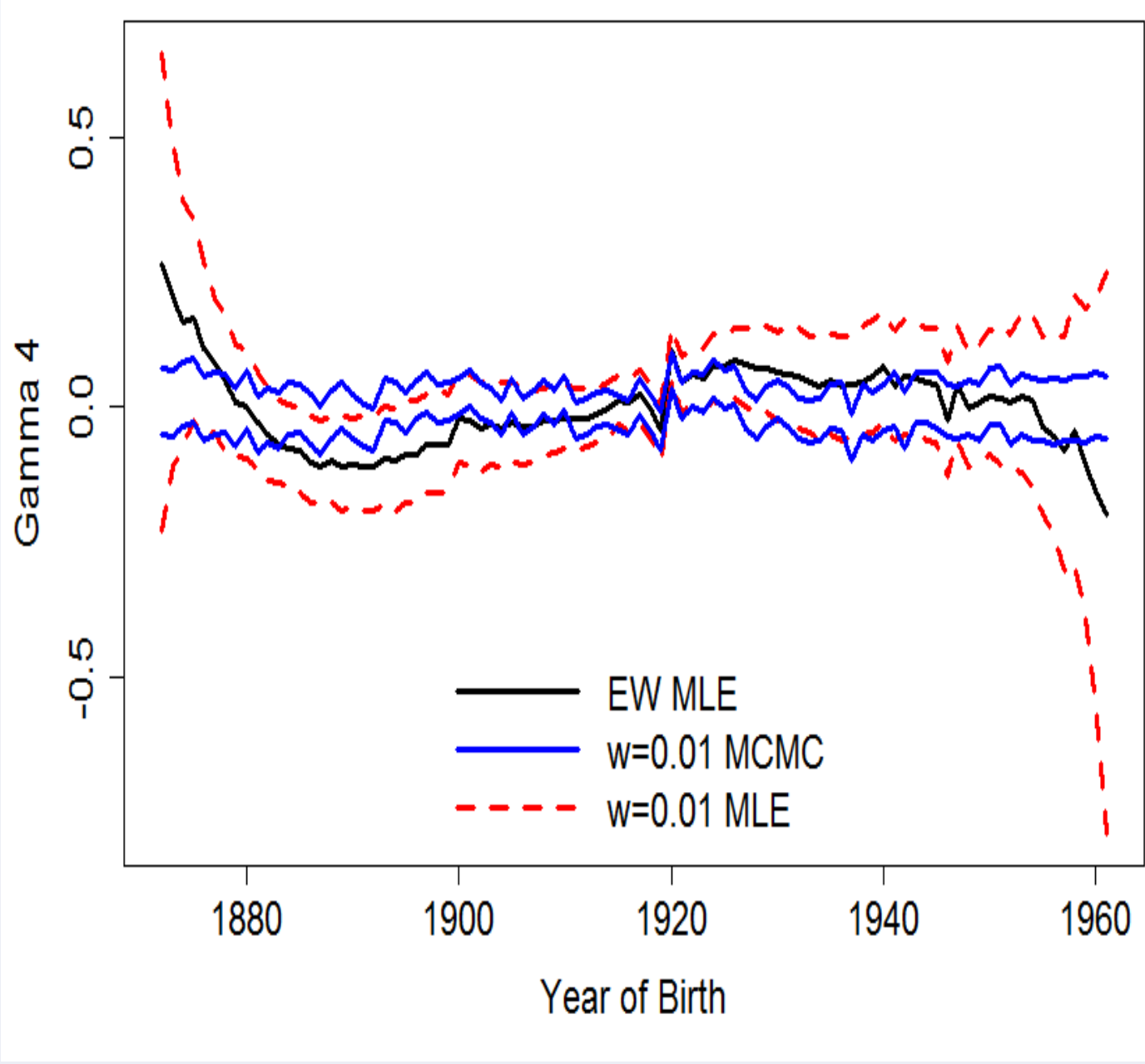
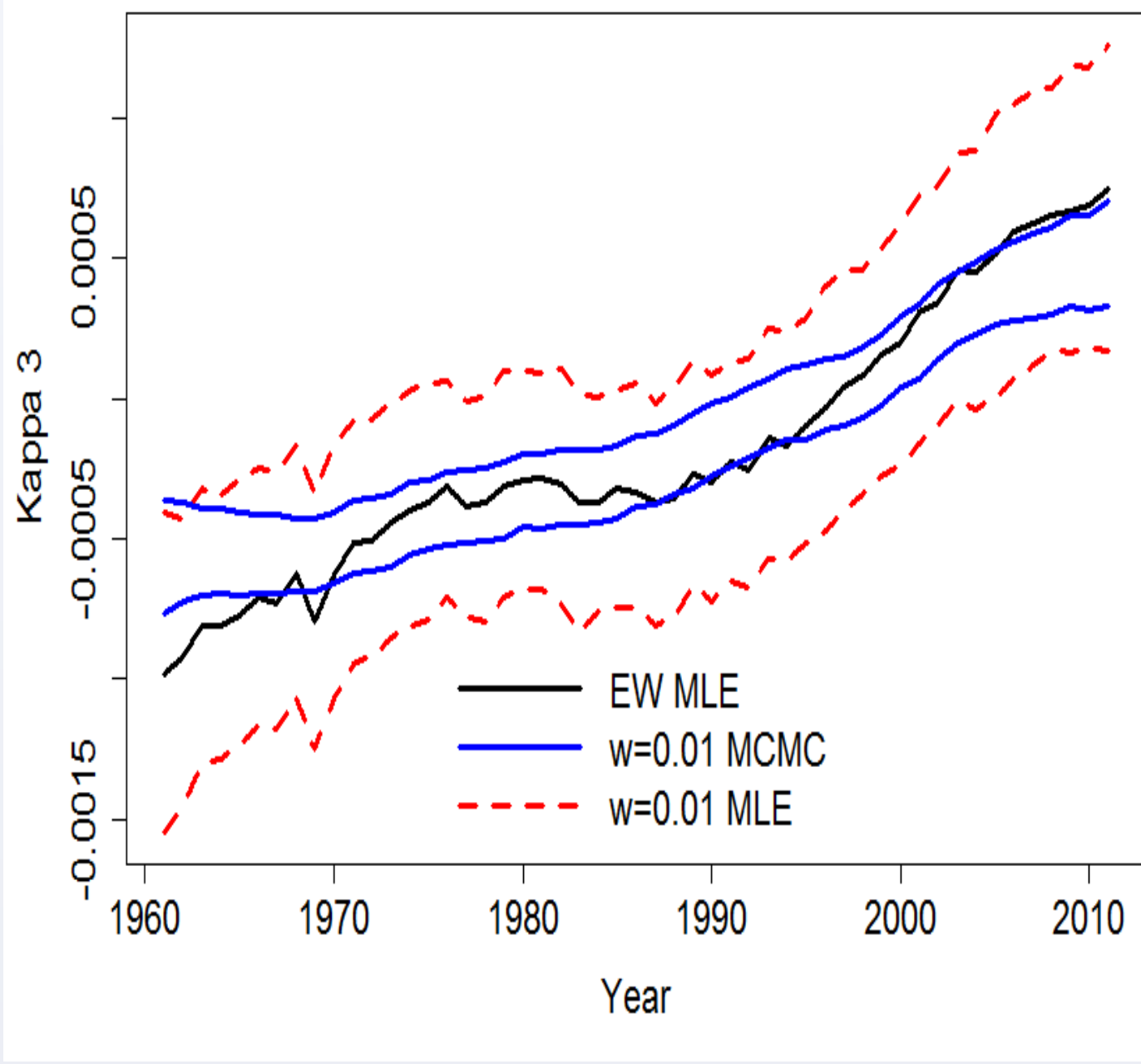
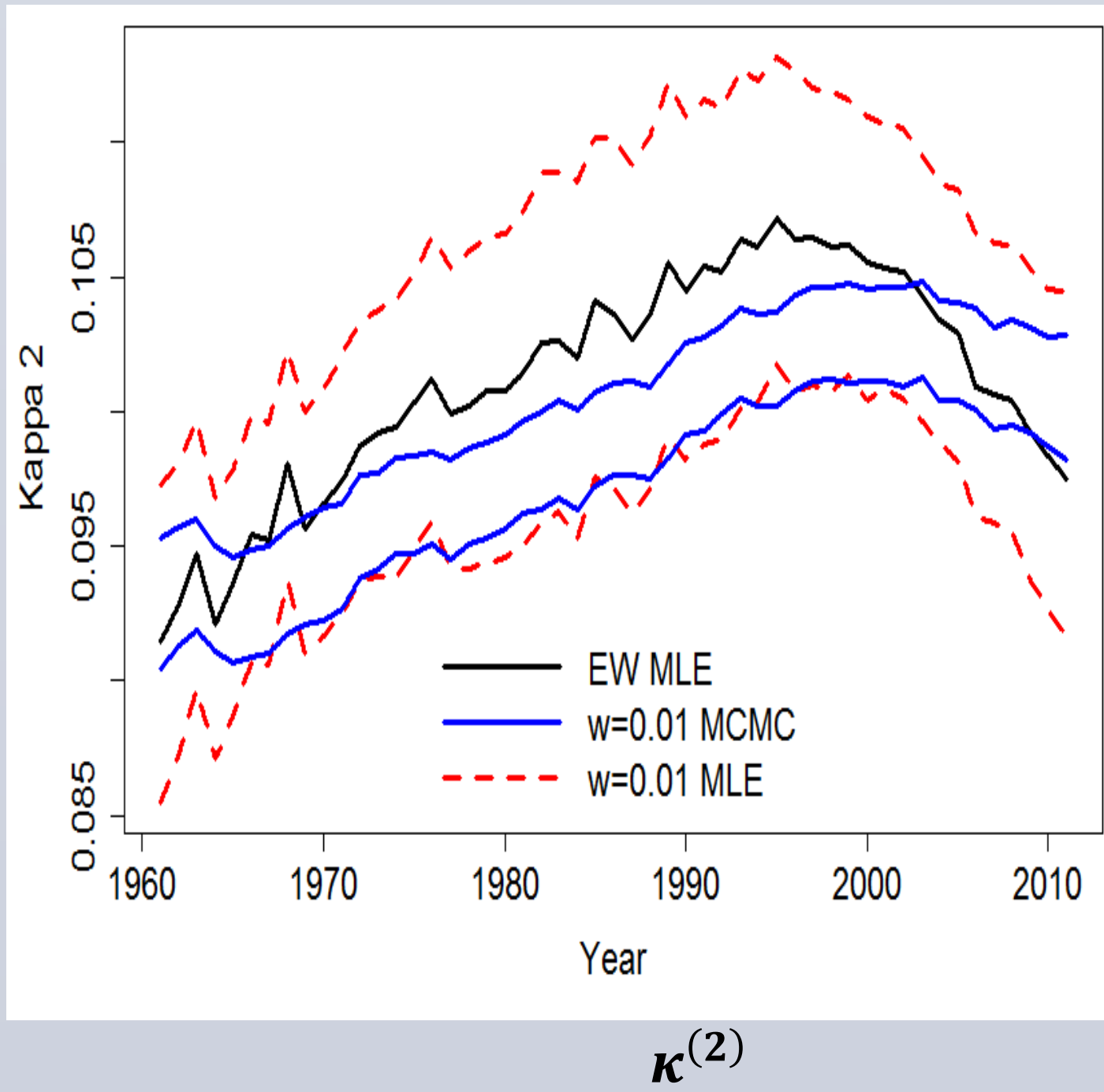
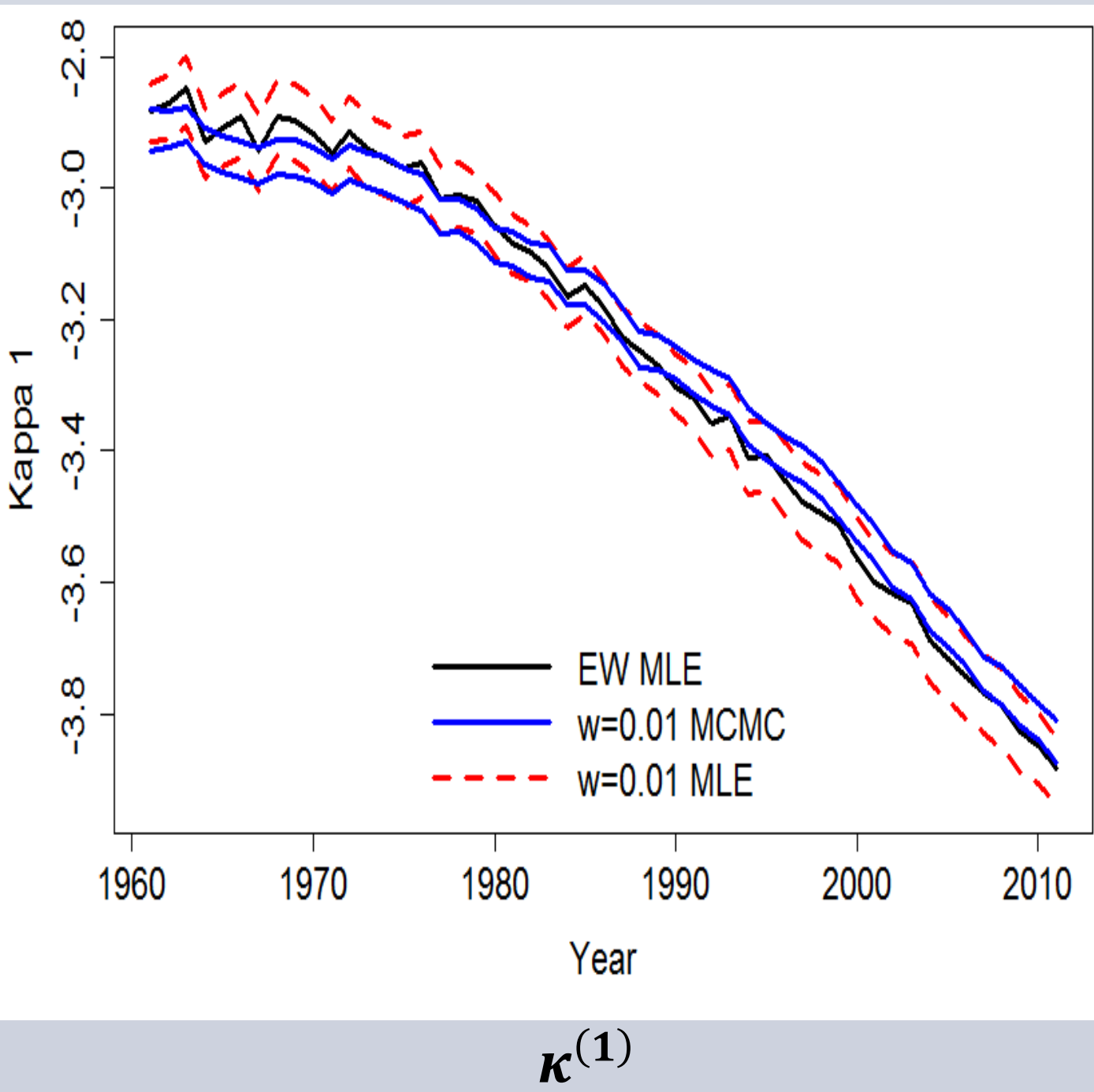
Data and Bootstrap Simulation

- Benchmark exposure $E_0(t, x)$ and corresponding death counts $D_0(t, x)$: Males in England and Wales (EW) from HMD, during 1961-2011, age 50-89 last birthday.
- $\hat{\theta}_0$: MLE for England and Wales
- $m(\hat{\theta}, t, x)$: fitted death rates given $\hat{\theta}_0$
- Simulate $D_w(t, x) | \hat{\theta}_0 \sim \text{Poi}(m(\hat{\theta}, t, x)wE_0(t, x))$, for $w = 0.01$
- $\hat{\theta}_w$: MLE for $D_w(t, x)$

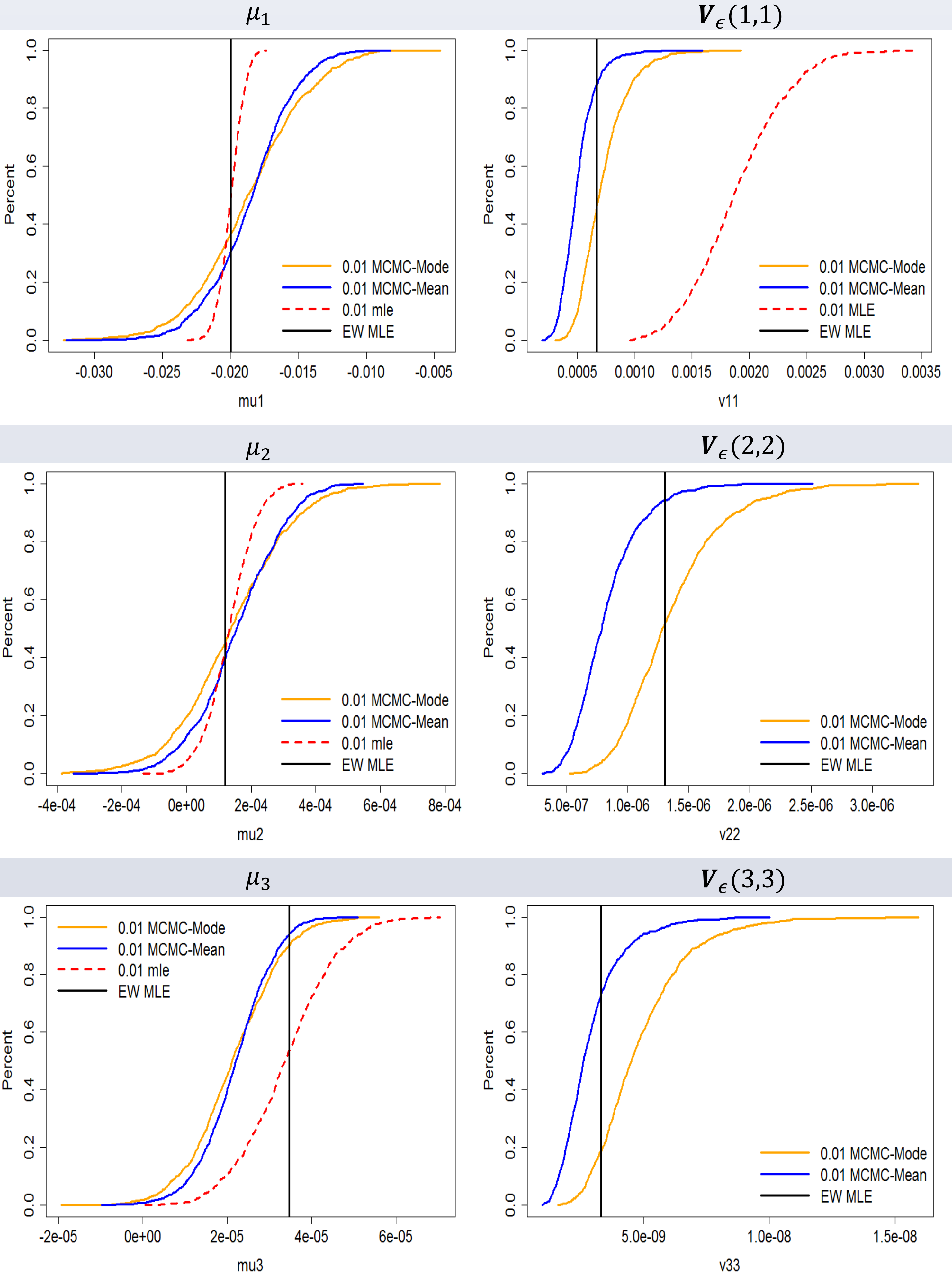
Prior Distributions for Latent and Hyper Parameters

- $\kappa_t = (\kappa_t^{(1)}, \kappa_t^{(2)}, \kappa_t^{(3)})^T$
- $\kappa_1 \propto$ Uniform Distribution
- $\kappa_t | \kappa_{t-1} \sim \text{MVN}(\mu + \kappa_{t-1}, V_\epsilon)$, for $t > 1$
- $\mu = (\mu_1, \mu_2, \mu_3)^T \propto$ Uniform Distribution
- $V_\epsilon \propto$ Inverse Wishart (ν, Σ), independent of t
- MCMC-Mean: Fix the mean of the prior for V_ϵ to the respected England and Wales' estimation
- MCMC-Mode: Fix the mode of the prior for V_ϵ to the respected England and Wales' estimation
- $\gamma_c^{(4)} | \gamma_{c-1}^{(4)} \sim N(\alpha_\gamma \gamma_{c-1}^{(4)}, \sigma_\gamma^2)$, for $c > 2$
- $\gamma_1^{(4)} \sim N(0, \frac{\sigma_\gamma^2}{1 - \alpha_\gamma^2})$
- $\alpha_\gamma \propto (1 - \alpha_\gamma^2)^g$ for $|\alpha| < 1$
- $\sigma_\gamma^2 \sim$ Inverse Gamma(a_γ, b_γ)

Credibility Interval for κ and γ



Posterior Distribution for μ and V_ϵ



Note: The MLEs of $V_\epsilon(2,2)$ and $V_\epsilon(3,3)$ for $w=0.01$ are also biased to the right of the true value (vertical line). We exclude them for a clear view of the posterior distribution.

Conclusions for Fitting Small Population Modelling

- The co-variance matrix estimated by MLE is significantly biased to the right of the assumed true value due to the Poisson model's over fitting.
- We combine the two stages into one by adding time series likelihood for the latent parameters and gained the posterior distribution with the MCMC procedure.
- The Bayesian method provides an improved fit to the hyper parameter V_ϵ .
- The low level information involved in short cohorts is balanced by the time series prior.
- The posterior distribution for small population is sensitive and fixing the mode of the prior for the co-variance matrix to the assumed true rates provides approximately unbiased fit to V_ϵ