## Minimum Reversion in Multivariate Time Series Application to Human Mortality Data

Torsten Kleinow<br>joint work with Michel Vellekoop

Heriot-Watt University, Edinburgh

Actuarial Research Centre, IFoA

Liverpool - 27 Nov 2018

## Fitted log mortality rates at age 70



Aim: projection of mortality for all populations simultaneously

## Scenario for projected log mortality rates at age 70



Projections based on multivariate random walk with common drift

## Scenario for projected log mortality rates at age 70



Projections based on our model

## Introduction

- There are a number of models for the mortality experience in multiple populations available
- Such models have typically population specific period effects in addition to some common period or age effects


## Introduction

- There are a number of models for the mortality experience in multiple populations available
- Such models have typically population specific period effects in addition to some common period or age effects
- Focus of this talk is on a model for projecting mortality rates and generating mortality scenarios simultaneously for many countries
- multivariate time series model for period effects

Actuarial

## Motivation

- improvements in survival probabilities are driven by similar changes that do not stop at the border
- medical innovations: reduced death rates from cardio-vascular diseases, ...
- life style factors: smoking ban, sugar tax, minimum price per unit of alcohol


## Motivation

- improvements in survival probabilities are driven by similar changes that do not stop at the border
- medical innovations: reduced death rates from cardio-vascular diseases, ...
- life style factors: smoking ban, sugar tax, minimum price per unit of alcohol
- countries will tend to copy the most successful innovations from other countries

Actuarial
Research Centre Institute and Faculty
of Actuaries

## Motivation

- improvements in survival probabilities are driven by similar changes that do not stop at the border
- medical innovations: reduced death rates from cardio-vascular diseases, ...
- life style factors: smoking ban, sugar tax, minimum price per unit of alcohol
- countries will tend to copy the most successful innovations from other countries
- if true, survival probabilities in any country will show a tendency to move towards those of the country with highest survival rates
- we include a term in our model to incorporate that tendency ...


## Motivation

- improvements in survival probabilities are driven by similar changes that do not stop at the border
- medical innovations: reduced death rates from cardio-vascular diseases, ...
- life style factors: smoking ban, sugar tax, minimum price per unit of alcohol
- countries will tend to copy the most successful innovations from other countries
- if true, survival probabilities in any country will show a tendency to move towards those of the country with highest survival rates
- we include a term in our model to incorporate that tendency ...
- ... and investigate whether such a "'learning" ' effect is signifcant


## Model

- The number of deaths, $D_{x t c}$, in population $c \in \mathcal{C}$ at age $x \in \mathcal{X}$ in calendar year $t \in \mathcal{T}$ has a Poisson distribution:

$$
D_{x t c} \sim \operatorname{Pois}\left(\mu_{x t c} E_{x t c}\right)
$$

- $\mu_{x t c}$ is the force of mortality
- $E_{x t c}$ refers to the central exposed to risk.


## Common Age Effects

- Our model for the force of mortality is a modification, Kleinow (2015), of the Lee-Carter model:

$$
\begin{equation*}
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c} \tag{1}
\end{equation*}
$$

- Common age effects, $\alpha_{x}$ and $\beta_{x}$, ensure that period effects are comparable across populations since they are all rescaled with the same (age-dependent) constant.


## Common Age Effects

- Our model for the force of mortality is a modification, Kleinow (2015), of the Lee-Carter model:

$$
\begin{equation*}
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c} \tag{1}
\end{equation*}
$$

- Common age effects, $\alpha_{x}$ and $\beta_{x}$, ensure that period effects are comparable across populations since they are all rescaled with the same (age-dependent) constant.
- The parameters in (1) are not identifiable
- impose constraints on $\alpha$ and $\beta$ :

$$
\begin{equation*}
\alpha_{x_{r}}=0 \text { and } \beta_{x_{r}}=1 \tag{2}
\end{equation*}
$$

for a fixed reference age $x_{r} \in \mathcal{X}$.
That means, fitted $\log$ mortality $\log \mu_{x t c}=\kappa_{t, c}$ for $x=x_{r}$ in every population $c \in \mathcal{C}$.

- In our empirical study we set $x_{r}=70$.
- mortality data for male populations in 20 countries: The Netherlands, Sweden, Denmark, Belgium, Finland, England \& Wales, France, Switzerland, Australia, Italy, Austria, Ireland, Norway, Japan, Canada, New Zealand, Portugal, Spain, USA, Iceland
- ages: 20-95 ( $\mathcal{X}=\{20,21, \ldots, 95\})$,
- years 1951-2011 $(\mathcal{T}=\{1951, \ldots, 2011\})$
- source: Human Mortality Database


## Common Age Effects - Empirical Results - alpha

$$
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c}
$$



## Common Age Effects - Empirical Results - beta

$$
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c}
$$



## Common Age Effects - Empirical Results - kappa

$$
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c}
$$



## Time Series Model for Period Effects

We propose the following model for the dynamics of the period effects $\kappa_{c}$ for any population $c \in \mathcal{C}$ :

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu_{c}+\zeta_{c}\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c} \quad\left(\operatorname{Corr}\left(Z_{t, c_{1}}, Z_{t, c_{2}}\right)=\rho_{c_{1}} \rho_{c_{2}}\right)
\end{aligned}
$$

where $\zeta_{c}, \rho_{c} \in(-1,1), \lambda_{c} \in[0,1), \sigma_{c}>0$ and $\left\{W_{t}, W_{t, c}\right\}_{c \in \mathcal{C}, t \in \mathcal{T}}$ are independent and identically distributed random variables with a standard normal distribution.

## Time Series Model for Period Effects

We propose the following model for the dynamics of the period effects $\kappa_{c}$ for any population $c \in \mathcal{C}$ :

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu_{c}+\zeta_{c}\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c} \quad\left(\operatorname{Corr}\left(Z_{t, c_{1}}, Z_{t, c_{2}}\right)=\rho_{c_{1}} \rho_{c_{2}}\right)
\end{aligned}
$$

where $\zeta_{c}, \rho_{c} \in(-1,1), \lambda_{c} \in[0,1), \sigma_{c}>0$ and $\left\{W_{t}, W_{t, c}\right\}_{c \in \mathcal{C}, t \in \mathcal{T}}$ are independent and identically distributed random variables with a standard normal distribution.
"Reversion" is to the minimum period effect at time $t \in \mathcal{T}$ as

$$
m_{t}:=\min _{c \in \mathcal{C}} \kappa_{t, c}
$$

## Time Series Model for Period Effects

Special case, $\mu_{c}, \zeta_{c}, \rho_{c}=0$

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c} \\
m_{t} & :=\min _{c \in \mathcal{C}} \kappa_{t, c}
\end{aligned}
$$

with $Z_{t, c}$ iid normal.

## Time Series Model for Period Effects

Special case, $\mu_{c}, \zeta_{c}, \rho_{c}=0$

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c} \\
m_{t} & :=\min _{c \in \mathcal{C}} \kappa_{t, c}
\end{aligned}
$$

with $Z_{t, c}$ iid normal.

- The minimum process $m_{t}$, and therefore, all $\kappa_{t, c}$ processes have a downward drift (despite $\mu_{c}=0$ ):

$$
\mathbb{P}\left(m_{t+1} \leq a \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right) \kappa_{t, c}-a+\lambda_{c} m_{t}}{\sigma_{c}}\right)
$$

## Time Series Model for Period Effects

Special case, $\mu_{c}, \zeta_{c}, \rho_{c}=0$

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c} \\
m_{t} & :=\min _{c \in \mathcal{C}} \kappa_{t, c}
\end{aligned}
$$

with $Z_{t, c}$ iid normal.

- The minimum process $m_{t}$, and therefore, all $\kappa_{t, c}$ processes have a downward drift (despite $\mu_{c}=0$ ):

$$
\mathbb{P}\left(m_{t+1} \leq a \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right) \kappa_{t, c}-a+\lambda_{c} m_{t}}{\sigma_{c}}\right)
$$

- Setting $a=m_{t}$ we obtain

$$
\mathbb{P}\left(m_{t+1} \leq m_{t} \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right)\left(\kappa_{t, c}-m_{t}\right)}{\sigma_{c}}\right)>\frac{1}{2}
$$

$\Phi$ is the $N(0,1)$ distribution function.


## Time Series Model for Period Effects

- Conditional probability for the minimum to decrease

$$
\mathbb{P}\left(m_{t+1} \leq m_{t} \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right)\left(\kappa_{t, c}-m_{t}\right)}{\sigma_{c}}\right)>\frac{1}{2}
$$

$\Phi$ is the $N(0,1)$ distribution function.

- The larger $\sigma_{c}$ the larger the probability of $m_{t}$ decreasing


## Time Series Model for Period Effects

- Conditional probability for the minimum to decrease

$$
\mathbb{P}\left(m_{t+1} \leq m_{t} \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right)\left(\kappa_{t, c}-m_{t}\right)}{\sigma_{c}}\right)>\frac{1}{2}
$$

$\Phi$ is the $N(0,1)$ distribution function.

- The larger $\sigma_{c}$ the larger the probability of $m_{t}$ decreasing
- The smaller the differences between countries the larger the probability of $m_{t}$ decreasing


## Time Series Model for Period Effects

- Conditional probability for the minimum to decrease

$$
\mathbb{P}\left(m_{t+1} \leq m_{t} \mid\left\{\kappa_{t, c}\right\}_{c \in \mathcal{C}}\right)=1-\prod_{c \in \mathcal{C}} \Phi\left(\frac{\left(1-\lambda_{c}\right)\left(\kappa_{t, c}-m_{t}\right)}{\sigma_{c}}\right)>\frac{1}{2}
$$

$\Phi$ is the $N(0,1)$ distribution function.

- The larger $\sigma_{c}$ the larger the probability of $m_{t}$ decreasing
- The smaller the differences between countries the larger the probability of $m_{t}$ decreasing
- The more countries the larger the probability of $m_{t}$ decreasing


## Time Series Model for Period Effect - Co-integration

$$
\kappa_{t+1, c}-\kappa_{t, c}=\mu_{c}+\zeta_{c}\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c}
$$

Individual components $\kappa_{t, c}$ are not stationary but they turn out to be co-integrated.

If all processes $\kappa_{t, c}$ (for all $c$ ) have a common minimum reversion parameter $\lambda$ and a common drift $\mu$ and there is no autoregressive term, so $\mu_{c}=\mu, \lambda_{c}=\lambda>0$ and $\zeta_{c}=0$ for all $c \in \mathcal{C}$, then the processes $\left\{\kappa_{\kappa, c}\right\}_{c \in \mathcal{C}}$ are co-integrated.

## Time Series Model for Period Effect - Co-integration, proof

$$
\kappa_{t+1, c}-\kappa_{t, c}=\mu_{c}+\zeta_{c}\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c}
$$

Fix a $c^{*} \in \mathcal{C}$ and define $\tilde{\kappa}_{t, c}:=\kappa_{t, c}-\kappa_{t, c^{*}}$ for any $c \in \mathcal{C}$. We then find for any $c \neq c^{*}$

$$
\begin{array}{rlr}
\tilde{\kappa}_{t, c} & =(1-\lambda)\left(\kappa_{t-1, c}-\kappa_{t-1, c^{*}}\right)+\tilde{Z}_{t} \\
& =(1-\lambda) \tilde{\kappa}_{t-1, c}+\tilde{Z}_{t}, \quad \tilde{Z}_{t}=\sigma_{c} Z_{t, c}-\sigma_{c^{*}} Z_{t, c^{*}} .
\end{array}
$$

Since $0<\lambda \leq 1$ we obtain that $\tilde{\kappa}_{t, c}$ is a stationary $\operatorname{AR}(1)$ process for all $c \neq c^{*}$.

## Time Series Model for Period Effect - Co-integration, proof

$$
\kappa_{t+1, c}-\kappa_{t, c}=\mu_{c}+\zeta_{c}\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda_{c}\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t+1, c}
$$

Fix a $c^{*} \in \mathcal{C}$ and define $\tilde{\kappa}_{t, c}:=\kappa_{t, c}-\kappa_{t, c^{*}}$ for any $c \in \mathcal{C}$. We then find for any $c \neq c^{*}$

$$
\begin{array}{rlr}
\tilde{\kappa}_{t, c} & =(1-\lambda)\left(\kappa_{t-1, c}-\kappa_{t-1, c^{*}}\right)+\tilde{Z}_{t} \\
& =(1-\lambda) \tilde{\kappa}_{t-1, c}+\tilde{Z}_{t}, & \tilde{Z}_{t}=\sigma_{c} Z_{t, c}-\sigma_{c^{*}} Z_{t, c^{*}} .
\end{array}
$$

Since $0<\lambda \leq 1$ we obtain that $\tilde{\kappa}_{t, c}$ is a stationary $\operatorname{AR}(1)$ process for all $c \neq c^{*}$.
Furthermore, we find that

$$
m_{t}=\min _{c \in \mathcal{C}}\left(\kappa_{t, c^{*}}+\tilde{\kappa}_{t, c}\right)=\kappa_{t, c^{*}}+\min _{c \in \mathcal{C}} \tilde{\kappa}_{t, c} \text { and therefore }
$$

$$
\begin{aligned}
\Delta \kappa_{t+1, c^{*}}:=\kappa_{t+1, c^{*}}-\kappa_{t, c^{*}} & =\mu_{c^{*}}+\lambda\left(m_{t}-\kappa_{t, c^{*}}\right)+\sigma_{c^{*}} Z_{t+1, c^{*}} \\
& =\lambda \min _{c \in \mathcal{C}} \tilde{\kappa}_{t, c}+\mu_{c^{*}}+\sigma_{c^{*}} Z_{t+1, c^{*}}
\end{aligned}
$$

The first term in the last expression is a minimum over stationary processes and the other terms are stationary too, hence $\Delta \kappa_{t, c^{*}}$ is stationary.

## Time Series Model - two-dimensional model

We take $C=2, \lambda_{1}=\lambda_{2}:=\lambda, \mu_{1}=\mu_{2}=\mu$ and $\operatorname{Corr}\left(Z_{t+1,1}, Z_{t+1,2}\right)=\rho$. And we assume that there is no AR term, that is, $\zeta_{1}=\zeta_{2}=0$

$$
\begin{align*}
\kappa_{t+1,1}-\kappa_{t, 1} & =\lambda\left(\min _{c \in\{1,2\}} \kappa_{t, c}-\kappa_{t, 1}\right)+\sigma_{1} Z_{t+1,1}+\mu,  \tag{3}\\
\kappa_{t+1,2}-\kappa_{t, 2} & =\lambda\left(\min _{c \in\{1,2\}} \kappa_{t, c}-\kappa_{t, 2}\right)+\sigma_{2} Z_{t+1,2}+\mu . \tag{4}
\end{align*}
$$

We define

$$
m_{t}=\min _{c \in\{1,2\}} \kappa_{t, c} \text { and } M_{t}=\max _{c \in\{1,2\}} \kappa_{t, c} .
$$

## Time Series Model - two-dimensional model

$$
\begin{align*}
& \kappa_{t+1,1}-\kappa_{t, 1}=\lambda\left(\min _{c \in\{1,2\}} \kappa_{t, c}-\kappa_{t, 1}\right)+\sigma_{1} Z_{t+1,1}+\mu  \tag{5}\\
& \kappa_{t+1,2}-\kappa_{t, 2}=\lambda\left(\min _{c \in\{1,2\}} \kappa_{t, c}-\kappa_{t, 2}\right)+\sigma_{2} Z_{t+1,2}+\mu \tag{6}
\end{align*}
$$

We then obtain

$$
\begin{equation*}
\lim _{t \rightarrow \infty} \mathbb{E}\left[M_{t+1}-M_{t}\right]=\lim _{t \rightarrow \infty} \mathbb{E}\left[m_{t+1}-m_{t}\right]=\mu-s \sqrt{\frac{\lambda}{2 \pi(2-\lambda)}} \tag{7}
\end{equation*}
$$

and

$$
\begin{equation*}
\lim _{t \rightarrow \infty} \mathbb{E}\left[M_{t}-m_{t}\right]=s \sqrt{\frac{2}{\lambda \pi(2-\lambda)}} \tag{8}
\end{equation*}
$$

with $s=\sqrt{\sigma_{1}^{2}+\sigma_{2}^{2}-2 \rho \sigma_{1} \sigma_{2}}$.

## Time Series Model - Extra drift from minimum reversion

Table: Generated drift for different minimum reversion parameters $\lambda$ according to (7) (second column) and using $10^{5}$ simulations (third to last column), $s=1$ and $\mu=0$.

| Exact |  |  |  |  |  | Simulation |
| :---: | ---: | ---: | ---: | ---: | ---: | :--- |
| $\lambda$ | $\|\mathcal{C}\|=2$ | $\|\mathcal{C}\|=2$ | $\|\mathcal{C}\|=3$ | $\|\mathcal{C}\|=4$ | $\|\mathcal{C}\|=8$ | $\|\mathcal{C}\|=16$ |
| 0.0125 | -0.0447 | -0.0448 | -0.0671 | -0.0817 | -0.1129 | -0.1401 |
| 0.025 | -0.0635 | -0.0635 | -0.0952 | -0.1158 | -0.1602 | -0.1987 |
| 0.05 | -0.0903 | -0.0903 | -0.1355 | -0.1649 | -0.2280 | -0.2828 |
| 0.1 | -0.1294 | -0.1294 | -0.1942 | -0.2362 | -0.3266 | -0.4052 |
| 0.2 | -0.1881 | -0.1881 | -0.2821 | -0.3432 | -0.4745 | -0.5887 |
| 0.4 | -0.2821 | -0.2821 | -0.4231 | -0.5147 | -0.7118 | -0.8830 |

## Time Series Model - Expected range

Table: Expectation of the stationary distribution for $M_{t}-m_{t}$ for different minimum reversion parameters $\lambda$ according to (8) (second column) and using $10^{5}$ simulations (third to last column).

|  | Exact | Simulation |  |  |  |  |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\lambda$ | $\|\mathcal{C}\|=2$ | $\|\mathcal{C}\|=2$ | $\|\mathcal{C}\|=3$ | $\|\mathcal{C}\|=4$ | $\|\mathcal{C}\|=8$ | $\|\mathcal{C}\|=16$ |
| 0.0125 | 7.1589 | 7.1593 | 10.7386 | 13.0627 | 18.0644 | 22.4106 |
| 0.025 | 5.0781 | 5.0774 | 7.6185 | 9.2656 | 12.8138 | 15.8962 |
| 0.05 | 3.6137 | 3.6132 | 5.4202 | 6.5931 | 9.1178 | 11.3112 |
| 0.1 | 2.5887 | 2.5886 | 3.8831 | 4.7229 | 6.5316 | 8.1031 |
| 0.2 | 1.8806 | 1.8806 | 2.8209 | 3.4313 | 4.7454 | 5.8866 |
| 0.4 | 1.4105 | 1.4104 | 2.1156 | 2.5735 | 3.5591 | 4.415 |

## Fitting the model to kappa - just a reminder

$$
\log \mu_{x t c}=\alpha_{x}+\beta_{x} \kappa_{t, c}
$$



## Summary Statistics

$$
\begin{aligned}
& \kappa_{t+1, c}-\kappa_{t, c}=\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
& Z_{t, c}=\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c} \\
& \text { Drift + AR }
\end{aligned}
$$

## Summary Statistics

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c}
\end{aligned}
$$

Drift + AR
$K V+A R$

| $\log L$ | $K$ | BIC | $\mu$ | $\lambda$ | $\zeta$ | $\bar{\sigma}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3785.59 | 42 | -7274.10 | -0.0202 | - | -0.3289 | 0.0334 |
| 3782.30 | 42 | -7267.52 | - | 0.0355 | -0.3263 | 0.0354 |

## Summary Statistics

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c}
\end{aligned}
$$

Drift + AR<br>KV + AR<br>KV + Drift

| $\log L$ | $K$ | BIC | $\mu$ | $\lambda$ | $\zeta$ | $\bar{\sigma}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3785.59 | 42 | -7274.10 | -0.0202 | - | -0.3289 | 0.0334 |
| 3782.30 | 42 | -7267.52 | - | 0.0355 | -0.3263 | 0.0354 |
| 3735.49 | 42 | -7173.91 | -0.0095 | 0.0187 | - | 0.0347 |

## Summary Statistics

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c}
\end{aligned}
$$

|  | $\log L$ | $K$ | BIC | $\mu$ | $\lambda$ | $\zeta$ | $\bar{\sigma}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Drift + AR | 3785.59 | 42 | -7274.10 | -0.0202 | - | -0.3289 | 0.0334 |
| KV + AR | 3782.30 | 42 | -7267.52 | - | 0.0355 | -0.3263 | 0.0354 |
| KV + Drift | 3735.49 | 42 | -7173.91 | -0.0095 | 0.0187 | - | 0.0347 |
| KV + Drift + AR | 3792.99 | 43 | -7281.83 | -0.0144 | 0.0201 | -0.3290 | 0.0334 |

Summary statistics for males aged 20-95 in years 1951-2011, CAE. $K$ is the number of parameters.

## Projections - Random Walk



Actuarial
Research Centre Institute and Faculty
of Actuarios

## Projections - Random Walk + AR



Actuarial
Research Centre Institute and Faculty
of Actuarios

## Projections - Random Walk + AR + KV



## Projections - Random Walk



## Projections - Random Walk + AR



## Projections - Random Walk + AR + KV



## Different dataset

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c}
\end{aligned}
$$

| Drift + AR | 3605.35 | 30 | -6996.87 | -0.0234 | - | -0.3447 | 0.0425 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| KV + AR | 3566.37 | 30 | -6918.91 | - | 0.0674 | -0.2664 | 0.0439 |
| KV + Drift | 3541.84 | 30 | -6869.84 | -0.0121 | 0.0391 | - | 0.0454 |
| KV + Drift + AR | 3610.95 | 31 | -7000.94 | -0.0194 | 0.0257 | -0.3321 | 0.0422 |

Summary statistics for females aged 20-95 in years 1921-2011, CAE, 14 countries

## Projections - Random Walk

Female population in 14 countries, aged 20-95, years 1921-2011


## Projections - Random Walk + AR

Female population in 14 countries, aged 20-95, years 1921-2011


## Projections - Random Walk + AR + KV

Female population in 14 countries, aged 20-95, years 1921-2011


## Conclusions

$$
\begin{aligned}
\kappa_{t+1, c}-\kappa_{t, c} & =\mu+\zeta\left(\kappa_{t, c}-\kappa_{t-1, c}\right)+\lambda\left(m_{t}-\kappa_{t, c}\right)+\sigma_{c} Z_{t, c} \\
Z_{t, c} & =\rho_{c} W_{t}+\sqrt{1-\rho_{c}^{2}} W_{t, c}
\end{aligned}
$$

- simultaneous projections of mortality in multiple populations
- changing improvement rates
- downward trend generated from random innovations
- learning (copying) from other populations
- better fit than model without KV term (for many datasets)
- improved scenario generation
paper on arXiv:
Kleinow, Vellekoop: Minimum reversion in multivariate time series

