SKELETON TABLES DERIVED FROM THE EXPERIENCE OF THE CONTINUOUS MORTALITY INVESTIGATION

By W. H. CHAPMAN, F.I.A.,

Assistant Actuary of the Gresham Life Assurance Society, Ltd.

THE method adopted in the calculation of the tables appended to this note was essentially that of the abridged mortality tables described by George King in J.I.A. Vol. XLVIII, pp. 294–300. Special formulae were used to suit the form in which the statistics were presented and to deal with the select data of durations 1–4.

FORMULAE FOR CENTRAL TERMS

King's first step was to obtain values of q_x at central ages of quinquennial groups by applying the formula $u_0 = 2w_{-2} - \cos \Delta^2 w_{-7}$ to both deaths and exposed to risk. In this formula $w_x = \sum_{0}^{4} u_{x+t}$ and differences are taken quinquennially. It is the particular case, for n=5, of the general formula for the central term of three consecutive groups of n terms each:

$$u_0 = \frac{1}{n} w_{-\frac{1}{2}(n-1)} - \frac{n^2 - 1}{24n^3} \Delta^2 w_{-\frac{1}{2}(3n-1)},$$

where $w_x = \sum_{t=0}^{n-1} u_{x+t}$ and differences are taken over successive *n*th values of w. The assumption underlying the formula is that u_x is of the form

$$a+bx+cx^2+dx^3$$

for 3n consecutive values of x.

The corresponding formula for the central term of four groups is

$$u_0 = \frac{1}{2n} \left(w_{\frac{1}{2}-n} + w_{\frac{1}{2}} \right) - \frac{4n^2 - 1}{48n^3} \left(\Delta^2 w_{\frac{1}{2}-2n} + \Delta^2 w_{\frac{1}{2}-n} \right),$$

or, when n=5,

$$u_0 = 1 \left(w_{-4\frac{1}{2}} + w_{\frac{1}{2}} \right) - 0165 \left(\Delta^2 w_{-9\frac{1}{2}} + \Delta^2 w_{-4\frac{1}{2}} \right),$$

which may also be expressed in the form

$$u_0 = \frac{1}{60} \left\{ 7 \left(w_{-4\frac{1}{2}} + w_{\frac{1}{2}} \right) - \left(w_{-9\frac{1}{2}} + w_{5\frac{1}{2}} \right) \right\} + \frac{1}{6000} \left(\Delta^2 w_{-9\frac{1}{2}} + \Delta^2 w_{-4\frac{1}{2}} \right).$$

If the last term is neglected as being insignificant and the fraction $\frac{1}{60}$ be ignored (since it will cancel out in obtaining q), the operation may be stated in summation form as 8 [2]-[4]. This has the advantage of avoiding differences of changing sign and is fairly simple to use. It was applied to the data for duration 0 and durations 5 and over respectively, summed in quinquennial age-groups $10\frac{1}{2}-14\frac{1}{2}$, $15\frac{1}{2}-19\frac{1}{2}$, $20\frac{1}{2}-24\frac{1}{2}$, etc., and gave q's at the quinquennial points 20, 25, etc.

King's formula for the central term of three groups was applied in summation form, viz. 28 [1]-[3], to the data for durations 1-4 added together. This produced q's at ages $22\frac{1}{2}$, $27\frac{1}{2}$, etc., which were assumed for the purposes of this investigation to relate to an average duration of $2\frac{1}{2}$ and therefore to ages 20, 25, etc., at entry.

ULTIMATE TABLES

The ultimate tables were completed on King's plan, obtaining successively:

- (1) $\log p_x$.
- (2) $\log_5 p_x$ by the formula $w_5 = 5u_0 + 7\Delta u_0 + 1.6\Delta^2 u_0 2\Delta^3 u_0$. Differences were avoided, however, by transforming the formula into

$$w_5 = -2u_0 + 32u_5 + 22u_{10} - 2u_{15}.$$

For the earliest age King's formula was

$$w_0 = 5u_0 + 2\Delta u_0 - 4\Delta^2 u_0 + 2\Delta^3 u_0,$$

but it is sufficient at this end of the table to stop at first differences, taking

$$w_0 = 3u_0 + 2u_5$$
.

At ages 95 and 100, the graduated values of $\log p_x$ and $\log p_x$ from the A 1924-29 table were used.

- (3) $\log l_x$ by summing $\log 5p_x$, taking $\log l_{20}$ as 5.00000. At this stage $x \log v$ at 2% was introduced to give $\log D_x$.
 - (4) l_x and D_x .
 - (5) $\sum_{1}^{5} l_{x+t}$ and $\sum_{1}^{5} D_{x+t}$ by the formula

$$w_6 = 5u_0 + 8\Delta u_0 + 2.6\Delta^2 u_0 - 2\Delta^3 u_0$$

which was transformed into

$$w_6 = -2u_0 + 2\cdot 2u_5 + 3\cdot 2u_{10} - 2u_{15}.$$

For the earliest age

$$w_1 = 2u_0 + 3u_5$$

was used instead of King's

$$w_1 = 5u_0 + 3\Delta u_0 - 4\Delta^2 u_0 + 2\Delta^3 u_0$$
.

At age 100, $\sum_{1}^{5} l_{x+t}$ was taken as $1 \cdot 175 l_{100}$, where $1 \cdot 175 = \frac{\sum_{1}^{5} l_{100+t}}{l_{100}}$ by the A 1924-29 ultimate table. Similarly $\sum_{1}^{5} D_{95+t}$ was taken as $1 \cdot 55 D_{95}$.

- (6) $\sum_{1}^{\omega-x} l_{x+t}$ and N_{x+1} .
- (7) e_x and a_x .

SELECT TABLES

For the select portions of the tables the steps were as follows:

- (1) $\log p_{[x]}$ and $\log p_{[x]+2\frac{1}{2}}$.
- (2) $\log l_{(x)} = \log l_{x+5} (\log p_{(x)} + 4 \log p_{(x)+2\frac{1}{2}}).$
- (3) $\log l_{[x]+2\frac{1}{2}} = \log l_{x+5} 2.5 \log p_{[x]+2\frac{1}{2}}$.
- (4) $l_{[x]}$, $l_{[x]+2\frac{1}{2}}$ and l_{x+5} .
- (5) $\sum_{1}^{5} l_{[x]+t}$ from (4) by the formula

$$w_1 = \cdot 4u_0 + 3 \cdot 2u_{2\frac{1}{2}} + 1 \cdot 4u_5.$$

- (6) $\sum_{1}^{\omega-x} l_{[x]+t} = \sum_{1}^{5} l_{[x]+t} + \sum_{6}^{\omega-x} l_{x+t}$.
- (7) $e_{[x]}$.

The above was the process actually used in the construction of the tables, values of $\log v^x$ being introduced at appropriate points in order to proceed to monetary functions.

The following alternative for steps (3), (4) and (5) was later tested, but was found to make very little difference to the final results:

- (3) $\log l_{[x]+1} = \log l_{[x]} + \log p_{[x]}$.
- (4) $l_{[x]}$, $l_{[x]+1}$ and l_{x+5} .
- (5) $\sum_{1}^{5} l_{|x|+t}$ from (4) by the formula

$$w_1 = -2u_0 + 5u_1 + 2u_5.$$

APPLICATION TO THE EXPERIENCE OF THE CONTINUOUS MORTALITY INVESTIGATION

The method outlined above was applied to the 1924-38 statistics of the Continuous Mortality Investigation in the following sections:

- (1) Whole-life with profits (medical and non-medical combined), 1924-38.
- (2) All classes, 1924-38.
- (3) All classes, 1934-38.

It is not possible to make a direct comparison of the results obtained with the graduated A 1924–29 table, either for the select or ultimate table, since the A 1924–29 ultimate table was constructed from the data of durations 3 and over, while the select portion, in addition to being restricted to the first three years' duration, was based on the statistics of 1927–29 only.

An attempt was therefore made to apply the method to the A 1924–29 data. Two ultimate tables were constructed, one from the data for durations 5 and over for comparison with the 1924–38 experiences and the other for durations 3 and over to compare with the graduated table. The small differences in the latter case are probably attributable to the overstatement of mortality in the graduated table which is apparent from Table VII on p. xxii of the Introduction to Vol. 1 of the *Monetary Tables*.

In calculating the select portion of the table the deaths and exposed for durations 1 and 2, which contained only 3 years' material, were doubled as a correction for the effect of aggregating them with those of durations 3 and 4. A further table was constructed without doubling the data for durations 1 and 2, by way of control, but the differences were insignificant, as will be seen from the extracts given below.

Age	Ţ	Iltimate Tab	le	Select Table Values of $e_{[x]}$			
		Values of e_x					
	A 1924-29 graduated	A 1924–29 data by abridged method		A 1924-29	A 1924-29 data by abridged method (5-year select period)		
	table (durations 3 and over)	(durations 3 and over)	(durations 5 and over)	graduated table	(durations 1 and 2 doubled)	(durations 1 and 2 flat)	
(1)	(2)	(3)	(4)	(5)	(6)	(7)	
20	49.23	49.24	49.08	49.29	49.19	49.19	
30	40.30	40.30	40.30	40.32	40.38	40.59	
40	31.34	31.34	31.52	31.42	31.44	31.43	
50	22.76	22.77	22.72	22.89	22.92	22.91	
60	15.00	15.01	14.98	15.53	15.30	15.29	
70 80	8·81 4·73	8·82 4·73	8·82 4·73				

The differences between the figures shown in cols. (3) and (4) and between those in cols. (5) and (6) are to be expected if selection persists into durations 3 and 4. The relative weights of the data for durations 3 and 4 in the different age groups are shown in the following table where the A 1924–29 exposed to risk for durations 3 and 4 are given as percentages of the total exposed for durations 3 and over.

Age-group	Percentage		
$ \begin{array}{r} 15\frac{1}{2}-24\frac{1}{2} \\ 25\frac{1}{2}-34\frac{1}{2} \\ 35\frac{1}{2}-44\frac{1}{2} \end{array} $	56 29 15		
$45\frac{1}{2}$ $-54\frac{1}{2}$ $55\frac{1}{2}$ $-64\frac{1}{2}$	7		

A comparison of col. (6) with col. (5) shows the effect of transferring this data from the ultimate part to the select part of the table. Part of the difference may however be due to the method employed and part to the fact that the graduation of the A 1924-29 select table on the whole overstated the mortality (see Table VIII on p. xxii of the Introduction to Vol. 1 of the Monetary Tables).

All comparisons have therefore been made with the abridged table constructed from the A 1924–29 data. The conclusion to be drawn is that there is a small margin in current premiums on the ground of improved mortality judged by the experience of 1934–38 or even by that of the whole 15 years 1924–38. At the present time, however, any such margin may be more than absorbed by the insufficiency of other factors.

SKELETON MORTALITY TABLES

	Select				Ultimate (durations 5 and over)					
Age x	Abridged table con- structed from A 1924-29 data	Life with profits 1924-38	All classes 1924–38	All classes 1934–38	Abridged table con- structed from A 1924–29 data	Life with profits 1924–38	All classes 1924–38	All classes 1934–38		
		100,0	00071			100,000q,				
20 25 30 35 40 45 50	145 149 175 187 234 321 468 706	156 110 121 179 235 420 564 778	145 138 134 160 225 312 474 713	138 133 126 146 207 301 449 652	262 239 251 296 401 541 769	232 226 222 253 373 547 806	224 218 222 263 358 504 747 1.182	196 188 199 239 319 461 716		
55 60 65 70 75 80	1,110 — — — —	1,147 — — — —	1,009 — — —	890 — — — —	1,963 3,242 5,340 8,485 12,748	2,076 3,281 5,259 8,332 12,655	1,903 3,163 5,228 8,332 12,644	1,849 3,015 4,948 8,137 12,480		
20	49.19	<i>e</i> [. 49•30 °	^{x1} 49·68	50.25	49.08	49·18	x 49.59	50.19		
25 30 35 40 45 50 65 70	44·79 40·28 35·81 31·44 27·10 22·92 18·92 15·30	44·87 40·31 35·74 31·22 26·90 22·71 18·90 15·41	45·22 40·68 36·14 31·66 27·31 23·09 19·11 15·51	45.73 41.15 36.58 32.01 27.62 23.39 19.43 15.77	44.68 40.20 35.71 31.27 26.94 22.72 18.70 14.98 11.66 8.82	44.73 40.21 35.65 31.15 26.79 22.59 18.62 14.98 11.72 8.90	45·13 40·60 36·05 31·56 27·17 22·92 18·88 15·14 11·78 8·92	45.65 41.07 36.49 31.95 27.51 23.22 19.17 15.42 12.03 9.09		
		2 %	a_x 2 %							
20 25 30 35 45 55 65 70	30·26 28·59 26·70 24·64 22·46 20·10 17·62 15·07 12·61	30·33 28·65 26·73 24·61 22·32 19·96 17·47 15·06 12·69	30·47 28·80 26·90 24·82 22·59 20·22 17·74 15·21 12·76	30·72 29·03 27·13 25·04 22·77 20·39 17·91 15·43	30·19 28·52 26·64 24·57 22·35 19·98 17·48 14·91 12·36 9·92 7·72	30·26 28·56 26·66 24·55 22·27 19·88 17·38 14·84 12·34 9·96 7·78	30·42 28·74 26·85 24·76 22·51 20·12 17·61 15·03 12·47 10·01 7·80	30·68 28·98 27·08 24·99 22·73 20·32 17·79 15·22 12·67 10·20 7·94		

	Sele	ct net pren	niums % at	Increase (+) or decrease (-) compared with A 1924-29 abridged table						
Age	Abridged table constructed from A 1924-29 data	Life with profits 1924–38	All classes 1924–38	All classes 1934–38	Life with profits 1924–38	All classes 1924–38	All classes 1934–38			
Whole-life Assurance, premiums throughout life										
20 30 40 50	1·238 1·649 2·302 3·410 5·387	1.231 1.645 2.327 3.453 5.344	1·217 1·623 2·278 3·375 5·307	1·192 1·594 2·246 3·327 5·208		·021 ·026 ·024 ·035 ·080	046 055 056 083 179			
\					ed premiums	,	-72			
				ars' paymen						
20 40 60	4·263 5·993 8·810	4·244 6·032 8·771	4·215 5·962 8·729	4·157 5·919 8·656	030 +.030 010	-·048 -·031 -·081	-·106 -·074 -·154			
			20 yea	ars' payment	ts					
20 40 50	2·368 3·402 4·278	2·355 3·430 4·334	2.339 3.381 4.339	2·305 3·353 4·215	+·028 +·056	-·029 -·021 -·027	•63 •63			
			30 yea	ars' paymen	ts		į			
20 40	1.753 2.630	1·741 2·659	1.250 1.20	1·702 2·584	013 +.029	•024 •020	•o51 •o46			
			Endown	nent Assura	nce					
				m 10 years						
20 40 50 60	9.053 9.137 9.363 10.055	9.043 9.155 9.417 10.028	9°044 9°133 9°358 9°992	9·034 9·126 9·350 9·954	-·010 +·018 +·054 -·027	-·009 -·004 -·063	013 011 011			
				m 20 years						
20 40 50	4·158 4·339 4·776	4·145 4·361 4·834	4·146 4·331 4·760	4·134 4·319 4·739	-·013 +·022 +·058	-·o12 -·o16	~·024 ~·020 ~·037			
	Term 30 years									
20 30 40	2·569 2·652 2·909	2·554 2·640 2·939	2·554 2·634 2·896	2·540 2·616 2·879	-·015 -·012 +·030	013 018 012	029 036 030			
	Term 40 years									
20	1.822	1.810	1.805	1.789	012	017	033			